

Extreme Precipitation Return Levels for Multiple Durations on a Global Scale

Gaby J. Gründemann¹, Enrico Zorretto², Hylke E. Beck³, Marc Schleiss⁴,
Nick van de Giesen¹, Marco Marani^{2,5}, Ruud J. van der Ent^{1,6}

¹Department of Water Management, Faculty of Civil Engineering and Geosciences, Delft University of Technology, Delft, Netherlands

²Division of Earth and Ocean Sciences, Duke University, Durham, NC, USA

³Department of Civil and Environmental Engineering, Princeton University, Princeton, USA

⁴Department of Geoscience and Remote Sensing, Faculty of Civil Engineering and Geosciences, Delft University of Technology, Delft, Netherlands

⁵Dipartimento di Ingegneria Civile, Edile ed Ambientale, Università degli Studi di Padova, Padova, Italy

⁶Department of Physical Geography, Faculty of Geosciences, Utrecht University, Utrecht, Netherlands

Key Points:

- Three methods are used to analyze global precipitation return levels for 3-hourly to 10-day durations
- The Metastatistical Extreme Value distribution shows the most coherent spatiotemporal patterns
- A globally consistent change from heavy to thin tails is observed with increasing duration

Corresponding author: Gaby J. Gründemann, g.j.gruendemann@tudelft.nl

Abstract

Quantifying the magnitude and frequency of extreme precipitation events is key in translating climate observations to planning and engineering design. Current efforts have mostly focused on the estimation of daily extremes using gauge observations. Due to the recent development of high-resolution global precipitation products it is now also possible to estimate extremes globally. This research aims to calculate extreme precipitation return levels for multiple durations on a global scale, using the global Multi-Source Weighted-Ensemble Precipitation (MSWEP) V2.2 dataset spanning the period 1979–2017 at 0.1° resolution. Both classical and novel extreme value distributions are used. The application of these different methods at the global scale provides valuable insight into the spatial patterns of extreme precipitation frequency. Our results show that estimates based on the traditional Generalized Extreme Value (GEV) distribution and Peak-Over-Threshold (POT) methods, which only use the largest events to estimate precipitation extremes, exhibit some spatial incoherence. The recently developed Metastatistical Extreme Value (MEV) distribution, on the other hand, includes all precipitation events and leads to smoother and more reliable spatial patterns of local extremes. Furthermore, for all three extreme value methods, a consistent shift from heavy to thin tails with increasing duration is observed and quantified. The generated extreme precipitation return levels and corresponding parameters are released as the Global Precipitation EXtremes (GPEX) dataset. We expect the dataset to be particularly relevant to serve as a benchmark, and to provide extreme precipitation estimates in otherwise data-sparse regions for local engineers and planners.

1 Introduction

Extreme precipitation events are a major contributor to natural disasters (CRED, 2019). Accurate estimates of the severity of such intense precipitation events are needed for an enhanced disaster risk understanding, such as that of floods and landslides. The urgency of this is indicated as the first priority of the Sendai Framework for Disaster Risk Reduction (UNSIDR, 2015). The accurate quantification of extremes is also necessary for infrastructure planning and design. Such spatiotemporal estimates of extreme precipitation, based on extreme value distributions (EVDs), are available, for example, for Australia (Ball et al., 2019), the Netherlands (Beersma et al., 2018), and the US (e.g., Perica et al., 2015, 2018). However, many countries and regions do not have sufficient local data available (van de Giesen et al., 2014; Kidd et al., 2017; Gründemann et al., 2018), such that spatially-distributed extreme precipitation estimates are not possible.

Several previous studies have developed global-scale datasets of extreme precipitation. Courty et al. (2019) calculated intensity-duration-frequency curves at the global scale and their scaling with different event durations using reanalysis data and the Generalized Extreme Value (GEV) distribution with fixed tail behavior. Donat et al. (2013) produced the HadEx-2 dataset, which contains 29 generic precipitation and temperature indices, although these indices are not based on EVDs. Furthermore, this dataset has a coarse 2.50° latitudinal \times 3.75° longitudinal resolution, with data-gaps due to insufficient available gauge data. Other global studies mostly focused on examining which type of distribution is most suitable to capture the tail behavior of extreme precipitation (Papalexiou et al., 2013; Cavanaugh & Gershunov, 2015; Cavanaugh et al., 2015). In addition, the spatial patterns of the parameter controlling the tail decay have been studied for the GEV distribution (Papalexiou & Koutsoyiannis, 2013; Ragulina & Reitan, 2017), and the Generalized Pareto (GP) distribution (Serinaldi & Kilsby, 2014). However, several issues remain to be addressed in order to obtain a comprehensive global scale comparative benchmark of extreme precipitation return levels: 1) the choice of dataset, 2) the focus on daily durations, 3) the choice of the time blocks over which block-maxima are determined, and

70 4) the exploration of possible alternatives to the classical EVDs and the associated un-
 71 certainty with respect to the tail behavior.

- 72 1. Several (quasi-)global gridded precipitation datasets have been developed in the
 73 recent years, each with strengths and weaknesses. See Sun et al. (2018) and Beck,
 74 Pan, et al. (2019) for recent overviews of available datasets. Most of these datasets
 75 are based on gauge, reanalysis, or satellite sensor data. Notable examples of gauge-
 76 based datasets include GPCP-FDR (Becker et al., 2013; Schneider et al., 2011)
 77 and REGEN (Contractor et al., 2020). However, gauges are extremely unevenly
 78 distributed across the globe (Schneider et al., 2014; Kidd et al., 2017), and the num-
 79 ber of active gauges has been declining in recent decades (Mishra & Coulibaly, 2009).
 80 Satellite-based products such as CMORPH (Joyce et al., 2004), GSMaP (Ushio
 81 et al., 2009), IMERG (Huffman et al., 2015), and PERSIANN (Hong et al., 2004)
 82 have a relatively high spatio-temporal resolution. However, they do not cover re-
 83 gions outside of 60°N/S, and are only available from 2000 onwards, which signif-
 84 icantly hinders their use for extreme value analyses. Precipitation products with
 85 a true global coverage and long records are reanalyses, such as ERA-5 (Hersbach
 86 et al., 2020), JRA-55 (Kobayashi et al., 2015), and MERRA-2 (Gelaro et al., 2017).
 87 However, reanalysis products tend to exhibit strong systematic biases in the mag-
 88 nitude and frequency of precipitation (Decker et al., 2012; Ménégoz et al., 2013;
 89 Liu et al., 2018).
- 90 2. Global-scale analyses of precipitation extremes are generally based on daily pre-
 91 cipitation records (Koutsoyiannis, 2004b, 2004a; Papalexiou & Koutsoyiannis, 2013;
 92 Papalexiou et al., 2013; Serinaldi & Kilsby, 2014; Cavanaugh et al., 2015; Ragulina
 93 & Reitan, 2017). In practice, however, multiple durations are needed for the de-
 94 sign of infrastructure (e.g., Nissen & Ulbrich, 2017) or urban drainage networks
 95 (e.g., Mailhot & Duchesne, 2009). It is known that precipitation extremes of dif-
 96 ferent durations scale differently with temperature (Wasko et al., 2015), but lit-
 97 tle is known about the variation of EVD properties (tail behavior) for different
 98 temporal resolutions. Studies that did derive extreme precipitation statistics for
 99 durations ranging from minutes to a few days have mostly focused on small re-
 100 gions (McGraw et al., 2019; Nissen & Ulbrich, 2017; Overeem et al., 2008).
- 101 3. Studies estimating return levels of extreme precipitation by using annual maxima
 102 typically use calendar years to delineate the annual periods from which maxima
 103 values are extracted (e.g., Villarini et al., 2011; Papalexiou & Koutsoyiannis, 2013;
 104 Marani & Zanetti, 2015; Ragulina & Reitan, 2017; De Paola et al., 2018). When
 105 the variable of interest is river discharge instead of precipitation, however, hydro-
 106 logical years are typically used instead of calendar years to determine the annual
 107 maxima (Ward et al., 2016). For discharge values this is important, since peak dis-
 108 charge and flooding could occur during 31 December to 1 January transition and
 109 one event would be included in two calendar years. Although not often considered,
 110 this could also happen for precipitation. The annual maxima method could pick
 111 multiple values from a single rainy season that may, for example, be highly influ-
 112 enced by the El Niño-/Southern Oscillation, which is known to impact precipita-
 113 tion extremes (Rasmusson & Arkin, 1993; Allan & Soden, 2008).
- 114 4. The Generalized Extreme Value distribution (GEV), the most widely used EVD,
 115 is typically fitted through one of two approaches: a) using annual maximum pre-
 116 cipitation series and maximum likelihood (Coles, 2001) or L-moment (Hosking,
 117 1990) estimation approaches, or b) using a Peak-Over-Threshold (POT) method
 118 to fit a Generalized Pareto Distribution to excesses above the threshold and a Pois-
 119 son process to the sequence of threshold exceedances (Coles, 2001). In contrast
 120 to GEV and POT, the recently developed Metastatistical Extreme Value (MEV)
 121 distribution is fitted using all events with recorded precipitation instead of only
 122 the most severe. The inclusion of more events reduces the uncertainty due to sam-
 123 pling effects, which is important when dealing with short time series (Marani &

124 Ignaccolo, 2015; Zorzetto et al., 2016; Marra et al., 2018; Zorzetto & Marani, 2019;
 125 Miniussi & Marani, 2020). This is particularly advantageous when analyzing short
 126 remote sensing precipitation products, as the commonly applied GEV requires many
 127 years of data to accurately estimate the tail of the distribution (Papalexiou & Koutsoyiannis, 2013).
 128 Additionally, GEV parameter estimation depends heavily on a
 129 few large values, which makes it very sensitive to the possible presence of outliers,
 130 a relatively common occurrence in remote sensing estimates of precipitation amounts
 131 (Zorzetto & Marani, 2020). The GEV tail behavior is mostly controlled by its shape
 132 parameter, which is very sensitive to sampling effects and the choice of the method
 133 used for estimation. To overcome these problems, some studies have suggested to
 134 use one universal value of the shape parameter that is applicable to the whole world
 135 Koutsoyiannis (2004b, 2004a), or a shape parameter value within a narrow range
 136 between exponential and heavy-tail behavior (Papalexiou & Koutsoyiannis, 2013),
 137 or one shape parameter per region, that is similar within climate types and ele-
 138 vation ranges (Ragulina & Reitan, 2017). The estimation of the shape parame-
 139 ter is particularly difficult with short data series, though crucial for the accurate
 140 estimation of extremes.

141 In this study we contribute to overcome these issues by 1) using a dataset that merges
 142 all three main sources of precipitation data, 2) estimating extremes for several event du-
 143 rations, 3) using hydrological years in our analyses, and 4) comparing results from three
 144 different extreme value methods (GEV, POT and MEV). Specifically, we are interested
 145 in exploring the global scale behavior of extreme precipitation and the spatiotemporal
 146 variation of extreme value distributional tails. Moreover, this study aims to provide a
 147 comparative benchmark of extreme precipitation return levels based on a state-of-the-
 148 art global precipitation data (MSWEP) for further research and application.

149 **2 Materials and Methods**

150 **2.1 Data**

151 For the present study we selected the global gridded Multi-Source Weighted-Ensemble
 152 Precipitation (MSWEP-V2.2) dataset, due to its global coverage, long temporal span (1979–
 153 2017), high spatial (0.1°) and temporal (3-hourly) resolution, and robust performance
 154 compared to other precipitation datasets in numerous evaluations (e.g., Alijanian et al.,
 155 2017; Beck et al., 2017; Sahlou et al., 2017; Bai & Liu, 2018; Casson et al., 2018; Beck,
 156 Pan, et al., 2019; Satgé et al., 2019; Zhang et al., 2019). We refer to Beck, Wood, et al.
 157 (2019) for a comprehensive description of the dataset. We used data from 1 January 1979
 158 to 31 October 2017 at a 0.1° latitude \times 0.1° longitude resolution at 3-hourly time steps.
 159 MSWEP was derived by merging five different satellite- and reanalysis-based precipita-
 160 tion datasets globally. The dataset is one of the few with daily (as opposed to monthly)
 161 gauge corrections, applied using a scheme that accounts for gauge reporting times (Beck,
 162 Wood, et al., 2019). We selected all land-cells between 90°N and 58°S for our analysis.

163 **2.1.1 Quality Control**

164 The integration of erroneous gauge observations into MSWEP-V2.2 has sometimes
 165 resulted in implausible precipitation values. Therefore, we implemented a three-step qual-
 166 ity control procedure of the 3-hourly data prior to the analysis. We first discarded nega-
 167 tive values, which are physically impossible. The second step was to discard outliers,
 168 which we defined as the values deviating from the mean more than 30 standard devia-
 169 tions. We also discarded data surrounding the outliers for the same time step using a
 170 11×11 grid-cell window, as erroneous gauge observations may have influenced surround-
 171 ing cells in the production of the MSWEP dataset. The third step was to discard years
 172 with > 30 missing days or < 5 ‘wet’ 3-hourly periods, identified using a threshold of

173 0.2 mm 3h⁻¹ following Wasko et al. (2015), and to discard cells with < 30 years of data
 174 remaining, as a minimum record length of 30 years is customary and recommended to
 175 obtain reliable results (Arguez & Vose, 2011; Kendon et al., 2018; Westra et al., 2013).

176 **2.1.2 Durations**

177 We aggregated the 3-hourly data to create longer duration time series. The dura-
 178 tions we selected for our analysis are 3, 6, 12, and 24 hours, and 2, 3, 5, and 10 days. To
 179 avoid any overlap in the aggregated data, we took the n -hour window that contains the
 180 highest precipitation volume among all n -hour running windows that overlap with it, dis-
 181 carding the overlapping windows with lower precipitation volumes.

182 **2.1.3 Hydrological Year**

183 A common challenge in global-scale assessments is the delineation of the hydrolog-
 184 ical year, given the regional variability in the climatological precipitation seasonality. We,
 185 therefore, developed an uniform way to define the hydrological year. To avoid splitting
 186 one rainy season over two different years, we computed the median of the monthly pre-
 187 cipitation for each grid-cell, and defined the start of the hydrological year to be the first
 188 day of the driest month. Supplemental Figure S1a shows the starting month of the hy-
 189 drological year as determined by this method. As MSWEP-V2.2 spans 1 January 1979
 190 to 31 October 2017, we discarded the data prior to the start of the hydrological year and
 191 kept the 38 complete years. Only if the hydrological year starts in December there are
 192 just 37 complete years, which occurs in 5.8% of the cells.

193 We also investigated whether there is a significant difference between the use of cal-
 194 endar and hydrological years for the estimated daily extremes for GEV and MEV. The
 195 POT method is based on the values over a high threshold, irrespective of when they oc-
 196 curred. Therefore, there is by definition no difference in calculating the extremes using
 197 hydrological or calendar years for the POT method. To determine the difference for GEV
 198 and MEV, we first calculated the daily return levels for normal calendar years, using the
 199 MSWEP data from 1979 to 2016. Second, we calculated the return levels for the same
 200 distributions and the same years, by removing the months before the start of the hydro-
 201 logical year from the year 1979 and adding them to the year 2016. We did this in order
 202 to use the exact same data, so the differences in the return level estimates are solely due
 203 to a different starting month.

204 **2.2 Extreme Value Distributions**

205 Three extreme value distributions were fitted to the MSWEP data to calculate ex-
 206 treme precipitation return levels: the GEV, POT, and MEV. Annual (hydrological year)
 207 maxima were used to estimate the three parameters of the GEV using the L-moments
 208 approach, because of its robust performance for small samples (Hosking, 1990). The GEV
 209 cumulative distribution function (CDF) is given by:

$$G(z) = \begin{cases} \exp \left\{ - \left[1 + \xi \left(\frac{z-\mu}{\sigma} \right) \right]^{-\frac{1}{\xi}} \right\}, & \xi \neq 0 \\ \exp \left\{ -\exp \left[- \left(\frac{z-\mu}{\sigma} \right) \right] \right\}, & \xi = 0 \end{cases} \quad (1)$$

210 with location parameter $\mu \in (-\infty, \infty)$, scale parameter $\sigma > 0$, and shape parameter
 211 $\xi \in (-\infty, \infty)$.

212 Another common approach for precipitation extremes is the POT method. Instead
 213 of annual maxima, values exceeding a high threshold are described using a GP distri-
 214 bution, while the frequency of threshold exceedances is modelled as a Poisson point pro-
 215 cess (Davison & Smith, 1990; Coles, 2001). This framework also yields GEV as the re-
 216 sulting extreme value distribution, which is then used to determine the quantile corre-

217 sponding to a given return period. The GP CDF is given by:

$$H(y) = \begin{cases} 1 - \left(1 + \frac{\xi y}{\beta}\right)^{-\frac{1}{\xi}}, & \xi \neq 0 \\ 1 - \exp\left(-\frac{y}{\beta}\right), & \xi = 0 \end{cases} \quad (2)$$

218 where the precipitation excess over threshold $y > 0$, and is defined by scale parameter $\beta > 0$ and shape parameter $\xi \in (-\infty, \infty)$. The POT-threshold could be a fixed precipitation magnitude, or a fixed percentile to account for local variations in precipitation climatologies. As we are interested in studying the distribution of precipitation extremes across many different climate regions, a fixed number is not suitable and the 98th percentile of all ‘wet’ 3-hourly periods at a given location was chosen as a threshold. For dry areas and long aggregation time periods, the number of observations above the 98th percentile can be very small. For such instances, we lowered the threshold to equal the number of events to the number of years. The method used to fit the parameters of the GP distributions was the Probability Weighted Moments (PWM; Hosking & Wallis, 1987).

The MEV distribution is a recently developed statistical distribution (Marani & Ignaccolo, 2015; Zorzetto et al., 2016; Hosseini et al., 2020; Miniussi, Marani, & Villarini, 2020; Miniussi, Villarini, & Marani, 2020). All “ordinary” (as opposed to just extreme) precipitation events, i.e. all events above a small threshold, are used to infer this extreme value distribution. The threshold we applied is $0.2 \text{ mm } 3\text{h}^{-1}$, coinciding with the earlier defined threshold for a wet event. Weibull parameters were estimated for each year based on all wet events using the PWM method (Greenwood et al., 1979), equivalent to the use of L-moments, following the same approach as Zorzetto et al. (2016). The MEV-Weibull CDF is given by:

$$\zeta_m(x) = \frac{1}{M} \sum_{j=1}^M \left\{ 1 - \exp \left[- \left(\frac{x}{C_j} \right)^{w_j} \right] \right\}^{n_j} \quad (3)$$

229 where j is the hydrological year ($j = 1, 2, \dots, M$), $C_j > 0$ is the Weibull scale parameter, $w_j > 0$ is the Weibull shape parameter, and n_j is the number of wet events in hydrological year j (Marani & Ignaccolo, 2015).

232 **2.2.1 Observed Return Period**

233 The MSWEP dataset analyzed here has 38 complete years of data. Therefore, the return period associated with the maximum value on record computed according to the Weibull empirical frequency estimate is $T_{\text{observed}} = 39$ years. However, only 91 % of all cells had 38 complete years of data, so the maximum observed return period is sometimes lower: for 7 % of the cells only 37 years were available, and for 2 % of the cells 36 years or less were available. However, for simplicity we still refer to this maximum return period as T_{39} in the results.

240 **2.2.2 Tail Behavior**

241 Both the GEV and MEV distributions are flexible and can describe different tail behaviors and can, therefore, be used to investigate how different locations may be more or less sensitive to large extremes. The tail behavior of the two distributions differs, as illustrated in Figure 1 for different combinations of scale and shape parameters. The shape parameter of the GEV distribution, ξ , obtained either through the annual maxima or POT approach, encodes the nature of the tail of the distribution. Based on the value of ξ , the GEV can take one of three forms: a positive GEV shape parameter ($\xi > 0$, “Fréchet”) corresponds to a power-law tail, i.e., to a slowly-decaying probability of large events. This heavy-tail behavior contrasts with the case of an exponential tail ($\xi = 0$, “Gumbel”),

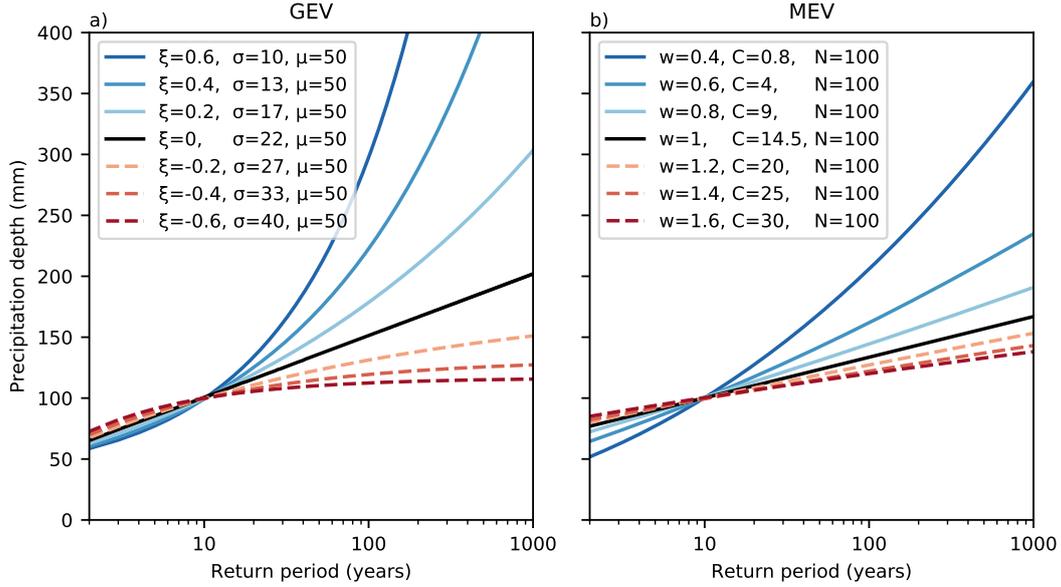


Figure 1. (a) Behavior of the shape (ξ) and scale (σ) parameters of the GEV distribution, with a constant location parameter (μ), and (b) behavior of the shape (w) and scale (C) parameters of the MEV-Weibull distribution, with a constant number of events (N). The results for MEV have been obtained with constant w , C and N parameters for each year. The values of the shape and scale parameter pairs have been chosen such that they all have a precipitation depth of approximately 100 mm for a 10-year return period.

250 and with the case of a distribution with an upper end point, which corresponds to neg-
 251 ative values of the shape parameter ($\xi < 0$, “inverse Weibull”).

252 The MEV distribution (Figure 1b) used here is based on the assumption of a Weibull
 253 distribution, which is not heavy-tailed, according to the traditional definition. However,
 254 when the shape parameter of the Weibull distribution w is less than one, it does induce
 255 a tail behavior which is intermediate between an exponential and a heavy-tail (algebraic
 256 decay), also known as sub-exponential, albeit with a characteristic scale (Laherrere &
 257 Sornette, 1998; Wilson & Toumi, 2005). When the Weibull shape parameter equals one,
 258 the exponential case arises. For a shape larger than one, the distribution becomes hyper-
 259 exponential with a fast decreasing tail, while still retaining an infinite upper end point.
 260 Hence, the shape parameter of the Weibull distribution can also be used to character-
 261 ize the propensity of a site to be subjected to large extreme events (Wilson & Toumi,
 262 2005; Zorzetto et al., 2016).

263 3 Results and Discussion

264 3.1 Hydrological Year

265 Figure 2 shows the frequency distribution of the difference between the values of
 266 T_{1000} estimated using calendar and hydrological years for GEV and MEV. The figure
 267 shows more values centered around $\pm 0.5\%$ difference for GEV than for MEV, indicat-
 268 ing that for GEV there are more cells that have the same T_{1000} estimate irrespective
 269 of which type of years is used. However, the figure also shows a larger spread of the 5th
 270 and 95th percentile as well as a wider decay of both tails for GEV compared to MEV.
 271 This implies that larger differences between hydrological and calendar years occur for

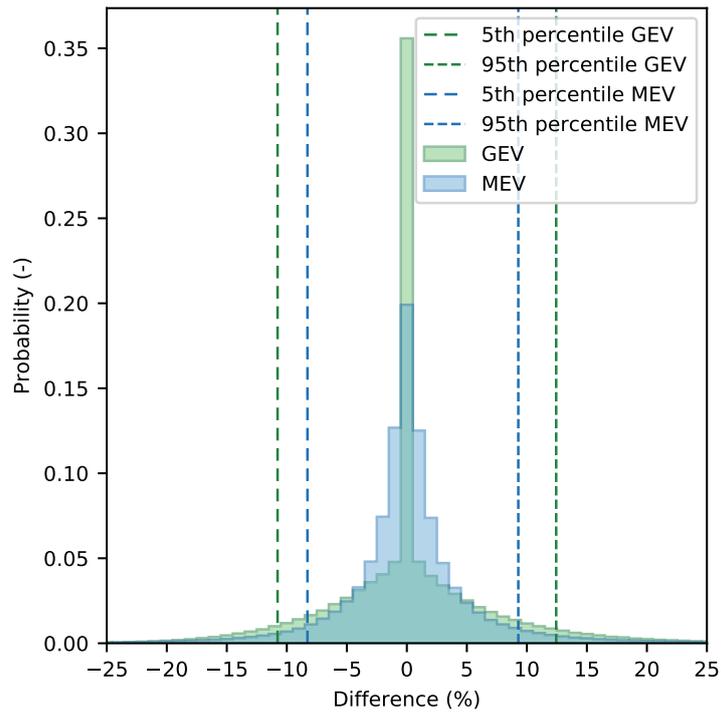


Figure 2. Weighted histogram showing the percentage difference in the values of $T1000$ calculated using calendar years and hydrological years. Included in the figure are all cells where the start of the hydrological year is different than the calendar year (i.e., the hydrological year does not start in January, see Supplement Figure 1a). A negative difference indicates that the $T1000$ estimate is larger using hydrological years, whereas a positive difference indicates that the $T1000$ estimate is larger using calendar years.

272 GEV. The spatial distribution of these differences is presented in Supplement Figure S1b
 273 for GEV and Figure S1c for MEV. The differences are most pronounced in the South-
 274 ern hemisphere, and at locations where the start of the hydrological year is around June
 275 (around the Mediterranean, in the Middle-East, in Brazil, around Indonesia, and in the
 276 western US).

277 When the hydrological year starts around June, the hydrological year is offset by
 278 approximately six months compared to a calendar year. As a result, there could be many
 279 different events included in the hydrological years compared to the calendar years, re-
 280 sulting in different annual maxima. GEV only uses annual maxima of this relatively short
 281 time series, so when there is a large difference in the annual maxima of the hydrologi-
 282 cal and calendar years, the difference in their estimated extremes are large. When the
 283 hydrological year starts in the winter months, the hydrological year is only shifted by
 284 a few months. In such instances, the annual maxima mostly stay the same between the
 285 calendar and hydrological years. For GEV this means that for many cells there is little
 286 difference in the $T1000$ estimates.

287 MEV, on the other hand, is fitted to all precipitation events in a year. Therefore,
 288 even if the hydrological year is only shifted by a couple months, there will be small dif-
 289 ferences in $T1000$ estimates. However, another implication of MEV being fitted to all
 290 precipitation events, is that it is not as sensitive to the inclusion or exclusion of very ex-
 291 treme events. When the hydrological year is offset by approximately six months, there
 292 are differences between the hydrological and calendar year estimates, but they are smaller
 293 than for GEV.

294 3.2 Extreme Precipitation Estimates

295 Figure 3 shows the 100-year precipitation return levels for a 24-hour duration. Other
 296 durations and return periods can be viewed by accessing the Global Precipitation EX-
 297 tremes (GPEX) dataset (Gründemann, 2020). The global spatial patterns of the extremes
 298 estimated by GEV and MEV are similar to Zorretto and Marani (2020, their Figure 9),
 299 while the spatial pattern of the underlying GEV parameters are consistent with Courty
 300 et al. (2019, their Figure 1). The overall spatial pattern of return levels of all three meth-
 301 ods is similar, although large regional differences can be observed. The GEV and POT
 302 results are similar in magnitude and show similar differences when compared to MEV.
 303 The estimated precipitation extremes are generally lower for both GEV and POT com-
 304 pared to MEV. MEV exhibits smooth spatial patterns, whereas the spatial patterns us-
 305 ing GEV and POT seem more irregular, consistent with the results of Zorretto and Marani
 306 (2020) for the conterminous US. The lack of spatial coherence in patterns of extremes
 307 for GEV and POT is particularly evident in the Great Plains of North America, and in
 308 Northern Russia, Southeast Asia, and Central Africa. Furthermore, the figure shows a
 309 large number of distinct circular areas with heavier extremes, particularly for GEV and
 310 POT, reflecting the locations of gauges used in MSWEP for precipitation correction (Beck,
 311 Wood, et al., 2019). MEV has a reduced sensitivity to the presence of outliers and the
 312 most extreme events, which may dominate the fitting of GEV.

313 In order to study the ability of the three distributions to capture the spatial co-
 314 herence of the estimated extremes, we selected several case studies. They collectively cover
 315 several climates, the locations of which can be found in Figure 3a. Within a single case
 316 study area, we expect the precipitation estimates to be similar because of their precip-
 317 itation generating mechanisms (Cavanaugh & Gershunov, 2015; Cavanaugh et al., 2015)
 318 or elevation (Ragulina & Reitan, 2017). Figure 4a shows the coefficient of variation (CV)
 319 for these case studies. The CV is the ratio of the standard deviation to the mean and
 320 is used to compare the relative variation between the study areas. The higher the CV,
 321 the higher the spread of the precipitation estimates within an area. This figure shows
 322 quite similar behavior for GEV and POT, though POT is slightly higher overall. The

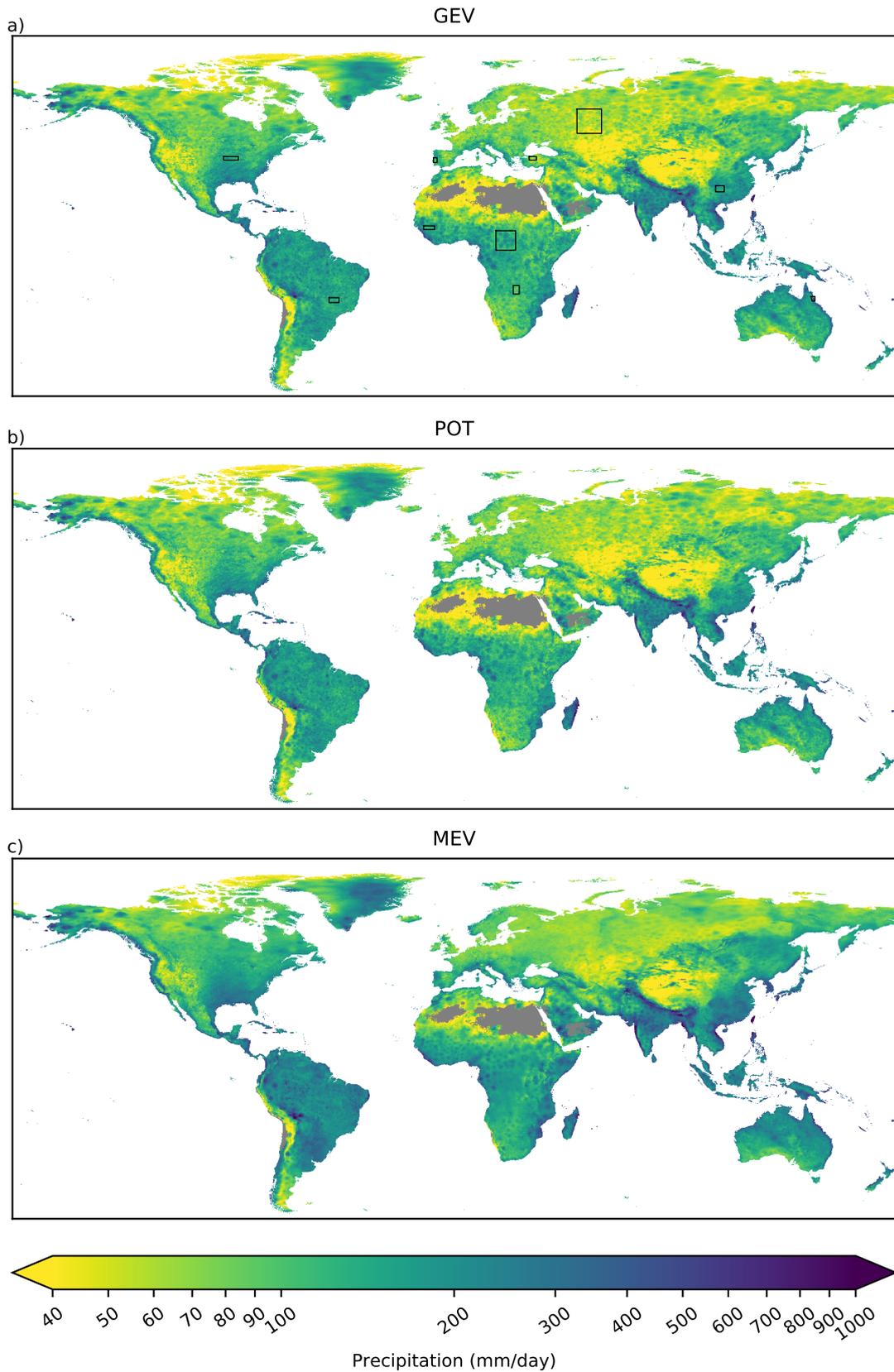


Figure 3. Precipitation return levels with a duration of 24-hours for a 100-year return period for different extreme value distributions: (a) the Generalized Extreme Value (GEV) distribution, (b) the Peak Over Threshold (POT) method, and (c) the Metastatistical Extreme Value (MEV) distribution. The black rectangles in panel a are the case studies corresponding to the areas in Figure 4.

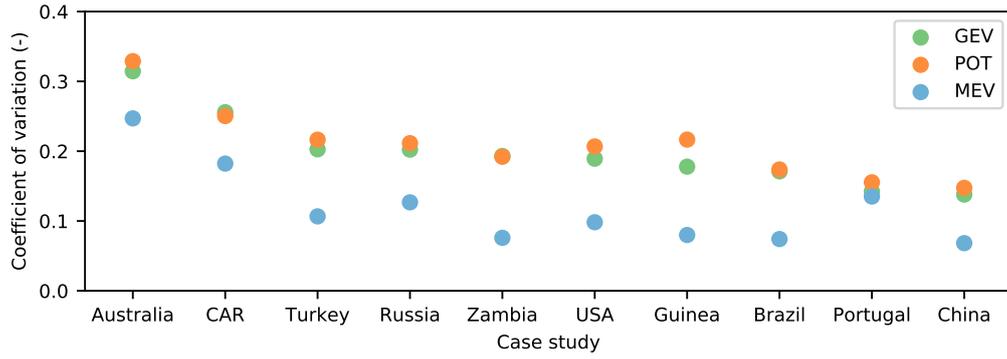


Figure 4. Coefficient of variation for the difference in estimated T_{100} quantiles for the three extreme value methods for 24-hour precipitation at selected case studies. The coefficient of variation is the standard deviation of the precipitation divided by the mean precipitation. The locations of the case studies are displayed in Fig 3a.

CV for MEV is lower. This implies that the internal variability within the study areas is lower using MEV, and that MEV gives a more coherent estimate of the extremes.

To further investigate the differences in magnitude between the three methods, we examine the extremes for each distribution using a spatially weighted mean over the global land surface. This is displayed for several return periods and durations as depth-duration-frequency curves (Figure 5). We first compare the maximum precipitation in the dataset to the predicted precipitation from each distribution. As there are 38 complete years of MSWEP data, the maximum observed return period is 39 years (T_{39} observed, the black dotted line in Figure 5). Doing this for individual grid cells would make little sense as the observed return level may be very different from the true return level, but globally averaged this should be close. For the lower duration events, the observed T_{39} is close to the T_{39} of all three distributions. Both GEV and POT show an underestimation and MEV overestimates in these cases. For increasing durations, the deviation between observed and estimated increases, particularly for MEV. This figure also shows again that the differences between GEV and POT are small. The global average estimated extremes for GEV and POT are notably lower than for MEV, as was already visible from Figure 3. This difference is more pronounced for larger return periods and longer durations.

One reason for the higher estimations of MEV may be that MEV improves significantly for return times that are much larger than the length of the observational time series. Zorzetto et al. (2016) and Marra et al. (2018, 2019) showed that MEV only has a 20% estimation error for a return period greater than 5 times the length of the time series. Another potential reason for this could be our use of a fixed threshold for defining a precipitation event. Zorzetto et al. (2016), Zorzetto and Marani (2019), and Zorzetto and Marani (2020) use a minimum of 1 mm day^{-1} for a precipitation event, and as we have 3-hour durations instead of solely daily we lowered this threshold. This results in a high number of sub-daily precipitation events per year. We also note that a running parameter based on auto-correlation to distinguish between separate events would have reduced the number of sub-daily events (Marra et al., 2018), but characterizing the temporal correlation of precipitation fields worldwide was outside the scope of our analysis. For a small number of events, occurring for long durations and dry regions, there are few events in each hydrological year to fit the distribution, which results in increased estimation uncertainty. It is, therefore, possible that this leads to an overestimation by MEV for particularly long durations and return periods. To overcome this, windows of two or

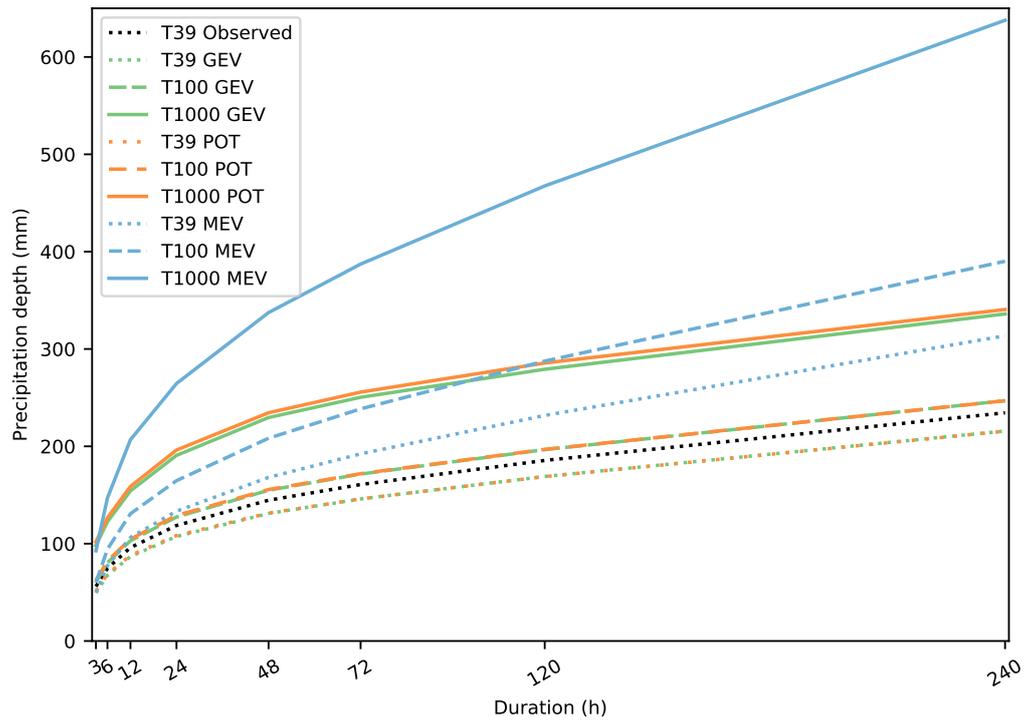


Figure 5. Area-weighted average depth-duration-frequency curves for the global land surface. T39 Observed is the mean spatially weighted maximum precipitation observed in the MSWEP-V2.2 dataset.

more years could result in a better parameter estimation (Miniussi & Marani, 2020). Therefore, the high MEV extreme precipitation estimates are not necessarily a characteristic of the MEV distribution, but rather our application of it.

3.3 Tail Behavior

To better understand the differences between extremes estimated using the three extreme value methods, we analyze their shape parameters in more detail as they provide information on the tail behavior (see section 2.2.2). Figure 6 presents the shape parameter for a 24-hour duration worldwide for each of the three distributions. Both GEV (Figure 6a) and POT (Figure 6b) exhibit a large spatial variability in addition to a low spatial coherence. This makes it difficult to discern spatial patterns, though some can be distinguished. In the Amazon, for instance, the shape parameter is relatively low, suggesting a tail with upper limit, while in Southern Australia it is quite high, denoting heavy tail behavior. This map roughly corresponds to the spatial patterns of the daily GEV shape parameter shown by Papalexioiu and Koutsoyiannis (2013, their Figure 13) and GP shape parameter shown by Ragulina and Reitan (2017, their Figure 4), although the actual values differ. We also find that for GEV and POT cells associated with a heavy tail may be adjacent to cells with thin tails. This conflicts with the expected spatial correlation of precipitation regimes, which should lead to spatial coherence (Ragulina & Reitan, 2017). Furthermore, GEV and POT do not always show the same type of tail, heavy or thin, at the same location. In 75 % of the cases the sign of the shape parameter agrees, while in 25 % of the cases the signs are different.

The mean value of the shape parameters underlying the MEV-Weibull distribution (Figure 6c) shows a clear and coherent spatial pattern. At virtually all grid cells the average MEV-Weibull shape parameter indicates a heavy sub-exponential tail (shape parameter less than 1), and the spatial patterns of the shape parameter of MEV-Weibull are much smoother and there is more consistency in different geographic regions. This predominantly heavy-tail behavior for daily precipitation was expected based on previous studies (Papalexioiu & Koutsoyiannis, 2013; Papalexioiu et al., 2013; Cavanaugh et al., 2015; Ragulina & Reitan, 2017), and is well captured by MEV. There are also clear topographical patterns visible, and Figure 6c has many similarities to a global elevation map. The mean shape values indicate thinner tails at higher elevations, such as the Rocky Mountains and Sierra Madres in North America, the Andes and large areas of the Brazilian Highlands in South America, the Ethiopian Highlands, the Scandinavian Mountains, and the Ural and Tibetan Plateau in Asia. These spatial patterns are in contrast with what Papalexioiu et al. (2018, their Figure 6) found for hourly Weibull tails in the USA, where the heaviest tails are in the mountainous areas, and the thin tails are in the southeast. However, our results correspond well to Ragulina and Reitan (2017, their Figure 4), who showed that the GP shape parameter decreases with elevation.

We find, however, that the relationship of the Weibull shape parameter with elevation is more complicated. Heavier tails are generally observed on the leeward side of large mountain ranges, and thinner tails on the windward side that is dominated by orographically enhanced frontal precipitation (Figure 6c). This is for example visible in the Rocky Mountains, Indonesia, and Norway, and corresponds to findings of Cavanaugh and Gershunov (2015, their Figure 5), who showed that exponential tails are observed in regions where extreme precipitation is predominantly generated by one type of system. There are also some similarities with the main climate types of the Köppen-Geiger classification (Peel et al., 2007; Beck et al., 2018). We observe thinner tails in Western Europe, which has a warm temperate climate. Heavier tails are located in both hot and dry Southern Europe and cold and snowy regions as parts of the Alps. Interestingly, in the USA this pattern is almost reversed, where the heavier tails are in the arid areas of the Great Plains and in the warm temperate climates in the southeast, whereas thinner tails are predominantly observed in the colder mountainous regions.

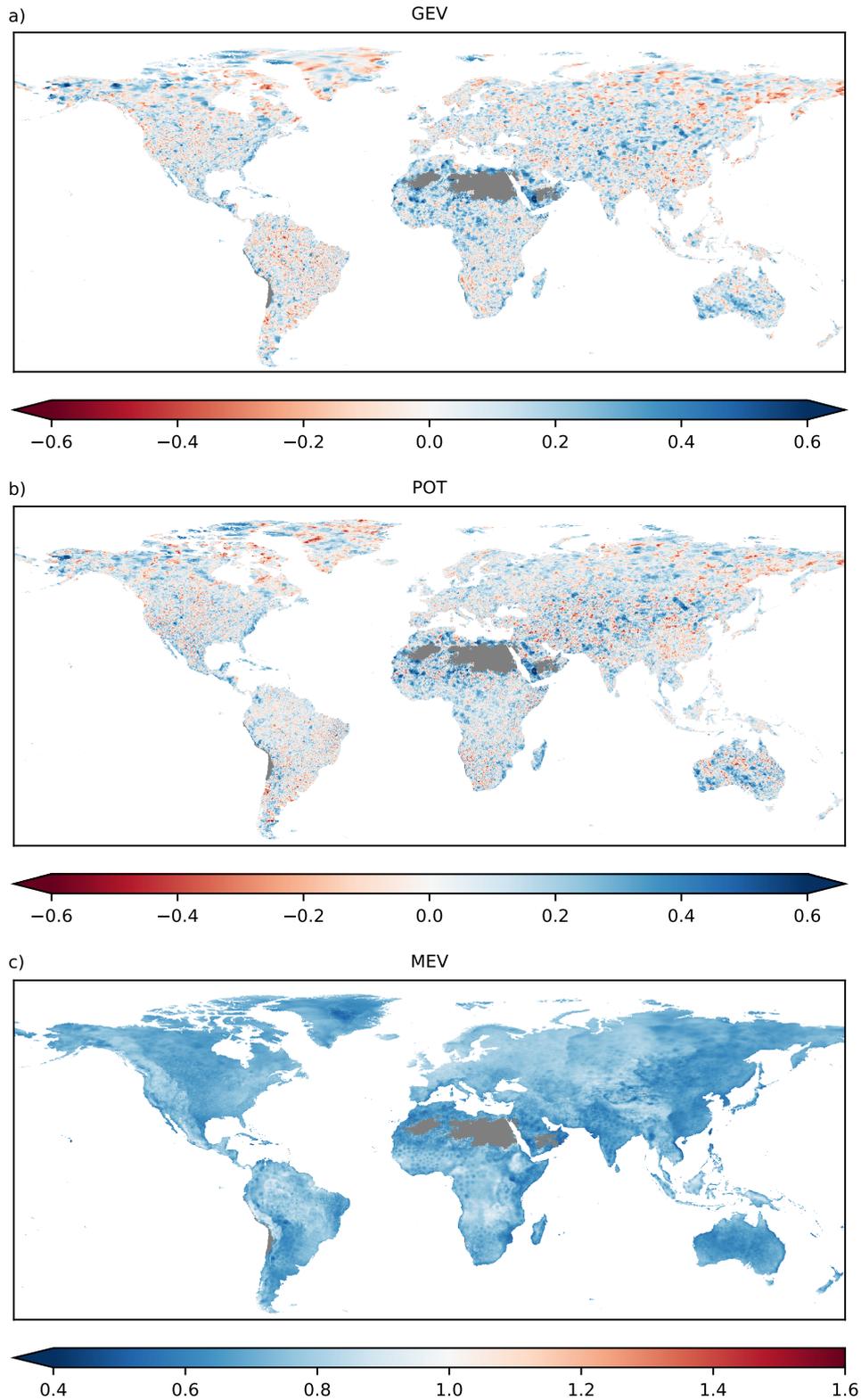


Figure 6. The shape parameter ($-$) for daily precipitation calculated for different extreme value methods: (a) GEV, equation 1 — ξ_{GEV} , (b) POT, equation 2 — ξ_{GP} and (c) MEV, equation 3 — w . For MEV, the mean shape parameter of all yearly Weibull distributions is displayed. Red indicates a thin tail (for GEV and POT with an upper limit, for MEV hyper-exponential), white an exponential tail, and blue a heavy tail (power-law for GEV and POT, sub-exponential for MEV, see section 2.2.2 and Figure 1).

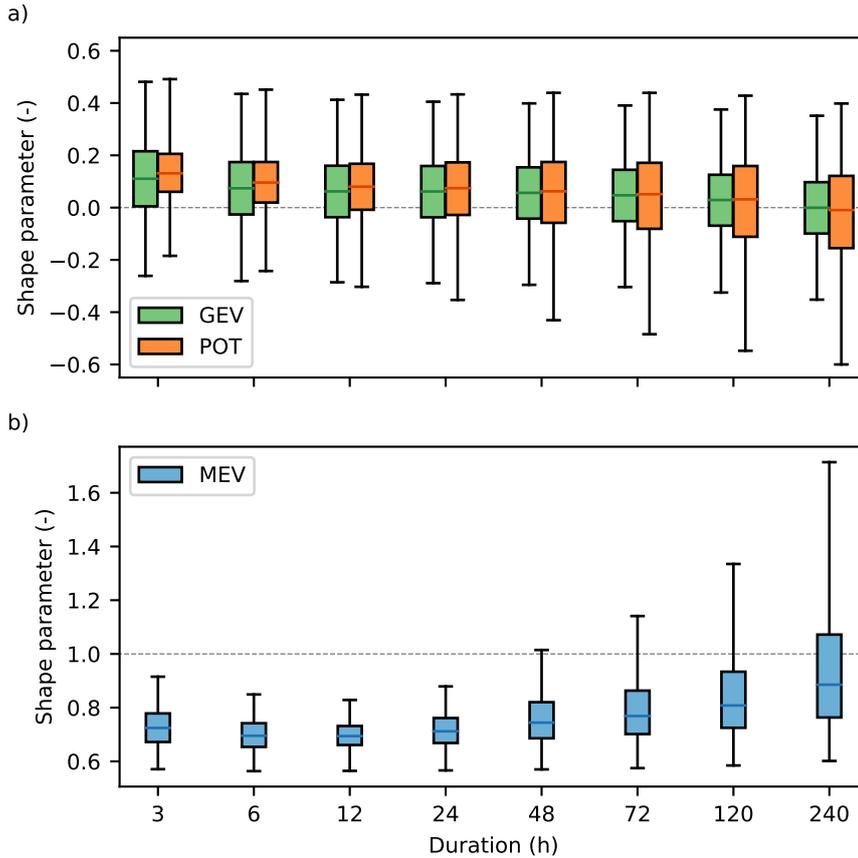


Figure 7. Boxplots showing the distribution of the shape parameter for different durations and extreme value methods: (a) GEV and POT, and (b) MEV. The whiskers denote the 1st and 99th percentiles. The top and bottom of the boxes represent the 75th and 25th percentiles, respectively. The dashed gray horizontal lines indicate exponential tails. Section 2.2.2 and Figure 1 explain the effect of the shape parameter on the tail behavior.

408 This demonstrates that the MEV method is able to produce spatially coherent tail
 409 behavior of precipitation extremes using a time series of just 38 years. This smoothness
 410 of MEV was obtained by doing a single-cell analysis, as opposed to the need to identify
 411 homogeneous regions using a clustering scheme (Demirdjian et al., 2018). This is a promis-
 412 ing result as MEV can thus be applied on a single grid-cell-basis using much shorter time
 413 series than GEV, enabling extreme value analysis in those regions of the world where only
 414 short precipitation time series are available.

415 A comparison of shape parameters for different durations is presented in Figure 7.
 416 Heavier tails are observed for short durations, and thinner tails for long durations. This
 417 is in line with the findings of Cavanaugh and Gershunov (2015), who found that longer
 418 duration extremes exhibit thinner tails. For GEV and POT the longer durations largely
 419 indicate tails with an upper limit. This occurs for instance in half of the cases for a du-
 420 ration of 10 days. MEV shows some thin tails for long durations, but not as many as for
 421 GEV and POT. Furthermore, there are clear patterns for the variability of the shape pa-
 422 rameter. POT shows an increasing spatial variability of the shape for an increasing du-
 423 ration, indicated by longer whiskers. For GEV, on the other hand, the variability remains

Table 1. Variables included in the GPEX dataset.

Variable	Description
GEV estimate	Extreme precipitation return levels estimated using GEV (mm)
POT estimate	Extreme precipitation return levels estimated using POT (mm)
MEV estimate	Extreme precipitation return levels estimated using MEV (mm)
Observed estimate	Observed extreme precipitation return levels (mm)
GEV location parameter	Location parameter of the GEV distribution
GEV scale parameter	Scale parameter of the GEV distribution
GEV shape parameter	Shape parameter of the GEV distribution
POT location parameter	Location parameter for a GEV distribution estimated by fitting the GP distribution
POT scale parameter	Scale parameter for a GEV distribution estimated by fitting the GP distribution
POT shape parameter	Shape parameter for a GEV distribution estimated by fitting the GP distribution
MEV scale parameter	Scale parameter of the MEV distribution for each hydrological year
MEV shape parameter	Shape parameter of the MEV distribution for each hydrological year
MEV number of events	Number of events per hydrological year, n parameter of the MEV distribution
Annual maxima	Annual maximum precipitation for each hydrological year (mm)
Start hydrological year	Number indicating the month in which the hydrological year starts
Land mask	Mask used for this study to indicate land cells and ocean cells

424 constant for different durations. For MEV, an increasing duration shows first a slight
 425 decrease of the variability with a heavier tails, and then after 12 hours a steady increase
 426 of the spread with thinner tails.

427 4 Dataset Usage Notes

428 The GPEX dataset created in this study is available at the 4TU repository (Gründemann,
 429 2020). It provides openly accessible and readily available hydrologically relevant return
 430 levels of extreme precipitation estimates worldwide. It contains the precipitation esti-
 431 mates of the three extreme values distributions, the observed estimates, the parameters
 432 of the three distributions, as well as a few other variables used in this study (Table 1).
 433 In this section we provide some possible uses of the dataset, and instructions and dis-
 434 claimers for proper use, both for large or regional-scale usage as well as for a single cell
 435 or point-scale.

436 4.1 Large-Scale Applications

437 The GPEX dataset contains global scale extreme precipitation estimates and its
 438 parameters at a spatial resolution of 0.1° , covering 3-hourly to 10-day durations. The
 439 dataset contains information about precipitation extremes for the entire Earth's land sur-
 440 face except Antarctica. The estimates of three distributions as well as the return lev-
 441 els as observed are included in the dataset. Among the three distributions, the traditional
 442 GEV and POT provide comparable large-scale average extremes, although differences
 443 can be substantial at smaller scales. When using the dataset at regional scales, we ad-
 444 vise taking the average of the precipitation estimates, as neighboring cells could differ.
 445 Note that since only 38 years of data were available, the fitted model parameters and
 446 associated return values are subject to considerable uncertainty. Furthermore, we acknowl-
 447 edge that the use of just one dataset does not represent the true uncertainty in the gen-
 448 eration of the dataset created. We do not think this affects our results for observed global
 449 spatial patterns significantly, but in a practical setting we recommend verifying the es-
 450 timates with local observations if available, and to reproduce the precipitation return
 451 level estimates with a full uncertainty range estimation.

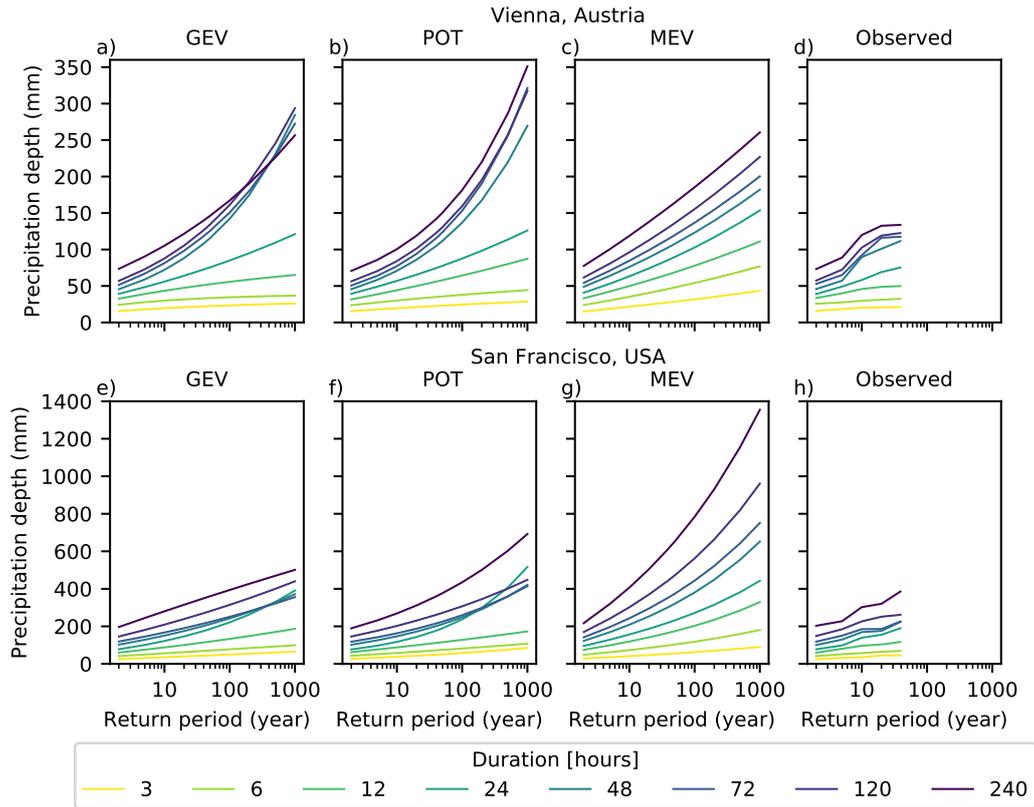


Figure 8. Return level plots for specific locations and different distributions. a–d) Vienna, Austria (48.234°N , 16.415°E) and e–g) San Francisco, California, USA (37.784°N , 122.400°W). Observed (d & h) are the annual maxima converted to return periods.

452 The novel MEV distribution provides more spatially coherent patterns of the ex-
 453 tremes. Its mean shape parameter for daily events captures the (heavy-)tail behavior,
 454 and follows orographic patterns. The extremes estimated by MEV are higher than those
 455 estimated by GEV and POT. However, for large return periods and long durations, MEV
 456 can overestimate the extremes, due to the small number of events available for the fit-
 457 ting. We, therefore, recommend analyzing the extremes of all three distributions to ob-
 458 tain an indication of the uncertainty.

459 4.2 Small-Scale Applications

460 The dataset is also suitable for small-scale applications either in comparative stud-
 461 ies or for direct use in data sparse regions, but one should be aware of the different sta-
 462 tistical characteristics of point-scale and grid-scale. Due to averaging effects in gridded
 463 datasets, precipitation extremes of point-scale observations are higher (Sivapalan & Blöschl,
 464 1998; De Michele et al., 2001; Ensor & Robeson, 2008; Cavanaugh & Gershunov, 2015;
 465 Zorzetto & Marani, 2019). Illustrative examples of two locations, Vienna and San Fran-
 466 cisco, are included in Figure 8. Analysis of the return level plots shows the estimates of
 467 the three distributions compared to the observed ones. We converted the annual maxi-
 468 mum precipitation to ‘observed’ return levels (Figure 8d+h). It should be kept in mind
 469 though that these ‘observed’ return levels are also different from the ‘true’ return lev-
 470 els. For (sub-)daily durations and low return periods, there is generally a good agree-
 471 ment between the observed return levels and the estimates of the three EVDs. For longer

472 durations and return periods, however, the estimated extremes deviate from the observed
 473 extremes. This is seen in San Fransisco (Figure 8e-h) where MEV overestimates and GEV
 474 and POT underestimate the extremes.

475 Furthermore, increasing event durations result in lower shape parameters, which
 476 was seen for all three distributions (Figure 7). An implication of this is that for long-
 477 durations the shape parameter indicates a finite endpoint (GEV and POT), or a very
 478 thin tail (MEV), while heavier tails are generally observed for short-durations. When
 479 estimating very large return periods (e.g., T_{500}), it is therefore possible for shorter du-
 480 ration estimates to be more intense than the corresponding quantiles computed for longer
 481 durations, which is physically impossible (see also Figure 8a,b,e and f). We should thus
 482 be careful when calculating very large return periods for multiple durations based on short
 483 records and interpreting the estimated extremes.

484 To get a better understanding of the range and uncertainty of a single cell loca-
 485 tion, we recommend to look at return level plots of the three distributions at the cell of
 486 interest in combination with its neighboring cells. This is particularly important for GEV
 487 and POT, due to the absence of coherent spatial patterns and the erratic manifestation
 488 of the tail behaviors. Previous results (Zorzetto et al., 2016) show that the benefits of
 489 MEV over GEV are greater for large return periods relative to the sample size available
 490 for the fit. Hence, for the estimation of large quantiles, MEV may be presumed to be
 491 more accurate. Depending on the practical application one could then choose to use the
 492 most extreme value, use the MEV value, or a spatial average of the GEV and POT es-
 493 timates.

494 5 Conclusions

495 In this study we have fitted three different extreme value methods (GEV, POT,
 496 and MEV) to a global precipitation dataset, MSWEP V2.2, to estimate extreme precip-
 497 itation return levels for several durations. The estimated precipitation extremes for the
 498 three approaches as well as their parameters have been published in the openly avail-
 499 able GPEX dataset (Gründemann, 2020). Instead of using calendar years to delineate
 500 between different years, we used hydrological years. We demonstrated that there is a sub-
 501 stantial difference in the extremes depending on the definition of yearly blocks used in
 502 the extreme value analysis (Figure 2). These differences were most notable in the South-
 503 ern hemisphere, and in locations where the driest month is around June (Figure S1). Al-
 504 though there is no systematic bias, we still recommend to apply the extreme value anal-
 505 yses for estimating extreme precipitation based on hydrological years in future studies.
 506 Our analysis indicates that this issue can be particularly relevant in the Southern hemi-
 507 sphere and in regions characterized by marked seasonal cycles.

508 It is well known that the traditional GEV and POT methods require very long data-
 509 series for accurate estimation of the tail parameter, and our study confirms that there
 510 is a low spatial coherence for the tail properties of both distributions (Figure 6a and b).
 511 The tail properties of the MEV distribution are spatially more coherent (Figure 6c) and
 512 hence the estimated return levels are spatially coherent as well (Figure 3c). In partic-
 513 ular, the analysis of the MEV-Weibull shape parameter reveals distinct spatial patterns.
 514 The shape parameter appears to be significantly controlled by orographic patterns, with
 515 the Weibull tail becoming thinner at higher elevations, as well as on the windward side
 516 of large mountain ranges. This behavior, consistent with previous results obtained over
 517 the conterminous US (Zorzetto & Marani, 2020), shows that the MEV distribution is able
 518 to capture spatially consistent tail behavior from short time series, a promising result
 519 for regions without long local precipitation records. Furthermore, our study shows that
 520 for all three distributions, the tail behavior decreases with increased event duration (Fig-
 521 ure 7). For GEV and POT for instance, about half of the cells have a thin tail with a
 522 finite endpoint for a duration of 10 days. We also conclude that both GEV and POT gen-

erally underestimate the observed extremes, whereas MEV overestimates them (Figure 5). This occurs particularly for long-duration extremes and large return periods. We do consider it likely, however, that the results of the MEV estimation could be improved, for instance by changing the event threshold, by using a running parameter to separate events (Marra et al., 2018), or by fitting the Weibull distribution over two or more years for dry areas and long durations.

The global gridded 0.1° GPEX dataset we created covers extreme precipitation return levels and its parameters for durations between 3 hours and 10 days. The dataset is freely available and could be used as a benchmark of precipitation extremes for data-scarce regions in particular. The dataset can be used for a single location, to obtain return level plots for planning and engineering design. For scientific purposes, not only the extreme precipitation estimates can be useful, but also the underlying parameters, allowing a more in-depth analysis of the local or regional circumstances. GPEX thus serves as a benchmark of global precipitation extremes with applications in global- and local-scale research as well as a reference for engineering practitioners in data sparse environments.

Acknowledgments

We acknowledge the use of the `mevpy` Python package (<https://github.com/EnricoZorzetto/mevpy>) for the extreme value analysis. This work was carried out on the Dutch National e-Infrastructure (DNI) with support from the SURF cooperative. Enrico Zorzetto acknowledges support from the NASA Earth and Space Science Fellowship 19-EARTH19R-27. Hylke Beck was supported in part by the U.S. Army Corps of Engineers' International Center for Integrated Water Resources Management (ICIWaRM). Nick van der Giesen acknowledges support of the European Commission's Horizon 2020 Programme under grant agreement number 776691 (TWIGA). Ruud van der Ent acknowledges funding from the Netherlands Organization for Scientific Research (NWO), project number 016.Veni.181.015.

Upon acceptance the GPEX dataset will be available at the 4TU repository (Gründemann, 2020). The data included are the extremes estimated using the different distributions, the observed extremes, and the parameters to estimate the extremes. These data are available for all durations included in this study. The resolution of the dataset is 0.1°, the resolution of the MSWEP-V2.2 dataset.

References

- Alijanian, M., Rakhshandehroo, G. R., Mishra, A. K., & Dehghani, M. (2017). Evaluation of satellite rainfall climatology using CMORPH, PERSIANN-CDR, PERSIANN, TRMM, MSWEP over Iran. *International Journal of Climatology*, *37*(14), 4896–4914. doi: 10.1002/joc.5131
- Allan, R. P., & Soden, B. J. (2008). Atmospheric warming and the amplification of precipitation extremes. *Science*, *321*(5895), 1481–1484. doi: 10.1126/science.1160787
- Arguez, A., & Vose, R. S. (2011). The definition of the standard WMO climate normal: The key to deriving alternative climate normals. *Bulletin of the American Meteorological Society*, *92*(6), 699–704. doi: 10.1175/2010BAMS2955.1
- Bai, P., & Liu, X. (2018). Evaluation of five satellite-based precipitation products in two gauge-scarce basins on the Tibetan Plateau. *Remote Sensing*, *10*(8). doi: 10.3390/RS10081316
- Ball, J., Babister, M., Nathan, R., Weeks, W., Weinmann, E., Retallick, M., & Testoni, I. (2019). *Australian Rainfall and Runoff: A guide to flood estimation*. Commonwealth of Australia (Geoscience Australia).
- Beck, H. E., Pan, M., Roy, T., Weedon, G. P., Pappenberger, F., van Dijk,

- 573 A. I. J. M., ... Wood, E. F. (2019). Daily evaluation of 26 precipitation
574 datasets using Stage-IV gauge-radar data for the CONUS. *Hydrology and*
575 *Earth System Sciences*, *23*, 207–224. doi: 10.5194/hess-23-207-2019
- 576 Beck, H. E., Vergopolan, N., Pan, M., Levizzani, V., van Dijk, A. I. M., Weedon,
577 G. P., ... Wood, E. F. (2017). Global-scale evaluation of 22 precipitation
578 datasets using gauge observations and hydrological modeling. *Hydrology and*
579 *Earth System Sciences*, *21*(12), 6201–6217. doi: 10.5194/hess-21-6201-2017
- 580 Beck, H. E., Wood, E. F., Pan, M., Fisher, C. K., Miralles, D. G., van Dijk,
581 A. I. J. M., ... Adler, R. F. (2019). MSWEP V2 Global 3-Hourly 0.1° Precip-
582 itation: Methodology and Quantitative Assessment. *Bulletin of the American*
583 *Meteorological Society*, *100*(3), 473–500. doi: 10.1175/BAMS-D-17-0138.1
- 584 Beck, H. E., Zimmermann, N. E., McVicar, T. R., Vergopolan, N., Berg, A.,
585 & Wood, E. F. (2018). Present and future köppen-geiger climate clas-
586 sification maps at 1-km resolution. *Scientific Data*, *5*(180214). doi:
587 10.1038/sdata.2018.214
- 588 Becker, A., Finger, P., Meyer-Christoffer, A., Rudolf, B., Schamm, K., Schneider,
589 U., & Ziese, M. (2013). A description of the global land-surface precipitation
590 data products of the global precipitation climatology centre with sample appli-
591 cations including centennial (trend) analysis from 1901-present. *Earth System*
592 *Science Data*, *5*(1), 71. doi: 10.5194/essd-5-71-2013
- 593 Beersma, J., Versteeg, R., & Hakvoort, H. (2018). *Neerslagstatistieken voor korte*
594 *duren actualisatie 2018* (Tech. Rep.). Amersfoort: STOWA.
- 595 Casson, D. R., Werner, M., Weerts, A., & Solomatine, D. (2018). Global re-analysis
596 datasets to improve hydrological assessment and snow water equivalent estima-
597 tion in a sub-arctic watershed. *Hydrology and Earth System Sciences*, *22*(9),
598 4685–4697. doi: 10.5194/hess-22-4685-2018
- 599 Cavanaugh, N. R., & Gershunov, A. (2015). Probabilistic tail dependence of intense
600 precipitation on spatiotemporal scale in observations, reanalyses, and GCMs.
601 *Climate Dynamics*, *45*(11-12), 2965–2975. doi: 10.1007/s00382-015-2517-1
- 602 Cavanaugh, N. R., Gershunov, A., Panorska, A. K., & Kozubowski, T. J. (2015).
603 The probability distribution of intense daily precipitation. *Geophysical Re-*
604 *search Letters*, *42*, 1560–1567. doi: 10.1002/2015GL063238
- 605 Coles, S. (2001). *An introduction to statistical modeling of extreme values*. Springer
606 Series in Statistics; Springer: London, UK.
- 607 Contractor, S., Donat, M., Alexandre, L. V., Ziese, M., Meyer-Christoffer, A.,
608 Schneider, U., ... Vose, R. S. (2020). Rainfall estimates on a gridded net-
609 work (regen)—a global land-based gridded dataset of daily precipitation from
610 1950 to 2016. *Hydrology and Earth System Sciences*, *24*(2), 919–943. doi:
611 10.5194/hess-24-919-2020
- 612 Courty, L. G., Wilby, R. L., Hillier, J. K., & Slater, L. J. (2019). Intensity-duration-
613 frequency curves at the global scale. *Environmental Research Letters*, *14*,
614 084045. doi: 10.1088/1748-9326/ab370a
- 615 CRED. (2019). *Natural disasters 2018* (Tech. Rep.). Brussels, Belgium: Institute
616 Health and Society UClouvain. Retrieved from [https://www.cred.be/sites/
617 default/files/CREDNaturalDisaster2018.pdf](https://www.cred.be/sites/default/files/CREDNaturalDisaster2018.pdf)
- 618 Davison, A. C., & Smith, R. L. (1990). Models for Exceedances over High Thresh-
619 olds. *Journal of the Royal Statistical Society. Series B (Methodological)*, *52*(3),
620 393–442. Retrieved from <https://www.jstor.org/stable/2345667>
- 621 De Paola, F., Giugni, M., Pugliese, M., Annis, A., & Nardi, F. (2018). GEV param-
622 eter estimation and stationary vs. non-stationary analysis of extreme rainfall in
623 African test cities. *Hydrology*, *5*(28). doi: 10.3390/hydrology5020028
- 624 Decker, M., Brunke, M. A., Wang, Z., Sakaguchi, K., Zeng, X., & Bosilovich, M. G.
625 (2012). Evaluation of the Reanalysis Products from GSFC, NCEP, and
626 ECMWF Using Flux Tower Observations. *Journal of Climate*, *25*(6), 1916-
627 1944. doi: 10.1175/JCLI-D-11-00004.1

- 628 De Michele, C., Kottegoda, N. T., & Rosso, R. (2001). The derivation of areal re-
629 duction factor of storm rainfall from its scaling properties. *Water Resources*
630 *Research*, *37*(12), 3247–3252. doi: 10.1029/2001WR000346
- 631 Demirdjjan, L., Zhou, Y., & Huffman, G. J. (2018). Statistical modeling of extreme
632 precipitation with trmm data. *Journal of applied meteorology and climatology*,
633 *57*(1), 15–30. doi: 10.1175/JAMC-D-17-0023.1
- 634 Donat, M. G., Alexander, L. V., Yang, H., Durre, I., Vose, R., Dunn, R. J., ...
635 Kitching, S. (2013). Updated analyses of temperature and precipitation
636 extreme indices since the beginning of the twentieth century: The HadEX2
637 dataset. *Journal of Geophysical Research: Atmospheres*, *118*, 2098–2118. doi:
638 10.1002/jgrd.50150
- 639 Ensor, L. A., & Robeson, S. M. (2008). Statistical characteristics of daily precipita-
640 tion: comparisons of gridded and point datasets. *Journal of Applied Meteorol-*
641 *ogy and Climatology*, *47*(9), 2468–2476. doi: 10.1175/2008JAMC1757.1
- 642 Gelaro, R., McCarty, W., Suárez, M. J., Todling, R., Molod, A., Takacs, L., ...
643 Zhao, B. (2017). The modern-era retrospective analysis for research and ap-
644 plications, version 2 (merra-2). *Journal of Climate*, *30*(14), 5419–5454. doi:
645 10.1175/JCLI-D-16-0758.1
- 646 Greenwood, J. A., Landwehr, J., Matalas, N., & Wallis, J. (1979). Probability
647 weighted moments: definition and relation to parameters of several distribu-
648 tions expressible in inverse form. *Water Resources Research*, *15*(5), 1049–
649 1054. doi: 10.1029/WR015i005p01049
- 650 Gründemann, G. J. (2020). *Global Precipitation EXtremes dataset*. doi: 10.4121/
651 uuid:12b5c941-cd54-45db-8d7b-efaaacecaa69
- 652 Gründemann, G. J., Werner, M., & Veldkamp, T. I. E. (2018). The potential of
653 global reanalysis datasets in identifying flood events in southern africa. *Hydro-*
654 *logy and Earth System Sciences*, *22*(9), 4667–4683. doi: 10.5194/hess-22-4667
655 -2018
- 656 Hersbach, H., Bell, B., Berrisford, P., Hirahara, S., Horanyi, A., Muñoz-Sabater, J.,
657 ... Thepaut, J.-N. (2020). The ERA5 global reanalysis. *Quarterly Journal of*
658 *the Royal Meteorological Society*, 1-51. doi: 10.1002/qj.3803
- 659 Hong, Y., Hsu, K.-L., Sorooshian, S., & Gao, X. (2004). Precipitation estimation
660 from remotely sensed imagery using an artificial neural network cloud clas-
661 sification system. *Journal of Applied Meteorology*, *43*(12), 1834–1853. doi:
662 10.1175/JAM2173.1
- 663 Hosking, J. R. M. (1990). L-Moments: Analysis and estimation of distributions us-
664 ing linear combinations of order statistics. *Journal of the Royal Statistical So-*
665 *ciety. Series B (Methodological)*, *52*(1), 105–124.
- 666 Hosking, J. R. M., & Wallis, J. R. (1987). Parameter and Quantile Estimation
667 for the Generalized Pareto Distribution Parameter and Quantile Estima-
668 tion Generalized Pareto Distribution. *Technometrics*, *23*(3), 339–349. doi:
669 https://doi.org/10.1080/00401706.1987.10488243
- 670 Hosseini, S. R., Scaioni, M., & Marani, M. (2020). Extreme Atlantic Hur-
671 ricane Probability of Occurrence Through the Metastatistical Extreme
672 Value Distribution. *Geophysical Research Letters*, *47*(1), 1–9. doi:
673 10.1029/2019GL086138
- 674 Huffman, G. J., Bolvin, D. T., Braithwaite, D., Hsu, K., Joyce, R., Kidd, C., ...
675 Xie, P. (2015). *NASA Global Precipitation Measurement (GPM) Integrated*
676 *Multi-satellitE Retrievals for GPM (IMERG): Algorithm Theoretical Basis*
677 *Document (ATBD) Version 4.5* (Tech. Rep.). NASA/GSFC, Greenbelt, MD
678 20771, USA.
- 679 Joyce, R. J., Janowiak, J. E., Arkin, P. A., & Xie, P. (2004). CMORPH: A
680 method that produces global precipitation estimates from passive microwave
681 and infrared data at high spatial and temporal resolution. *Journal of Hy-*
682 *drometeorology*, *5*(3), 487–503. doi: 10.1175/1525-7541(2004)005<0487:

- 683 CAMTPG>2.0.CO;2
- 684 Kendon, E. J., Blenkinsop, S., & Fowler, H. J. (2018). When will we detect changes
685 in short-duration precipitation extremes? *Journal of Climate*, *31*(7), 2945–
686 2964. doi: 10.1175/JCLI-D-17-0435.1
- 687 Kidd, C., Becker, A., Huffman, G. J., Muller, C. L., Joe, P., Skofronick-Jackson, G.,
688 & Kirschbaum, D. B. (2017). So, how much of the Earth’s surface is covered
689 by rain gauges? *Bulletin of the American Meteorological Society*, *98*, 69–78.
690 doi: 10.1175/BAMS-D-14-00283.1
- 691 Kobayashi, S., Ota, Y., Harada, Y., Ebata, A., Moriya, M., Onoda, H., . . . oth-
692 ers (2015). The jra-55 reanalysis: General specifications and basic charac-
693 teristics. *Journal of the Meteorological Society of Japan*, *93*(1), 5–48. doi:
694 10.2151/jmsj.2015-001
- 695 Koutsoyiannis, D. (2004a). Statistics of extremes and estimation of extreme rainfall:
696 II. Empirical investigation of long rainfall records. *Hydrological Sciences Jour-
697 nal*, *49*(4), 591–610. doi: 10.1623/hysj.49.4.591.54424
- 698 Koutsoyiannis, D. (2004b). Statistics of extremes and estimation of extreme rainfall:
699 I. Theoretical investigation. *Hydrological Sciences Journal*, *49*(4), 575–590.
700 doi: 10.1623/hysj.49.4.575.54430
- 701 Laherrere, J., & Sornette, D. (1998). Stretched exponential distributions in na-
702 ture and economy: “fat tails” with characteristic scales. *The European Physical
703 Journal B-Condensed Matter and Complex Systems*, *2*(4), 525–539.
- 704 Liu, Z., Liu, Y., Wang, S., Yang, X., Wang, L., Baig, M. H. A., . . . Wang, Z. (2018).
705 Evaluation of Spatial and Temporal Performances of ERA-Interim Precip-
706 itation and Temperature in Mainland China. *Journal of Climate*, *31*(11),
707 4347–4365. doi: 10.1175/JCLI-D-17-0212.1
- 708 Mailhot, A., & Duchesne, S. (2009). Design criteria of urban drainage infrastructures
709 under climate change. *Journal of Water Resources Planning and Management*,
710 *136*(2), 201–208. doi: 10.1061/(asce)wr.1943-5452.0000023
- 711 Marani, M., & Ignaccolo, M. (2015). A metastatistical approach to rainfall extremes.
712 *Advances in Water Resources*, *79*, 121–126. doi: 10.1016/j.advwatres.2015.03
713 .001
- 714 Marani, M., & Zanetti, S. (2015). Long-term oscillations in rainfall extremes in a
715 268 year daily time series. *Water Resources Research*, *51*, 639–647. doi: 10
716 .1002/2014WR015885
- 717 Marra, F., Nikolopoulos, E. I., Anagnostou, E. N., & Morin, E. (2018). Metastatis-
718 tical Extreme Value analysis of hourly rainfall from short records: Estimation
719 of high quantiles and impact of measurement errors. *Advances in Water Re-
720 sources*, *117*, 27–39. doi: 10.1016/j.advwatres.2018.05.001
- 721 Marra, F., Zoccatelli, D., Armon, M., & Morin, E. (2019). A simplified MEV
722 formulation to model extremes emerging from multiple nonstationary un-
723 derlying processes. *Advances in Water Resources*, *127*, 280–290. doi:
724 10.1016/j.advwatres.2019.04.002
- 725 McGraw, D., Nikolopoulos, E. I., Marra, F., & Anagnostou, E. N. (2019). Pre-
726 cipitation frequency analyses based on radar estimates: An evaluation over
727 the contiguous United States. *Journal of Hydrology*, *573*, 299–310. doi:
728 10.1016/j.jhydrol.2019.03.032
- 729 Ménégoz, M., Gallée, H., & Jacobi, H. W. (2013). Precipitation and snow cover in
730 the himalaya: from reanalysis to regional climate simulations. *Hydrology and
731 Earth System Sciences*, *17*(10), 3921–3936. doi: 10.5194/hess-17-3921-2013
- 732 Miniussi, A., & Marani, M. (2020). Estimation of daily rainfall extremes through
733 the metastatistical extreme value distribution: Uncertainty minimization
734 and implications for trend detection. *Water Resources Research*, *56*(7),
735 e2019WR026535. doi: 10.1029/2019WR026535
- 736 Miniussi, A., Marani, M., & Villarini, G. (2020). Metastatistical Extreme Value
737 Distribution applied to floods across the continental United States. *Advances*

- 738 *in Water Resources*, 136, 103498. doi: 10.1016/j.advwatres.2019.103498
- 739 Miniussi, A., Villarini, G., & Marani, M. (2020). Analyses Through the Metas-
740 tatistical Extreme Value Distribution Identify Contributions of Tropical
741 Cyclones to Rainfall Extremes in the Eastern United States *Geophysical*
742 *Research Letters*. *Geophysical Research Letters*, 47, e2020GL087238. doi:
743 10.1029/2020GL087238
- 744 Mishra, A. K., & Coulibaly, P. (2009). Developments in hydrometric network design:
745 A review. *Reviews of Geophysics*, 47(2). doi: 10.1029/2007RG000243
- 746 Nissen, K. M., & Ulbrich, U. (2017). Increasing frequencies and changing char-
747 acteristics of heavy precipitation events threatening infrastructure in Europe
748 under climate change. *Natural Hazards and Earth System Sciences*, 17(7),
749 1177–1190. doi: 10.5194/nhess-17-1177-2017
- 750 Overeem, A., Buishand, A., & Holleman, I. (2008). Rainfall depth-duration-
751 frequency curves and their uncertainties. *Journal of Hydrology*, 348, 124–134.
752 doi: 10.1016/j.jhydrol.2007.09.044
- 753 Papalexiou, S. M., Aghakouchak, A., & Foufoula-Georgiou, E. (2018). A diagnostic
754 framework for understanding climatology of tails of hourly precipitation ex-
755 tremes in the United States. *Water Resources Research*, 54(9), 6725–6738. doi:
756 10.1029/2018WR022732
- 757 Papalexiou, S. M., & Koutsoyiannis, D. (2013). Battle of extreme value distribu-
758 tions: A global survey on extreme daily rainfall. *Water Resources Research*,
759 49, 187–201. doi: 10.1029/2012WR012557
- 760 Papalexiou, S. M., Koutsoyiannis, D., & Makropoulos, C. (2013). How extreme
761 is extreme? An assessment of daily rainfall distribution tails. *Hydrology and*
762 *Earth System Sciences*, 17, 851–862. doi: 10.5194/hess-17-851-2013
- 763 Peel, M., Finlayson, B., & McMahon, T. (2007). Updated world map of the
764 Köppen–Geiger climate classification. *Hydrology and Earth System Sciences*,
765 11(5), 1633–1644. doi: 10.5194/hess-11-1633-2007
- 766 Perica, S., Pavlovic, S., St. Laurent, M., Trypaluk, C., Unruh, D., Martin, D., &
767 Wilhite, O. (2015). *NOAA Atlas 14 Volume 10, Precipitation-Frequency At-*
768 *las of the United States, Northeastern States*. (Vol. 11; Tech. Rep.). NOAA,
769 National Weather Service.
- 770 Perica, S., Pavlovic, S., St. Laurent, M., Trypaluk, C., Unruh, D., & Wilhite, O.
771 (2018). *NOAA Atlas 14: Precipitation-frequency atlas of the United States,*
772 *Texas* (Vol. 11; Tech. Rep.). NOAA, National Weather Service.
- 773 Ragulina, G., & Reitan, T. (2017). Generalized extreme value shape parameter and
774 its nature for extreme precipitation using long time series and the Bayesian
775 approach precipitation. *Hydrological Sciences Journal*, 62(6), 863–879. doi:
776 10.1080/02626667.2016.1260134
- 777 Rasmusson, E. M., & Arkin, P. A. (1993). A global view of large-scale precipitation
778 variability. *Journal of Climate*, 6(8), 1495–1522.
- 779 Sahlu, D., Moges, S. A., Nikolopoulos, E. I., Anagnostou, E. N., & Hailu, D. (2017).
780 Evaluation of high-resolution multisatellite and reanalysis rainfall products
781 over East Africa. *Advances in Meteorology*, 2017. doi: 10.1155/2017/4957960
- 782 Satgé, F., Ruelland, D., Bonnet, M.-P., Molina, J., & Pillco, R. (2019). Consis-
783 tency of satellite-based precipitation products in space and over time compared
784 with gauge observations and snow-hydrological modelling in the Lake Tit-
785 icaca region. *Hydrology and Earth System Sciences*, 23(1), 595–619. doi:
786 10.5194/hess-23-595-2019
- 787 Schneider, U., Becker, A., Finger, P., Meyer-Christoffer, A., Rudolf, B., & Ziese, M.
788 (2011). *GPCC full data reanalysis version 6.0 at 0.5°: Monthly land-surface*
789 *precipitation from rain-gauges built on GTS-based and historic data*. doi:
790 10.5676/DWD_GPCC/FD.M.V6.050
- 791 Schneider, U., Becker, A., Finger, P., Meyer-Christoffer, A., Ziese, M., & Rudolf,
792 B. (2014). GPCC’s new land surface precipitation climatology based on

- 793 quality-controlled in situ data and its role in quantifying the global wa-
 794 ter cycle. *Theoretical and Applied Climatology*, 115(1-2), 15–40. doi:
 795 10.1007/s00704-013-0860-x
- 796 Serinaldi, F., & Kilsby, C. G. (2014). Rainfall extremes: Toward reconciliation af-
 797 ter the battle of distributions. *Water Resources Research*, 50, 336–352. doi: 10
 798 .1002/2013WR014211
- 799 Sivapalan, M., & Blöschl, G. (1998). Transformation of point rainfall to areal rain-
 800 fall: Intensity-duration-frequency curves. *Journal of Hydrology*, 204(1-4), 150–
 801 167.
- 802 Sun, Q., Miao, C., Duan, Q., Ashouri, H., Sorooshian, S., & Hsu, K. (2018). A re-
 803 view of global precipitation data sets: Data sources, estimation, and intercom-
 804 parisons. *Review of Geophysics*, 56(1), 79–107. doi: 10.1002/2017RG000574
- 805 UNSIDR. (2015). *Sendai Framework for Disaster Risk Reduction 2015 - 2030* (Tech.
 806 Rep.). Retrieved from www.unisdr.org/we/inform/publications/43291
- 807 Ushio, T., Sasashige, K., Kubota, T., Shige, S., Okamoto, K., Aonashi, K., ...
 808 Kawasaki, Z. I. (2009). A kalman filter approach to the Global Satellite
 809 Mapping of Precipitation (GSMaP) from combined passive microwave and in-
 810 frared radiometric data. *Journal of the Meteorological Society of Japan*, 87 A,
 811 137–151. doi: 10.2151/jmsj.87A.137
- 812 van de Giesen, N., Hut, R., & Selker, J. (2014). The Trans-African Observatory
 813 (TAHMO). *WIREs Water*, 1, 341–348. doi: 10.1002/wat2.1034
- 814 Villarini, G., Smith, J. A., Lynn, M., Vitolo, R., Stephenson, D. B., & Krajew-
 815 ski, W. F. (2011). On the frequency of heavy rainfall for the Midwest
 816 of the United States. *Journal of Hydrology*, 400(1-2), 103–120. doi:
 817 10.1016/j.jhydrol.2011.01.027
- 818 Ward, P. J., Kumm, M., & Lall, U. (2016). Flood frequencies and durations and
 819 their response to El Niño Southern Oscillation: Global analysis. *Journal of Hy-
 820 drology*, 539, 358–378. doi: 10.1016/j.jhydrol.2016.05.045
- 821 Wasko, C., Sharma, A., & Johnson, F. (2015). Does storm duration modulate the
 822 extreme precipitation-temperature scaling relationship? *Geophysical Research
 823 Letters*, 42(20), 8783–8790. doi: 10.1002/2015GL066274
- 824 Westra, S., Alexander, L. V., & Zwiers, F. W. (2013). Global increasing trends in
 825 annual maximum daily precipitation. *Journal of Climate*, 26(11), 3904–3918.
 826 doi: 10.1175/JCLI-D-12-00502.1
- 827 Wilson, P. S., & Toumi, R. (2005). A fundamental probability distribu-
 828 tion for heavy rainfall. *Geophysical Research Letters*, 32(14812). doi:
 829 10.1029/2005GL022465
- 830 Zhang, D., Liu, X., Bai, P., & Li, X.-H. (2019). Suitability of satellite-based precip-
 831 itation products for water balance simulations using multiple observations in a
 832 humid catchment. *Remote Sensing*, 11(2). doi: 10.3390/rs11020151
- 833 Zorzetto, E., Botter, G., & Marani, M. (2016). On the emergence of rainfall ex-
 834 tremes from ordinary events. *Geophysical Research Letters*, 43(15), 8076–8082.
 835 doi: 10.1002/2016GL069445
- 836 Zorzetto, E., & Marani, M. (2019). Downscaling of rainfall extremes from
 837 satellite observations. *Water Resources Research*, 55, 156–174. doi:
 838 10.1029/2018WR022950
- 839 Zorzetto, E., & Marani, M. (2020). Extreme value metastatistical analy-
 840 sis of remotely sensed rainfall in ungauged areas: Spatial downscaling
 841 and error modelling. *Advances in Water Resources*, 135, 103483. doi:
 842 10.1016/j.advwatres.2019.103483