

International Journal of Forecasting

Ensemble Calibration and Uncertainty Quantification of Precipitation Forecasts for a Risk-based UAS Navigation --Manuscript Draft--

Manuscript Number:	
Full Title:	Ensemble Calibration and Uncertainty Quantification of Precipitation Forecasts for a Risk-based UAS Navigation
Short Title:	
Article Type:	Full Length Article
Keywords:	Ensemble calibration; Uncertainty quantification; CRPS learning; aerospace weather; probabilistic forecasting
Corresponding Author:	Mounir CHRIT University of North Dakota UNITED STATES
Corresponding Author Secondary Information:	
Corresponding Author's Institution:	University of North Dakota
Corresponding Author's Secondary Institution:	
First Author:	Mounir Chrit
First Author Secondary Information:	
Order of Authors:	Mounir Chrit
Order of Authors Secondary Information:	
Abstract:	<p>Uncertainty on precipitation forecasts results in major high cancellation rate in UAS operations and reduces the benefits of BVLOS operations in terms of risk-based contingency planning. Hence, quantifying and reducing the uncertainty on precipitation forecasts will reduce mission uncertainties and avoid accidents . To achieve this goal, the Member-By-Member post-processing technique is used to calibrate a probabilistic forecast composed of 20 members of precipitation rate over South Florida during summer period. The CRPS of the ensemble is minimised to achieve the optimal regression between ensemble members without any assumption on the forecasted parameter. The MRMS observations is used to correct ensemble spread every 10 min and reduce forecasting uncertainty . A multi-physics ensemble was used to generate high-resolution, convection resolving/allowing forecasts . The comparison between the raw and calibrated ensemble from unseen data is presented in terms of bias correction and ensemble reliability . The calibration was able to correct the bias found in raw probabilistic forecasts relative to MRMS data. The comparison with precipitation data from tipping buckets over four airports revealed that the calibrated ensemble tends to overestimate the precipitation rates mainly because of the particles evaporation that is taking place under radar beam.</p>
Suggested Reviewers:	
Opposed Reviewers:	
Additional Information:	
Question	Response

Declaration of interests

The authors declare that they have no known competing financial interests or personal relationships that could have appeared to influence the work reported in this paper.

The authors declare the following financial interests/personal relationships which may be considered as potential competing interests:

Ensemble Calibration and Uncertainty Quantification of Precipitation Forecasts for a Risk-based UAS Navigation

Mounir Chrit¹

1. Department of Atmospheric Sciences, University of North Dakota, Grand Forks, ND 4149, USA.

Abstract

Uncertainty on precipitation forecasts results in major high cancellation rate in Unmanned Aircraft Systems operations and reduces the benefits of BVLOS operations in terms of risk-based contingency planning. Hence, quantifying and reducing the uncertainty on precipitation forecasts will reduce mission uncertainties, avoid accidents and make the integration of UAS into the National Airspace System more efficient and reliable. To achieve this goal, the Member-By-Member post-processing technique is used to calibrate a probabilistic forecast composed of 20 members of precipitation rate over South Florida during summer period. The Continuous Ranked Probability Score (CRPS) of the ensemble is minimized to achieve the optimal regression between ensemble members without any assumption on the forecasted parameter. The radar data from the Multi-Radar/Multi-Sensor (MRMS) is used to correct ensemble spread every 10 min and reduce forecasting uncertainty. A multi-physics ensemble was used to generate high-resolution, convection resolving/allowing 48-hours forecasts. The calibration was obtained over a learning process over the 48-h simulated period over 3 years. The comparison between the raw and calibrated ensemble from unseen data is presented in terms of bias correction and ensemble reliability. The calibration was able to correct the bias found in raw probabilistic forecasts relative to MRMS data. The comparison with precipitation data from tipping buckets over four airports over South Florida revealed that the calibrated ensemble tends to overestimate the precipitation rates mainly because of the particles evaporation that is taking place under radar beam.

Introduction

Weather events, such as heavy rain accompanied by strong winds, hail, lightning or thunderstorms present a non-negligible threat to the small Unmanned Aircraft Systems (sUAS) navigation (Campbell et al. 2017). In addition, a primary concern of UAS operators conducting BVLOS missions is lost-link contingency planning. The current operational philosophy regarding flying in wet conditions is to be very conservative in go/no-go decision making even if there is a small chance of precipitation on either the flight plan or contingency routing. This results in a relatively high cancellation rate compared to conventional aviation, and significantly decreases the efficiency and potential benefit of BVLOS operations.

In addition, precipitation uncertainty significantly impacts lost-link contingency planning due to the need to avoid the safety problems that arise with lost-link coupled with the need to avoid

1
2
3
4 overly conservative assessments of weather impact. More effective contingency planning would
5 be possible if there were precipitation uncertainty information that could be used for a risk-based
6 contingency planning. Campbell 2017 recommended to 1) explore concepts to provide UAS
7 operators with weather impact uncertainty information to aid in contingency planning and 2)
8 reduce uncertainty of precipitation forecasts. The intent of these recommendations is to decrease
9 mission cancellation rates by increasing user confidence in forecasts of route-based weather
10 impacts.
11
12
13

14 Within this risk-based planning approach, ensemble forecasting, and more generally probabilistic
15 forecasting has been widely covered in the meteorological community (Gneiting & Katzfuss,
16 2014) to provide more accurate forecasts and uncertainty information. Moreover, the ensemble
17 mean is the quantity usually disseminated while the ensemble spread is a measure of the flow-
18 dependent forecast uncertainty. However, it is well-known that, for state-of-the-art weather
19 forecasts, the uncertainty measure is not very accurate (Schulz and Lerch 2022) as models'
20 accuracy strongly degrades as a function of lead time (Nicolis et al. 2009). Moreover, at the
21 surface, experiments show that ensemble forecasts are consistently under-dispersive (or
22 overconfident) for long lead times (Leutbecher and Palmer, 2008). This feature can be partly
23 traced back to systematic errors, relevant to the model at hand, that could be partly corrected by
24 calibration or post-processing. That is why post-processing methods that are based on different
25 regression techniques were introduced for the improvement of both scalars (Vannitsem, 2009;
26 Van Schaeybroeck and Vannitsem, 2011, 2012) but also for vector variables such as wind
27 (Pinson, 2012).
28
29
30
31
32

33 Two approaches exist today to correct the ensemble members. The first method is 'statistical'
34 calibration approaches which assume specific ensemble distributions and have predictive
35 distributions as output, rather than an ensemble of discrete size. For example, logistic distribution
36 has been successfully applied in the context of post-processing of precipitation forecasts (Wilks,
37 2009; Schmeits and Kok, 2010; Roulin and Vannitsem, 2012). For temperature, one of the most
38 competitive approaches is the Non-homogeneous Gaussian Regression (Gneiting et al., 2005;
39 Hagedorn et al., 2008). NGR uses Gaussian predictive distributions with mean and spread that
40 depend linearly on the corresponding quantities of the raw forecast. However, applying
41 statistical post-processing methods to reconstruct ensembles by random sampling do not take
42 into account the correlations between the values of nearby stations, lead times and
43 meteorological variables as shown in Van Schaeybroeck et al. 2015. The second approach,
44 adopted in this work, is member by member (MBM) correction in which each ensemble member
45 is corrected individually by a linear mapping, thereby retaining rank correlations. Therefore,
46 each member retains to a large extent correlation structures in the case of multiple independent
47 calibrations (Van Schaeybroeck and Vannitsem 2015). Moreover, in terms of skill our MBM
48 approach can be as high as NGR as shown in Van Schaeybroeck and Vannitsem 2015.
49
50
51
52
53
54

55 Different loss functions and fitting procedures exist today to conduct MBM correction: Bayesian
56 model averaging (BMA) techniques provide a mixture of parametric distributions, usually a
57 Gaussian sum (Raftery et al. 2005) or gamma distributions sum for wind and precipitation
58 applications (Sloughter et al. 2010, Sloughter et al. 2007). Non-homogeneous regression fits the
59
60
61
62
63
64
65

parameters of a parameterized distribution using characteristics of the ensemble of forecasts Gneiting et al. (2005), Thorarinsdottir and Gneiting (2010), Wilks (2009). For instance, a Gaussian distribution is fitted using a linear model between the mean of the distribution and the mean of the forecasts. Besides, likelihood maximization with the logarithm loss is not an appropriate tool in our setting since it fails to produce satisfactory scores for a discrete probability distribution. A discussion on local scores such as the logarithm loss is addressed by Bröcker and Smith (2007b). However, these techniques do not offer theoretical guarantees of robustness and usually resort to strong assumptions on the distributions. The continuous ranked probability score (CRPS) is the squared difference between the cumulative distribution functions of the ensemble forecast and the observation was used by Thorey et al. 2018, Gneiting et al. 2005, Gebetsberger et al. 2017 as a cost function to obtain calibrated probabilistic forecasts as it does not need a theoretical assumption regarding parameters distribution.

The goal of this paper is to show how ensemble-spread correction using CRPS minimization relative to the Multi-Radar Multi-Sensor (MRMS) precipitation data yield to an improvement of the predictions and evaluate the performance of probabilistic forecasts of precipitation by comparison to airport precipitation observations. In this study, we start with 20-members ensemble of precipitation forecasts and apply a MBM calibration approach developed by Schaeybroeck et al. 2015 to improve the probabilistic forecasts of a precipitation event in South Florida.

This paper is structured as follows: section 1 describes the calibration method. Section 2 discusses the simulated use case, the simulation setup and ensemble building, and the datasets used in the calibration and evaluation. Section 3 explains the evaluation method. Section 4 discusses the results and evaluation findings.

1. Ensemble Calibration

1.1. MBM post-processing method

Following Schaeybroeck and Vannitsem 2015, the calibrated ensemble of M members at time n $X_{C,n} = (X_{C,n}^m)_{1 \leq m \leq M}$ can be expressed as a function of the raw ensemble $X_n = (X_n^m)_{1 \leq m \leq M}$ as shown in Equation 1 where \bar{X}_n is the ensemble-mean values, regression coefficient α is the bias parameter while the coefficient β is the ensemble-mean scale parameter. The parameter τ_n defined in Equation 2 adjusts the spread of the new ensemble and the deviation from the ensemble mean ϵ_n is defined in equation 4. The fact that τ_n depends on the ensemble index n comes from its dependence on ensemble spread δ_n defined in Equation 3. $\langle \cdot \rangle_m$ denotes the ensemble average.

$$X_{C,n} = \alpha + \beta \bar{X}_n + \tau_n \epsilon_n \quad (1)$$

$$\tau_n = \gamma_1 + \gamma_2 \delta_n^{-1} \quad (2)$$

$$\delta_n = \left\langle \left\langle |X_n^{m_1} - X_n^{m_2}| \right\rangle_{m_1} \right\rangle_{m_2} \quad (3)$$

$$\epsilon_n = X_n - \bar{X}_n \quad (4)$$

The correlation between the corrected ensemble mean and the observation is equal to the correlation between the uncorrected ensemble mean and the observation (Johnson and Bowler, 2009). The standard deviation of the corrected ensembles is used as a spread measure of the corrected forecasts to quantify the uncertainty of the forecasts.

1.2. CRPS minimization

The parameters $(\alpha, \beta, \gamma_1, \gamma_2)$ are estimated through regression learning through 3 years by the minimization of the associated Continuous Ranked Probability Score (CRPS) which is the squared difference between the Cumulative Distribution Functions (CDFs) of the ensemble forecast and the observation.

The loss function defined as the CRPS corresponding to the observations $X_{o,n}$ and the corrected-forecast members $X_{c,n}^m$ can be written as shown in Equation 5 (Gneiting and Raftery, 2007). The minimization and hence the correction is used every 10 min during the two simulated summer days for 3 years: 2019, 2020 and 2021. The forecast ensemble used here covers three years 2019, 2020 and 2021 and the CDF of the observations was based on the radar MRMS observations over the same years. Data from 2022 will be used as an independent test for the calibration. A short training period was chosen which 48 hours. In fact, there is a trade-off in selecting the length of the training period. Shorter training periods can adapt rapidly to seasonally varying model biases, changes in the performance of the ensemble member models, and changes in environmental conditions. On the other hand, longer training periods reduce the statistical variability in the estimation of the different coefficients and hence the calibrated PR.

$$CRPS(\alpha, \beta, \gamma_1, \gamma_2) = \left\langle \left\langle |X_{c,n}^m - X_{o,n}| \right\rangle_m - \frac{\delta_n}{2} \right\rangle_n \quad (5)$$

2. Materials and Methods

2.1. Use case description

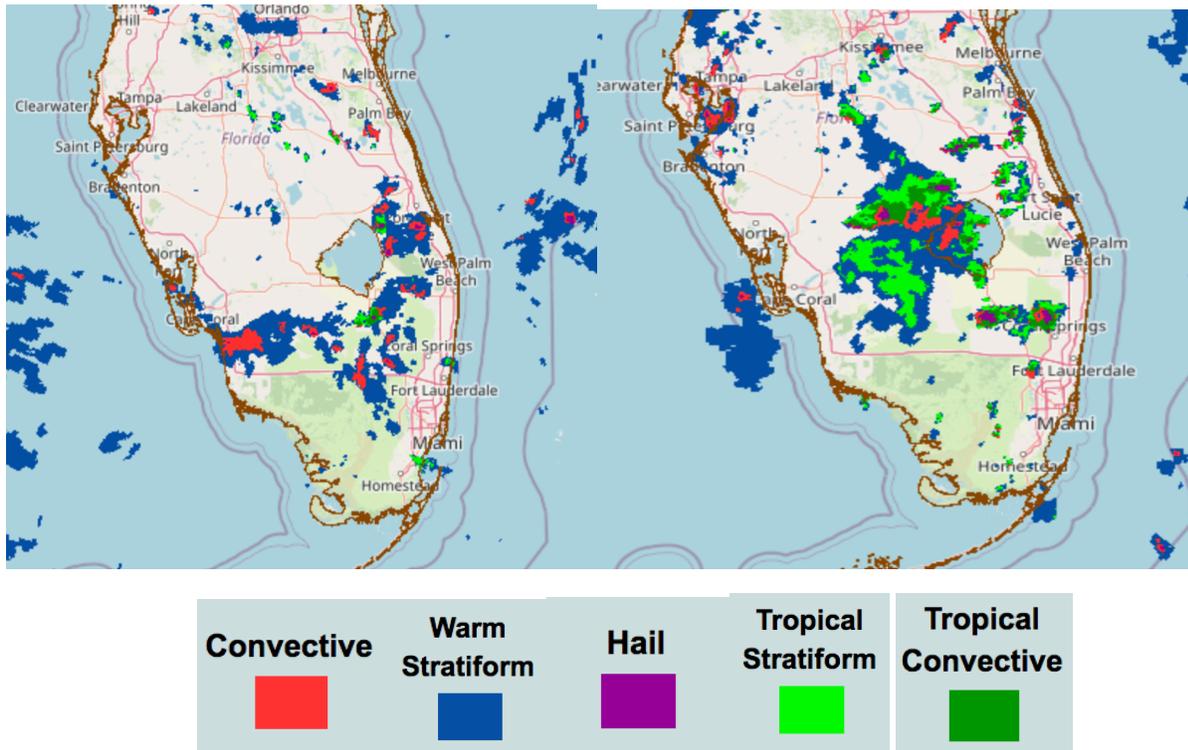
The simulated event is precipitation event that took place in South Florida that was visible in the MRMS data with scales of 200 km and small scales of 1-50 km as shown Figures 1 and 3. These events fall under Meso- β and Meso- γ features.

In South Florida, particularly during the summer, mesoscale weather features (e.g., land-sea breezes, thermal troughs, outflow boundaries, etc.) have a significant impact on day-to-day weather forecasting, as they frequently represent the primary forcing for convection. During the simulated period,

These mesoscale features necessitate the use of high-resolution forecast tools in order to provide the detailed information needed to improve local forecasts and warnings. Moreover, Florida has

1
2
3
4 recently emerged as a leader in autonomous vehicles including UAS through different
5 investments in its Department of Transportation. Therefore, South Florida is a suitable area to
6 study precipitation forecasting and its impact on UAS contingency planning.
7
8

9 During the simulation summer period, precipitations were of different types: mainly convective
10 because sea breezes are often form on the west and east sides of Florida, and due to differences
11 in temperature between the land (which heats quickly) and the ocean (which heats up more
12 slowly) which enhance the convective lift and induce intense rainfall and thunderstorms.
13 Convective and tropical convective precipitation are often embedded in areas of warm stratiform
14 precipitation. Warm stratiform precipitations are also present in South Florida that result from
15 frontal systems where the growth of hydrometeor particles occurs.
16
17
18



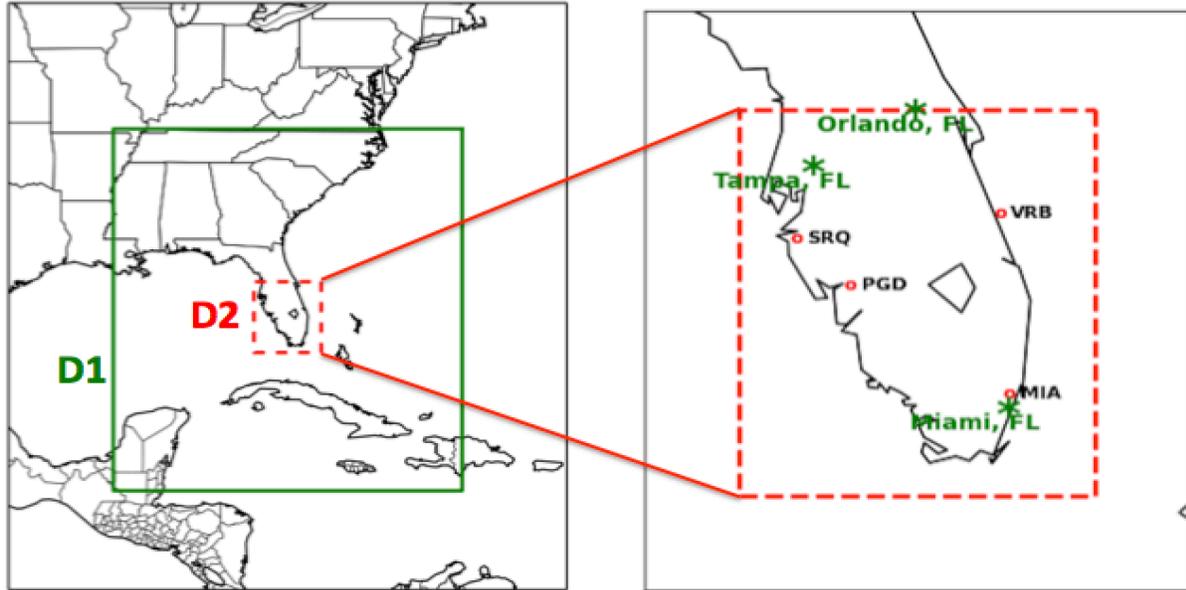
39
40
41
42
43
44
45 Figure 1: MRMS Precipitation type over South Florida on July 16th at 5:40pm (left
46 panel) and on July 17th at 6:44 pm (right panel
47 (https://mrms.nssl.noaa.gov/qvs/product_viewer/) .
48

49 2.2.Ensemble Forecasts

50 2.2.1. Simulations Setup

51
52 WRF (Sharmarock et al. 2005) was widely used in both academic research and industry (Chrit et
53 al. 2022, Chrit et al. 2018, Chrit et al. 2017). A fully compressible and non-hydrostatic dynamic
54 framework is adopted in the ARW module. The outermost and innermost simulated domains D1,
55 D2 respectively are shown in Figure 2. The horizontal resolutions of D1, D2 are 3km and 1km.
56
57 The two domains are centered on 80.74332 °W, 26.40334 °N.
58
59
60
61
62
63
64
65

1
2
3
4 The outermost D1 and innermost D2 domains have 560 x 720 and 460 x 400 grid points
5 respectively in the south-north and east-west directions. The WRF model contains 80 vertical
6 levels in the vertical and the lowest 30 levels are below 1-km. The adaptive time stepping is used
7 to guarantee the numerical stability of the WRF model. Table [1] of Appendix A shows the
8 configuration and the parameterizations used in the simulations over D1 and D2.
9
10



11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33 Figure 2: Left panel: Map of the simulated outermost and nested domains D1 and D2 delimited
34 with green solid and red dashed rectangles, respectively. Right panel: Simulation domain is
35 delimited with red dashed rectangle with the four ASOS stations used for evaluation shown with
36 red points. The Three major cities in Florida (Miami, Orlando and Tampa) are shown in green
37 stars.
38

39 40 2.2.2. Ensemble design

41
42 The ensemble used in the present study is a multi-physics ensemble with forecasts initialized
43 with different initial and boundary conditions. In fact, multi-physics schemes have been very
44 successful in generating reliable probabilistic forecasts, especially for mesoscale prediction
45 systems. Although the maintenance of these forecasts is resource-intensive when deployed
46 operationally, the ensemble will result in members with physical interpretation rather than
47 members generated with perturbed initial conditions, which poses difficulties for physical
48 interpretation and statistical post-processing. On the other hand, precipitation forecasting is
49 sensitive to details of the cumulus convection scheme (Vitart et al. 2001; Biswas et al. 2014),
50 microphysics scheme (Liu et al. 2020), boundary layer parameterization (Taraphdar and Pauluis
51 2021) and radiations schemes (Li et al. 2014).
52
53
54

55
56 20 different combinations of physics packages for parameterizing the microphysics (MP
57 scheme), cumulus (C), Short Waves (SW) and Long Waves (LW) parameterization, planetary
58 boundary layer (PBL), and land-surface models, (Table 1) are used to build four ensembles: three
59 ensembles for the training and the fourth for testing. To maximize ensemble diversity, different
60
61
62
63
64
65

boundary and initial conditions were used based on 4 models: the North American Model (NAM), Rapid Refresh (RAP), North American Regional Reanalysis (NARR) and Global Forecast System (GFS). A total of 20 WRF simulations were performed to build the ensemble. Similar ensemble design were built for three years 2019, 2020 and 2021.

Two MP parameterizations used are Microphysics schemes used are Thompson (Thom.; Thompson et al. 2008), WRF single-moment 6-class (WSM6; Hong and Lim . 2006). The C schemes used here are: Kain–Fritsch (Kain and Fritsch, 1993) cumulus parameterization, and Betts–Miller–Janjic cumulus parameterization (Betts & Miller, 1993). Two PBL parameterizations were used: Mellor–Yamada–Janjic (MYJ; Janjic 1994), Yonsei University (YSU; Noh et al. 2003). Two Land-Surface models were used: Rapid Update Cycle (RUC; Benjamin et al. 2004) or NOAH (NCEP–Oregon State University–Air Force–NWS Office of Hydrology; Ek et al. 2003). The SW parameterizations are Goddard (Tao et al. 2003) and Dudhia (Dudhia 1989), the LW radiations schemes are RRTM (Mlawer et al. 1997) and GFDL (Fels and Schwarzkopf 1981).

Member number	ICs and LBCs	MP scheme (Thom and WSM6)	PBL parameterization (MYJ and YSU)	Land-Surface model (NOAH and RUC)	SW parameterization (GFDL and DUDHIA)	LW parameterization (GFDL and RRTM)	C parameterization (KAIN FRTISCH and BMJ)
1	NAM	Thom	MYJ	NOAH	DUDHIA	RRTM	Kain
2	NAM	WSM6	MYJ	NOAH	DUDHIA	RRTM	Kain
3	NAM	Thom	YSU	NOAH	DUDHIA	RRTM	Kain
4	NAM	Thom	MYJ	RUC	DUDHIA	RRTM	Kain
5	NAM	Thom	MYJ	NOAH	GFDL	RRTM	Kain
6	NAM	Thom	MYJ	NOAH	DUDHIA	GFDL	Kain
7	NAM	Thom	MYJ	NOAH	DUDHIA	RRTM	BMJ
8	NAM	Thom	YSU	RUC	DUDHIA	RRTM	Kain
9	NAM	Thom	YSU	RUC	GFDL	RRTM	Kain
10	NAM	WSM6	YSU	RUC	DUDHIA	GFDL	BMJ
11	RAP	WSM6	YSU	RUC	DUDHIA	RRTM	Kain
12	NARR	Thom	MYJ	NOAH	DUDHIA	RRTM	Kain
13	GFS	Thom	MYJ	NOAH	DUDHIA	RRTM	Kain
14	NARR	Thom	MYJ	NOAH	DUDHIA	RRTM	Kain
15	RAP	Thom	MYJ	NOAH	DUDHIA	RRTM	Kain
16	RAP	Thom	YSU	RUC	GFDL	RRTM	Kain
17	NAM	WSM6	YSU	RUC	DUDHIA	GFDL	Kain
18	RAP	Thom	MYJ	RUC	DUDHIA	RRTM	Kain
19	GFS	Thom	MYJ	RUC	DUDHIA	RRTM	Kain
20	GFS	WSM6	MYJ	NOAH	DUDHIA	GFDL	Kain

1
2
3
4 Table 1: Physics packages for multi-physics ensemble: Parameterizations and
5 schemes used for every ensemble member.
6
7
8
9

10 **2.2.3. MRMS radar data**

11 The Multi-Radar/Multi-Sensor (MRMS) system was created at the NOAA National Severe
12 Storms Laboratory (NSSL) to produce severe weather and precipitation products for decision-
13 making capabilities to improve severe weather forecasts and warnings, hydrology, aviation, and
14 Numerical Weather Prediction. MRMS currently integrates about 180 operational radars and
15 creates a seamless 3D radar mosaic across the contiguous United States (CONUS) and southern
16 Canada at very high spatial (1 km) and temporal (2 min) resolution.
17
18
19

20 The performance of the MRMS system over single radar-based Quantitative Precipitation
21 Estimates (QPE) across CONUS was reasonable (Zhang et al., 2016). Chen et al. (2020)
22 evaluated the MRMS and Global Precipitation Measurement Mission (GPM) products at 1-hr
23 temporal resolution across Harris County and Spring Basin Texas. Their results showed that
24 remote sensing technologies could detect and estimate the unprecedented extreme rainfall
25 associated with Hurricane Harvey. Among the remote sensing products they used in their study,
26 MRMS had the best agreement with the network rain gauge observations.
27
28
29

30 The MRMS surface precipitation rate used in this paper is currently calculated using multiple R-
31 Z relationships. Polarimetric variables are not used because various polarimetric radar QPE
32 schemes are still under evaluation across CONUS and an optimal approach for all seasons and all
33 geographic regions has yet to be developed. The following empirical R-Z relationships are used
34 in MRMS to compute surface precipitation rate for each precipitation type: convective rain, hail,
35 warm and cold stratiform rain, snow and tropical stratiform mixed rain. More information about
36 the MRMS system can be found at NSSL's MRMS webpage (ASOS user guide), the MRMS
37 Fact Sheet (https://www.nssl.noaa.gov/news/factsheets/MRMS_2015.March.16.pdf), and
38 Kirstetter et al., 2012. The MRMS data for the two simulated days were re-gridded to the same
39 WRF grid over D2 with a 1-km resolution for every year of the learning and testing years.
40
41
42
43

44 **2.2.4. ASOS data**

45 The Automated Surface Observing System (ASOS) network provides most of the basic
46 hydrometeorological observations at different airports, including 1-hour accumulated
47 precipitation. The data is reported every 5 min in the majority of the stations. One hour
48 precipitation for the period from the observation time to the time of the previous hourly
49 precipitation reset. The precipitation accumulation algorithm obtains precipitation accumulation
50 data from the Heated Tipping Bucket (HTB) precipitation gauge once each minute (ASOS user
51 guide). The trace reports are considered as 0.1 mm. The detection threshold specified for the
52 ASOS HTB is 0.01 inch per hour (0.254 mm per hour), and the precipitation rate accuracy is the
53 larger of 10 percent or 0.01 inches per hour (0.254 mm per hour).
54
55
56
57

58 For this study, four METAR observation sites located over South Florida were used for the
59 evaluation of the different forecasts, and these sites are shown in Figure 1. Table [1] of Appendix
60
61
62
63
64
65

B shows the characteristics of the four stations that will be used for comparison and evaluation. Additional stations are available, but they are either no precipitation is recorded, or most data is missing.

3. Evaluation method

The probabilistic evaluation will be based on the rank histogram score and the reliability diagram. The rank-histogram score δ defined in Equation (6) is a tool used to measure the spread and hence the reliability of the ensemble.

$$\delta = \frac{N+1}{NM} \sum_{j=0}^N (r_j - \bar{r})^2 \quad (6)$$

$$\bar{r} = \frac{M}{N+1} \quad (7)$$

The rank-histogram score is used to measure the deviation from flatness of a rank histogram (Talagrand et al., 1999; Candille and Talagrand, 2005). In Equation (6), N is the number of members (i.e., models), M is the number of observations, r_j the number of observations of rank j, and \bar{r} is the expectation of r_j defined in Equation (7). In theory, the optimal ensemble has a score of 1 when enough members are available. A score lower than 1 would indicate overconfidence in the results, with an ensemble matching the observed variability better than statistically expected.

From a frequentist perspective, assuming that we test the occurrence of an event with probability p on several occasions, the proportion of trials where the event occurs is approximately equal to p. Based on this interpretation, a desirable property of a probabilistic prediction system is reliability. For p in [0, 1], we define the relative frequency f(p) as the proportion of the events that occurred among the events for which the system assigned a forecast probability p. A prediction system is reliable if, $\forall p \in [0, 1], f(p) = p$. For instance, with a reliable prediction system, among the events that obtained a probability of p = 0.2 according to the prediction system, exactly f(p) = 20% of them actually occurred (i.e., were observed). Reliability is assessed with a reliability diagram, which is simply the plot of f(p) against p. In an ideally reliable case, we obtain a curve that is overlaid on the first bisector.

The statistical evaluation of the forecasted PR against the ASOS data was based on a set of performance statistical indicators: the simulated mean (\bar{s}), the Root Mean Square Error (RMSE), the correlation coefficient (R) and the Mean Bias Error (MBE). These metrics are defined in Table 1 in Appendix D.

4. Results and discussions

4.1. Performance evaluation

In this section, we evaluate the two ensembles: the ‘‘Raw Ensemble’’ and the ‘‘Calibrated Ensemble’’ against the MRMS observations. Figure 3 compares the PR measured by MRMS data and the simulated data using the Raw and calibrated ensembles on July 16th, 2022 at 10 pm.

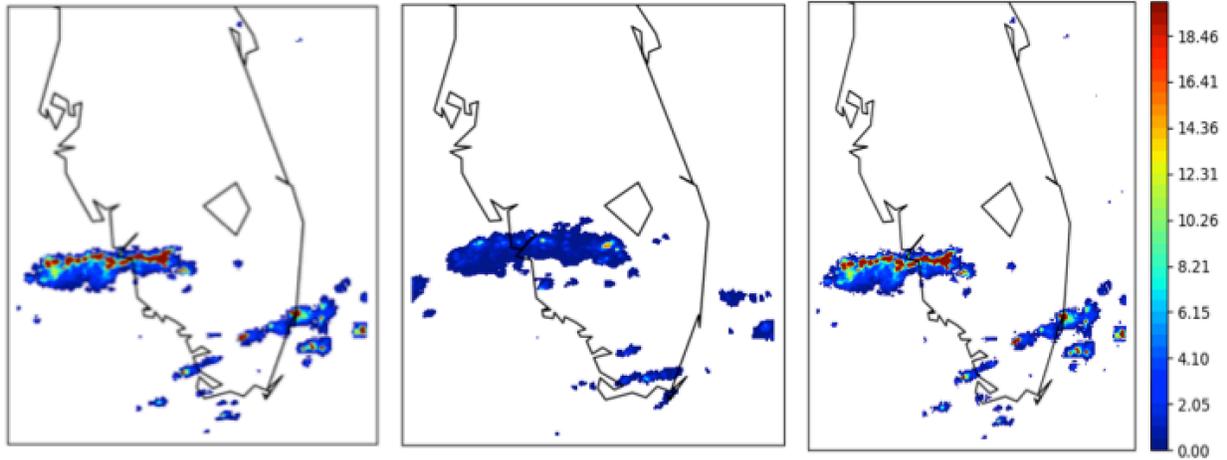


Figure 3: Left panel: PR from MRMS data on July 16th at 10pm. Middle panel: Simulated PR with Raw Ensemble mean at the same time and date as the left panel. Right panel: Simulated PR with Calibrated Ensemble mean at the same time and date as the left panel. The white area represents areas with zero PR.

Figure 3 shows clear discrepancies between the means of the Raw and Calibrated Ensembles. The Raw Ensemble was able to predict the location and timing of the meso- β precipitation system but was not able to reproduce the meso- γ precipitation systems over the south-eastern part of the simulation domain. However, the raw prediction of the PR is underestimated by a factor of 2. In fact, 75 % of the raw ensemble members underestimate the PR mainly because 75% of the simulated members use the Thompson microphysical scheme as the Thompson scheme produces less liquid condensate which results in lower precipitation amount. Similar results were found by Guo et al. 2019 by comparing four MP parameterizations over Eastern China over a six-year summer period (2009-2014). They concluded that the Thompson scheme creates more snow particles than other schemes which produces less graupel and precipitations during warm times. The prediction of PR using the calibrated ensemble substantially improved the PR forecasts as the predicted mean is closer to the MRMS observations. Furthermore, the calibration improved the timing and the location of this simulated precipitation event.

1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60
61
62
63
64
65

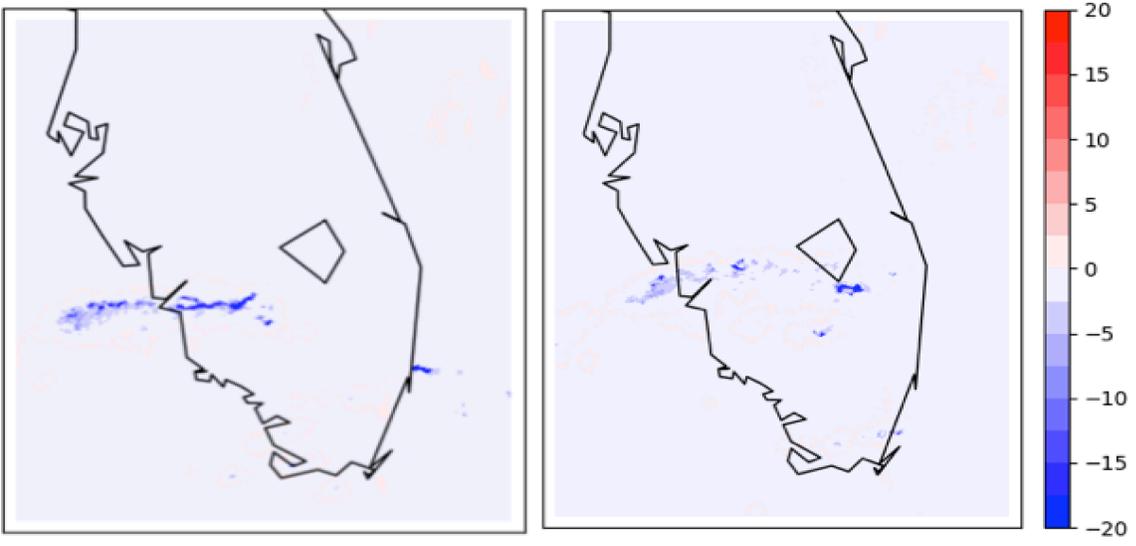


Figure 4: Left panel: Absolute difference (mm.h^{-1}) between the CRPS of the raw ensemble and the calibrated ensemble on July 16 at 10pm. Right panel: Similar to the left panel on July 17th at 00 am.

Figure 4 shows the impact of the calibration of the CRPS of the PR forecasts. The calibration was successful in reducing the CRPS of the calibrated ensemble by a 90 % approximately over the high PR areas, hence improving accuracy relative to MRMS observations. This improvement was guaranteed by the MBM method as it was based on learning the minimization of the CRPS. This is indicative that the weighting coefficients were able to accurately learn temporal features and correct the raw forecasts.

Figure 5 shows the bias of the means of the Raw and Calibrated Ensembles relative to the MRMS data at two different times. The mean of the raw ensemble has a high bias significant over the precipitation areas that can be as high as 20% against the MRMS data. Figure 5 shows also the impact of ensemble calibration on bias and CRPS of the probabilistic forecasts. The calibration had a significant impact over the forecasted PR as the bias of the calibrated mean decreased by 20% relative to the MRMS observed PR.

1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60
61
62
63
64
65

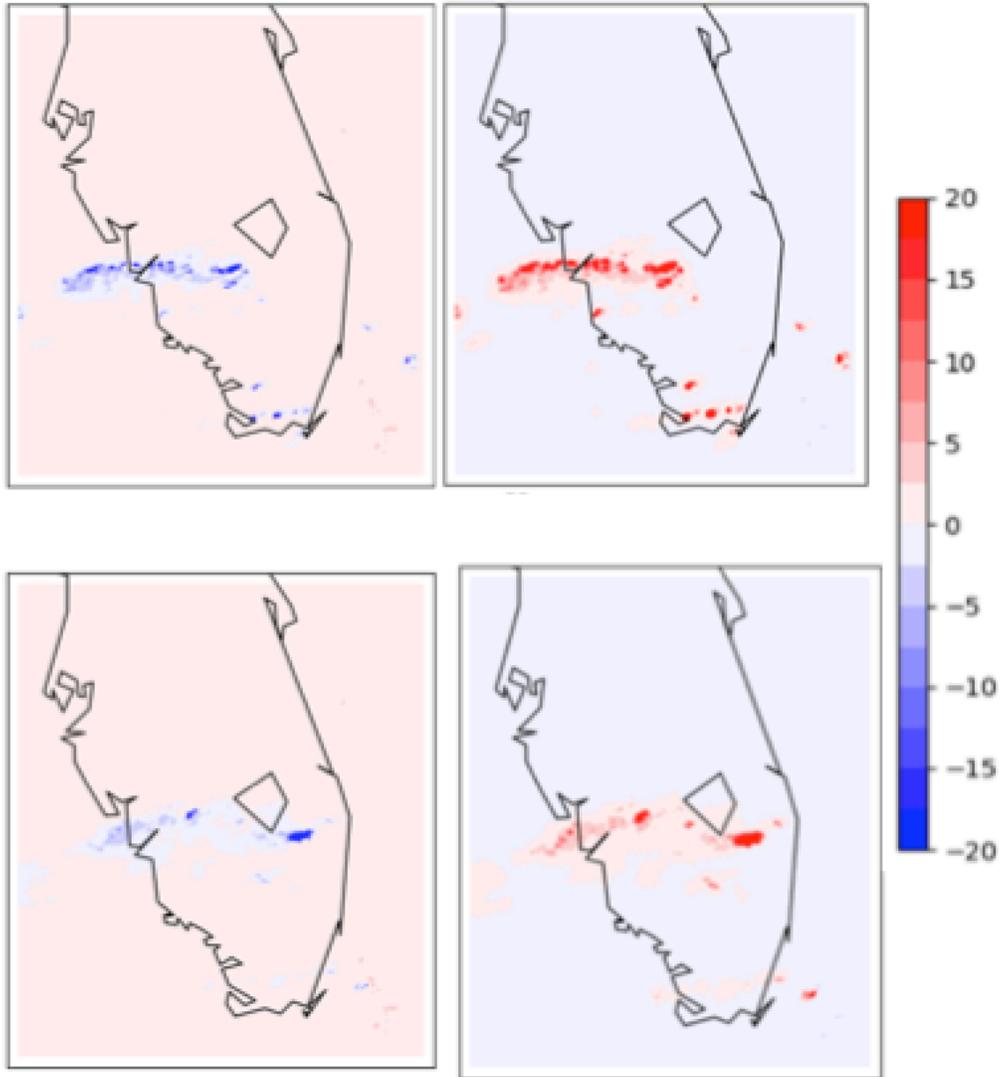
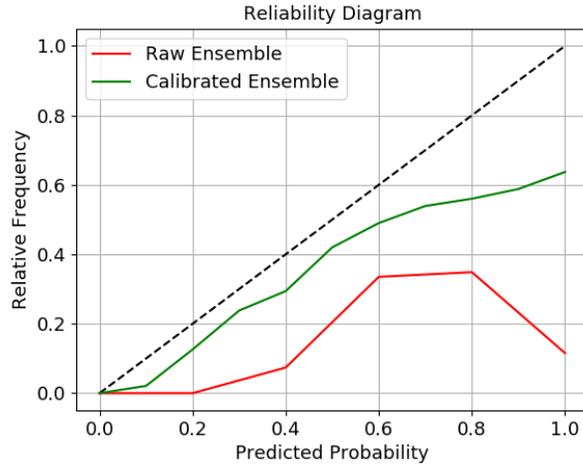


Figure 5: Top left panel: Bias error (mm.h^{-1}) of the Raw Ensemble mean at July 16th 2022 at 10pm . Top right panel: Absolute difference between Bias errors of the Calibrated and Raw ensemble means at 8pm. Bottom left panel: Similar to top left panel at July 17th 00pm . Bottom right panel: Similar to top right panel at July 17th at 00 am.

The reliability diagram of the Raw and Calibrated ensembles are shown in Figure 6. Raw PR forecasts tend to over-forecast both high and low probability events. When considering the calibrated ensemble, the reliability increased for both low and high frequency events. In addition, there is a better reliability for low frequency precipitation events, but the calibrated ensemble is still over forecasting the high-frequency precipitation events. The calibrated ensemble was not able to reproduce the high-frequency event because of biases related to the location and spatial extent of the precipitation events of different scales. The rank-histogram scores of the raw and calibrated ensembles are 15.9 and 4.1 respectively. The rank-histogram score decreased but still

1
2
3
4 more than the optimal score confirming that the calibration improved the spread of the ensemble
5 but still do not have optimal spread in our ensemble.
6
7
8
9



10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27 Figure 6: Reliability diagram of the Raw and Calibrated ensembles over the simulated time and
28 over the precipitation areas of the D2 domain.
29
30
31

32 **4.2.Comparison with ASOS data**

33
34 The calibration is evaluated against the measured PR over the four ASOS stations shown in
35 Figure 7. Table 2 shows the statistical scores of both raw and calibrated means. Tables 1, 2, 3
36 and 4 in Appendix C show the statistical evaluation of PR over the stations PGD, MIA, SRQ and
37 VRB respectively.
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60
61
62
63
64
65

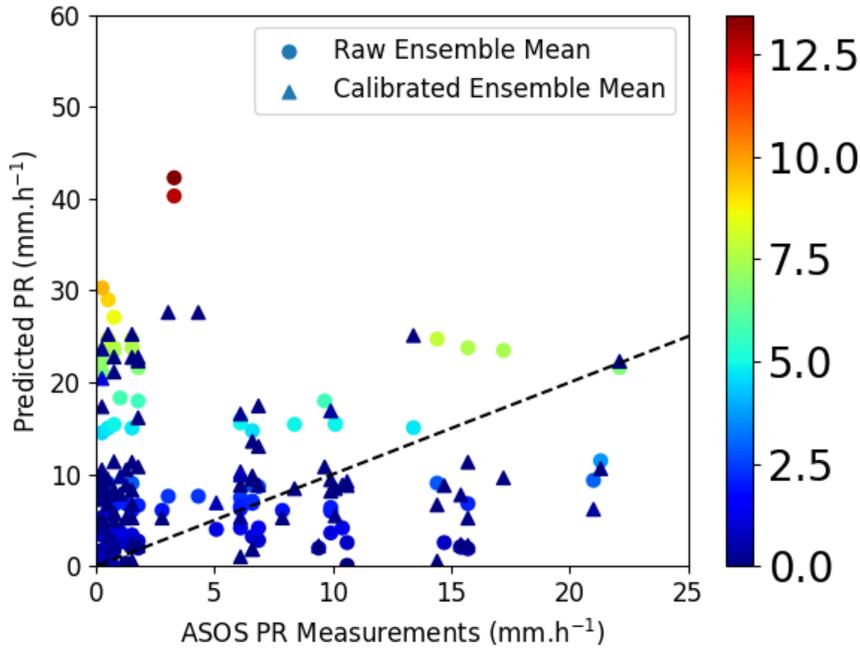


Figure 7: Scatter plot of the simulated PR using the means Raw and Calibrated ensembles. The colors are the uncertainty of the forecasts.

$\bar{o} = 4.82 \text{ mm.h}^{-1}$	$\bar{s} \text{ (mm.h}^{-1}\text{)}$	RMSE (mm.h ⁻¹)	R (%)	MBE (%)
Raw Ensemble Mean	19.07	31.60	16.70	1807.09
Calibrated Ensemble Mean	10.31	11.07	23.15	615.91

Table 2: Statistics of the means of the Raw and Calibrated ensembles against data over the four ASOS.

The scatter plot in Figure 5 shows that both raw and calibrated means overestimate the observed PR over the four ASOS stations with a simulated means of 19.07 and 10.31 mm.h⁻¹ for raw and calibrated ensemble respectively against 4.82 mm.h⁻¹. The slopes of the lines of best fit are 3.87 and 1.57 for the raw and calibrated means respectively. The calibration improved the forecasts as the RMSE decreased from 31.60 mm.h⁻¹ to 11.07 mm.h⁻¹ and the MBE decreased from 1807.09% to 615.91%. The calibrated ensemble still has high bias and significantly overestimates the PR by a factor of 2. This overestimation may be due to the overestimation of PR during summertime by the MRMS data compared to ground based ASOS data because of the evaporation process occurring under the radar beam. In fact, both raw and calibrated PR forecasts overestimate the light precipitation (particularly PR less than 2 mm.h⁻¹ because the MRMS data also overestimates the light precipitations. Similar result was found by Gao et al. 2018 by evaluating the MRMS data against the NEXt generation weather RADar (NEXRAD)

1
2
3
4 data over TEXAS, USA and a dense rain gauge network covering the whole Harris County,
5 Texas, USA. Santer and Grams 2020 evaluated the MRMS QPE and PR during 18-months
6 period relative to rain gauges from the National Centers for Environmental Prediction
7 Meteorological Assimilation Data Ingest System (MADIS) over CONUS and showed that, under
8 warm conditions, a non-negligible systematic overestimation exist because of sub-radar beam
9 evaporation. They also quantified the uncertainty of a MRMS radar measurement based on
10 distance from the radar and partial radar beam blockage. The uncertainty of the forecasted PR
11 was reduced because of the calibration as the uncertainty of the calibrated mean decreased from
12 14 to 4 mm.h⁻¹.
13
14
15
16
17

18 **Conclusion**

19
20
21 In this study, we have applied the MBM calibration technique by minimizing CRPS in order to
22 improve the probabilistic forecasting of precipitation as part of a risk-based approach to integrate
23 UAS into the NAS. The algorithm does not depend on any assumptions on distributions such as
24 gaussianity or uniformity and comes with theoretical guarantee of performance.
25

26
27 The case study examined the impact of ensemble calibration on precipitation forecasts accuracy
28 and uncertainty over South Florida. The MRMS radar data was used to calibrate a 20-members
29 ensemble that was underestimating the PR. This paper showed that CRPS minimization brings
30 improvement on classical scores for the ensemble mean and probabilistic diagnostic tools.
31 Indeed, the forecasting capability measured by classical scores (RMSE, MBE and bias) are
32 improved by the algorithm used during the two simulated summer days. Besides, this spread
33 correction provides a bias correction, improved the reliability of the ensemble and reduced
34 forecasts' uncertainty although the comparison with ASOS data shows a persistent
35 overestimation because of the inherent bias of the MRMS data.
36
37
38

39
40 In addition, the selection of more predictors such as relative humidity, cloud cover and vertical
41 wind velocity may further enhance the skill of probabilistic post-processing for near-real-time
42 precipitation estimates. Besides, using satellite data along with radar data as used here may also
43 improve the evaluation against ground-based validation. The use of deep learning methods such
44 as distributional regression network, Bernstein quantile network and histogram estimation
45 network is a promising as demonstrated in Schulz and Lerch 2022.
46
47

48
49 Future work should investigate the validation of the impact of the calibration and weights on
50 other use cases and the assessment of the performance of the calibrated ensemble over longer
51 lead times and different testing periods. The validation against denser rain gauges network is also
52 necessary as it will show the accuracy of the calibration over off-airport areas which is important
53 for weather-risk assessment and contingency planning during BVLOS operations.
54

55 **Acknowledgments**

56
57 This work was funded by the University of North Dakota.
58

59 The authors declare no conflict of interest.
60
61

Data Availability Statement

The MRMS data used in this paper are publicly available in <https://www.nssl.noaa.gov/projects/mrms/> . The WRF outputs are available upon request from the corresponding author. The code used to calibrate the ensemble is available in the opensource python library available “pythie” here: <https://github.com/Climdyn/pythie>. The ASOS data are publicly available in https://mesonet.agron.iastate.edu/request/download.phtml?network=JP_ASOS.

References

- Automated Surface Observing System : ASOS User's Guide. [Washington, D.C.] :U.S. Dept. of Commerce, National Oceanic and Atmospheric Administration : Federal Aviation Administration : U.S. Navy : U.S. Dept. of the Air Force, 1998.
- Benjamin, S. G., G. A. Grell, J. M. Brown, T. G. Smirnova, and R. Bleck, 2004: Mesoscale weather prediction with the RUC hybrid isentropic-terrain-following coordinate model. *Mon. Wea. Rev.*, 132, 473–494.
- Betts, A., & Miller, M. (1993). The Betts-Miller scheme. In *The representation of cumulus convection in numerical models* (pp. 107–121). American Meteorological Society. https://doi.org/10.1007/978-1-935704-13-3_9
- Biswas, M., Bernardet, L., & Dudhia, J. (2014). Sensitivity of hurricane forecasts to cumulus parameterizations in the HWRF model. *Geophysical Research Letters*, 41, 9113–9119. doi:10.1002/2014GL062071
- Bröcker, J., Smith, L.A., 2007b. Scoring probabilistic forecasts: The importance of being proper. *Weather and Forecasting* 22, 382–388.
- Campbell, S. D., Clark, D. A., Evans, J. E., 2017, Preliminary UAS Weather Research Roadmap, Project Report ATC-438, MIT Lincoln Laboratory, Lexington, MA.
- Candille, G., and O. Talagrand, 2005: Evaluation of probabilistic prediction systems for a scalar variable. *Quart. J. Roy. Meteor. Soc.*, 131, 2131–2150, doi:10.1256/qj.04.71.
- Chen, M.; Nabih, S.; Brauer, N.S.; Gao, S.; Gourley, J.J.; Hong, Z.; Kolar, R.L.; Hong, Y. Can Remote Sensing Technologies Capture the Extreme Precipitation Event and Its Cascading Hydrological Response? A Case Study of Hurricane Harvey Using EF5 Modeling Framework. *Remote Sens.* 2020, 12, 445. <https://doi.org/10.3390/rs12030445>
- Chrit, M.; Majdi, M. Using Objective Analysis for the Assimilation of Satellite-Derived Aerosol Products to Improve PM2.5 Predictions over Europe. *Atmosphere* 2022, 13, 763. <https://doi.org/10.3390/atmos13050763>

1
2
3
4 Chrit, M.; Sartelet, K.; Sciare, J.; Pey, J.; Nicolas, J.B.; Marchand, N.; Freney, E.; Sellegri, K.;
5 Beekmann, M.; Dulac, F. Aerosol sources in the western Mediterranean during summertime: A
6 model-based approach. *Atmos. Chem. Phys.* 2018, 18, 9631–9659.
7
8 Chrit, M.; Sartelet, K.; Sciare, J.; Majdi, M.; Nicolas, J.; Petit, J.E.; Dulac, F. Modeling organic
9 aerosol concentrations and properties during winter 2014 in the northwestern Mediterranean
10 region. *Atmos. Chem. Phys. Discuss.* 2018, 2018, 1–28.
11
12 Dudhia, J., 1989: Numerical study of convection observed during the winter monsoon
13 experiment using a mesoscale two-dimensional model. *J. Atmos. Sci.*, 46 , 3077–3107.
14
15 Ek, M. B., K. E. Mitchell, Y. Lin, P. Grunmann, E. Rogers, G. Gayno, and V. Koren, 2003:
16 Implementation of upgraded Noah land-surface model advances in the National Centers for
17 Environmental Prediction operational mesoscale Eta model. *J. Geophys. Res.*, 108, 8851,
18 doi:10.1029/2002JD003296.
19
20 Fels, S.B.; Schwarzkopf, M.D. An efficient, accurate algorithm for calculating CO₂ 15 micron
21 band cooling rates. *J. Geophys. Res. Ocean* 1981, 86, 1205–1232.
22
23 Gao, S., Zhang, J., Li, D., Jiang, H., and Fang, N. Z., “Evaluation of Multi-Radar Multi-Sensor
24 (MRMS) and Stage IV Gauge-adjusted Quantitative Precipitation Estimate (QPE) During
25 Hurricane Harvey”, vol. 2018, 2018.
26
27 Gebetsberger, M.; Messner, J. W.; Mayr, G. J.; Zeileis, A. (2017) : Estimation methods for non-
28 homogeneous regression models: Minimum continuous ranked probability score vs. maximum
29 likelihood, Working Papers in Economics and Statistics, No. 2017-23, University of Innsbruck,
30 Research Platform Empirical and Experimental Economics (eeecon), Innsbruck
31
32 Gneiting, T., and M. Katzfuss (2014): “Probabilistic Forecasting,” *Annual Review of Statistics*
33 *and Its Application*, 1, 125–151.
34
35 Gneiting T, Raftery AE, Westveld A, Goldman T. 2005. Calibrated probabilistic forecasting
36 using ensemble model output statistics and minimum CRPS estimation. *Mon. Weather Rev.* 133:
37 1098–1118.
38
39 Gneiting T, Raftery AE. 2007. Strictly proper scoring rules, prediction, and estimation. *J. Am.*
40 *Statist. Assoc.* 102: 359–378.
41
42 Hagedorn R, Hamill TM, Whitaker JS. 2008. Probabilistic forecast calibration using ECMWF
43 and GFS ensemble reforecasts. Part I: Two-meter temperatures. *Mon. Weather Rev.* 136: 2608–
44 2619.
45
46 Hong, S-Y., and J-O. J. Lim, 2006: The WRF Single-Moment 6-Class Microphysics Scheme
47 (WSM6). *J. Korean Meteor. Soc.*, 42 , 129–151.
48
49 Janjic, Z. I., 1994: The step-mountain eta coordinate model: Further development of the
50 convection, viscous sublayer, and turbulence closure schemes. *Mon. Wea. Rev.*, 122, 927-945.
51
52
53
54
55
56
57
58
59
60
61
62
63
64
65

- 1
2
3
4 Johnson C, Bowler N. 2009. On the reliability and calibration of ensemble forecasts. *Mon. Weather Rev.* 137: 1717–1720.
- 5
6
7 Kain, J.S., Fritsch, J.M., 1993. Convective parameterization for mesoscale models: the Kain–
8 Fritsch scheme. *The Representation of Cumulus Convection in Numerical Models*, Meteor.
9 Monogr., No. 46. Amer. Meteor. Soc., pp. 165–170
- 10
11
12 Kirstetter, P. E., Y. Hong, J. J. Gourley, S. Chen, Z. Flamig, J. Zhang, M. Schwaller, W.
13 Petersen, and E. Amitai (2012), Toward a framework for systematic error modeling of
14 spaceborne precipitation radar with NOAA/NSSL ground radar-based national mosaic QPE, *J.*
15 *Hydrometeorol.*,13(4), 1285–1300
- 16
17
18 Leutbecher M, Palmer TN. 2008. Ensemble forecasting. *J. Comput. Phys.* 227: 3515–3539.
- 19
20
21 Lin Liu, Chunze Lin, Yongqing Bai, Dengxin He, "Assessing the Effects of Microphysical
22 Scheme on Convective and Stratiform Characteristics in a Mei-Yu Rainfall Combining WRF
23 Simulation and Field Campaign Observations", *Advances in Meteorology*, vol. 2020, Article ID
24 8231320, 17 pages, 2020. <https://doi.org/10.1155/2020/8231320>
- 25
26
27 Li J-L F, Forbes R M, Waliser D E, Stephens G and Lee S W 2014b Characterizing impacts of
28 precipitating snow hydrometeors in the radiation using the ECMWF IFS global model *J.*
29 *Geophys. Res. Atmos.* 119 10981–85
- 30
31
32 Mlawer, E.J.; Taubman, S.J.; Brown, P.D.; Iacono, M.J.; Clough, S.A. Radiative transfer for
33 inhomogeneous atmospheres: RRTM, a validated correlated-k model for the longwave. *J.*
34 *Geophys. Res. Atmos.* 1997, 102, 16663–16682.
- 35
36
37 Noh, Y., W. G. Cheon, S-Y. Hong, and S. Raasch, 2003: Improvement of the K-profile model
38 for the planetary boundary layer based on large eddy simulation data. *Bound.-Layer Meteor.*, 107
39 , 401–427.
- 40
41
42 Pinson P. 2012. Adaptive calibration of (u, v) wind ensemble forecasts. *Q. J. R. Meteorol. Soc.*
43 138: 1273–1284.
- 44
45
46 Raftery, A.E., Gneiting, T., Balabdaoui, F., Polakowski, M., 2005. Using Bayesian model
47 averaging to calibrate forecast ensembles 133, 1,155–1,174.
- 48
49
50 Roulin E, Vannitsem S. 2012. Postprocessing of ensemble precipitation predictions with
51 extended logistic regression based on hindcasts. *Mon. Weather Rev.* 140: 874–888.
- 52
53
54 Santer, H. M. and Grams H. M., Evaluation and Uncertainty of MRMS v12 Dual-polarized
55 Radar Quantitative Precipitation Estimation Product, National Weather Center Research
56 Experience for undergraduates, summer 2020, [https://caps.ou.edu/reu/reu20/finalpapers/Santer-](https://caps.ou.edu/reu/reu20/finalpapers/Santer-finalpaper.pdf)
57 [finalpaper.pdf](https://caps.ou.edu/reu/reu20/finalpapers/Santer-finalpaper.pdf).
- 58
59
60 Schmeits MJ, Kok KJ. 2010. A comparison between raw ensemble output, (modified) Bayesian
61 model averaging, and extended logistic regression using ECMWF ensemble precipitation
62 reforecasts. *Mon. Weather Rev.* 138: 4199–4211.
- 63
64
65

- 1
2
3
4 Schulz, B., and Lerch, S. (2022). Machine Learning Methods for Postprocessing Ensemble
5 Forecasts of Wind Gusts: A Systematic Comparison. *Monthly Weather Review* 150, 1, 235-257,
6 <https://doi.org/10.1175/MWR-D-21-0150.1>
7
8
9 Skamarock, W. C., Klemp, J. B., Dudhia, J., Gill, D. O., Barker, D. M., Wang, W., & Powers, J.
10 G. (2005), A description of the Advanced Research WRF version 2 (No. NCAR/TN-468+ STR).
11 National Center For Atmospheric Research Boulder Co Mesoscale and Microscale Meteorology
12 Div.
13
14
15 Sloughter, J.M., Gneiting, T., Raftery, A.E., 2010. Probabilistic wind speed forecasting using
16 ensembles and bayesian model averaging. *Journal of the American Statistical Association* 105,
17 25–35.
18
19 Sloughter, J.M.L., Raftery, A.E., Gneiting, T., Fraley, C., 2007. Probabilistic quantitative
20 precipitation forecasting using bayesian model averaging. *Monthly Weather Review* 135, 3209–
21 3220.
22
23
24 Talagrand, O., R. Vautard and B. Strauss, 1999, Evaluation of Probabilistic Prediction Systems,
25 in Proceedings of Workshop on Predictability (October 1997), ECMWF, Reading, England, 1-
26 25, available at the address <http://www.ecmwf.int/publications/library/do/references/list/16233>.
27
28
29 Tao, Wei, Joanne Simpson, Deborah Baker, Scott A. Braun, Ming Dah Chou, Brad S. Ferrier,
30 Daniel E. Johnson, Alexander Khain, Stephen E Lang, Barry H. Lynn, Chung-lin Shie, David
31 Starr, C-H. Sui, Yansen Wang and Peter J. Wetzel. “Microphysics, Radiation and Surface
32 Processes in the Goddard Cumulus Ensemble (GCE) Model.” *Meteorology and Atmospheric*
33 *Physics* 82 (2003): 97-137.
34
35
36 Taraphdar, S., & Pauluis, O. M. (2021). Impact of planetary boundary layer and cloud
37 microphysics on the sensitivity of monsoon precipitation using a gray-zone regional model.
38 *Earth and Space Science*, 8, e2020EA001535. <https://doi.org/10.1029/2020EA001535>
39
40
41 Thompson, G., P. R. Field, R. M. Rasmussen, and W. D. Hall, 2008: Explicit forecasts of winter
42 precipitation using an improved bulk microphysics scheme. Part II: Implementation of a new
43 snow parameterization. *Mon. Wea. Rev.*, 136 , 5095–5115.
44
45
46 Thorarinsdottir, T.L., Gneiting, T., 2010. Probabilistic forecasts of wind speed: ensemble model
47 output statistics by using heteroscedastic censored regression. *Journal of the Royal Statistical*
48 *Society: Series A (Statistics in Society)* 173, 371–388.
49
50
51 Thorey, J., Chaussin, C., Mallet, V.. Ensemble forecast of photovoltaic power with online CRPS
52 learning. *International Journal of Forecasting*, Elsevier, 2018, 34 (4), pp.762-773.
53 [ff10.1016/j.ijforecast.2018.05.007ff](https://doi.org/10.1016/j.ijforecast.2018.05.007ff).
54
55
56 Van Schaeybroeck, B., and S. Vannitsem, 2015: Ensemble post-processing using member-by-
57 member approaches: Theoretical aspects. *Quart. J. Roy. Meteor. Soc.*, 141, 807–818,
58 <https://doi.org/10.1002/qj.2397>.
59
60
61
62
63
64
65

1
2
3
4 Van Schaeybroeck, B., and S. Vannitsem, 2011: Post-processing through linear regression.
5 Nonlinear Processes Geophys., 18, 147–160, <https://doi.org/10.5194/npg-18-147-2011>.

6
7 Vannitsem S. 2009. A unified linear Model Output Statistics scheme for both deterministic and
8 ensemble forecasts. Q. J. R. Meteorol. Soc. 135: 1801–1815.

9
10 Vannitsem S, Hagedorn R. 2011. Ensemble forecast post-processing over Belgium: Comparison
11 of deterministic-like and ensemble regression methods. Meteorol. Appl. 18: 94–104.

12
13 Vitart, F., Anderson, J.L., Sirutis, J., Tuleya, R.E.: Sensitivity of tropical storms simulated by a
14 general circulation model to changes in cumulus parameterization. Quart. J. Roy. Meteor. Soc.
15 127, 25–51 (2001)

16
17
18 Wilks, D.S., 2009. Extending logistic regression to provide full-probability-distribution MOS
19 forecasts. Meteorological Applications 16, 361–368.

20
21 Zhang, J., Howard, K., Langston, C., Vasiloff, S., Kaney, B., Arthur, A., Van Cooten, S.,
22 Kelleher, K., Kitzmiller, D., Ding, F., Seo, D., Wells, E., & Dempsey, C. (2011). National
23 Mosaic and Multi-Sensor QPE (NMQ) System: Description, Results, and Future Plans, Bulletin
24 of the American Meteorological Society, 92(10), 1321-1338.
25
26
27
28
29
30
31

32 Appendix

33 34 35 Appendix A

36
37
38
39

40 Model parameter	41 Used configuration
42 Model and domains	
43 Model version	44 ARWv4.0 (Skarmarock et al. 2008)
45 Time step	46 Adaptative time step (36 s for D1)
47 Map projection	48 Lambert
49 Pressure top	50 50 hPa
51 Vertical levels	52 80 (*)
53 Time integration scheme	54 Third order Runge-Kutta scheme
55 Time integration scheme for 56 acoustic and gravity-wave 57 modes	58 Second order scheme

59
60
61

Horizontal/vertical advection	Fifth order upwind
Scalar advection	Positive definite
Upper-level damping (for vertical propagating gravity waves)	Rayleigh damping
Computational horizontal diffusion	6th-order numerical diffusion
Forecast period	60 h (from July 15 th , 2018 at 12 pm UTC to July 18 th , 2021 at 12 am UTC)

Table [1]: WRF model configuration and input physics parameterizations. * η levels are 1, 0.99938147, 0.9918859506, 0.9860143, 0.9835575, 0.97480931, 0.9691238, 0.95061912, 0.938789424, 0.91847208, 0.89114445, 0.87771024, 0.8344125, 0.807124586, 0.76820505, 0.71652851, 0.6848121, 0.615978875, 0.5720332, 0.5472062, 0.5233661, 0.5004734, 0.4784906, 0.4573815, 0.4371113, 0.4176468, 0.3989559, 0.3810079, 0.3637731, 0.3472234, 0.3313315, 0.316071, 0.3014172, 0.2873457, 0.2738335, 0.2608584, 0.2483989, 0.2364347, 0.2249459, 0.2139138, 0.2033201, 0.1931475, 0.1833792, 0.173999, 0.1649918, 0.1563425, 0.1480369, 0.1400615, 0.132403, 0.1250489, 0.1179871, 0.111206, 0.1046944, 0.09844154, 0.09243726, 0.08667168, 0.08113512, 0.07581868, 0.07071351, 0.06581128, 0.06110381, 0.0565835, 0.05224282, 0.04807468, 0.04407217, 0.04022875, 0.0365381, 0.03299413, 0.02959097, 0.02632311, 0.0231851, 0.02017184, 0.01727832, 0.0144998, 0.01183172, 0.00926967, 0.006809457, 0.004447003, 0.002178475, 0.

Appendix B

Station ID	Latitude(°N)	Longitude(°W)	Height ASL (m)
VRB	27.6556	80.4179	8.00
PGD	26.9172	81.9914	8.00
MIA	25.7880	80.3169	4.00
SRQ	27.4014	82.5586	9.00

Table 1: List of the four ASOS stations in South Florida and their corresponding latitude, longitude and above sea-level (ASL) height.

Appendix C

$\bar{o}=5.02 \text{ mm.h}^{-1}$	$\bar{s} \text{ (mm.h}^{-1}\text{)}$	RMSE (mm.h ⁻¹)	R (%)	MBE (%)
Raw Ensemble Mean	29.71	53.92	14.49	2469.24

Calibrated Ensemble Mean	10.77	14.24	20.81	575.37
---------------------------------	-------	-------	-------	--------

Table 1: Statistics of the raw and calibrated means over the PGD ASOS station

$\bar{o} = 4.34 \text{ mm.h}^{-1}$	$\bar{s} \text{ (mm.h}^{-1}\text{)}$	RMSE (mm.h ⁻¹)	R (%)	MBE (%)
Raw Ensemble Mean	14.57	16.91	4.96	1023.87
Calibrated Ensemble Mean	13.79	15.09	8.39	940.71

Table 2: Statistics of the raw and calibrated means over the MIA ASOS station

$\bar{o} = 2.54 \text{ mm.h}^{-1}$	$\bar{s} \text{ (mm.h}^{-1}\text{)}$	RMSE (mm.h ⁻¹)	R (%)	MBE (%)
Raw Ensemble Mean	15.57	17.07	8.30	1302.87
Calibrated Ensemble Mean	8.60	9.77	29.36	605.77

Table 3: Statistics of the raw and calibrated means over the SRQ ASOS station

$\bar{o} = 7.39 \text{ mm.h}^{-1}$	$\bar{s} \text{ (mm.h}^{-1}\text{)}$	RMSE (mm.h ⁻¹)	R (%)	MBE (%)
Raw Ensemble Mean	16.43	15.61	19.72	904.30
Calibrated Ensemble Mean	8.06	7.04	30.06	30.06

Table 4: Statistics of the raw and calibrated means over the VRB ASOS station

Appendix D

Statistical indicator	Definition
\bar{s}	$\sqrt{\frac{1}{n} \sum_{i=1}^n s_i}$
\bar{o}	$\sqrt{\frac{1}{n} \sum_{i=1}^n o_i}$
RMSE	$\sqrt{\frac{1}{n} \sum_{i=1}^n (c_i - o_i)^2}$

1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60
61
62
63
64
65

Correlation	$\frac{\sum_{i=1}^n (s_i - \bar{s})(o_i - \bar{o})}{\sqrt{\sum_{i=1}^n (s_i - \bar{s})^2} \sqrt{\sum_{i=1}^n (o_i - \bar{o})^2}}$
MB	$\frac{1}{n} \sum_{i=1}^n (c_i - o_i)$

Table 1: Definition of the statistics used in this work. o_i and s_i are the observed and simulated wind speeds at time i . n is the number of data.