

## **To what extent are changes in flood magnitude related to changes in precipitation extremes?**

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### **Key points**

- In contrast to conventional beliefs, we find weak causal linkages between precipitation extremes and floods
- The spatial pattern of changes in precipitation extremes explains less than 20% of the spatial pattern of changes in floods
- Most catchments have a co-variation of less than 0.5 between annual precipitation extremes and annual floods

## Abstract

Despite increasing evidence of intensification of extreme precipitation events associated with a warming climate, the magnitude of extreme river flows is decreasing in many parts of the world. To better understand the range of relationships between precipitation extremes and floods, we analyzed annual precipitation extremes and flood events over the Contiguous United States from 1980 to 2014. A low spatial correlation (less than 0.2) between changes in precipitation extremes and changes in floods was found, attributable to a weak causal relationship. The co-variation between precipitation extremes and floods is also substantially low, with a majority of catchments having a coefficient of determination of less than 0.5, even among the catchments with a relatively strong causal relationship. The findings indicate a need for more investigations into causal mechanisms driving a non-linear response of floods to intensified precipitation extremes in a warming climate.

## 1 Introduction

Among the most important implications of global climate change is the intensification of the hydrologic cycle [Huntington, 2006], including the intensification of rainfall extremes [Westra *et al.*, 2014]. As air temperature rises, the water vapor held in the atmosphere also increases following the Clausius-Clapeyron relation [Clausius, 1850]. This relationship has been documented extensively in the climate literature [Donat *et al.*, 2013; Guerreiro *et al.*, 2018; Papalexiou and Montanari, 2019; Westra *et al.*, 2013], and has led to concerns of a future characterized broadly by an increase in the magnitude of global flood events.

Large scale investigations into changes in floods, however, indicate a broader range of global flood response, with many studies documenting sites with a decrease in flood magnitude [Do *et al.*, 2017; Do *et al.*, 2020b; Gudmundsson *et al.*, 2019; Hodgkins *et al.*, 2017; Kundzewicz *et al.*, 2004; Lins and Slack, 1999]. These somewhat unexpected relationships between trends in extreme precipitation and trends in extreme discharge can be attributed to the influence of other flood generation mechanisms such as soil moisture [Ivancic and Shaw, 2015; Wasko *et al.*, 2020; Ye *et al.*, 2017] and snow dynamics [Berghuijs *et al.*, 2016; Blöschl *et al.*, 2017; Do *et al.*, 2020a; Ledingham *et al.*, 2019; Stein *et al.*, 2020]. Even when floods are triggered by precipitation extremes, the relationship between precipitation magnitude and flood magnitude is likely non-linear [Sharma *et al.*, 2018], owing to the complex interactions of many variables which have undergone substantial changes such as land cover [Keenan *et al.*, 2015; Lambin *et al.*, 2003], river channels [Slater *et al.*, 2015; Yamazaki *et al.*, 2014] and evapotranspiration [Bosilovich *et al.*, 2005; Gronewold and Stow, 2014].

However, there is still a lack of quantitative understanding of the relationship between precipitation extremes and floods, partially due to unavailable discharge observations in many parts of the world [Do *et al.*, 2018; Do *et al.*, 2020b]. Even for regions with relatively good streamflow records, empirical investigations have primarily focused on the consistency between the timing of precipitation extremes and that of floods [Berghuijs *et al.*, 2019; Blöschl *et al.*, 2017; Do *et al.*, 2020a; Ivancic and Shaw, 2015; Stein *et al.*, 2020; Wasko *et al.*, 2020] rather than co-variation between precipitation extreme intensity and flood magnitude. As a result, it is difficult to promote generalized statements about global relationships between changes in precipitation extremes and changes in floods, which is essential to the design of robust flood prevention and mitigation strategies in a warming climate [Milly *et al.*, 2008].

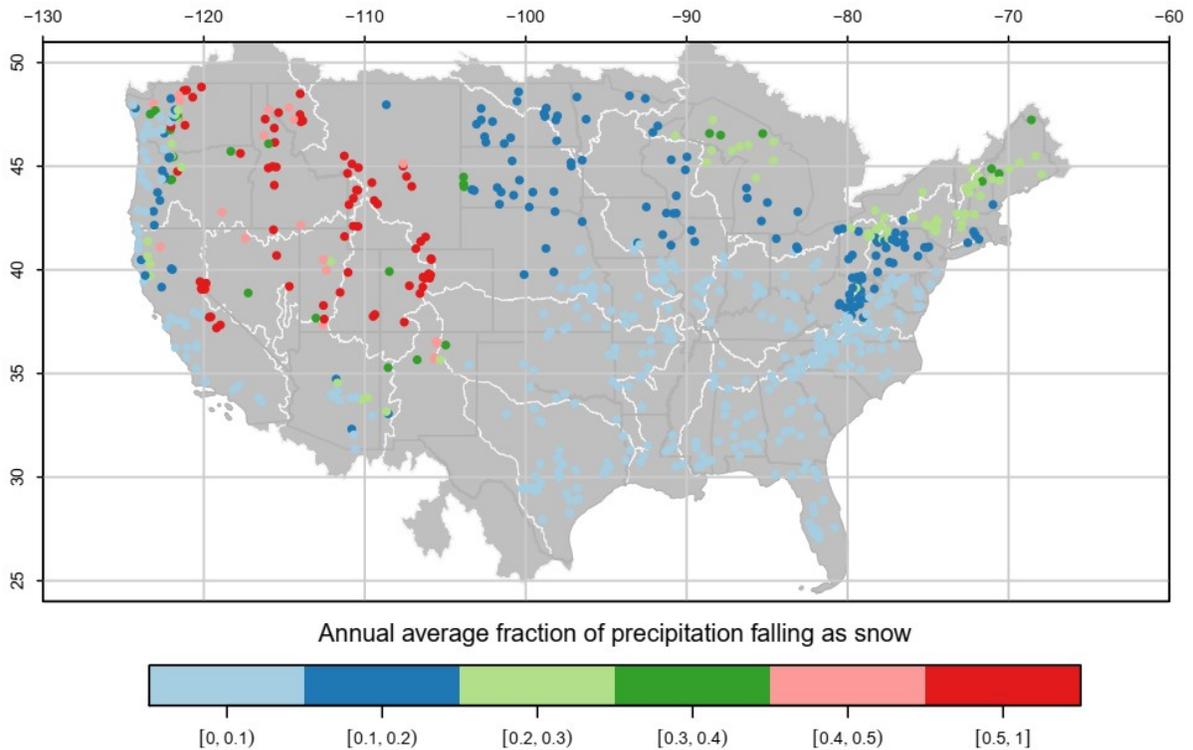
We aim to fill in this gap through an empirical assessment of the co-variation of precipitation extremes and flood magnitude using a large sample (671) of catchments across the Contiguous United States (CONUS) (Section 2.1). We used annual maxima streamflow from 1980 to 2014 from these catchments as the flood population, and we used three metrics of annual maxima precipitation to represent precipitation extremes (Section 2.2). Temporal changes in floods and precipitation extremes were then estimated at each catchment and the correlation between the spatial patterns of these trends was assessed (Section 2.3). The ordinal date of precipitation extreme events was then compared to that of annual flood events (Section 2.4) to assess potential causal relationships between these hydro-climatic extremes. Finally, the co-variation between the intensity of precipitation extremes and flood magnitude across catchments was assessed (Section 2.5) to evaluate the appropriateness of using changes in extreme precipitation as a proxy for changes in floods.

## 2 Data and Methods

### 2.1 Data

Data for our analysis was derived from the Catchment Attributes and Meteorology for Large-sample Studies (CAMELS) dataset [Addor *et al.*, 2017a; Newman *et al.*, 2015]. The CAMELS database aggregates a variety of hydrometeorological variables (primarily derived from other studies) for 671 catchments across the CONUS (the outlets of CAMELS catchments are shown in Figure 1). The catchments in the CAMELS database are intended to reflect relatively natural hydrologic conditions (the impervious surface area of each catchment is less than 5% of the total catchment area; see Newman *et al.* [2015] for more information). These catchments have a relatively small size (the median catchment area is 340.7 km<sup>2</sup>) and cover a range of climatic conditions (e.g., dry, temperate, and continental climates) as well as geographic features (e.g., mountains and deserts). Other variables in the CAMELS database include daily streamflow (originally obtained from the United States Geological Survey), catchment-average daily precipitation and temperature (derived from the Daymet dataset [Thornton *et al.*, 1997]), and daily evapotranspiration, simulated by the conceptual SAC-SMA model [Burnash *et al.*, 1973].

In addition to the hydro-meteorological data available through CAMELS, we also obtained soil moisture data from the NOAA Climate Prediction Center [Van den Dool *et al.*, 2003]. This dataset provides monthly soil moisture water height equivalent, simulated by a leaky bucket model, with a spatial resolution of 0.5-degree latitude x 0.5-degree longitude (i.e., a cell-size of about 2,500 km<sup>2</sup>). We used monthly soil moisture from the cell containing each catchment outlets as a proxy for catchment-wide to obtain soil moisture conditions from 1980 to 2014 for each of the CAMELS catchments. We judged this approach to be appropriate for the CAMELS catchments, which have a relatively small size.



**Figure 1.** Location of CAMELS catchment outlets across the CONUS. Each catchment was classified into one of the six groups based on the annual average fraction of annual precipitation falling as snow ( $f_{snow}$ ; categories were defined at 0.1 intervals, with all catchments having  $f_{snow}$  of at least 0.5 grouped into one category). The 18 major water resources regions over the CONUS are also shown (grey polygons bounded by white lines) for reference (see also Figure S1 in Supporting Information).

## 2.2 Identifying streamflow and precipitation extremes

Our approach to quantifying rainfall and streamflow extremes is based on the annual maxima (AMAX) index, one of the most common indices for assessing temporal changes in hydro-climatic extremes [Do *et al.*, 2017; Kundzewicz *et al.*, 2004; Ledingham *et al.*, 2019; Villarini and Smith, 2010; Westra *et al.*, 2013]. We first processed streamflow data to obtain the magnitude ( $Q_{MAX}$  index) and the timing ( $Q_{DOYMAX}$  index; defined as the ordinal date of annual floods) for each AMAX streamflow event. To reduce the chance of misattributing flood events, we omitted from our data set any years missing more than 15% of daily values (to have a compromise between data coverage and data quality).

We then processed daily precipitation to derive three sets of variables, each representing a different approach to quantifying precipitation extremes (a threshold of 15% missing data was also used to remove years with insufficient data.). The first variable representing precipitation extremes is AMAX precipitation (P), which is defined using the same approach to that of AMAX streamflow. The second precipitation variable allows us to assess whether constraining the timing of precipitation extremes to seasons where the catchments are wet, and when floods are more likely to occur [Ivancic and Shaw, 2015], could improve the relationship between

precipitation extremes and floods. This second variable is calculated as the annual maximum precipitation based only on months in which soil moisture was above-average. The third precipitation variable is intended to take into account catchment saturation and snow dynamics, and how they affect the propagation of precipitation into *effective* precipitation [Berghuijs *et al.*, 2016; Berghuijs *et al.*, 2019]. We calculated this variable using a coupled soil-snow routine [Berghuijs *et al.*, 2016; Hock, 2003; Stein *et al.*, 2020; Woods, 2009] based on daily precipitation, temperature, and evapotranspiration (all readily available in the CAMELS dataset). Details of this routine is provided in the Supporting Information; for further reading, see Stein *et al.* [2020].

Finally, we calculate the intensity and timing of each of the three precipitation AMAX variables, leading to a total of six new precipitation indices;  $P_{MAX}$  and  $P_{DOYMAX}$  (for the first precipitation variable),  $P_{sm,MAX}$  and  $P_{sm,DOYMAX}$  (for the second), and  $P_{eff,MAX}$  and  $P_{eff,DOYMAX}$  (for the third). Note that evapotranspiration is only available from October 1980 onward, thus  $P_{eff,MAX}$  and  $P_{eff,DOYMAX}$  are not available for 1980.

### 2.3 Assessing the correlation between the spatial pattern of changes in precipitation extremes and the spatial pattern of changes in floods

We calculated temporal changes in the magnitude of AMAX streamflow ( $Q_{MAX}$ ) and changes in precipitation extreme intensity ( $P_{MAX}$ ,  $P_{sm,MAX}$ , and  $P_{eff,MAX}$ ) using normalized Theil-Sen slope [Gudmundsson *et al.*, 2019; Stahl *et al.*, 2012] as follows:

$$\tau_c = \text{median} \left( \frac{x_j - x_i}{j - i} \right) \quad (1)$$

$$T_c = \frac{\tau_c \times 10 \text{ years}}{\hat{x}_c} \times 100 \quad (2)$$

where  $\tau_c$  is the Theil-Sen slope estimator for catchment  $c$ , which is defined as the median of the average annual difference in AMAX values ( $x$ ) between all possible pairs of years. The indices  $i$  and  $j$  represent year numbers such that  $i \in [1, n_c - 1]$ ,  $j \in [2, n_c]$ ,  $i < j$ , and  $n_c$  is the number of years in the data record (after the screening process described above) for each catchment.  $T_c$  is the normalized trend, expressed as a percentage of change per decade relative to the mean of all AMAX values in a catchment ( $\hat{x}_c$ ). This approach leads to four  $T_c$  values for each catchment, one for  $Q_{MAX}$  and three for precipitation intensity ( $P_{MAX}$ ,  $P_{sm,MAX}$ , and  $P_{eff,MAX}$ ).

To evaluate whether the spatial pattern of changes in floods can be explained by the spatial pattern of changes in precipitation extremes, we calculated the coefficient of determination  $R^2$  [Rao, 1973], which is the square of the correlation between the  $T_c$  values of  $Q_{MAX}$  and the  $T_c$  values of a precipitation extreme metric (e.g.,  $T_c$  of  $P_{MAX}$ ). The value of  $R^2$  ranges from 0 to 1, and a high  $R^2$  indicates a strong correlation.

### 2.4 Assessing the causal relationship between precipitation extremes and floods

To quantify the strength of a potential relationship between precipitation extremes and floods, we calculated the probability that an annual precipitation extremes would be followed closely (in time) by an annual flood extreme in each catchment (referred to as the co-occurrence probability hereafter). We matched the timing of AMAX precipitation events (represented by  $P_{DOYMAX}$ ,  $P_{sm,DOYMAX}$ , and  $P_{eff,DOYMAX}$  indices) and the timing of annual floods (represented by

$Q_{\text{DOYMAX}}$  index) and derive the fraction of annual flood events that can be directly attributed to a preceding precipitation extremes. To account for travel time required for precipitation to reach a catchment outlet, we adopted a previous approach [Ivancic and Shaw, 2015] and allowed a lag of up to 5 days. Specifically, we presume that there is a causal relationship between AMAX precipitation and AMAX streamflow if  $0 \leq Q_{\text{DOYMAX}} - P_{\text{DOYMAX}} \leq 5$ . If precipitation extremes and floods were independent, the random chance of a match, on average, would be less than 2%, which is the random chance of  $Q_{\text{DOYMAX}}$  having a value between  $P_{\text{DOYMAX}}$  and  $P_{\text{DOYMAX}} + 5$  (six days) of all possible days in a non-leap year (365 days).

Given the important relationship between the precipitation type (i.e. snow or rain) over much of the CONUS [Berghuijs *et al.*, 2016], as well as other parts of the world [Blöschl *et al.*, 2017; Do *et al.*, 2020a], we also assessed whether relationships between precipitation extremes and floods vary by precipitation type. To achieve this objective, we used the annual average proportion of precipitation that falls as snow, readily available in Addor *et al.* [2017a], and is denoted as  $f_{\text{snow}}$ . Each catchments was classified into one of the six categories; the first five are defined by  $f_{\text{snow}}$  values between 0.0 and 0.5 at intervals of 0.1; the sixth category includes all catchments with an  $f_{\text{snow}}$  value between 0.5 and 1.0 (see Figure 1). We then assessed whether there are significant differences in the co-occurrence probability of precipitation extremes and floods across  $f_{\text{snow}}$  classification categories.

## 2.5 Assessing temporal correlation between the intensity of precipitation extremes and flood magnitude

We identified catchments with similar causal relationship between precipitation extremes and flood by grouping catchments into seven groups according to the co-occurrence probability at 0.1 intervals (note that all catchments with at least 0.6 co-occurrence probability grouped into one category). We then measured the co-variation between the intensity of precipitation extremes (e.g.,  $P_{\text{MAX}}$  index) and the magnitude of floods ( $Q_{\text{MAX}}$  index) at each catchment using the coefficient of determination  $R^2$ . The  $R^2$  values were then analyzed alongside the co-occurrence probability to quantify the extent of which changes in precipitation extremes are useful to infer changes in flood magnitude.

## 3 Results and Discussions

### 3.1 A low correlation between the spatial pattern of changes in floods and the spatial pattern of changes in precipitation extremes

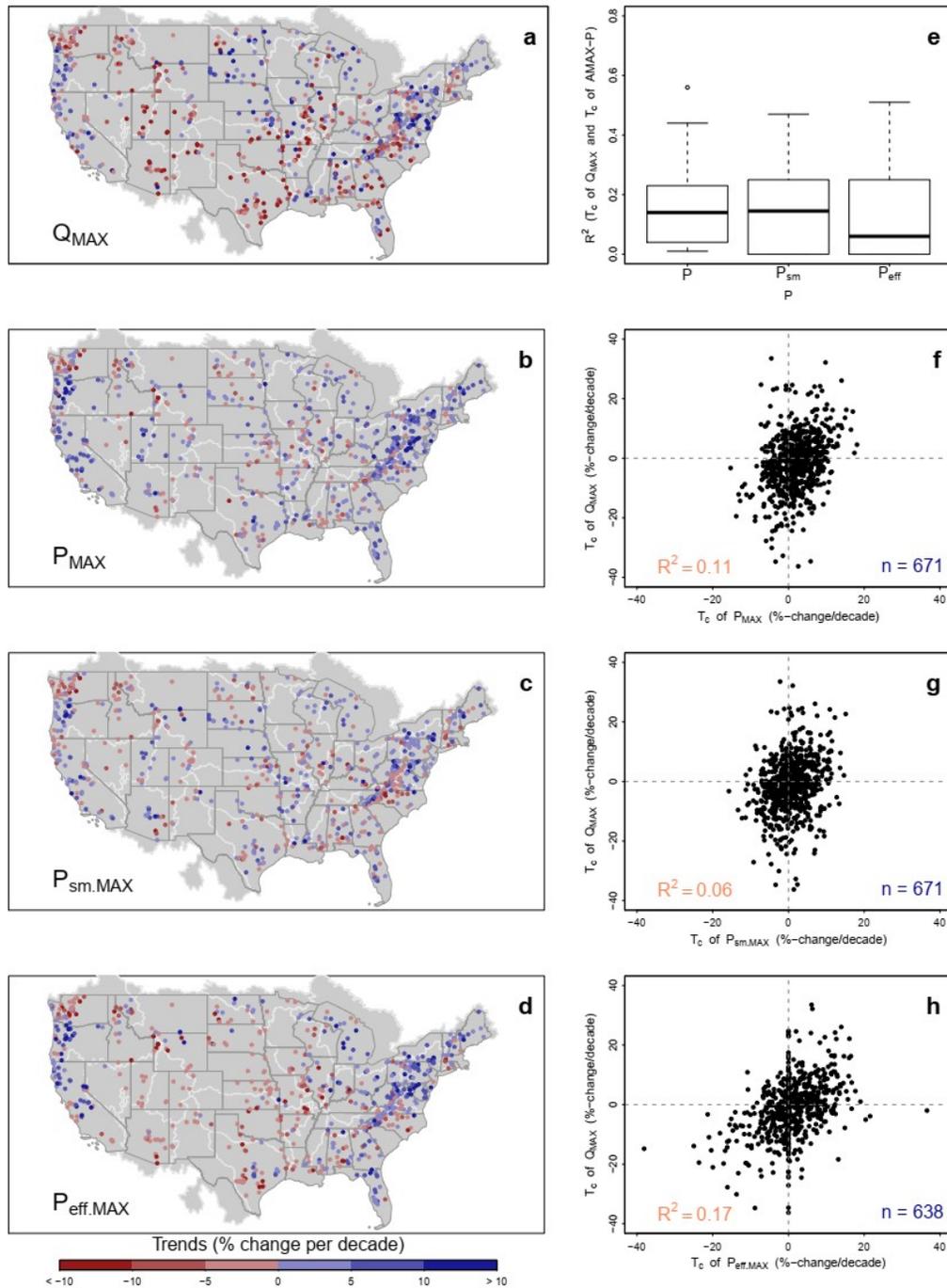
Figure 2 shows temporal changes in the magnitude of annual floods and precipitation extremes across the CAMELS catchments. We note that the effective precipitation ( $P_{\text{eff}}$ ) could be equal to zero throughout the year wherever precipitation could not make the catchment fully saturated. Specifically, there are 110 catchments having zero  $P_{\text{eff,MAX}}$  over more than 20 years, leading to a zero Thiel-Sen slope estimated as shown in Figure 2h (see also Supplementary Figure S2). We also removed 33 catchments that have  $P_{\text{eff,MAX}}$  equal to zero across all years from our analyses, leading to a sample size of 638 catchments for  $P_{\text{eff,MAX}}$  assessment.

Over the reference period, more CAMELS catchments (53%) experienced a decrease in  $Q_{\text{MAX}}$  index (Figure 2a), consistent with recent investigations [Do *et al.*, 2017; Do *et al.*, 2020b; Gudmundsson *et al.*, 2019; Hodgkins *et al.*, 2019; Hodgkins *et al.*, 2017]. On the contrary,  $P_{\text{MAX}}$  index shows an increasing trend (Figure 2b) over the majority of catchments (67%), although

some interior water resources regions exhibited a more prominent decreasing trend (e.g. Missouri Region; see Figure S2 in the Supporting Information for trends in annual floods and precipitation extremes over individual regions). There is a low correlation between the spatial pattern of changes in  $P_{MAX}$  and the spatial pattern of changes in  $Q_{MAX}$  (Figure 2f) with an  $R^2$  of 0.11, indicating that only 11% of the spatial variation of trends of  $Q_{MAX}$  can be explained using trends of  $P_{MAX}$ .

The spatial pattern of  $P_{sm,MAX}$  trends (Figure 2c) is generally consistent with that of  $P_{MAX}$  trends, while the spatial pattern of  $P_{eff,MAX}$  trends (Figure 2d) shows a substantial difference relative to that of  $P_{MAX}$  trends, and appears to be more consistent with the spatial pattern of  $Q_{MAX}$  trends. The coefficient of determination between precipitation extreme trends and  $Q_{MAX}$  trends support this finding, with an  $R^2$  of 0.06 and 0.17 for  $P_{sm,MAX}$  trends and  $P_{eff,MAX}$  trends respectively (Figure 2g and Figure 2h). These results are generally expected, as the snow-soil routine underlying  $P_{eff,MAX}$  can be seen as a simple conceptual model that takes into account several catchment processes.

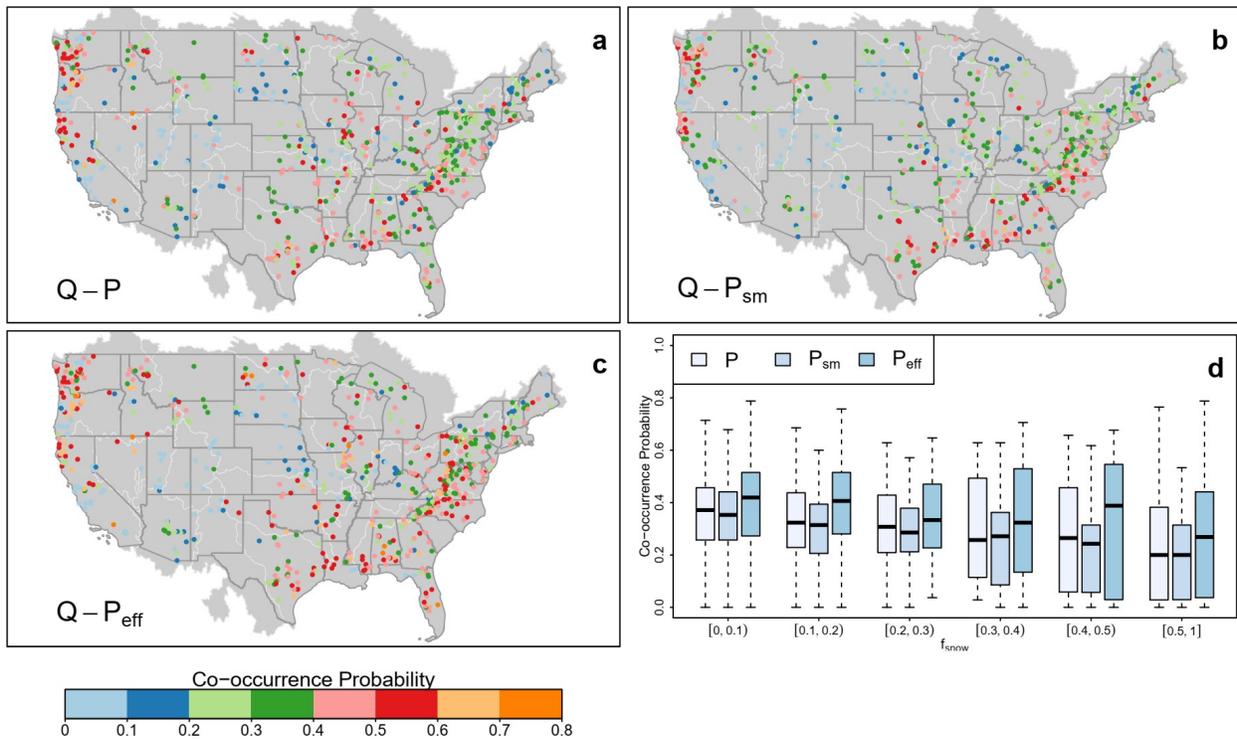
More importantly, the  $R^2$  values between  $Q_{MAX}$  trends and precipitation extreme trends are less than 0.2 across all precipitation extreme metrics. This result means that the spatial variation of precipitation extreme trends can explain less than 20% of the spatial variation of  $Q_{MAX}$  trends across the CONUS from 1980 to 2014. The  $R^2$  calculated over individual water resources regions (Figure 2e; also see Figure S2 of the Supporting Information) also indicates a low correlation. Specifically, more than 60% of the regions having an  $R^2$  value of less than 0.2, indicating the limitation of using trends of precipitation extremes to infer trends of floods.



**Figure 2.** Trends in (a)  $Q_{MAX}$ , (b)  $P_{MAX}$ , (c)  $P_{sm.MAX}$  and (d)  $P_{eff.MAX}$  across each of the 671 CAMELS catchments based on the normalized Thiel-Sen slope ( $T_c$ ). Scatter plots between  $T_c$  values of  $Q_{MAX}$  and  $T_c$  values of (f)  $P_{MAX}$ , (g)  $P_{sm.MAX}$ , (h)  $P_{eff.MAX}$  are also shown. (e) Boxplot of the  $R^2$  between  $T_c$  values of annual floods and  $T_c$  values of annual precipitation extremes (see Figure S2 in the Supporting Information for disaggregation of results across water resources regions).

### 3.2 Co-occurrence probability of precipitation extremes and floods across the CONUS

The low correlation between the spatial pattern of changes in precipitation extremes and that of floods (discussed in Section 3.1) is potentially attributable to a weak causal relationship, as there are other mechanisms that could trigger floods [Blöschl *et al.*, 2019; Merz and Blöschl, 2003]. To assess this hypothesis, we assessed the co-occurrence probability between precipitation extremes and floods over individual catchments and the results are shown in Figure 3 (See Figure S3-S5 in the Supporting Information for  $Q_{DOYMAX}$ ,  $P_{DOYMAX}$ ,  $P_{sm,DOYMAX}$ , and  $P_{eff,DOYMAX}$  across all catchments). The averaged co-occurrence probability across all catchments is 32%, 30%, and 37% for  $P_{DOYMAX}$ ,  $P_{sm,DOYMAX}$ , and  $P_{eff,DOYMAX}$  extremes respectively. This number is consistent with a previous investigation [Ivancic and Shaw, 2015], indicating that annual precipitation extremes are only responsible for about one-third of the annual flood population.



**Figure 3.** Co-occurrence probability between AMAX streamflow and (a) AMAX precipitation, (b) AMAX wet-month precipitation, (c) AMAX effective precipitation across all CAMELS catchments; and (d) when grouped into six categories using the fraction of precipitation falling as snow ( $f_{snow}$ ). Note that  $P_{eff}$  were available for only 638 catchments.

The vast majority of catchments (more than 95%) have a co-occurrence probability that is much higher than random chance (i.e., 2%), which is generally expected. Catchments with a relatively high co-occurrence probability are mostly located in coastal regions (e.g., South Atlantic-Gulf Region, Texas Gulf Region, California Region, and Pacific Northwest Region) while co-occurrence probability tends to be low whenever the fraction of precipitation falling as snow ( $f_{snow}$ ) is high (e.g., Upper Colorado Region and Great Basin Region). More importantly, only a small fraction (14%) of all catchments having a co-occurrence probability between

$Q_{\text{DOYMAX}}$  and  $P_{\text{DOYMAX}}$  (Figure 3a; see Figure S6 for regional results) of at least 0.5, indicating a weak causal linkage.

The co-occurrence probability between  $Q_{\text{DOYMAX}}$  and  $P_{\text{sm.DOYMAX}}$  (Figure 3b; see Figure S7 for regional results) is higher than or equal to 0.5 over 7% of all catchments, indicating a weaker causal linkage relative to that between  $Q_{\text{DOYMAX}}$  and  $P_{\text{DOYMAX}}$ . A possible reason is that soil moisture has a relatively strong seasonal cycle [Eltahir, 1998; Findell and Eltahir, 1997], contrasting to a weak seasonality of short-duration precipitation extremes [Do et al., 2020a]. Using  $P_{\text{sm}}$  has potentially masked out many short-duration flood-induced rainfall events that spread throughout the years, leading to a lower co-occurrence probability. This finding suggests that precipitation extremes constrained over wet months has a potentially lower predictability for changes in annual floods.

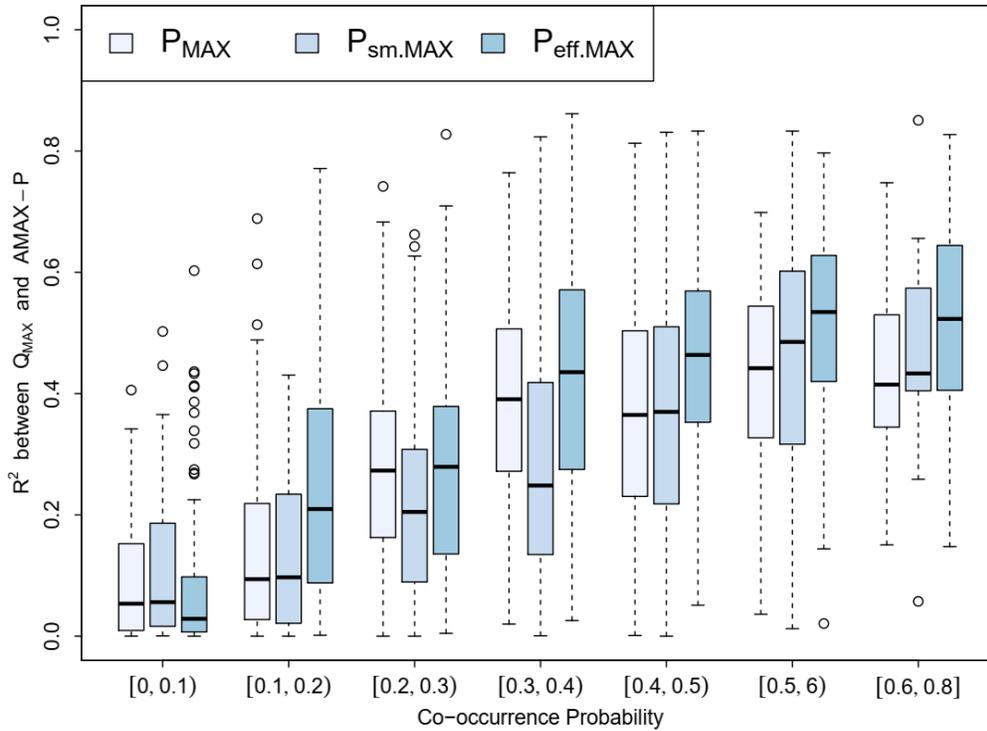
Among the three precipitation extreme metrics, effective precipitation (Figure 3c; see Figure S8 for regional results) is the variable with the strongest causal relationship to floods. Specifically, 26% of all catchments have a co-occurrence probability between  $P_{\text{eff.DOYMAX}}$  and  $Q_{\text{DOYMAX}}$  of at least 0.5. A simple approach to take into account snow-soil interaction has led to a substantial increase in co-occurrence probability, suggesting that catchment processes potentially play a more important role in modulating floods relative to precipitation intensity.

When catchments are divided into different categories using  $f_{\text{snow}}$  (Figure 3d), there is a notable decrease of co-occurrence probability when  $f_{\text{snow}}$  increases. We note that the co-occurrence probability between precipitation extremes and floods is not consistently low across all catchments with a high  $f_{\text{snow}}$ . For instance, of all 73 catchments with an  $f_{\text{snow}}$  of higher than or equal to 0.5, six catchments (located in the Pacific Northwest Region) have a co-occurrence probability between  $P_{\text{DOYMAX}}$  and  $Q_{\text{MAX}}$  of at least 0.5. As a result, catchments with a high snow-to-rain ratio are likely to have floods not driven by precipitation extremes, but there are exceptions such as catchments strongly influenced by atmospheric rivers which are responsible for flood-induced rainfall events.

### 3.3 To what extent are changes in precipitation extremes useful to explain changes in floods?

The co-variation between precipitation extremes and  $Q_{\text{MAX}}$  is relatively low, with 81%, 85%, and 66% of all catchments having an  $R^2$  of less than 0.5 for  $P_{\text{MAX}}$ ,  $P_{\text{sm.MAX}}$ , and  $P_{\text{eff.MAX}}$  respectively. When catchments are grouped into different categories according to co-occurrence probability, a strong positive correlation between co-variation and co-occurrence probability is observed (Figure 4). Of all catchments with co-occurrence probability of less than 0.5, the averaged  $R^2$  is 0.28 (for  $P_{\text{MAX}}$ ), 0.24 (for  $P_{\text{sm.MAX}}$ ) and 0.33 (for  $P_{\text{eff.MAX}}$ ) respectively, indicating that only about 30% of the temporal variability of floods can be explained by precipitation extremes.

Focusing on the catchments with the strongest causal relationship (i.e., a co-occurrence probability of at least 0.6), a low-to-moderate correlation is observed, with the median of  $R^2$  between  $Q_{\text{MAX}}$  and precipitation extremes is 0.41, 0.43 and 0.52 for  $P_{\text{MAX}}$ ,  $P_{\text{sm.MAX}}$  and  $P_{\text{eff.MAX}}$  respectively. The causal relationship between  $P_{\text{eff.MAX}}$  and  $Q_{\text{MAX}}$  is the strongest, with 34 out of 63 catchments (54%) have an  $R^2$  of above 0.5. The relationship between  $P_{\text{MAX}}$  and  $Q_{\text{MAX}}$  is the lowest, with 10 out of 28 catchments (36%) have an  $R^2$  value of above 0.5, further confirming the need for considering catchment processes (e.g., snow-soil interaction) to explain changes in annual floods.



**Figure 4.** Coefficient of determination ( $R^2$ ) between AMAX discharges and AMAX precipitation across all CAMELS catchments, grouped by the co-occurrence probability. Results are shown for precipitation ( $P$ ), wet-month precipitation ( $P_{sm}$ ) and effective precipitation ( $P_{eff}$ ). Note that  $P_{eff}$  were available for only 638 catchments.

#### 4 Summary and Conclusions

Using annual maxima precipitation and streamflow across a large sample of catchments, this study has empirically assessed the relationship between temporal changes in precipitation extremes and changes in annual flood magnitude. The spatial pattern of trends detected from precipitation extremes is weakly correlated to the spatial pattern of trends detected from AMAX streamflow over 671 CONUS catchments, with a coefficient of determination of less than 0.2.

A weak linkage between annual precipitation extremes and annual floods is apparent across the CAMELS catchments, with the vast majority of catchments have less than 50% of annual flood events directly linked to precipitation extremes (85%, 90%, and 73% of all catchments for AMAX precipitation, AMAX wet-month precipitation and AMAX effective precipitation respectively). Catchments with a high snow-to-rain ratio (indicated by  $f_{snow}$  value) generally have a low causal relationship between precipitation extremes and floods, but the impact of snow presence is not uniform. The co-variation between extreme precipitation intensity and flood magnitude is also low, with more than 60% of catchments having an  $R^2$  of less than 0.5, regardless of which precipitation extreme metrics being used. Using a snow-soil routine to correct the actual amount of precipitation modulating floods has led to a substantially improved predictability for changes in floods, suggesting that future trend detection studies should focus more on the catchment attributes such as soil profile and impervious area.

Notwithstanding the complex processes driving floods, this study has quantitatively assessed the limitation of using changes in precipitation as a proxy for potential changes in floods. The findings indicate that the intensity of daily precipitation extremes is a weak predictor for temporal changes in annual maxima of daily streamflow, even for catchments with a relatively high causal relationship. This study highlights the need for additional efforts to investigate the non-linear responses of floods to climate changes using a larger sample of catchments, which would hopefully achieve a universal understanding of how floods might evolve. For instance, the approach presented in this study can be applied for other large sample datasets [Addor *et al.*, 2019; Alvarez-Garreton *et al.*, 2018; Coxon *et al.*, 2020; Gudmundsson *et al.*, 2018] to quantify the contribution of extreme precipitation to historical changes in floods for other parts of the world.

### **Acknowledgments and Data**

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