

1 **Coherent streamflow variability in Monsoon Asia over**
2 **the past eight centuries—links to oceanic drivers**

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8 **Key Points:**

- 9 • Climate-informed dynamic streamflow reconstruction is skillful over most of Mon-
- 10 soon Asia
- 11 • Spatial coherence of streamflow suggests water management be coordinated be-
- 12 tween basins
- 13 • Mekong and Chao Phraya are most sensitive among rivers to anomalies in sea sur-
- 14 face temperature

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Abstract

The Monsoon Asia region is home to ten of the worlds biggest rivers, supporting the lives of 1.7 billion people who rely on streamflow for water, energy, and food. Yet, a synoptic understanding of multi-centennial streamflow variability for this region is lacking. To fill this gap, we produce the first large scale streamflow reconstruction over Monsoon Asia (48 stations in 16 countries), spanning the past eight centuries. In making this reconstruction, we develop a novel automated, climate-informed, and dynamic reconstruction framework that is skillful for 46/48 stations. We show that streamflow in Monsoon Asia is spatially coherent, owing to common drivers from the Pacific, Indian, and Atlantic Oceans. We also show that these drivers exert their greatest influence over the Mekong and Chao Phraya basins. We suggest that future water management in the region should be coordinated between basins, taking into account the states of the oceans.

Plain Language Summary

Ten of the world's biggest rivers are located entirely within the Asian Monsoon region. They provide water, energy, and food for 1.7 billion people. To manage these critical resources, we need a better understanding of river discharge—how does it change over a long time? Are there common variation patterns among rivers? To answer these questions, we use information derived from tree rings to reconstruct river discharge history at 48 gauges in 16 Asian countries. Our reconstruction reveals the riparian footprint of megadroughts and large volcanic eruptions over the past eight centuries. We show that simultaneous droughts and pluvials have often occurred at adjacent river basins in the past, because Asian rivers share common influences from the Pacific, Indian, and Atlantic Oceans. We also show that the oceans exert their greatest influences on the Mekong and Chao Phraya basins. From these findings, we suggest that future water management in the region should be coordinated between basins, taking into account the states of the oceans. Our findings can benefit the riparian people of the Asian Monsoon region.

1 Introduction

Of the world’s 30 biggest rivers, ten are located within Monsoon Asia, and two others originate from this region (Figure 1). These river basins are home to 1.7 billion people (Best, 2019). With high population densities, even smaller basins support the livelihood of millions—e.g., Chao Phraya (Thailand): 25 million, Angat (the Philippines): 13 million, and Citarum (Indonesia): 10 million (Nguyen & Galelli, 2018; Libisch-Lehner et al., 2019; D’Arrigo et al., 2011). River discharge, or *streamflow*, provides water for domestic and industrial uses, irrigation, and hydropower. It sustains aquatic life (including fish yield), carries sediment and nutrients, and enables navigation. Streamflow is an important link in both the water-energy-food nexus and the ecological cycle. To manage this resource, we need a good understanding of hydrologic variability. Such understanding is often derived from streamflow measurements; however, these instrumental data span typically only a few decades, too short to capture long-term variability and changes in streamflow.

When compared against instrumental data, longer streamflow records reconstructed from climate proxies—such as tree rings—often reveal striking insights. A reconstructed pre-dam variability of the Yellow River (Li et al., 2019) shows that streamflow in 1968–2010 was only half of what should have been; in other words, human activities depleted half of the available water! A reconstruction of the Citarum River (Indonesia) (D’Arrigo et al., 2011) shows that the period 1963–2006 contained an increasing trend of low flow years but no trend in high flow years, compared with the previous three centuries. This finding suggests that 10 million inhabitants of Jakarta may be facing higher drought risks than what is perceived from the instrumental record. The Mongolian “Breadbasket”, an agricultural region in north-central Mongolia (Pederson et al., 2013), experienced an unusually wet twentieth-century, and the recent dry epoch is not rare in the last four centuries (Davi et al., 2006; Pederson et al., 2013; Davi et al., 2013). Consequently, agricultural planning cannot take the twentieth century to be the norm, lest history repeats the lesson of the Colorado River Basin: observations over abnormally wet years (Stockton & Jacoby, 1976; Woodhouse et al., 2006; Robeson et al., 2020) led to water rights over-allocation, and the Colorado no longer reaches the Pacific Ocean.

Compelling evidence calls for more streamflow reconstructions in Monsoon Asia. Tremendous efforts, booming in the last four years (Figure S1), have partly addressed this need, but the hydrological knowledge gained was limited to individual catchments, more than half of which are in China (Figure S1 and Table S1). A synoptic understanding is lacking. Here, we produce the first large scale streamflow reconstruction for Monsoon Asia, covering 48 stations in 16 countries, unraveling eight centuries of annual stream-

77 flow variability. To achieve this task, we develop a novel automated framework with three
78 main components: (1) a climate-informed proxy selection procedure, (2) a dynamic state-
79 space reconstruction model, and (3) a rigorous cross-validation routine for parameter tun-
80 ing to achieve optimal skills. We also use the Monsoon Asia Drought Atlas version 2 as
81 the paleoclimate proxy instead of a tree ring network, as the former offers computational
82 advantages (supported with strong physical and statistical foundations) for this large scale
83 reconstruction. With this work, 44 stations are reconstructed for the first time while the
84 other four (Citarum, Yerru, Ping, and Indus Rivers) are extended back in time compared
85 to previous works (D'Arrigo et al., 2011; Pederson et al., 2013; Nguyen & Galelli, 2018;
86 Rao et al., 2018). This data set allows us to assess both local historical water availabil-
87 ity and regional streamflow patterns, revealing the spatial coherence of streamflow and
88 its links to the oceans. This understanding may improve interbasin water resources man-
89 agement and coordination.

90 **2 Data**

91 **2.1 Streamflow Data**

92 We obtained streamflow data from the Global Streamflow Indices and Metadata
93 Archive (GSIM) (Do et al., 2018; Gudmundsson et al., 2018), using stations having at
94 least 41 years of data, with less than 3% missing daily values, and with mean annual flow
95 of at least 50 m³/s. We also received streamflow data from our colleagues for some coun-
96 tries where public streamflow records are not available (see Acknowledgment).

97 Many stations in our collection have upstream reservoirs that may interfere with
98 the proxy-streamflow relationship. This interference is stronger for seasonal streamflow
99 than annual streamflow: reservoirs transfer water from the wet season to the dry sea-
100 son, but not all reservoirs retain water from year to year. Reservoirs that are filled and
101 emptied within a year do not change the annual water budget downstream. To minimize
102 reservoir interference, we reconstructed annual streamflow, and we removed stations that
103 have upstream retention time longer than a year. We identified upstream reservoirs by
104 overlaying the Global Reservoirs and Dams (GRanD) data (Lehner et al., 2011) on the
105 river network (Lehner & Grill, 2013; Barbarossa et al., 2018). The upstream retention
106 time was calculated as the total upstream reservoir capacity (million m³) divided by the
107 mean annual flow volume (million m³/year). For stations having over-year reservoirs con-
108 structed towards the end of their records, we also truncated the corresponding years, keep-
109 ing only the streamflow data before dam construction.

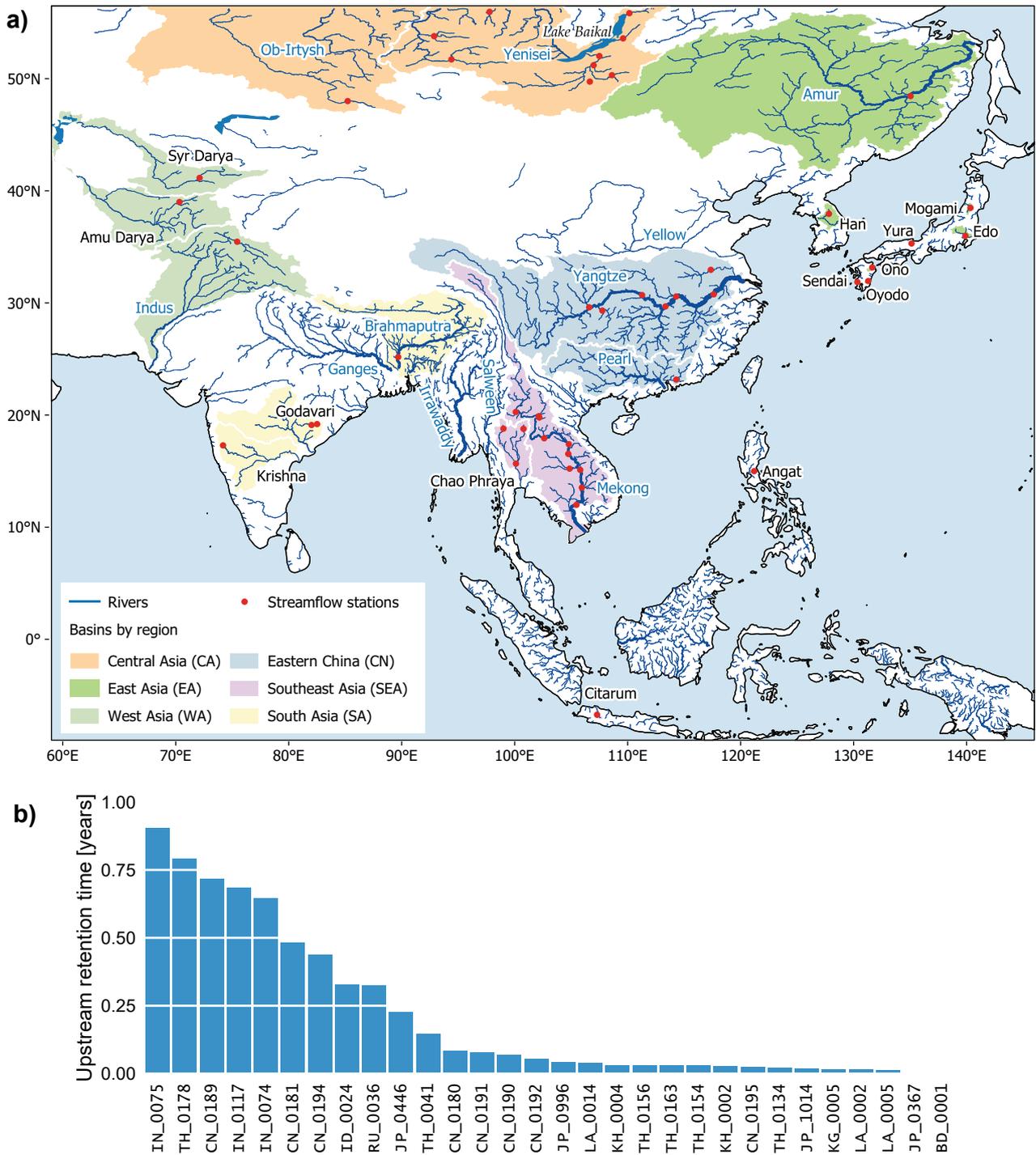


Figure 1. Study region. a) The Monsoon Asia region (Cook et al., 2010); river basins involved in this study are highlighted by sub-region, rivers belonging to the world's 30 biggest (Best, 2019) shown with blue names. b) Upstream retention time of the 30 stations that have upstream reservoirs. Refer to Table S2 for station details.

110 Our collection and quality control effort resulted in an annual streamflow data set
111 of 48 stations in 16 countries. We used the calendar year (January to December) as there
112 is not a common water year across the study domain (Knoben et al., 2018). The stations’
113 locations and upstream retention times (for those having upstream reservoirs) are shown
114 in Figure 1. More metadata are supplied in Table S2.

115 2.2 Proxy Data

116 Our paleoclimate proxy is the Monsoon Asia Drought Atlas version 2 (MADA v2)
117 (Cook, 2015), built upon the original MADA of Cook et al. (2010). The MADA is a grid-
118 ded data set of the Palmer Drought Severity Index (PDSI) (Palmer, 1965) over the Asian
119 monsoon region; each grid cell contains an annual time series of the mean June-July-August
120 PDSI, reconstructed from tree rings, and calibrated with the instrumental data set of
121 Dai et al. (2004). The MADA proves to be a reliable long-term record of monsoon strength,
122 having revealed the spatiotemporal extents of the four Asian megadroughts in the last
123 millennium, and linking variations in monsoon strength to sea surface temperature pat-
124 terns. MADA v2 improves over its predecessor by incorporating more tree ring chronolo-
125 gies (453 versus 327), and targeting the self-calibrating PDSI (scPDSI), which addresses
126 several limitations of the standard PDSI (Wells et al., 2004; van der Schrier et al., 2013).
127 We use the MADA v2 portion between 1200–2012 as this is the common period of most
128 grid points in the atlas (Figure S4), and is also the stable portion with sufficient sam-
129 ple depth in the source tree ring network.

130 Drought atlases (reconstructed from tree rings) have been shown to be good pa-
131 leoclimate proxies for streamflow reconstruction: since both streamflow and PDSI can
132 be modeled as functions of ring width, one can also build a model to relate streamflow
133 to PDSI (Ho et al., 2016, 2017; Nguyen & Galelli, 2018). Drought atlases enhance the
134 spatial expression of the underlying tree ring data—by incorporating the modern PDSI
135 field in its calibration—and are also more uniform in space and time than the tree ring
136 network itself (see Cook et al., 2010, Figure 1), making them better suited to large scale
137 studies. We now elaborate these points as we describe the methodology.

138 3 Methods

139 3.1 Using a Drought Atlas as Paleoclimate Proxy

140 3.1.1 Physical basis

141 The main physical processes that involve climate and tree growth are depicted in
142 Figure 2a. The climate at a given location can be characterized by precipitation and tem-

143 perature, among others. These climatic inputs control soil moisture on land. Except for
 144 losses (such as groundwater recharge, evaporation, and surface runoff), the net soil mois-
 145 ture storage then follows two main paths: one goes out of the catchment as streamflow,
 146 the other is taken up by the trees and transpired back into the atmosphere, influencing
 147 tree growth along the way. Thus, tree growth and streamflow are connected via land-
 148 atmosphere interactions—this is the basis for streamflow reconstruction from tree ring.
 149 Note, however, that tree growth does not directly control streamflow, and neither does
 150 streamflow control tree growth; we can infer a relationship between them only because
 151 they are both influenced by soil moisture. On the other hand, soil moisture directly con-
 152 trols streamflow and is, in principle, a reasonable predictor for streamflow.

153 It would thus be ideal to have a “natural” soil moisture proxy record, but of course
 154 that is not the case. We can instead rely on a surrogate—a soil moisture record recon-
 155 structed from tree rings, such as the MADA.

156 3.1.2 Statistical basis

157 The physical discussion above yields three types of paleoclimate reconstruction: stream-
 158 flow from tree rings, soil moisture from streamflow, and streamflow from soil moisture.
 159 We now derive mathematically the relationships between these reconstruction types.

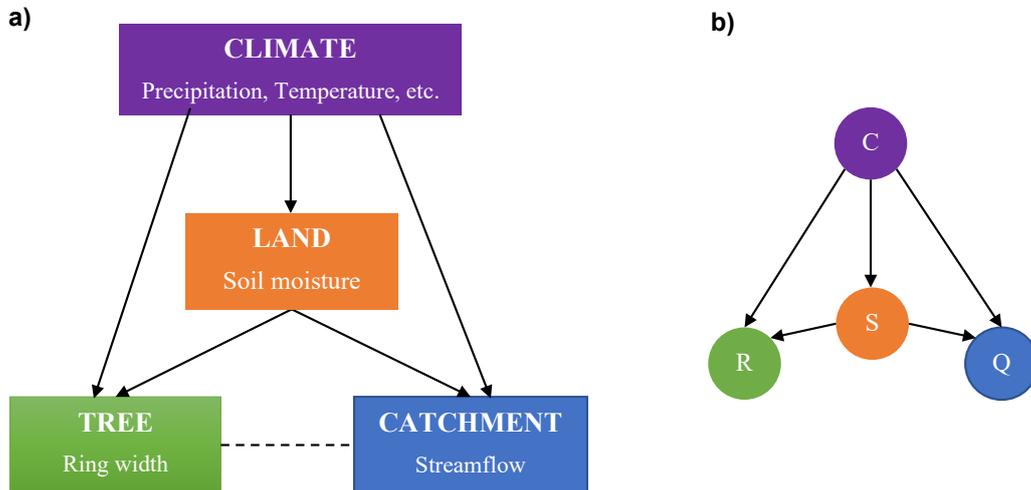


Figure 2. a) Relationships between hydroclimatic variables and tree growth. b) A probabilistic graphical model representing the relationships in a), where C is a vector of climate variables, S the soil moisture, R the ring width index, and Q streamflow. The arrows represent the conditional dependence among variables.

160 Each reconstruction is a conditional distribution of one variable (e.g. streamflow)
 161 given that we have observed another variable (e.g. soil moisture), and given the histor-
 162 ical climate. We represent these conditional distributions with a probabilistic graphical
 163 model (Koller & Friedman, 2009) as shown in Figure 2b. There are four random vari-
 164 ables involved: climate (C), soil moisture (S), ring width (R), and streamflow (Q). Each
 165 of these variables can be multivariate, i.e., C includes precipitation and temperature, among
 166 others, and all variables can include multiple sites or grid points. As a convention, let
 167 $f_X(x)$ be the probability density function (PDF) of the random variable X , $f_{XY}(x, y)$
 168 be the joint PDF of X and Y , and $f_{X|Y}(x|y)$ be the conditional PDF of X given that
 169 $Y = y$.

170 Reconstructing streamflow from tree ring is essentially deriving the distribution of
 171 Q given R and C , i.e. $f_{Q|R,C}(q|r, c)$, where r is the measured ring width index, and c is
 172 the historical climate. We can decompose this distribution as follows:

$$\begin{aligned}
 f_{Q|R,C}(q|r, c) &= \int f_{Q,S|R,C}(q, s|r, c) ds \\
 &= \int f_{Q|S,R,C}(q|s, r, c) f_{S|R,C}(s|r, c) ds.
 \end{aligned}
 \tag{1}$$

174 The first equality comes from the relationship between marginal and joint distributions.
 175 The second equality comes from Bayes' theorem. Now, Q is independent of R given S
 176 and C (Figure 2b), so $f_{Q|S,R,C}(q|s, r, c) = f_{Q|S,C}(q|s, c)$. Consequently,

$$f_{Q|R,C}(q|r, c) = \int f_{Q|S,C}(q|s, c) f_{S|R,C}(s|r, c) ds.
 \tag{2}$$

178 Observe that $f_{Q|S,C}$ is the streamflow reconstruction from the MADA, and $f_{S|R,C}$
 179 is the MADA reconstruction from tree rings. Thus we have established mathematically
 180 the reasoning that tree-ring-based streamflow reconstruction is possible based on the link
 181 through soil moisture. $f_{Q|R,C}$ is the marginal distribution without observing the soil mois-
 182 ture. Instead of constructing $f_{Q|R,C}$, we can infer S from R , then Q from S , by construct-
 183 ing $f_{S|R,C}$ and $f_{Q|S,C}$.

184 **3.1.3 Computational advantages of using the MADA, and a caveat**

185 The MADA can be thought of as a transformation from the tree ring network, ir-
 186 regular in both space and time, to a regular grid with homogeneous temporal coverage—
 187 analogous to transforming meteorological station data to gridded temperature and pre-
 188 cipitation products. This transformation brings several advantages to reconstructing stream-
 189 flow using the MADA, compared to using the underlying tree ring network.

190 First, in a typical reconstruction study, one must detrend and standardize the tree
 191 ring data to remove non-climate signals (cf. Cook and Kairiukstis (1990)). For a large

192 scale study like ours, such a task is complex. Instead, we can leverage the effort that has
193 been devoted to detrending and standardizing the chronologies in making the MADA,
194 and use the MADA as proxy, having built the physical and statistical foundations to do
195 so.

196 Second, the tree ring sites often cluster, with vast empty space between clusters
197 (see e.g. Cook et al. (2010), Figure 1). When taking a subset of them for reconstruction
198 at a station, there can be cases where none or very few sites are within a search radius.
199 The MADA helps “bridging” the space, bringing climate signals from further-away tree
200 sites to grid points nearer to the station. The high resolution grid ($1^\circ \times 1^\circ$ for version
201 2) makes automated grid point selection easier. (The automated grid point selection pro-
202 cedure is described in Section 3.2.1.)

203 Third, when reconstructing streamflow from tree rings, nested models are often nec-
204 essary because tree ring chronologies have different time spans. One starts with the short-
205 est nest, using the common time span of all chronologies to build a model, then drop-
206 ping the shortest chronology to build a second model with longer time span but less ex-
207 plained variance than the first, and repeating the process, dropping more chronologies
208 to achieve longer time spans until the final nest with the longest time span, but with the
209 lowest explained variance. The nests’ outputs are then corrected for their variance and
210 averaged to obtain the final reconstruction (see e.g. D’Arrigo et al. (2011)). This nest-
211 ing step was carried out for the MADA, such that most grid points have the same time
212 span (Figure S4). This lets us use a single common period (1200-2012), and eliminates
213 our need to build nested models back in time. This is particularly desirable for our dy-
214 namic state-space reconstruction model, as averaging the nests breaks the link between
215 the catchment state and streamflow. (The reconstruction model is described in Section
216 3.2.2).

217 The computational advantages of using the MADA are thus threefold: (1) no de-
218 trending and standardization, (2) easier grid point selection, and (3) no nesting. How-
219 ever, these come with a cost: uncertainty. When reconstructing streamflow from the MADA,
220 we treat the MADA (i.e, the model input) as constant. But in fact, the MADA is a re-
221 gression product and has its own uncertainty. Incorporating this uncertainty is difficult
222 and is out of the scope of this paper, but it is an interesting topic for further research.

223 **3.2 Climate-informed Dynamic Streamflow Reconstruction**

224 When reconstructing a climate field, such as a PDSI grid or a streamflow station
225 network, it is desirable to preserve the field covariance structure. However, building a

226 large-scale spatial regression model is challenging. Instead, one can reconstruct each point
 227 in the field independently, and rely on the proxy network to capture the spatial patterns.
 228 This is the premise of the Point-by-Point Regression (PPR) method (Cook et al., 1999),
 229 and this principle has led to the successful reconstruction of many drought atlases (Cook
 230 et al., 1999, 2010, 2015). Our reconstruction framework is inspired by PPR—we recon-
 231 struct station by station—but diverges in several ways. The two key differences are, first,
 232 in the way proxy points are selected, and second, in the regression model.

233 *3.2.1 Climate-informed Input Selection*

234 The PPR procedure selects proxy points (tree ring chronologies) within a search
 235 radius. Given that geographical proximity does not imply hydroclimatic similarity, we
 236 selected our proxies (MADA grid points) by hydroclimatic similarity directly. The hy-
 237 droclimate at location i (a MADA grid point or a streamflow station) is characterized
 238 by three indices: aridity a_i , moisture seasonality s_i , and snow fraction f_i , following Knoben
 239 et al. (2018) (hereafter referred to as the KWF system, after the three authors). The hy-
 240 droclimatic similarity between two locations i and j is then defined as their Euclidean
 241 distance in the hydroclimate space. This distance is termed the KWF distance and its
 242 mathematical definition is

$$243 \quad d_{KWF}(i, j) = \sqrt{(a_i - a_j)^2 + (s_i - s_j)^2 + (f_i - f_j)^2}. \quad (3)$$

244 The KWF distance lets us screen out MADA grid points that are geographically close
 245 to the station of interest but hydroclimatically different—a climate-informed grid point
 246 selection scheme. Whereas previous PPR implementations varied the search radius, we
 247 fixed the radius to 2,500 km—the scale of regional weather systems (Boers et al., 2019)—
 248 and varied the KWF distance between 0.1 and 0.3 in 0.05 increments. For reference, the
 249 maximum KWF distance between any two points in Monsoon Asia is 1.424. Each KWF
 250 distance yielded a search region encompassing a set of MADA grid points surrounding
 251 the streamflow station of interest. In our search regions, PDSI often correlates signifi-
 252 cantly and positively with streamflow (Figure 3); indeed hydroclimatic similarity is a phys-
 253 ical basis for correlation.

254 Next, we performed weighted principal component analysis (PCA) to remove mul-
 255 ticollinearity among the MADA grid points. Following PPR, we weighted each grid point
 256 by its correlation with the target streamflow, using equation (4):

$$257 \quad z_i = x_i \rho_i^p. \quad (4)$$

258 Here, x_i is the scPDSI time series at grid point i , ρ_i the correlation between x_i and the
 259 target streamflow, p the weight exponent, and z_i the weighted version of x_i . We used

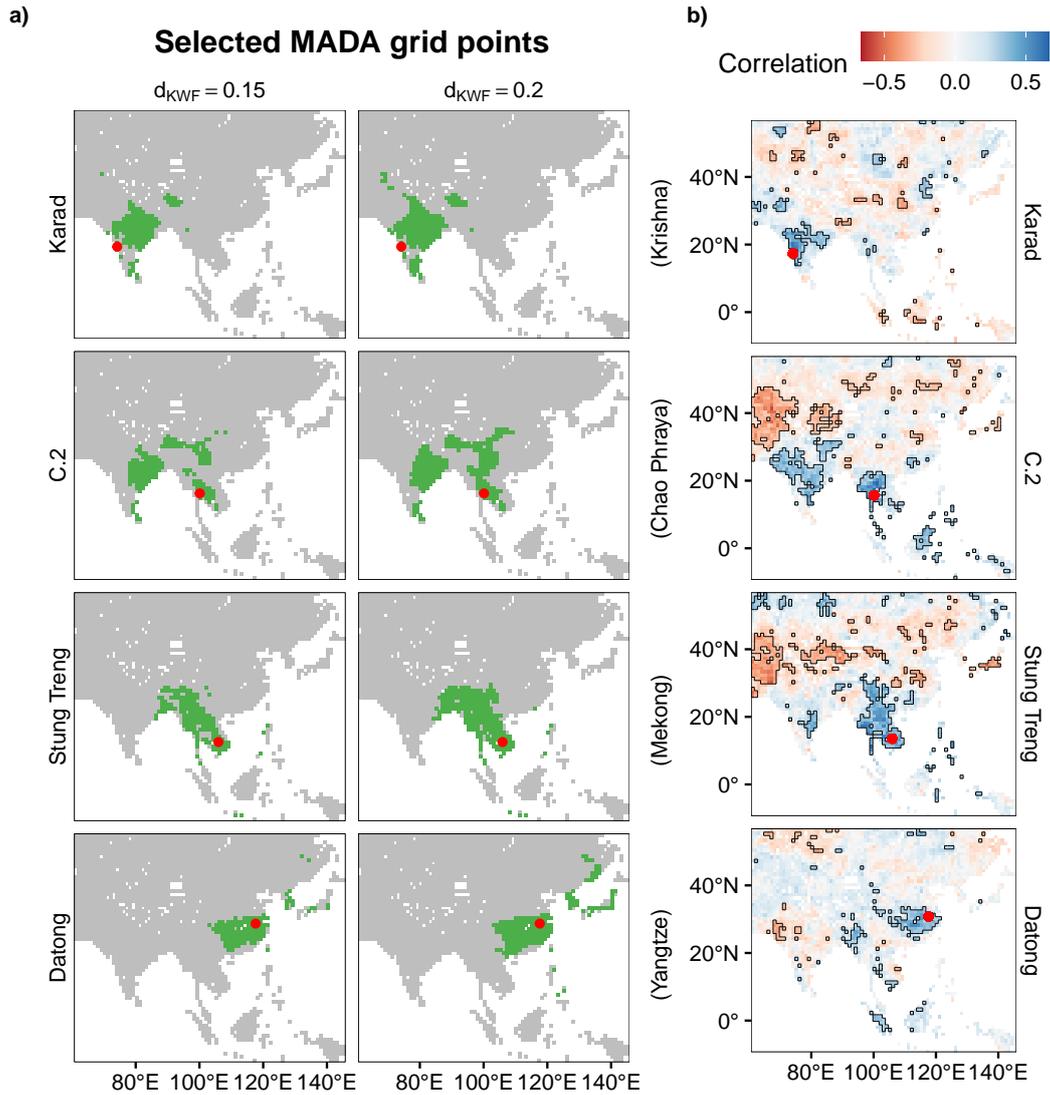


Figure 3. a) Examples of the climate-informed grid point selection: selected MADA grid points (green) based on two KWF distances (columns) at four stations (rows). b) Correlations between streamflow at the same four stations and the MADA, significant correlations ($\alpha = 0.05$) enclosed in black boundaries. The selection regions in (a) generally have significant positive correlation with streamflow. The areas with significant negative correlations need further investigation.

260 $p = 0, 0.5, 2/3, 1, 1.5,$ and $2,$ the same as those used by Cook et al. (2010). We then per-
 261 formed PCA on z_i 's, and retained only those principal components having eigenvalue at
 262 least 1.0 (Hidalgo et al., 2000). We further reduced this subset using the VSURF (Vari-
 263 able Selection Using Random Forest) algorithm (Genuer et al., 2010). So, for each com-
 264 bination of KWF distance and PCA weight, we arrived at a subset of principal compo-

265 nents for reconstruction. Each streamflow station has an ensemble of 30 such subsets,
 266 the best of which was identified using cross-validation (Section 3.2.3) and used for the
 267 final reconstruction.

268 **3.2.2 Linear Dynamical System**

269 Having obtained the climatic inputs, the next step was to model the relationship
 270 between these inputs and the catchment output (streamflow). Here, this relationship was
 271 not modeled with linear regression (as with original PPR, and as typical with previous
 272 reconstruction studies), but as a linear dynamical system (LDS), following equations (5)
 273 and (6):

$$274 \quad x_{t+1} = Ax_t + Bu_t + w_t \quad (5)$$

$$275 \quad y_t = Cx_t + Du_t + v_t \quad (6)$$

276 where t is the time step (year), y the catchment output (streamflow), u the climatic in-
 277 put (an ensemble member from the climate-informed grid point selection), w and v white
 278 noise, and x the system state, which can be interpreted as the catchment’s flow regime,
 279 i.e, wet or dry (Nguyen & Galelli, 2018). By modeling the flow regime and its transition,
 280 the LDS model accounts for both regime shifts (Turner & Galelli, 2016) and catchment
 281 memory (Pelletier & Turcotte, 1997). These behaviors are not modeled in linear regres-
 282 sion.

283 The LDS model assumes that the initial state and the noise processes are normally
 284 distributed:

$$285 \quad w_t \sim \mathcal{N}(0, Q) \quad (7)$$

$$286 \quad v_t \sim \mathcal{N}(0, R) \quad (8)$$

$$287 \quad x_1 \sim \mathcal{N}(\mu_1, V_1). \quad (9)$$

288 It follows that the catchment state and output must also be normally distributed. But
 289 some of our streamflow records are skewed. These were log-transformed to reduce skew-
 290 ness (Text S2 and Figure S3).

291 The LDS model is trained using a variant of the Expectation-Maximization algo-
 292 rithm. In the E-step, we fix the model parameters and learn the hidden state. In the M-
 293 step, we fix the hidden state and learn the model parameters. Iterations are repeated
 294 between the E- and M-steps until convergence. The reconstruction algorithm is imple-
 295 mented in the R package *ldsr* (Nguyen, 2020a).

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3.2.3 Cross-validation

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Consistent with the literature, we assessed reconstruction performance using the metrics Reduction of Error (RE) and Nash-Sutcliffe Coefficient of Efficiency (CE or NSE) (Nash & Sutcliffe, 1970; Fritts, 1976). Mathematically,

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$$RE = 1 - \frac{\sum_{t \in \mathcal{V}} (Q_t - \hat{Q}_t)^2}{\sum_{t \in \mathcal{V}} (Q_t - \bar{Q}_c)^2} \quad (10)$$

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$$CE = 1 - \frac{\sum_{t \in \mathcal{V}} (Q_t - \hat{Q}_t)^2}{\sum_{t \in \mathcal{V}} (Q_t - \bar{Q}_v)^2} \quad (11)$$

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where t is the time step, \mathcal{V} the validation set, Q the observed streamflow, \hat{Q} the reconstructed streamflow, \bar{Q}_c the calibration period mean, and \bar{Q}_v the verification period mean. Both RE and CE are based on squared error; they can be sensitive to outliers, especially the CE. To address this limitation, Gupta et al. (2009) proposed another metric, which assesses a model output based on its correlation with observation, as well as its bias and variability (equation (12)):

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$$KGE = 1 - \sqrt{(\rho - 1)^2 + \left(\frac{\hat{\mu}}{\mu} - 1\right)^2 + \left(\frac{\hat{\sigma}}{\sigma} - 1\right)^2}. \quad (12)$$

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Here, ρ is the correlation between model output and observation, $\hat{\mu}$ and μ the modeled and observed mean of the streamflow time series, and $\hat{\sigma}$ and σ the modeled and observed standard deviation of the streamflow time series. This metric is now known as the Kling-Gupta Efficiency (KGE). Compared to the CE, the KGE is more robust to outliers, hence we chose the KGE as the criterion for model selection.

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Conventionally, reconstruction skills are often calculated in a split-sample (i.e., two-fold) cross-validation scheme: the model is calibrated with the first half of the data and validated with the second half, then calibrated with the second half and validated with the first half (see e.g. D'Arrigo et al. (2011)). The contiguous halves aim to test a model's ability to capture a regime shift (Briffa et al., 1988). Unfortunately, this scheme is not practical for many stations in our record, where it would leave us only 20–25 data points for calibration (Figure S2). In addition, a two-fold cross-validation scheme provides only two point estimates for each skill score, and they may be notably different (for example, D'Arrigo et al. (2011) reported CE values of 0.21 and 0.73 for the two folds.) As a result, the mean skill score may not be robust. A number of recent works have instead used the leave- k -out cross-validation scheme (e.g. Ho et al. (2016); Li et al. (2019); Chen, Shang, Panyushkina, Meko, Li, et al. (2019)). In this scheme, a random chunk of k data points is withheld for validation while the model is calibrated with the remaining data

327 points, then calibration and validation are repeated over as many as 100 chunks of k . This
 328 scheme provides a more robust estimate of the mean skill score, but it may not correctly
 329 assess the model’s ability to capture a regime shift, because the withheld points are not
 330 contiguous like in the split-sample scheme.

331 We sought a balanced approach. In each cross-validation run for each station, we
 332 withheld a *contiguous* chunk of 25% of the data points for validation and trained the model
 333 on the remaining 75%. This way, we maintain the goal of the split-sample scheme while
 334 still having enough data for calibration and getting a reasonably robust mean skill es-
 335 timate. We could not get as many contiguous chunks as if they were random, so we re-
 336 peated the procedure 30 times instead of 100, and calculated the mean KGE over these
 337 30 runs. The ensemble member (cf. Section 3.2.1) that resulted in the highest mean KGE
 338 across the 30 cross-validation runs was used for the final reconstruction of each station.
 339 The cross-validation procedure is also available in the *ldsr* package.

340 4 Results and Discussion

341 4.1 Reconstruction Skills

342 Reduction of Error (RE) is positive at all stations (Figure 4a and b); Coefficient
 343 of Efficiency (CE) is positive at all but two: Nowrangpur (India) on the Godavari, and
 344 Ubon (Thailand) on the Nam Mun, a tributary of the Mekong (Figure 4c and d). Both
 345 negative values are larger than -0.08. The tree ring network used to build the MADA
 346 has lower density in India (Cook, 2015) so CE values here are understandably lower. Ubon,
 347 on the other hand, is located in an area of high quality tree ring chronologies (Buckley
 348 et al., 2007; Sano et al., 2009; Buckley et al., 2010), yet its variability is not captured
 349 by the MADA as well as nearby stations. We suspect there are data errors at this gauge.
 350 The histogram of CE resembles that of RE but shifts slightly left—this is expected as
 351 CE is a more stringent metric than RE (Cook & Kairiukstis, 1990). Much lower CE than
 352 RE implies overfitting; we do not observe that here.

353 The Kling-Gupta Efficiency (KGE) is all positive, and its histogram leans toward
 354 the higher end (Figure 4e and f). It should be noted that if one wishes to benchmark
 355 a model against the verification period mean (as is with the CE), the threshold value is
 356 $1 - \sqrt{2}$, i.e., $KGE > 1 - \sqrt{2}$ is analogous to $CE > 0$ (Knoben et al., 2019). The KGE
 357 scores in Figure 4 suggest that our reconstruction model performs well in terms of key
 358 characteristics: correlation, bias, and variability.

359 All three metrics have similar spatial distributions (Figure 4a, c, and e). As expected,
 360 lower (but positive) skills are seen in most of Central and West Asia, which lie outside

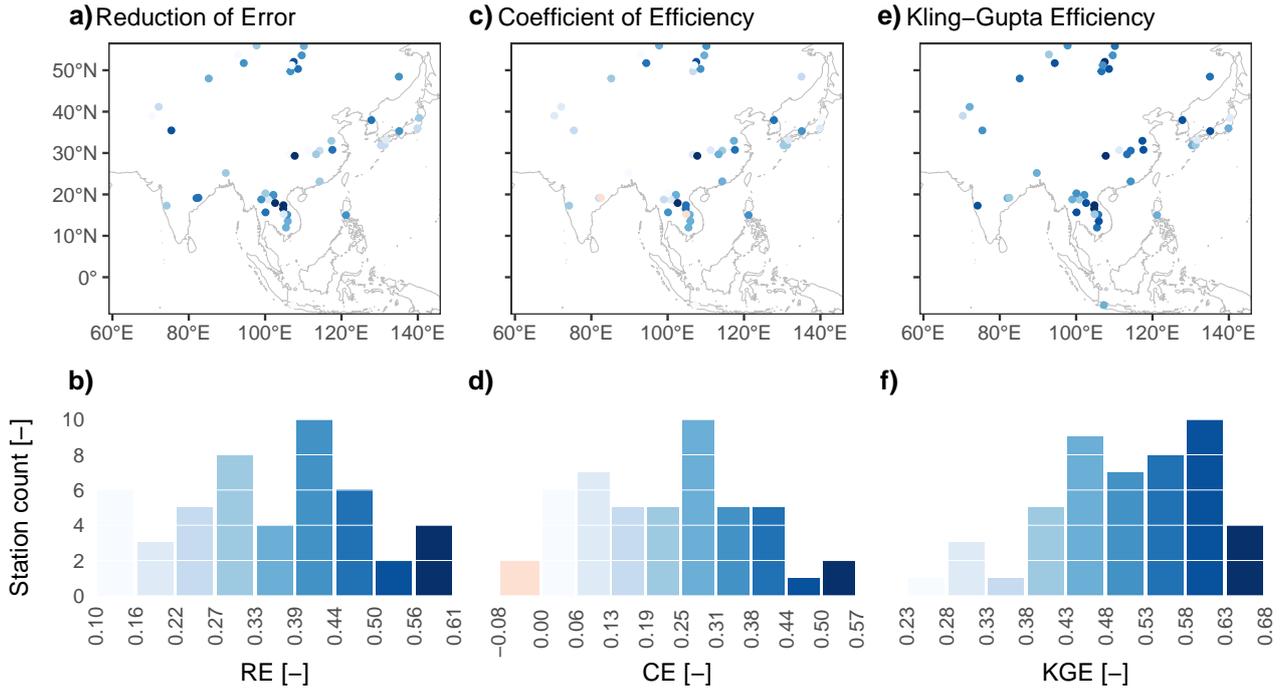


Figure 4. Distribution of model performance scores. Panels a, c, and e show the scores of each station following the color legends encoded with the histograms in panels b, d, and f.

361 the active monsoon area. An exception is the upper reach of the Selenge River (Mon-
 362 golia), upstream of Lake Baikal, where model skill is high, owing to high quality tree ring
 363 records from Mongolia (Davi et al., 2006; Pederson et al., 2013; Davi et al., 2013; Ped-
 364 erson et al., 2014). In Japan, where the small catchments are sensitive to local climate,
 365 model skill is reduced. In all other regions, model skill is homogeneous. The consistent
 366 performance of our model suggests that the MADA is a good proxy for streamflow re-
 367 construction in Asia, and our climate-informed dynamic reconstruction is skillful.

368 As an additional validation exercise, we compared the spatiotemporal variability
 369 of reconstructed streamflow against instrumental data for the period 1950–2012 (Fig-
 370 ure S5). Our reconstruction captures well the spatial variation patterns of streamflow
 371 in this period, as well as the timing, duration, and magnitude of extreme droughts and
 372 pluvials.

373 4.2 Spatiotemporal Variability of Monsoon Asia’s Streamflow

374 Having obtained good skill scores, we now present eight centuries of spatiotempo-
 375 ral streamflow variability in Monsoon Asia (Figure 5). This reconstructed history cap-

376 tures the riparian footprint of major historical events (large volcanic eruptions, megadroughts,
377 and pluvials). We first discuss the impact of the three largest eruptions of the past eight
378 centuries (Sigl et al., 2015): Samalas (1257) (Lavigne et al., 2013), Kuwae (1452-53) (Gao
379 et al., 2006), and Tambora (1815) (Stothers, 1984).

380 Assuming that Kuwae erupted in 1452 (consistent with tree ring records, see e.g.
381 Briffa et al., 1998), these three eruptions saw a persistent streamflow pattern across South-
382 east Asia, eastern China, and West Asia. In the eruption year t ($t = 1257, 1452, 1815$),
383 abnormally high streamflow occurred in all three regions. In year $t+1$, streamflow re-
384 mained high in Southeast Asia but abruptly turned low in West Asia and parts of east-
385 ern China. This is unexpected given the results of Li et al. (2013). They found that in
386 year t , PDSI (captured by the MADA) was negative in all three regions; in year $t+1$,
387 PDSI remained negative in Southeast Asia but turned positive in West Asia and east-
388 ern China. Based on their findings, one would expect streamflow to be low in all three
389 regions in year t , then remain low in Southeast Asia but turn high in West Asia and east-
390 ern China in year $t+1$. We observe the opposite. Interestingly, Anchukaitis et al. (2010),
391 also using PDSI, found in year t wet conditions in Southeast Asia (similar to our results)
392 but mixed wet and dry conditions in eastern China and West Asia (more similar to Li
393 et al., 2013). The disparity in these studies are attributed to the different sets of erup-
394 tions used—Anchukaitis et al. (2010) demonstrated this with three sets of events. Our
395 divergence from their results are partly because they used Superposed Epoch Analysis
396 while we analyze individual events, but we argue that the main cause is streamflow ver-
397 sus PDSI. With our streamflow results, we offer a reconciling explanation: during and
398 immediately after the eruptions, PDSI was more driven by temperature than precipita-
399 tion, and while low temperature may have caused negative PDSI, it reduced evapora-
400 tion and consequently, increased streamflow. This mechanism is particularly relevant in
401 midlatitude eastern China and West Asia. In Southeast Asia, however, reduced temper-
402 ature, from warm to cool, could increase soil moisture (Anchukaitis et al., 2010), result-
403 ing in high streamflow. Not disagreeing with previous works, our results offer a look at
404 another aspect of past climate using streamflow instead of PDSI.

405 As a drought/pluvial indicator, streamflow may differ from PDSI in individual years,
406 as discussed above, but on longer terms, our reconstructed streamflow agrees well with
407 reconstructed PDSI. For example, our record fully captures the Angkor Droughts (1345–
408 1374 and 1401–1425) (Buckley et al., 2010, 2014) with prolonged low flow throughout
409 the Mekong and Chao Phraya basins (Southeast Asia). Heavy monsoon rain interrupted
410 these megadroughts; such abrupt alterations to the flow regime proved difficult for Angkor’s
411 water managers (Buckley et al., 2014). After the first Angkor Drought, they altered the

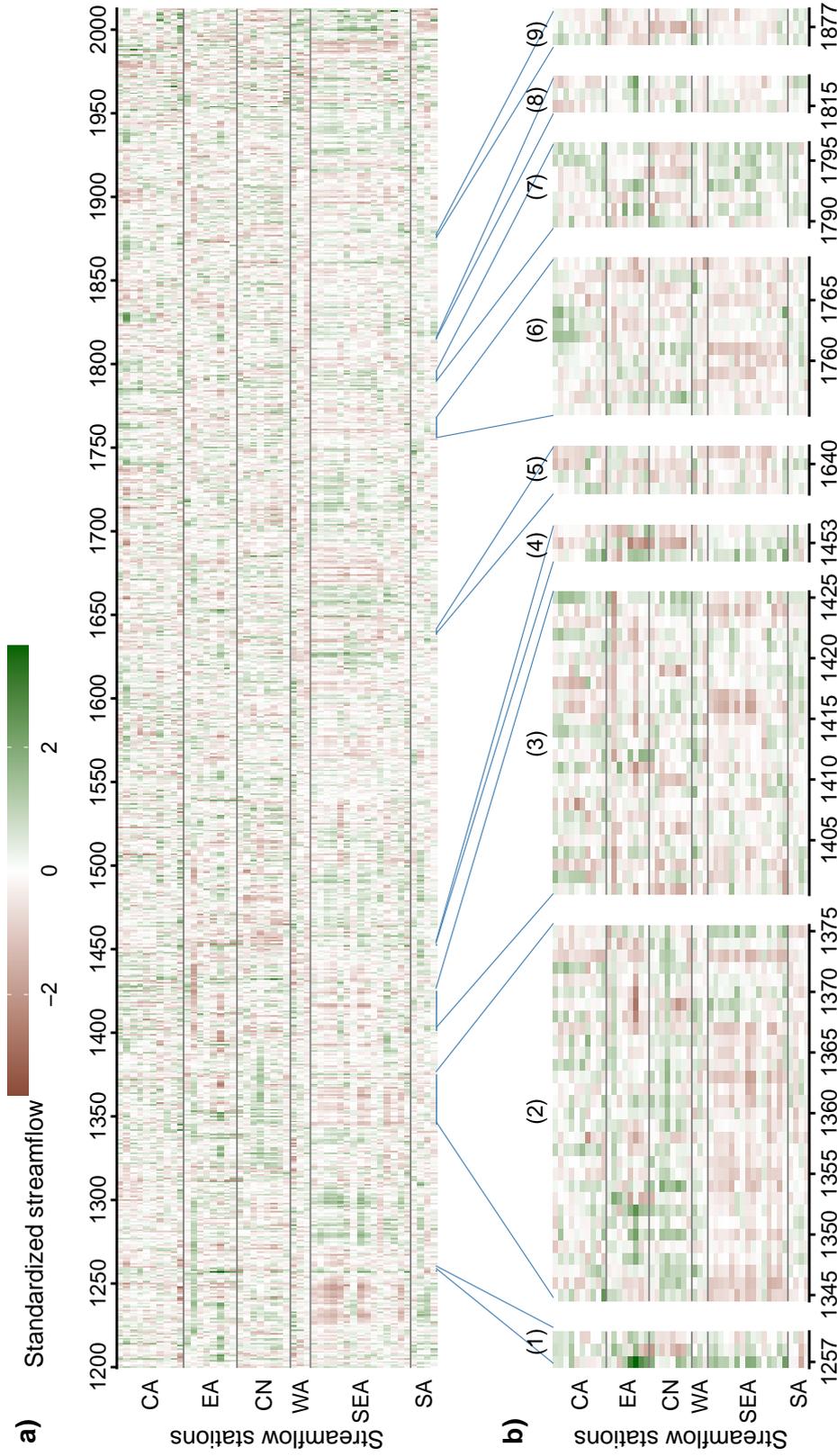


Figure 5. Spatiotemporal variability of streamflow in Asia. a) Variations over time (x-axis) and space (y-axis) of the standardized streamflow index (i.e., z-score of streamflow, or z-score of log-transformed streamflow when log-transformation was used in the reconstruction). The stations are arranged approximately north to south (top down on y-axis) and divided into five regions as delineated in Figure 1: CA (Central Asia), EA (East Asia), WA (West Asia), CN (eastern China), SEA (Southeast Asia), and SA (South Asia). b) Historic events captured in the reconstruction: (1) Samalas eruption, (2) and (3) Angkor Droughts I and II, (4) Kuwae eruption, (5) Ming Dynasty Drought, (6) Strange Parallels Drought, (7) East India Drought, (8) Tambora eruption, and (9) Victorian Great Drought.

412 inflow/outflow functions of their *barays* (reservoirs) in an attempt to preserve water. Heavy
 413 rains and flooding subsequently destroyed the reduced-capacity hydraulic infrastructure.
 414 This flood likely occurred in 1375 (Figure 5b, event 2).

415 By the second Angkor Drought, the hydraulic city had insufficient water storage
 416 and could not recover. Four more megadroughts that severely affected Asian societies
 417 (Cook et al., 2010) are also captured in our reconstruction, along with other major droughts
 418 and pluvials. Central Asia observed a six-decade drought between 1260–1320 and sus-
 419 tained pluvials during 1740–1769. East Asia experienced extended drought in the sec-
 420 ond half of the fifteenth century. Most notably, Southeast Asia suffered a drought be-
 421 tween 1225–1255 that was comparable in length to Angkor Drought I, but more severe
 422 in magnitude. Following this drought was a multi-decadal pluvial in 1271–1307. The drought
 423 is prominent in the speleothem record of Wang et al. (2019), and the pluvial can also
 424 be traced from there.

425 **4.3 Links to Oceanic Drivers**

426 To exemplify the spatial variation of how the oceans influence streamflow, we se-
 427 lected four river basins from west to east: Krishna, Chao Phraya, Mekong, and Yangtze
 428 (Figure 1), and selected one station from each basin. Because of over-year storage in the
 429 Krishna, the only station that met our data quality criteria (Section 2.1) lies in the up-
 430 per reach of the river. For the other basins, we were able to choose stations in the down-
 431 stream that are more representative of the basin. The names and locations of these sta-
 432 tions are shown in Figure 3.

433 We calculated the correlation between reconstructed annual streamflow at each basin
 434 and the seasonal averages of global sea surface temperature (SST) for the period 1855–
 435 2012. Correlation patterns vary both seasonally and spatially, with differences among
 436 rivers and among oceans.

437 **4.3.1 Pacific Ocean**

438 For the Krishna, correlations are weak, and small significant correlation areas are
 439 observed in the tropical Pacific, mainly from summer to winter of the current year (Fig-
 440 ure 6a). Tropical Pacific SST—a manifestation of the El Niño-Southern Oscillation (cf.
 441 McPhaden et al., 2006)—correlates negatively with streamflow. The hydroclimate of South
 442 Asia tends to be drier during El Niños and wetter during La Niñas. These tendencies
 443 have also been observed from tree ring records (Borgaonkar et al., 2010), reconstructed
 444 PDSI (Yu et al., 2018) and precipitation (Shi & Wang, 2018). The seasonality of cor-

445 relation suggests that annual streamflow responds more strongly to an ongoing ENSO
 446 event than to a decaying one.

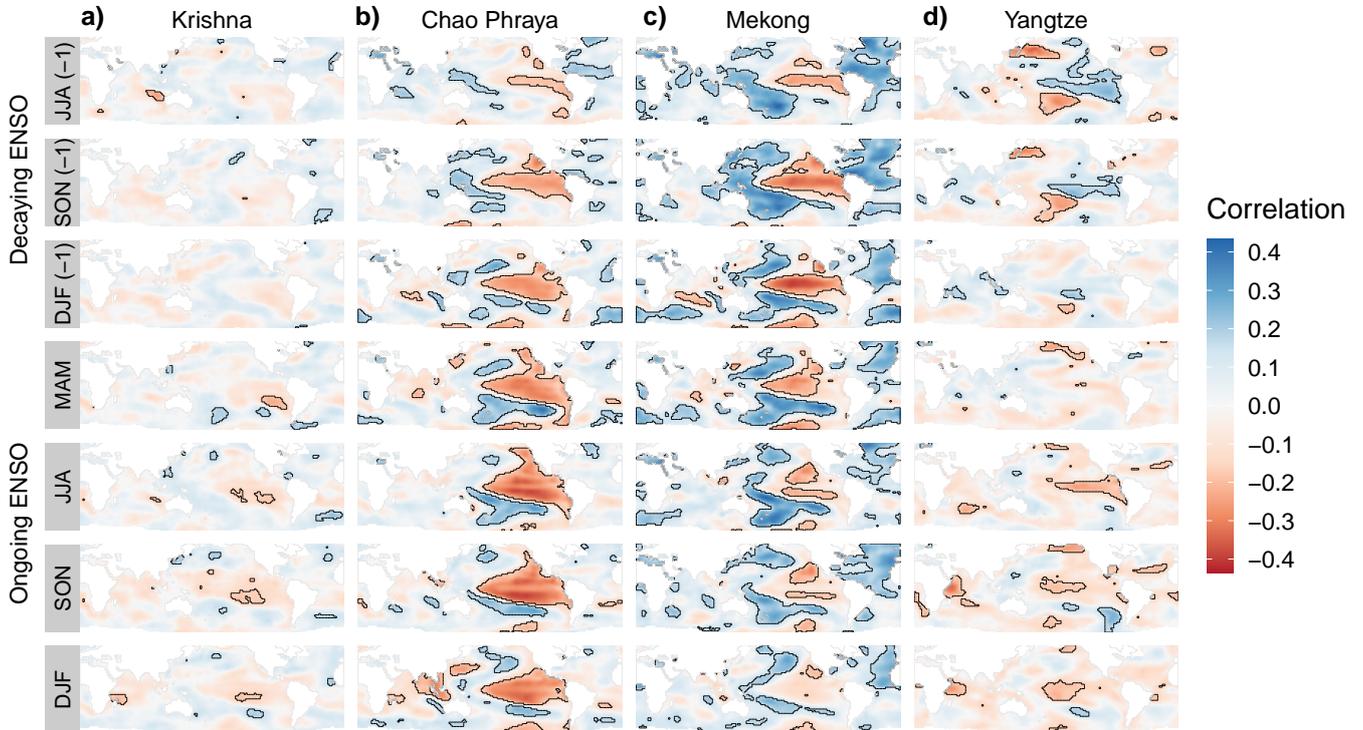


Figure 6. Correlation between reconstructed mean annual streamflow at four river basins (this work) and seasonal averages of global sea surface temperature (SST) from the NOAA.ERSST.v5 data set (Huang et al., 2017) for the period 1855–2012; significant correlations ($\alpha = 0.05$) enclosed in black boundaries. The stations used are shown in Figure 3. “(-1)” denotes previous year.

447 The Yangtze has a similar current summer–winter Pacific SST correlation pattern
 448 to that of the Krishna (i.e, related to ongoing ENSO events), with larger significant cor-
 449 relation areas. It also responds to decaying ENSO events (prior summer–winter) much
 450 more strongly than does the Krishna. Correlation with decaying ENSO events takes the
 451 opposite sign to that of ongoing ENSO events (Figure 6d). These opposite ENSO influ-
 452 ences on eastern China have been shown in a similar seasonal correlation analysis using
 453 reconstructed precipitation (Shi & Wang, 2018) but not in the annual composite anal-
 454 yses of Yu et al. (2018) and Li et al. (2013). The latter two works showed wetter ten-
 455 dencies during El Niño and drier tendencies during La Niña, likely capturing only the
 456 decaying phase.

457 Unlike in the Krishna and Yangtze, streamflow in the Chao Phraya and Mekong
458 correlates significantly with SST over most of the Pacific Ocean, and the correlation per-
459 sists across all seasons, reflecting equal influences from decaying and ongoing ENSO events
460 (Figure 6b and c). The basin-wide correlation pattern and its lack of seasonality sug-
461 gest influences from a driver at longer time scales, likely the Pacific Decadal Variabil-
462 ity (PDV)—decadal variations of Pacific SST resulted from complex tropical-extratropical
463 ocean-atmosphere interactions (Henley, 2017). The North Pacific component of PDV is
464 known as the Pacific Decadal Oscillation (PDO) (Mantua & Hare, 2002), its southern
465 counterpart the South Pacific Decadal Oscillation (Shakun & Shaman, 2009); basin-wide
466 SST variation patterns have also been termed Interdecadal Pacific Oscillation (Folland
467 et al., 1999). These modes are closely related (Henley, 2017). The PDO has been shown
468 to influence hydroclimatic variability in Monsoon Asia, in conjunction with ENSO (Yu
469 et al., 2018). Specifically for the Chao Phraya, PDO also modulates ENSO’s influence
470 on peak flow (C. Xu et al., 2019). Here, by juxtaposing the correlation maps, our anal-
471 yses reveal that ENSO and PDV exert their greatest influence on the Mekong and Chao
472 Phraya.

473 **4.3.2 Indian Ocean**

474 We observe negative correlations between streamflow and Indian Ocean SST in current-
475 year winter in the Chao Phraya, and to a lesser extent in the Yangtze and Krishna. These
476 basin-wide correlation patterns follow closely after peak ENSO correlations in summer
477 and fall, consistent with the Indo-Pacific coupling described by Saji et al. (1999): an ENSO
478 event in the Pacific leads to SST anomalies of the same sign Indian Ocean. This mode
479 accounts for about 30% of Indian Ocean SST variability [*ibid*]. These authors also pro-
480 posed another mode—the Indian Ocean Dipole (IOD) mode, the positive phase of which
481 is characterized by cool eastern Indian Ocean around Sumatra, and warm western In-
482 dian Ocean around East Africa. Positive IOD events often occur around June–July, peak
483 in October and abruptly end in November, a phenomenon called seasonal locking (Saji
484 et al., 1999; Ummenhofer et al., 2017). Positive IOD events have been linked to droughts
485 in Southeast Asia but this relationship is not robust (Ummenhofer et al., 2013). Con-
486 sistent with their results, we observe a weak east-positive–west-negative correlation pat-
487 tern between Indian Ocean SST and Southeast Asia streamflow (Mekong and Chao Phraya)
488 in the fall (the peak IOD season), both for prior- and present-year, with small areas of
489 significant correlation. This pattern becomes stronger in prior-year winter, suggesting
490 a lag between peak IOD and its effect to Southeast Asia, but not so for present-year win-
491 ter, likely because it is dominated by the basin-wide ENSO mode (the IOD mode only
492 accounts for 12% of Indian Ocean SST variability (Saji et al., 1999)).

4.3.3 *Atlantic Ocean*

The Chao Phraya and Mekong streamflow correlates positively with tropical and northern Atlantic SST. Significant and consistent correlations are observed throughout the seasons for the Mekong, but less consistent for the Chao Phraya. Wang et al. (2019) proposed a mechanism to explain relationship: increased tropical Atlantic SST leads to changes in zonal moisture transport, causing depression over tropical Indian Ocean, reducing rainout over the basin, leaving more moisture available to be transported to mainland Southeast Asia, ultimately strengthening Indian Monsoon rain over the region. This mechanism is consistent with their speleothem record and also with our streamflow reconstruction.

4.3.4 *Temporal variability*

Figure 6 shows how teleconnection between Monsoon Asia’s streamflow and global SST varied among river basins. To see if this teleconnection also changed through time, and how, we repeated the same analysis in three sub-periods: 1855–1904, 1905–1954, and 1955–2004 (see Figure S6). We observe the following. First, our reconstruction captures the SST correlation patterns in the instrumental period (1955–2004), thereby further validating the quality of our reconstructions. Second, the SST correlation patterns changed through time for all four rivers, but more interestingly, teleconnection weakened remarkably for the Chao Phraya, Mekong, and Yangtze during 1905–1954 compared to the other two time windows.

5 **Conclusions**

In this work, we produce the first large-scale and long-term record of streamflow variability for Monsoon Asia, covering 48 stations in 16 countries. In making this record, we also develop a novel automated, climate-informed, and dynamic streamflow reconstruction framework that leveraged the computational advantages offered by our climate proxy—the Monsoon Asia Drought Atlas (MADA) version 2. Our framework achieves good skills for 96% of the stations, and skill distribution is spatially homogeneous. Our results provide a synoptic understanding of Monsoon Asia’s streamflow variability over the past eight centuries, and reveal how the teleconnection between streamflow and its oceanic drivers varied over space and time.

From our reconstruction, streamflow in Monsoon Asia appears coherent: high and low flows often occur simultaneously at nearby stations and adjacent basins. This coherence is attributed to common oceanic drivers—the El Niño–Southern Oscillation (ENSO),

526 the Pacific Decadal Variability (PDV), the Indian Ocean Dipole, and tropical Atlantic
527 sea surface temperature variations. Coherence emerges even though we reconstructed
528 each station individually, demonstrating the merits of Point-by-Point Regression. More
529 importantly, this coherence implies that water management in Asia should be coordi-
530 nated among basins. For example, Thailand is increasingly purchasing Mekong-generated
531 hydropower from Laos, and when that is insufficient, complements its energy needs with
532 thermal power from plants that use water from the Chao Phraya for cooling (Chowdhury
533 et al., 2019). Thailand’s energy system is at risk when a prolonged drought occurs at
534 both rivers—our record shows such events have happened several times in the past.

535 We showed that the Pacific, Indian, and Atlantic Oceans contribute to streamflow
536 variability. Therefore, water management in Monsoon Asia should take into account the
537 ocean states. A case study of the Angat River (the Philippines) showed that reservoir
538 operating policies informed by ENSO states are more robust than conventional policies
539 that only account for local hydrological conditions (Libisch-Lehner et al., 2019). Oper-
540 ating policies may be improved further if, say, the PDV is also considered. There is prob-
541 ably even more potential for improving water resources management in the Mekong and
542 Chao Phraya River Basins, as the oceanic drivers exert very strong influences on these
543 rivers.

544 **Acknowledgments**

545 Hung Nguyen is supported by the President’s Graduate Fellowship from the Singapore
546 University of Technology and Design. We thank Edward Cook, Caroline Ummenhofer,
547 Nerilie Abram, Nathalie Goodkin, Xun Sun, Murray Peel, and Rory Nathan for insight-
548 ful comments. We are grateful to Thanh Dang, Mukund Rao, Rosanne D’Arrigo, Donghoon
549 Lee, and Caroline Leland for streamflow data of the Mekong, Brahmaputra, Citarum,
550 Han, and Yeruu Rivers. Chao Phraya River data were obtained from the Thai Royal Ir-
551 rigation Department at www.hydro1.net, Indus River from Rao et al. (2018, Support-
552 ing Information), other streamflow data from GSIM (Do et al., 2018; Gudmundsson et
553 al., 2018), reservoir data from GRanD v1.3 (Lehner et al., 2011), MADA v2 data from
554 Marvel et al. (2019) at www.dropbox.com/s/n21o99h9qn17prg/madaV2.nc, river net-
555 work data from FLO1K (Barbarossa et al., 2018) with help postprocessing by Valerio
556 Barbarossa, basin boundary data from HydroSHEDS (Lehner & Grill, 2013) at [hydrosheds](http://hydrosheds.org)
557 [.org](http://hydrosheds.org), SST data from NOAA ERSST v5 (Huang et al., 2017) provided by the NOAA/OAR/ESRL
558 PSD, Boulder, Colorado, USA, at www.esrl.noaa.gov/psd/. This work was conducted
559 with open-source software: analysis and visualization performed in R (R Core Team, 2019),
560 maps made in QGIS, and manuscript written in LaTeX. We thank the open-source soft-
561 ware community, especially the R package creators and maintainers, for their contribu-

562 tions to open science. We provide all data and code, together with a step-by-step guide
 563 to reproduce the results, in a GitHub repository (Nguyen, 2020b) available at [https://](https://github.com/ntthung/paleo-asia)
 564 github.com/ntthung/paleo-asia; exceptions are instrumental data of the Mekong, Yangtze,
 565 and Pearl rivers due to restrictions. Lamont contribution number XXXX.

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