

1                   **Accurate Simulation of Both Sensitivity and Variability for**  
2                   **Amazonian Photosynthesis: Too Much to Ask**

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9                   **Key Points:**

- 10                   • Regression logic is cause to doubt predictions whose variability is unrealistically  
11                   high.
- 12                   • A suite of models poorly reproduce tower estimates of Amazon rainforest gross pri-  
13                   mary productivity.
- 14                   • Highly seasonal models predict stronger GPP reactivity to meteorology than is likely  
15                   to be true.

16 **Abstract**

17 Causes of climate predictions' uncertainty include wide spread in modeled gross primary  
18 productivity (GPP) for evergreen broadleaf forests. Deterministic predictions inherently lack  
19 the portion of variability that a regression's error term summarizes. Omitted predictors' con-  
20 tribution to error represent simulations' necessary underestimation of real variability. Earth  
21 system model outputs with high variability relative to reference data warrant skeptical ex-  
22 amination. We compare three statistical and 15 process models to site-level means, seasonal  
23 amplitude and driver responsiveness of GPP as calculated at six Amazon eddy covariance  
24 (EC) towers. Current month's weather determines only 12% of the variability in EC GPP,  
25 implying that models whose predicted GPP's variability approaches that of EC GPP probably  
26 are substantially hypersensitive to weather drivers. Roughly half the models have stronger  
27 seasonal GPP variability than ECs show, and inaccurately identify the timing of annual min-  
28 imum GPP. Responses to temperature and light for some highly seasonal models are of the  
29 opposite sign as EC GPP's. Strongly seasonal models' deepest dip in photosynthesis both  
30 occurs later in the dry season and is more severe than EC estimates. Excessive reactivity to  
31 drivers appears to cause the high simulated variability of the strongly seasonal models.

32 **1 Key Words**

33 Amazon

34 model benchmarking

35 gross primary productivity (GPP)

36 Multi-Scale Synthesis and Terrestrial Model Intercomparison Project (MsTMIP)

37 outliers

38 tropical rainforest

39 regression

40 seasonality

41 Simple Biosphere Model (SiB)

42 variability

## 2 Plain Language Summary

Global climate models must accurately represent many processes, including the pace at which plants convert sunlight, water and CO<sub>2</sub> into sugar. The Amazon rainforest is enormous and extremely biologically productive, so the region strongly influences the world's cycling of CO<sub>2</sub>. Measurements from instruments on towers in the Amazon, despite imperfections, seem to be the most accurate estimates of rainforest plant productivity rates that exist. We compare "tower" estimates to 18 global models, focusing on rainforests' subtle dry v. wet seasons.

Modeled monthly plant productivity poorly matches tower estimates. About half the models have more seasonal variation than the towers and half have less. Simple equations that use current month's temperature, light, and rainfall describe model output quite closely, especially for the weakly seasonal models.

Reality is more variable than are mathematical model predictions that accurately describe the results of a particular change in inputs. Why, then, do some models have stronger seasonal swings than tower estimates? One cause is using a descriptive equation that overlooks models' non-linear responses to weather. But the main reason is that when weather changes, tower estimates of plant productivity change less than it do predictions from models with strong seasonal swings.

## 3 Introduction

Modeling tropical plant productivity accurately is important to the accuracy of global climate predictions because rainforests are so large and productive that their gross primary productivity (GPP) represents about 34% of the terrestrial total [Beer *et al.*, 2010]. Rainforest productivity largely drives interannual variability in global CO<sub>2</sub> concentrations [Bousquet *et al.*, 2000; Rödenbeck *et al.*, 2003; Wenzel *et al.*, 2014]. Positive feedbacks to change in rainforest productivity amplify change in CO<sub>2</sub> [Christoffersen *et al.*, 2014; Harper *et al.*, 2014; Zemp *et al.*, 2017].

The spread in simulations of current rainforest productivity [Ardö, 2015; Malhi *et al.*, 2009] is wide, and typically greater than for other biomes [Anav *et al.*, 2015; Beer *et al.*, 2010; Cavalieri *et al.*, 2015; Friedlingstein *et al.*, 2006; Jung *et al.*, 2020; Mystakidis *et al.*, 2016], shows of the need for better modeling for tropical rain forests. Cross-model differences represent material uncertainty about future rainforest productivity. Large differences in

74 tropical GPP persist even after removing model differences in simulated precipitation [*Malhi*  
75 *et al.*, 2009; *Poulter et al.*, 2010a].

### 76 **3.1 Predictions tend to have lower variance than source data.**

77 For assessing a non-stochastic model's responses to climate change, it is helpful to  
78 consider a trade-off between accurate responsiveness to drivers and accurate variance of pre-  
79 dicted outcomes. A model's sensitivity, or responsiveness, can be characterized as marginal  
80 change in outcome per unit change in a predictor [*Friedlingstein et al.*, 2006; *Hamby*, 1994].  
81 In a regression, responsiveness corresponds to the slope coefficients. For models of GPP,  
82 large slopes imply stronger responses to changing climate. Tropical GPP sensitivity is crit-  
83 ical because it describes the extent to which climate will continue to alter rainforest activity  
84 and even its viability.

85 In a model, a simplification by definition, omitted drivers cause some of any outcome's  
86 real variability. Predictions cannot include the variability that this portion of random error  
87 contributes because the model has no information about the missing drivers. For a regres-  
88 sion, if modeled responses to included drivers, or sensitivities, are accurate and other sources  
89 of model error modest, omitted variables will still make variability of the predictions unreal-  
90 istically low. We label the inherent tendency for predicted outcomes to have lower variability  
91 than true outcomes as "flattening".

92 An idealized illustration explains why predictions have low variability. Posit a regres-  
93 sion that predicts its outcome  $y_i$  as a response to one predictor,  $x_i$ . The model is perfectly ac-  
94 curate, meaning that its responsiveness,  $\beta$ , exactly equals the true responsiveness of  $y_i$  to  $x_i$ .  
95 The model's only source of error is omitted predictors, which account for the random error  
96 term  $\epsilon_i$ . There is no uncertainty due to imperfections of measurement, sampling, or specifi-  
97 cation. Variance of the predictions from this nearly perfect regression is necessarily smaller  
98 than the source data's variance.

$$\text{sampled real world observation} : y_i = \beta x_i + \epsilon \quad (1)$$

$$\text{regression prediction} : \hat{y}_i = \beta x_i$$

99 The error term adds variance only to the source data but not to predicted values (Equa-  
100 tion 1, [*Greene*, 2012, Chapter 3]). Text S1 expresses the argument formally.

101 The regression could describe for a particular model the responsiveness of rainforest  
102 GPP to temperature.

$$\widehat{GPP}_i = \beta * temperature_i + intercept \quad (2)$$

103 If the GPP model is based on enzyme kinetics, there is no internal  $\beta$  for temperature,  
104 but instead a variety of other calculations involving temperature. For example, one param-  
105 eter could be Q10, the exponent for a rate multiplier to rubisco carboxylation per increase  
106 of 10°C. The value of Q10 may be derived from bench or field research. The parameter is  
107 not estimated directly for rainforest due to lack of source data, and because in theory Q10 is  
108 constant for all chlorophyll [but see *Alster et al., 2020*]. Other steps within the model may  
109 further affect GPP's temperature responsiveness. The temperature to which Q10 is applied  
110 may be modified from ambient to account for degree of shading. There may be adjustments  
111 for each plant functional type's (PFT's) optimum temperature. Temperature may affect GPP  
112 indirectly through vapor pressure deficit. The descriptive regression summarizes as a linear  
113 approximation the effects of the process model's more complex underlying calculations.

114 As a thought experiment, assume that the net result of imperfections in the GPP model  
115 is that effective rainforest temperature sensitivity, or the descriptive regression's only  $\beta$ , is  
116 twice the true sensitivity. A consequence is higher variance of the  $\hat{y}_i$ 's. As shown in Equa-  
117 tion 3, exaggerating the regression coefficient by a factor of two quadruples the predictions'  
118 variance.

$$\hat{y} = 2\beta x + intercept \quad (3)$$

$$\sigma_{\hat{y}}^2 = \frac{1}{n-1} \sum_{n=1}^i (2\beta x_i - \overline{2\beta x})^2 = \frac{4}{n-1} \sum_{n=1}^i (\beta x_i)^2$$

119 More generally, when the slope of a regression with one predictor changes by a multi-  
120 plied constant, the variance of predictions made from the same  $x$ -values used to fit the model  
121 changes by the squared amount of the multiplier. A slope multiplier larger than one increases  
122 predictions' variance, and a multiplier smaller than one decreases variance.

123 If only because models by definition simplify reality, every regression model has a  
124 non-zero error term. While error due to omitted variables embodies real-world variability,

125 measurement error, sampling error, and misspecification of mathematical form represent the  
 126 data and model's uncertainty about included aspects of the real world [Vicari *et al.*, 2007].  
 127 More statistical noise from any source increases  $\epsilon$  and causes more flattening, or less vari-  
 128 ability of predicted outputs.

129 There is a trade off between accurate responsiveness of  $\hat{y}_i$  to  $x_i$  and accurate variance  
 130 of predicted outcomes. Either deliberately or inadvertently, changing the regression slope  
 131 can adjust variance to any level including to equal observed variance. But deviations from  
 132 the optimal regression fit to the data carry a cost of less accurate modeled responsiveness  
 133 to change in the observed driver(s). The right answer as measured by accurate variability of  
 134 outcomes may result from the wrong reason of excessive model sensitivity.

135 A model can predict only what it "knows" about. Numeric calculations simulate the  
 136 processes for which equations are included, and the consequences of the influences for which  
 137 driver data are provided. If the model's sensitivities are accurate, outputs have only as much  
 138 variability as the included processes and drivers create. The maximum portion of true out-  
 139 come variance that a model whose responsiveness to drivers is perfectly accurate can simu-  
 140 late is the portion that the included processes and drivers in fact determine. The more com-  
 141 pletely a model with accurate driver responsiveness includes all true determinants of its out-  
 142 come, the larger and closer to correct will be its predictions' variance.

143 The variability that model errors contribute and that otherwise is missing from deter-  
 144 ministic predictions can be added back in directly. If random statistical noise is added, then  
 145 predictions can have both realistic variance and accurate reactivity to predictors. If instead  
 146 the introduced noise is correlated with drivers, such as by drawing predictions from a prob-  
 147 ability distribution at calculated percentiles, then effective driver slope(s) will be altered as  
 148 described in Equation 3. Stochastic modeling has computational and other complications,  
 149 and is rare in full earth system models (ESMs).

### 150 **3.2 Flattening applies to Earth System Models.**

151 Flattening occurs within parameterized ESM calculations. Many hard-coded  
 152 model parameters are "essentially a smaller model within the larger model" [Dahan,  
 153 2010]. For example, in the Community Land Model (CLM5.0) each PFT's stomatal  
 154 resistance parameter originated in a regression fitted to a global database of conductances  
 155 [[https://escomp.github.io/ctsm-docs/doc/build/html/tech\\_note/index.html](https://escomp.github.io/ctsm-docs/doc/build/html/tech_note/index.html) section 2.9.3,

156 Table 2.9.1 of values from *De Kauwe et al., 2015*]. Real variability in resistances caused  
 157 by variables omitted from the source equation and subsequently from the ESM remains  
 158 unexplained and unmodeled.

159 Climate models' inner workings are decisively more intricate than a single linear re-  
 160 gression. But flattening is a tendency of any deterministic numeric prediction, including  
 161 non-linear equations, transformations of variables (Text S1) and, like entire ESMs, com-  
 162 plex combinations of equations with feedbacks. Inclusion of processes may be indirect, such  
 163 as a model forced with satellite data that is a proxy for deciduous leaves' annual cycle. The  
 164 driver data itself may be simulated, as weather is in fully-coupled ESM runs. In all of these  
 165 situations, predictions from the model will lack the portion of real variance determined by  
 166 omitted drivers and processes.

167 ESMs simulate systems so complex that omitted processes and drivers loom large. Be-  
 168 cause significant determinants of real variability are missing, the connection between the  
 169 accuracy of predictions' variance and the accuracy of driver sensitivity provides a diagnos-  
 170 tic tool. If the variance of a predicted outcome is higher than or even close to a benchmark's  
 171 variance, offsetting excessive sensitivity to drivers could be a cause.

### 172 **3.3 The Amazon is likely to become warmer, with more variable rainfall.**

173 Change in tropical GPP as represented in ESMs depends largely on four environmental  
 174 drivers: ambient CO<sub>2</sub>, precipitation, temperature, and top of canopy insolation. Prediction  
 175 accuracy depends on correctly simulating rainforest GPP responses to changes in the forc-  
 176 ings. Like the rest of the world, rainforests are experiencing consistently increasing ambient  
 177 CO<sub>2</sub>. Models concur that the region's precipitation will become more variable [*Bathiany*  
 178 *et al., 2018; Chadwick et al., 2015; Feng et al., 2013*], a trend for which there already are  
 179 observational indications at least on the drying side [*Fu et al., 2013; Gloor et al., 2013; Li*  
 180 *et al., 2008; Lopes et al., 2016*]. The direction of change in rainforest mean rainfall rather  
 181 than in its increasing variability is uncertain, however [*Gloor et al., 2012; Li et al., 2006;*  
 182 *Poulter et al., 2010a*].

183 Rainforest temperature is projected to rise and may already have changed measurably  
 184 [*Corlett, 2011; Jiménez-Muñoz et al., 2013*]. Temperature increases are likely to vary re-  
 185 gionally within the basin [*Gloor et al., 2012*]. Deforestation and temperature may have an  
 186 amplifying feedback on GPP at landscape scales, with lower plant productivity causing more

187 sunlight to become sensible rather than latent heat, and higher temperatures further reducing  
188 photosynthesis rates.

189 In the sometimes-cloudy tropics light can limit photosynthesis, especially for lower  
190 leaves in thick canopies. Amazonian surface insolation has increased slightly in recent  
191 decades due to reduced cloudiness [*Barkhordarian et al.*, 2017], though the trend's robust-  
192 ness is unclear [*Wielicki et al.*, 2002]. One likely cause of increased surface light, seasonal  
193 changes in the sizes of both Pacific and Atlantic Ocean warm pools, has an uncertain anthro-  
194 pogenic signal [*Arias et al.*, 2011]. Pollution and tropical fires on the other hand, reduce top  
195 of canopy insolation. Both reflect human behavior, an influence whose uncertainty increases  
196 the spread in trend predictions for most weather parameters.

197 This paper explores the fidelity of modeled Amazonian GPP to eddy covariance (EC)  
198 flux tower data, with an emphasis on the accuracy trade-offs that flattening presents. Credible  
199 representations of responsiveness to change in weather are especially important for climate  
200 models because they describe the direction and strength of trends in GPP in response to in-  
201 creasing concentrations of greenhouse gasses. If responsiveness is too weak, forecasts will  
202 be unreasonably reassuring. Excessively strong weather responsiveness will predict faster  
203 and more dramatic dieback of the rainforest [*Cox et al.*, 2013].

#### 204 **4 Methods**

205 We compare EC GPP to 15 process models and three statistical models. Methods sum-  
206 marized in this section are described further in Text S2. The statistical models, Fluxcom,  
207 Wecann and VPM, each have fared well in global accuracy intercomparisons. SG3 runs for  
208 Multi-scale synthesis and Terrestrial Model Intercomparison Project's [MsTMIP; *Huntzinger*  
209 *et al.*, 2014; *Wei et al.*, 2014] 14 process models have common initial land cover maps, land  
210 use and land cover change, spin-up procedures, and atmospheric CO<sub>2</sub> and weather inputs. To  
211 MsTMIP we added SiB4 [*Haynes et al.*, 2019a,b], a recent major revision to the participating  
212 SiB3 model and which now has prognostic phenology.

213 To compare process models to statistical models that rely on satellite data, the evalua-  
214 tion period is limited to 2000 - 2010. Wecann starts in 2007 with the earliest satellite dataset  
215 for solar-induced fluorescence. Wecann is included only in basinwide comparisons because  
216 its shorter period for approximating seasonal cycles at tower sites would increase Wecann's  
217 apparent variability compared to all other models.

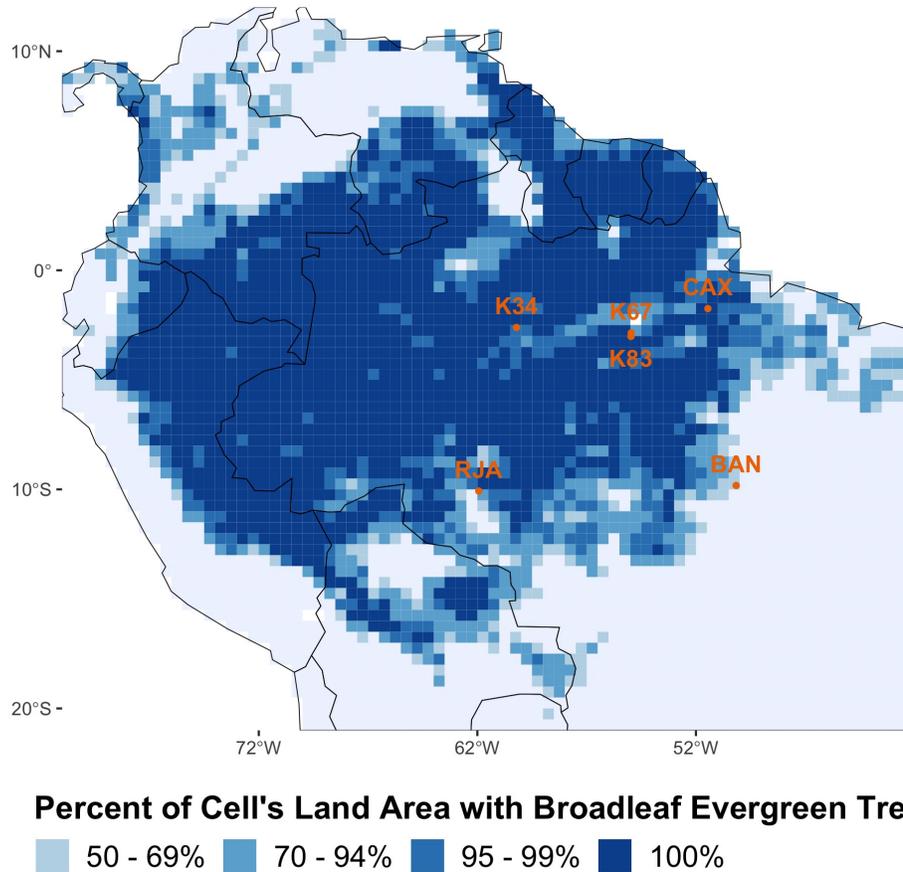


Figure 1: Study area, which includes all shaded cells. Orange dots are locations of eddy covariance towers. Based on MsTMIP's PFT classifications, a high proportion of the study area is almost pure rainforest. The land PFT distribution in 86% of study cells is at least 90% evergreen broadleaf forest.

218 For basinwide comparisons, the study area is northern South America, the world's  
 219 largest rainforest that we refer as the Amazon although we do not use a strict watershed  
 220 boundary. The study cells are limited to 42° - 81°W and 12°N - 21°S, excluding Central  
 221 America both north of 7°N and west of 77.5°W. Selecting grid cells whose MsTMIP  
 222 tiled PFTs are at least 50% evergreen broadleaf forest (EBF) limits the study to rainforest  
 223 vegetation. Fig. 1 shows the portion of each study cell that MsTMIP codes as EBF.

224 We compare the process and statistical models to six EBF eddy covariance sites from  
 225 the Large-scale Biosphere-Atmosphere Experiment in Amazonia: Rio Javaés-Bananal  
 226 (BAN), Caxiuanã (CAX), Manaus Kilometer 34 (K34), Tapajos Kilometer 67 (K67), Tapajos

227 Kilometer 83 (K83), and Reserva Jaru (RJA) [Restrepo-Coupe *et al.*, 2013]. Fig. 1 maps  
 228 the towers' locations. Each process or statistical model's GPP estimates for the grid cell  
 229 containing a site are matched to the months for which data exists at each tower.

230 Eddy covariance towers are flawed benchmarks for GPP. Measured net ecosystem ex-  
 231 change is a small residual whose much larger offsetting components of GPP and ecosystem  
 232 respiration must be modeled. A particularly thorny issue is lack of closure in energy budgets.  
 233 Calculated energy fluxes leaving a site do not equal measured energy entering [da Rocha  
 234 *et al.*, 2009; Jung *et al.*, 2019; von Randow *et al.*, 2004]. Where GPP seasonal variation is  
 235 smaller than average, as in the tropics, closure corrections introduce more noise [Clark *et al.*,  
 236 2017; Tramontana *et al.*, 2016]. These weaknesses are serious. Nevertheless, and partly on  
 237 faith, we take tower estimates to be the best reference data available, and their GPP respon-  
 238 siveness to individual drivers as true to the extent of being qualitatively strongly positive,  
 239 strongly negative, or weak.

240 We define 'site' as an eddy covariance location. 'EC' refers more specifically to mea-  
 241 surements made at a tower site and their derivatives. Unless noted, the weather driver data  
 242 used to assess modeled GPP's responsiveness is MsTMIP's. Precipitation is monthly total in  
 243 mm. Light is monthly mean top of canopy short-wave radiation under all sky conditions, in  
 244  $\text{Wm}^{-2}$ . Mean monthly temperature is measured in  $^{\circ}\text{C}$ , and GPP in  $\text{gCm}^{-2}\text{d}^{-1}$ .

## 245 **5 Results**

### 246 **5.1 Modeled GPP mean and variance grossly differ from EC estimates.**

247 An optimistic hypothesis that each model's simulated GPP mean and variance match  
 248 EC estimates is easily rejected. The lines in Fig. 2 enclosing the EC mean and variance mark  
 249 wide 99<sup>th</sup> percentile bootstrapped confidence intervals. Averaged across the six sites and all  
 250 months of each tower's operation, nearly all model estimates are outside the ECs' confidence  
 251 intervals. The EC variance reflects only calculated mean monthly GPP, however, and does  
 252 not include the considerable additional uncertainty from EC modeling and measurements.

253 For individual sites, at least one model severely underestimates mean GPP and at least  
 254 one model's mean is more than twice EC GPP (Fig. S7). On average one or two models'  
 255 means are credible matches to EC data. The variance of two process models' simulated GPP  
 256 is higher than EC variance for all sites, while one is too low for every site. No model's vari-  
 257 ance is within target range for every site. One statistical model is credible for five of six sites.

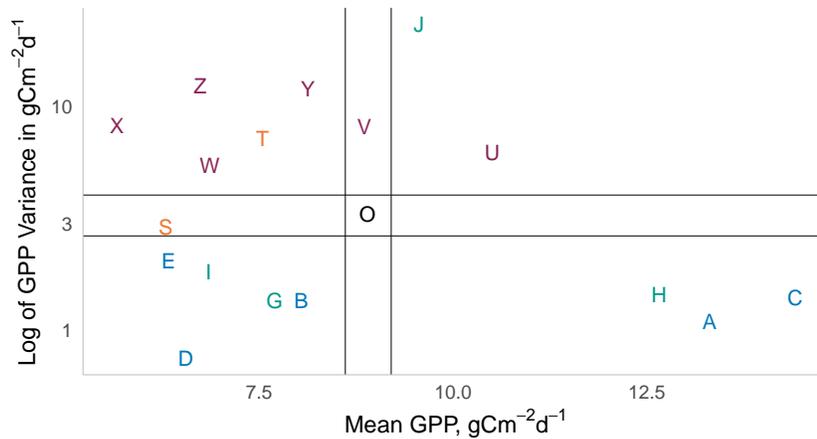


Figure 2: Comparison of EC GPP to process and statistical model means and variances across all site-months. Variance, on the y-axis, has a  $\log_{10}$  scale. Model "O" is EC GPP. Names of other lettered models are shown in Fig. 3. Lines bracketing EC estimates are 99<sup>th</sup> percentile confidence bounds. For all but two models both means and variances fall outside the confidence bounds.

258 Which models are outliers differs across sites, and some models have very different relative  
 259 performance among sites. For example, model J's variance is an outlier in Fig. 2 due to its  
 260 puzzling GPP close to zero for several grid cells near K67, while its GPP is well above aver-  
 261 age at most other sites.

262 Correlations with EC GPP are remarkably low. Individual models' average across all  
 263 sites ranges from -0.16 to 0.45, with a grand mean of only 0.12 (range across sites: -0.32 -  
 264 0.66, Fig. S8). Overall correlations with EC GPP for four models are statistically indistin-  
 265 guishable from zero. CAX and K67 have especially weak matches, with negative correla-  
 266 tions for 12 and 14 of the 17 models respectively. The most closely simulated site is RJA,  
 267 where EC GPP is especially variable and average correlation across all models is 0.66.

268 Data with larger magnitudes tend to have larger variance than data with smaller mag-  
 269 nitudes, which tends to make GPP variance of mildly responsive models lower for strongly  
 270 responsive models. But most models in Fig. 2 show the opposite pattern, with higher vari-  
 271 ance and lower mean GPP than the EC towers or the opposite. Means with large absolute  
 272 values do not correspond to large variances. The opposing tendencies mean that for these  
 273 GPP models, whatever causes differences in means does not explain variability. The causes  
 274 of differences in variance need to be considered directly. Based on the logic of flattening, we

275 hypothesize that high variance models may be overly sensitive to drivers. Before an explo-  
276 ration of model responsiveness, in the next section a more robust descriptor of model vari-  
277 ability is assigned.

## 278 **5.2 Seasonal cycle amplitude characterizes a model's GPP variability.**

279 To explore the connection between flattening and sensitivity, and to compare mod-  
280 els to EC GPP based on the relative variance of their predictions, one possibility is to rank  
281 the models based on variances in EC estimates. An alternative metric of variability that is  
282 more closely related to the biology being simulated is the amplitude of a model's seasonal  
283 cycle. Unlike a variance calculation, it takes into account the sequencing of observations. A  
284 Fourier series approximation smooths a site's GPP across outliers, uneven numbers of obser-  
285 vation years per month and missing data. Earth's annual insolation cycle is sinusoidal, giving  
286 Fourier transformations inherent good fit for some ecological cycles. Characterizing seasonal  
287 cycles of GPP with four pairs of Fourier terms is a compromise between overfitting versus  
288 forcing unrealistic simplification. The first pair can be thought of as creating an annual cy-  
289 cle, the second allows for asymmetric shoulders, and the third and fourth provide for limited  
290 shaping of the annual peak and trough.

291 We label the difference between maximum and minimum months of a site's mean an-  
292 nual Fourier cycle as seasonal amplitude. In Fig. 3 and elsewhere, models are listed in in-  
293 creasing order of their seasonal amplitude averaged across the EC sites and indicated with  
294 black dots. Colors indicate how a model's amplitude compares to the EC amplitude, both av-  
295 eraged across all sites. The nine 'mild' models with weaker mean seasonal cycles than ECs  
296 are shown in blue or green. The eight 'lively' models whose cycles are stronger are colored  
297 red or orange. The intensity distinctions for dark blue or red break at one standard deviation  
298 from the EC mean. Most notable is how widely the seasonal swings differ across models, by  
299 a factor of 8.2. The difference means roughly that model Z's simulated trees vary in produc-  
300 tivity eight times as much during a year as do model A's. The only model that would switch  
301 between the categories of mild versus lively if rankings were determined by variance instead  
302 is Model J, whose very high variance (Fig. 2) is due to anomalous GPP at one site.

303 ECs are a benchmark for model seasonality, albeit one with arguable accuracy. The  
304 mildest model varies during the year a third (0.35) as much as does EC GPP. The most  
305 strongly responsive model's mean site amplitude is almost triple (2.9 times) that of the ECs.

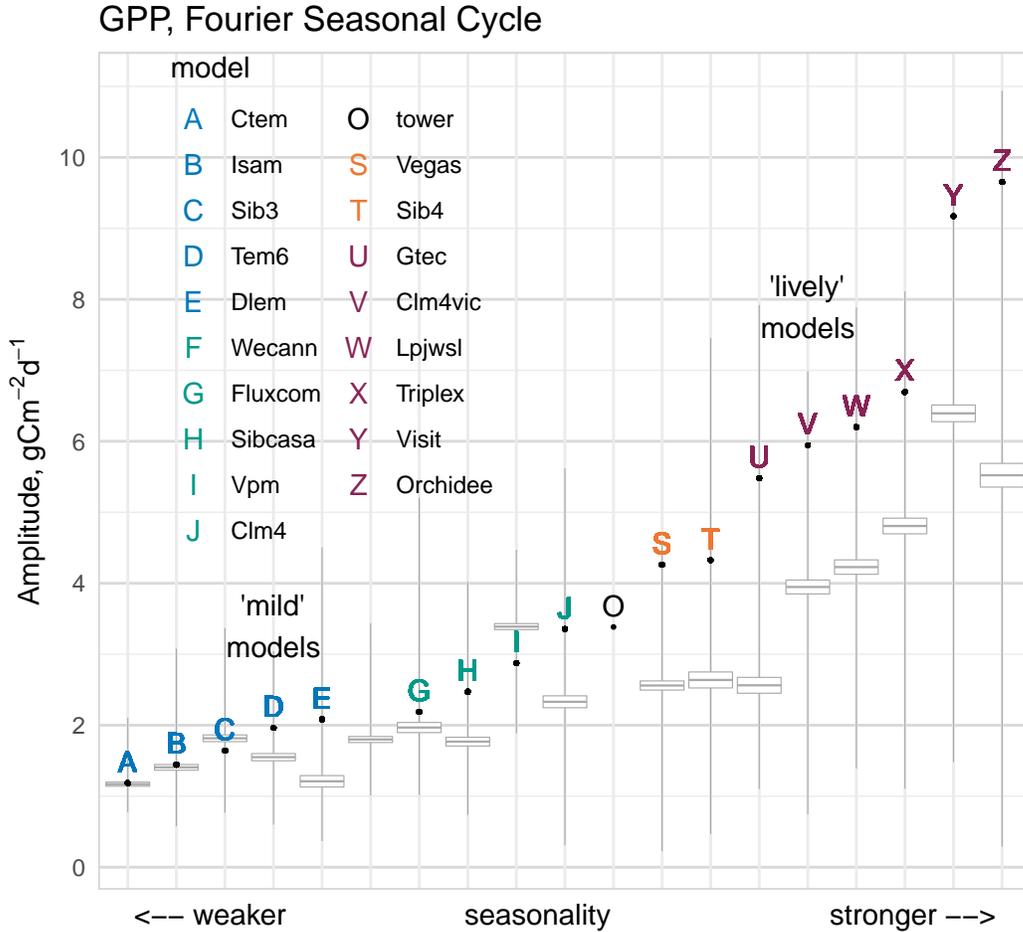


Figure 3: GPP’s seasonal amplitude, with models’ means across six sites ranked on the x-axis and marked by black dots labeled with colored letters. Grey boxes show a model’s amplitude tendency across all rainforest cells in the Amazon. For both sites and basin-wide, the range in seasonal amplitude across models approaches an order of magnitude.

306 The models’ degree of seasonality differs strongly also at individual sites. At no site is a  
 307 mild model’s mean amplitude larger than four  $\text{gCm}^{-2}\text{d}^{-1}$  (Fig. S9), while few of the most  
 308 responsive models, shown in red, have a seasonal amplitude below four  $\text{gCm}^{-2}\text{d}^{-1}$  at any  
 309 site. As climate parameters that affect productivity shift over time, a lively model is likely to  
 310 predict greater change in rainforest carbon fixation per unit area than a mild model.

311 Each mean amplitude summarizes only six data points, so the mean EC GPP, 3.4  
 312  $\text{gCm}^{-2}\text{d}^{-1}$  has large uncertainty. Based on a t-test, only the liveliest model’s amplitude is  
 313 outside a 95% confidence interval around the EC mean. K34 and K67 are located close

314 enough to each other to have somewhat similar climate. With four degrees of freedom rather  
 315 than five to reflect possible pseudoreplication, no model is outside the credible interval.  
 316 There are too few data points to tighten the confidence interval by bootstrapping.

317 ECs' differing periods of operation preclude a temporally exact comparison between a  
 318 model's basinwide and site tendencies. Fig. 3 shows in grey a box plot of a 95% confidence  
 319 interval around the mean seasonal amplitude for all Amazon rainforest cells in all months.  
 320 Whiskers on the grey boxes represent the tenth and ninetieth percentiles of cell amplitudes.  
 321 The 20% of each model's cells that are outliers are not shown. As explained in the methods  
 322 section, there is no EC mean for model F, whose ranking is an approximation. For the most  
 323 mild models, shown in blue, EC amplitudes are roughly representative of the entire basin.  
 324 For all the strongly lively red models but one, EC sites have moderately stronger seasonal-  
 325 ity than do basin-wide means. The extent to which the ECs are typical of the Amazon as a  
 326 whole decreases with model seasonal amplitude.

327 Comparing seasonal amplitudes to interannual variability (Fig. S10) reinforces how  
 328 strong the consequences of seasonal cycles are for simulated GPP. The difference between  
 329 highest and lowest year's mean GPP from 2000 to 2010 for individual models ranges from  
 330 0.1 to 1.2  $\text{gCm}^{-2}\text{d}^{-1}$ . The models' mean basin-level seasonal amplitude ranges are several  
 331 times larger, from 1.3 to 6.0  $\text{gCm}^{-2}\text{d}^{-1}$ . The range in seasonal cycle amplitudes across mod-  
 332 els, 1.2 to 9.7  $\text{gCm}^{-2}\text{d}^{-1}$ , approximately equals the grand mean of monthly GPP across all  
 333 models, 8.9. For GPP in the Amazon, understanding what drives variation within a year ex-  
 334 plains much more about a model's tendencies than do determinants of its interannual vari-  
 335 ability.

### 336 **5.3 EC GPP barely responds to current weather.**

337 Compared to many temperate locations, rainforests are always moist, always warm and  
 338 always green. But even the wettest tropical forests have subtle annual cycles in rain, temper-  
 339 ature and light. Our null hypothesis is that a simple linear combination of these three drivers  
 340 largely describes monthly mean GPP. If so, weather's cycles might logically also set the tim-  
 341 ing of the simulated annual GPP cycle. Importantly, the drivers' individual influences on  
 342 GPP could be parsed and evaluated [Hamby, 1994]. With differing trends expected for each  
 343 driver, a model with retrospectively credible responsiveness to each is more likely to predict  
 344 reliably.

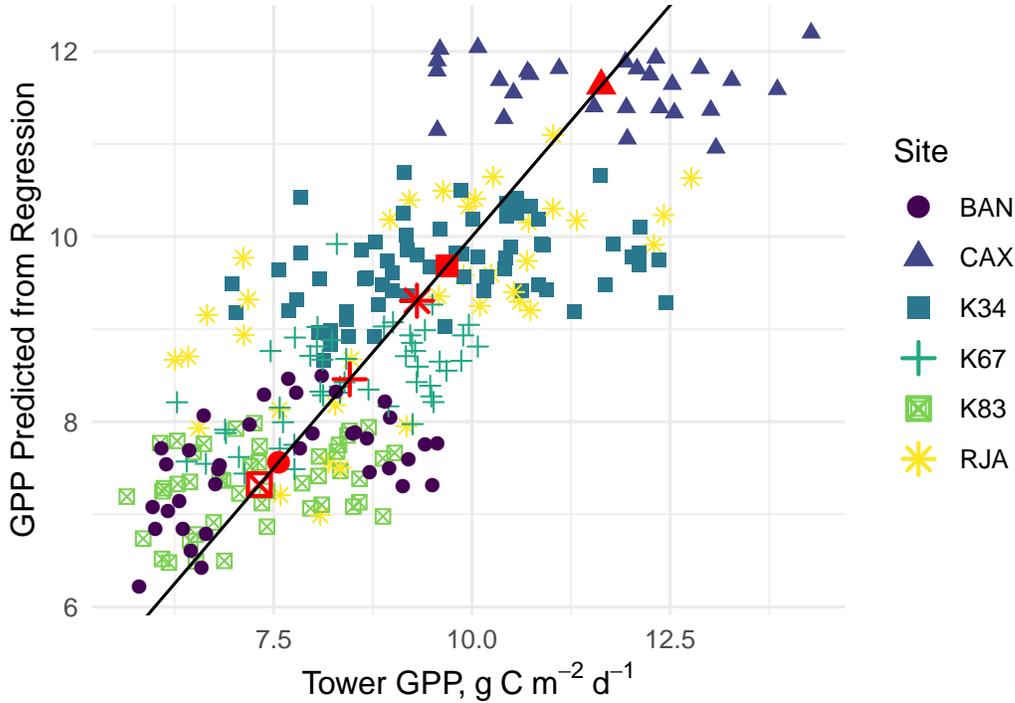


Figure 4: Monthly GPP for each eddy covariance tower compared to paired values predicted from a regression on MsTMIP rain, light, and temperature with site intercepts. Individual intercepts force each site’s mean predicted value to equal the mean EC value, indicated with red symbols. Differences between sites are the main source of the prediction’s power, with minimal EC GPP responsiveness to weather.

345 To judge models’ responsiveness requires EC benchmarks. A linear regression with  
 346 only current month’s rain, temperature and light explains a small 12% of the variability in  
 347 EC GPP (not shown). Rain’s coefficient but no other is statistically significant.

348 GPP varies substantially between sites (Fig. 4). Individual cell intercepts, or fixed ef-  
 349 fects, segregate out the undetermined sources of location-specific differences that cause a  
 350 particular site’s outcome to differ by a consistent increment over time. For process models of  
 351 GPP, potential underlying causes of site differences include soil depth and fertility, species  
 352 assemblage in the spectacularly diverse tropics, herbivory, tree age distribution that reflects  
 353 disturbance history, local geology that affects flooding and subsurface hydrology, and others.  
 354 Most of these causes for site differences are impossible to parameterize globally, and there-  
 355 fore from an ESM perspective constitute impenetrable statistical noise. Site intercepts repre-

356 sent differences between locations that may or may not be modeled accurately, and if not, are  
 357 apt to involve omitted variables. But separate intercepts allow a focus on how consistently  
 358 the three weather drivers cause GPP to change at all locations even if site characteristics sig-  
 359 nificantly influence long-term baseline productivity.

360 Equation 4 predicts monthly mean EC GPP from MsTMIP weather, with site intercepts  
 361 added.

$$GPP = Intercepts + 0.0054 * Rain + 0.019 * Light + 0.52 * Temperature \quad (4)$$

$$P\text{-values} : Rain = 0.00; Light = 0.01; Temperature = 0.00$$

$$Adjusted R^2 = 0.59; Residual standard error = 1.2; n = 260$$

362 Fig. 4 compares EC GPP on the x-axis to paired predictions from Equation 4 on the  
 363 y-axis. Site intercepts account for most of the spread in the EC dataset. Due to the weak  
 364 predictive power of the weather variables, paired values for individual sites in Fig. 4 do not  
 365 otherwise cluster near the 1:1 black line of perfect prediction. As measured by the adjusted  
 366  $r^2$ , the regression terms determine 59% of the variability. Site-level differences account for  
 367 81% of the regression's predictive power [Chevan and Sutherland, 1991], leaving 19% of  
 368 explained variability, or only 12% of total variability, predicted by the current month's envi-  
 369 ronmental attributes. Contrary to the initial hypothesis, current weather has little influence  
 370 on EC GPP.

371 Each site's predicted values are flattened, with less variability on the y-axis than the  
 372 source values on the x-axis. While the EC GPP data points have a variance of 3.3, the vari-  
 373 ance of matched but flattened predictions is 2.0, or 61% as large.

374 If EC meteorology for the comparable cell is used rather than MsTMIP weather, the  
 375 regression fit with site effects degrades slightly ( $r^2 = 0.54$ ). The only terms whose p-value  
 376 is  $\leq .10$  are four site intercepts and the slope for rain. GPP is statistically unrelated to ei-  
 377 ther light or temperature. Why site-specific weather should be less predictive than regional  
 378 weather is unclear. Perhaps soil moisture is influential, reflects regional recharge, and over-  
 379 whelms highly localized rainfall differences. Also, conceivably a geographically broader  
 380 weather summary more accurately represents conditions across ECs' full footprints than does  
 381 weather at point locations chosen to represent the upwind area. The unexpectedly weaker fit  
 382 with site weather is convenient, however. Errors in representing the true values of the drivers

383 cause attenuation bias in regression coefficients, or weaker sensitivity. If MsTMIP weather  
 384 were a worse fit than site weather, comparisons of EC driver sensitivities to model sensitivi-  
 385 ties would be less straightforward.

386 **5.4 Lively models respond more strongly to current weather than mild models.**

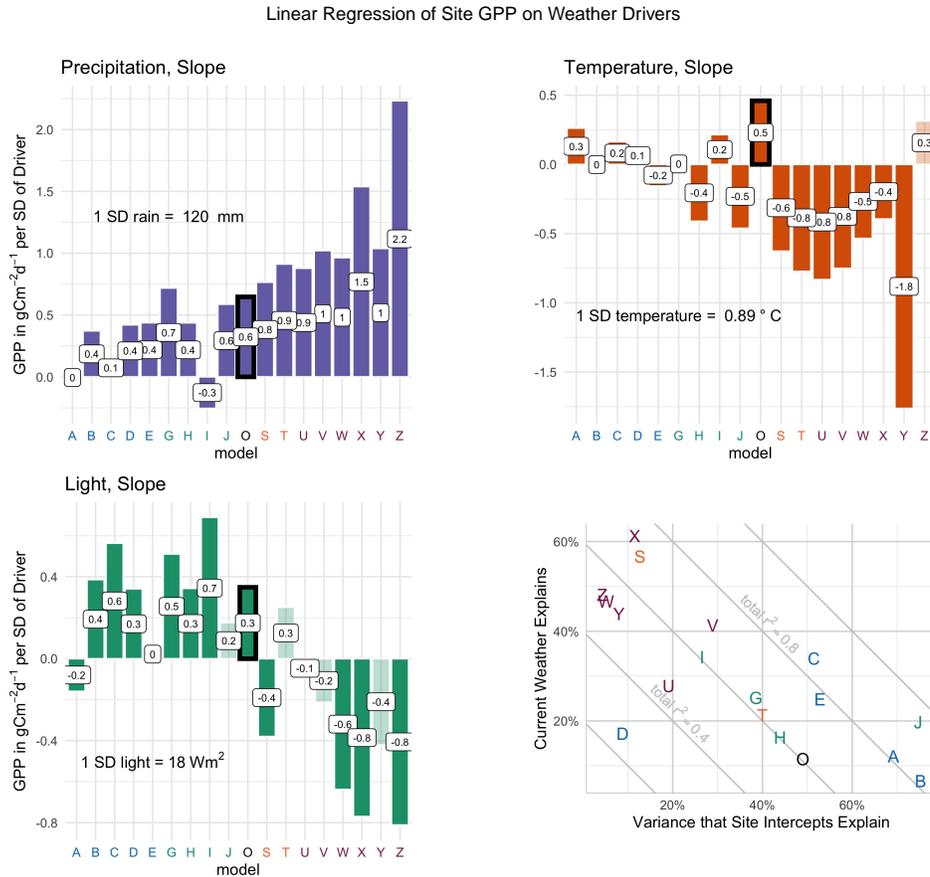


Figure 5: Linear regression slopes of GPP on current month’s temperature, radiation, and precipitation for each model across six sites. Drivers are in units of standard deviation across all sites and months. Pale bars are coefficients whose p-value exceeds .05. EC slopes are outlined in black. The lower right panel shows the residual standard error and adjusted  $r^2$  for each regression. Slopes, or driver responsiveness, range among models across an order of magnitude.

387 Whether flattening is as strong an influence on a model’s simulations as on EC predic-  
 388 tions depends on (a) whether current weather describes similarly little of the model’s GPP  
 389 variability, and (b) how reasonably a linear sum describes the mathematical form of current

390 weather's relationship to modeled GPP. The three driver variables included in Eq. 4 describe  
391 EC GPP's weak but statistically significant responses to each. Based on the Akaike Infor-  
392 mation Criterion fit, all three drivers should be retained. All would be kept even if the full  
393 combination were less than ideal, however, in order to explore next each model's potentially  
394 differing emphases among the weather elements.

395 The vigor with which some models respond to rain, light and temperature contrasts  
396 sharply with EC GPP's weak responses. The regressions whose responsiveness slopes appear  
397 in Fig. 5 are simple additive models with site-specific intercepts, parallel to the characteri-  
398 zation above of EC GPP. To facilitate comparison across drivers, environmental variables  
399 are shown in units of standard deviations. A regression slope of +0.5 in Fig. 5 indicates a  
400 tendency for a half  $\text{gCm}^{-2}\text{d}^{-1}$  increase in GPP to result from an increase of one standard de-  
401 viation in the driver. Among the models, statistically significant coefficients for rain range  
402 from -0.26 to 2.2, for temperature from -1.8 to 0.26, and for light from -0.81 to 0.69. Am-  
403 bient  $\text{CO}_2$  is an insignificant predictor of historical site GPP in nearly all models (Fig. S11)  
404 and is not included in Fig. 5's regressions.

405 Among the GPP drivers, rain is the most consistent predictor across models. Its sign  
406 is positive for all but one model, and its influence statistically significant ( $p \leq 0.05$ ) for all but  
407 one. The magnitude of rain responsiveness varies substantially, from 0.1 to  $2.2 \text{ gCm}^{-2}\text{d}^{-1}$  of  
408 GPP per 120 mm increase in a month's precipitation. GPP increases with rain in all but 2  
409 models. Models' rank for rain slopes almost matches that of seasonal amplitudes. The ECs'  
410 responsiveness to rain, outlined in black, sits solidly in the middle. For rain, the main differ-  
411 ence among models is the response strength.

412 No environmental variable has a monopoly on GPP. Mean absolute slopes differ by  
413 less than a factor of two: 0.75 for rain, 0.46 for temperature, and 0.40 for light. Responses to  
414 temperature differ more than they do to rain. Temperature is statistically insignificant for four  
415 models. For the mildest models, temperature's influence is positive, as it is for EC GPP, and  
416 very weak. For the liveliest, GPP strongly declines. For light, statistically unreliable slopes  
417 exist across the range of slope magnitudes. On average mild models have nearly as strong a  
418 GPP response to light as do lively models but light's effect is in the opposite direction. For  
419 temperature and light, there is as much disagreement between models in what direction GPP  
420 responds as how strongly.

421 The descriptive regressions largely characterize GPP for all models. For the mild mod-  
422 els as a group, residual standard error (RSE) averages 8% of site GPP and mean  $r^2$  is 0.69  
423 (Fig. S12). For the lively models, average RSE is 24% and mean  $r^2$  is 0.58. GPP and current  
424 weather have an even closer linear connection for seasonally mild models than for livelier  
425 models.

426 Differences in mean cell GPP, or among intercepts, are substantial and influential. Ex-  
427 cept for one model, the range in intercepts as a percent of mean site GPP is 12 - 55%. The  
428 exception, model J, has intercepts that vary by more than 100% of mean GPP due to one  
429 outlier site. Given that site intercepts are largely a proxy for omitted variables, it is not sur-  
430 prising that they are relatively less influential on process models than on EC GPP. The wide  
431 range in site GPP compared to relatively modest slopes for driver values means that site  
432 means constitute most of the descriptive regression's explanatory ability, as they do for the  
433 ECs.

434 As shown in Fig. 5's lower right panel, for mild models site intercepts explain more  
435 of GPP's variance (mean = 49%) than does weather (mean = 21%). For lively models, site  
436 intercepts explain a smaller share (mean = 16%) than does weather (mean = 43%)

437 A model's responses to drivers in the six cells with eddy covariance towers are gen-  
438 erally similar to its responses across the Amazon (Figs. S11 and S12), suggesting that as-  
439 sessing model responses for the Amazon by comparing them with EC estimates of GPP is a  
440 reasonable application of scarce benchmarking data. But the similar mean tendencies smooth  
441 across considerable spatial differences (Fig. S13). For percent of variance that a simple re-  
442 gression explains, the most striking spatial pattern is that for almost every model there are  
443 areas where a linear relationship of current environmental conditions plus site intercepts de-  
444 scribes change in GPP almost completely, and other places where it explains little.

445 Weather's stronger influence on responsive model variability compared to the mild  
446 models is consistent with lively models' generally steeper slopes for each weather driver. The  
447 principle of flattening suggested, and the data summarized in Fig. 5 confirm, that models  
448 with high variance, the strongly seasonal models, are on average overly responsive to drivers.

## 449 **5.5 Models' differing rainforest non-linearities are not benchmarked.**

450 Compared to regressions that characterize reality, those that describe model output are  
 451 unusually clean. All of the modeled values for the study period and area typically are acces-  
 452 sible, yielding a census with no sampling errors nor errors in measuring outputs. Some or all  
 453 of the inputs to the model's calculations may also be known exactly, as is MsTMIP weather.  
 454 Only two sources of stochastic noise remain in the regressions that describe modeled GPP:  
 455 omitted drivers and misspecification. This section considers alternative specifications, first  
 456 interactions between weather drivers then non-linear responses.

457 Adding interaction terms to the regression that describes EC GPP minimally increases  
 458 its total explanatory power, while diluting evidence of individual environmental drivers'  
 459 influence. The same predictors as Equation 4 were used plus all possible crosses for the  
 460 three weather drivers: rain times temperature, etc. The resulting regression has 5% more ex-  
 461 planatory power than the model without interactions ( $r^2 = 0.62$ ). But no single or combined  
 462 weather driver has a significant slope.

463 Unfortunately, there are too few months of noisy EC data to resolve a non-linear re-  
 464 gression form or assess the accuracy of flex points. It is possible that daily GPP estimates  
 465 from the towers would better resolve non-linear responsiveness, or might share with monthly  
 466 means having too much random uncertainty. Fig. 6 highlights differences in model responses  
 467 to extreme temperatures. On the x-axis monthly mean temperatures are grouped by deciles  
 468 basinwide. GPP on the y-axis is displayed as z-scores to remove model differences in mean  
 469 and variability for each cell. What remains is the degree to which a model's response to ex-  
 470 tremes of temperature are anomalous compared to its responses to currently more typical  
 471 temperatures.

472 The descriptive regressions with simple linear forms of the drivers do reliably indicate  
 473 mean tendencies across the range of driver values for the EC months sampled. But as climate  
 474 shifts, the models' responsiveness to more extreme values will be most relevant. For most  
 475 models, responses have particularly strong deviation at high temperature from their mean re-  
 476 sponsiveness. GPP rises continuously with temperature in about a third of models. The rest  
 477 eventually flex into declining plant productivity. Most of the mild models simulate that rain-  
 478 forest is more productive in the warmest of months, while lively models reach the opposite  
 479 conclusion. Mild models H and J resemble the lively models in this respect, simulating their  
 480 very lowest GPP at peak temperatures.

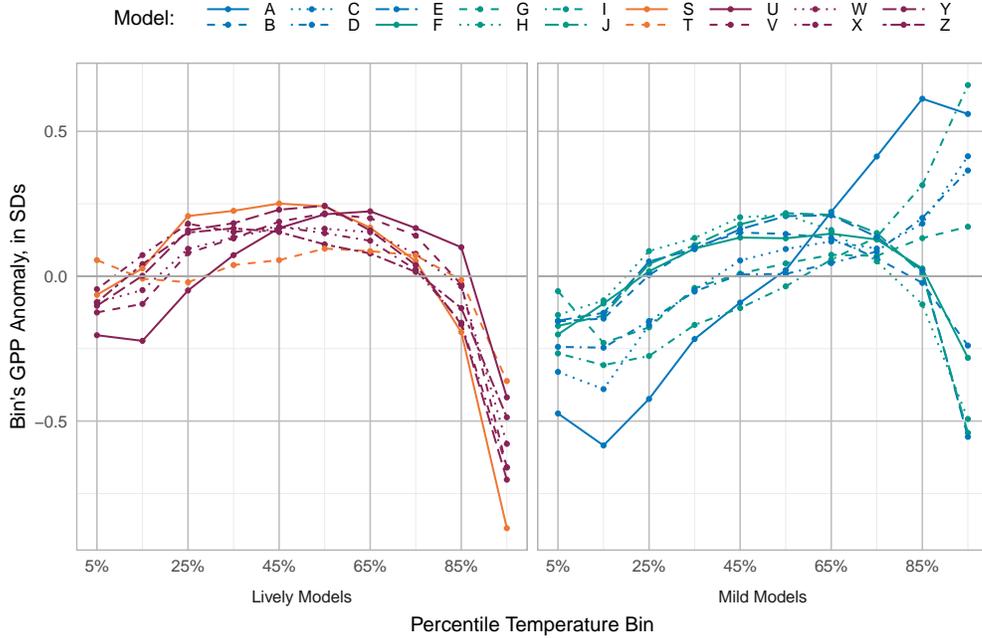


Figure 6: Non-linearity in modeled cell-level GPP responses to temperature across the Amazon basin. Monthly cell average temperature, on each panel's x-axis, is binned by deciles for all cells and months. GPP is scaled as the number of standard deviations from a particular model and cell's mean. Mild models are in the left panel and lively models in the right. At high temperature, GPP falls markedly in lively models, while the response of mild models varies widely.

481 Models' responses to rain and light also are non-linear (Text S4). Models with the  
 482 strongest, steepest GPP response to increasing rain tend to have only modest response to in-  
 483 creasing light and vice versa. Most lively models respond strongly to rain but have below-  
 484 average GPP in the brightest months. In contrast, most of the mild models simulate their  
 485 highest rainforest GPP with typical, middle decile rain amounts, and below-average GPP in  
 486 the wettest months.

487 **5.6 Lively models simulate strong, rapid drops in dry season GPP.**

488 Responsiveness is a critical characteristic of GPP models because it describes how the  
 489 model represents the consequences of climate change. An overly responsive model will pre-  
 490 dict more change than is realistic. The phase, or timing, of modeled GPP's seasonality is a  
 491 corroborative assessment of responsiveness' accuracy that can be benchmarked. Seasonal  
 492 timing may also suggest which model processes cause any mismatches. A slightly more re-

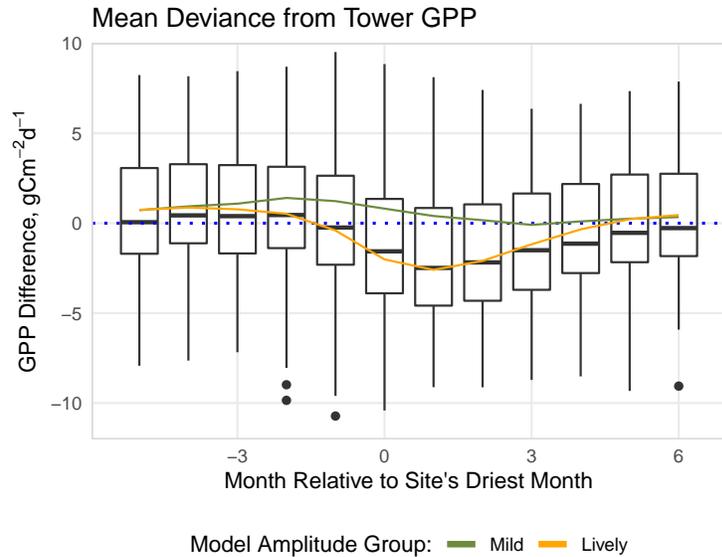


Figure 7: Seasonal deviations of GPP from EC estimates for models grouped by seasonal amplitude. On the y-axis, zero represents an exact match to EC GPP. Green and orange lines are mean deviances for mild and lively models respectively. Zero on the x-axis is the month at each EC tower with lowest average rainfall. Boxes show tendencies of all models as a group compared to ECs, with crossbars marking overall median deviations. Divergences are especially large for lively models in the latter half of the dry season.

493 cent version of Model J illustrates, for example, that an ESM's seasonal GPP timing can be  
 494 correct in most of the world but have major inaccuracies in the world's wettest regions [*Col-*  
 495 *lier et al.*, 2018, Fig. 5d].

496 One way to characterize seasonal timing focuses on the month of lowest GPP, as show-  
 497 ing when the modeled forest experiences its greatest stress. Across all sites, the month with  
 498 lowest modeled GPP is on average 2.6 months different from EC estimates (Fig. S3). Every  
 499 model matches at least one site's time of minimum EC GPP to within one month. But with  
 500 one exception, every model also is at least 5 months different from the EC estimate at one or  
 501 more sites, or essentially has an opposite seasonal cycle.

502 Fig. 7 summarizes seasonal timing tendencies for models generalized by responsive-  
 503 ness group. Zero on the figure's x-axis is each site's long-term average driest month, with  
 504 one month before the driest month shown as -1, two months after as +2, etc. Boxes for each  
 505 month enclose the 25<sup>th</sup> and 75<sup>th</sup> percentiles of all models' GPP deviances from the EC esti-

506 mates. Dots indicate outlier models. In months whose median value of the box plot is above  
507 zero, most models at most sites simulate higher GPP than EC estimates. Taller boxes late in  
508 the dry season show when the widest spread among models occurs.

509 Mild models on average, shown with the green line in Fig. 7, have relatively little sys-  
510 tematic difference from EC estimates over the course of a year. Mildness is defined as a  
511 dampened seasonal cycle, not by similarity to EC GPP, so this result is not inevitable. Mild  
512 models tend to exceed EC estimates slightly in the two months before the driest month, or  
513 early in the dry season. A possible mechanism is insufficient modeled water stress. In con-  
514 trast, lively models, whose average the orange line tracks, simulate lower GPP for 5 months  
515 starting with the driest month. For lively models, lack of plant available water during the dry  
516 season may more strongly curtail GPP than EC data suggest.

517 At individual sites, the mean differences between mild and lively models are sharper,  
518 with more variation in timing relative to the dry season (Fig. S14). But the overall pattern-  
519 ing at individual sites is similar to the means (Fig. S4), with lively models on average dif-  
520 fering more from EC estimates than mild models except at the K83 site. During transition  
521 months on either side of the dry season, boxes overlap the zero line, showing that on average  
522 the models simulate GPP close to EC estimates during the shoulder seasons. At the RJA site  
523 there is little seasonal pattern in differences between models and EC, and mild models match  
524 EC GPP more closely throughout the year. At other sites, both groups of models estimate  
525 lower GPP during the dry season than do ECs, and higher during the wettest months. The  
526 tendency for all models to simulate GPP that on average is lowest relative to EC estimates  
527 during the late dry season suggests challenges in modeling soil moisture.

## 528 **6 Analysis**

529 Flattening describes the tendency for the variance of predictions made from otherwise  
530 accurate equations with significant omitted variables or other statistical noise to be reduced.  
531 In light of flattening, this paper addresses the extent of modeled GPP's seasonal variability  
532 in Amazonian rainforest, how strongly current weather variables determine GPP at six eddy  
533 covariance sites, and the fidelity of seasonal timing. lively models are defined as those with  
534 higher seasonal amplitude than EC GPP, while mild models are less seasonal (Fig. 3).

## 6.1 Summary of Findings

Both process and statistical models struggle to reproduce EC estimates of Amazonian rainforest gross primary productivity. Mean and/or variance of all models' GPP falls outside of 99<sup>th</sup> percentile confidence intervals (Fig. 2). Flattening helps interpret benchmark comparisons of the variances. Flattening's degree of influence is a function of how completely a model's drivers determine predicted outcomes, leaving only a small error term in the descriptive regression. Model outcome variances that are similar to benchmark variance indicate model skill only if included drivers explain most of the variability in the reference outcomes. Otherwise, or worse if modeled variance exceeds its reference equivalent, excessive model sensitivity to drivers has overwhelmed flattening. Multiple metrics in this study suggest that lively models are overly responsive, while the mild models appear most likely to represent accurately the mean GPP consequences of climate shifts for rainforests.

The regression that predicts EC GPP might appear strong enough that process and statistical models' predicted GPP variance should be nearly as high. The regression for EC GPP that includes both weather and site-specific intercepts explains a total of 59% of variability (Equation 4). However, the descriptive regression treats intercepts as the fixed outcome of unspecified variables. With reference data for only six intercepts, this paper does not explore the important component of GPP accuracy that resides in site means and their drivers.

Separated from site effects, a linear combination of current month's rain, temperature and light explains only about an eighth (12%) of EC GPP's total variance, equal to  $0.38 \text{ gCm}^{-2}\text{d}^{-1}$ . Although there are severe difficulties in "measuring" GPP at a flux tower, the portion of variability explained is so low that qualitative conclusions seem warranted. Current month's weather is a weak linear determinant of rainforest GPP, and the amount of GPP variability due to weather is small. In terms of comparing modeled GPP variance to the eddy covariance estimates, flattening is a strong influence because included drivers do not largely explain EC GPP.

For the lively models, the weight of evidence favors excessive sensitivity to weather drivers. The models pass flattening's indirect test of hypersensitivity; GPP variability of all the lively models as measured by both simple variance (Fig. 2) and seasonal amplitude (Fig. 3) exceeds EC GPP variability. Direct comparisons of descriptive regression slopes are even stronger evidence of excessive sensitivity. Responsiveness to rain is stronger in every lively model than for EC GPP (1.17 average v. 0.48, Fig. 5). Perhaps in counterbalance, all statis-

567 tically significant slopes for temperature and light for each lively model are of the opposite  
 568 sign from EC GPP's. Finally, mismatched seasonal cycles also imply that lively models have  
 569 excessive responsiveness to at least current rain. For highly seasonal models, the annual min-  
 570 imum in photosynthesis tends to be both later in the dry season than EC estimates, and more  
 571 severe (Fig. 7).

572 Whether mild models are overly sensitive is less clear. Net flattening does occur, which  
 573 makes excessive driver sensitivity less likely. Mild models' seasonal GPP amplitudes and  
 574 in most cases variances are below EC GPP's (Figs. 2, 3). However, it is possible that other  
 575 flattening influences are sufficiently strong to counteract excessive driver responsiveness.  
 576 Two main contributors to the flattening are likely. One is model misspecification noise due to  
 577 non-linearity in responses to weather (Figs. 6 and S2). Non-linearities were assessed only  
 578 qualitatively due to EC data limitations. The second likely cause of low variance in mild  
 579 models' GPP predictions is low spread in site intercepts. The range in site means for EC GPP  
 580 is  $4.0 \text{ gCm}^{-2}\text{d}^{-1}$ . Except for one outlier, the ranges of GPP site means for mild models all are  
 581 smaller, 0.8 to 3.5. Both model misspecification and low sensitivity to site mean differences  
 582 could flatten the GPP predictions.

583 Since excessive driver responsiveness in mild models could coexist with net flattening,  
 584 the sensitivity needs to be assessed more directly. One test is whether weather predictors as  
 585 a group explain an appropriate amount of mild model responsiveness. If model sensitivity  
 586 to weather perfectly matched the ECs', weather would explain the same absolute amount of  
 587 variability as it does for the towers,  $0.38 \text{ gCm}^{-2}\text{d}^{-1}$ . For mild models, the average variance  
 588 that weather explains is  $0.79 \text{ gCm}^{-2}\text{d}^{-1}$  (range across models = 0.09 - 4.60). Given the de-  
 589 gree of uncertainty in EC GPP, this check seems at most suggestive that some mild models  
 590 respond too weakly. Direct comparison of driver slopes indicates that the mild model group's  
 591 sensitivities to weather is reasonable overall, although there is quite a bit of spread among  
 592 models (Fig. 5). Average mild model responsiveness to rain and light is similar to that of EC  
 593 GPP, while temperature responsiveness is lower but at least of the same sign.

## 594 **6.2 Assessment of Findings**

595 Our results are specific to the scales of time and space for which they are calculated:  
 596 monthly means for 6 tower sites between 2000 and 2010. Driver strengths can vary with  
 597 time integration. At the K67 flux tower, for example, vapor pressure deficit and total and dif-

598 fuse light largely determined hourly averaged GPP, while a derived index that also included  
 599 leaf area index was better at explaining monthly averages [Wu *et al.*, 2017]. A model that re-  
 600 produces hourly photosynthetic fluxes well may still have substantial biases in annual totals  
 601 [Keenan *et al.*, 2012]. Spatial amalgamation even more strongly affects variability [Rödig  
 602 *et al.*, 2018]. Given the grossly finite mean annual amount of atmospheric water, even though  
 603 precipitation may drive local variability of GPP, temperature largely determines global vari-  
 604 ability in net land:atmosphere carbon exchange [Jung *et al.*, 2017].

605 Our analysis agrees with prior studies that have found rainforest GPP in global vegeta-  
 606 tion models reacts excessively to weather [Ahlström *et al.*, 2017; Baker *et al.*, 2008; Cleve-  
 607 land *et al.*, 2015; Huang *et al.*, 2016; Li *et al.*, 2017; Parazoo *et al.*, 2014; Piao *et al.*, 2013;  
 608 Poulter *et al.*, 2009; Restrepo-Coupe *et al.*, 2016; von Randow *et al.*, 2013; Zhu *et al.*, 2016].  
 609 Excessive rainforest GPP seasonality was reported for an earlier version of Model J at K67  
 610 [Sakaguchi *et al.*, 2011], and for Model I at a flux tower in Guyana [Zhu *et al.*, 2018]. How-  
 611 ever, we found also that a few mild models have weak responses to weather.

612 While future trends in Amazon light and rain are uncertain, temperature are expected  
 613 to rise [Jiménez-Muñoz *et al.*, 2013]. This makes model responses to temperatures that cur-  
 614 rently are outliers particularly important [Cavaleri *et al.*, 2015]. Ten of the study models  
 615 have a statistically significant negative response, consistent with Huntingford *et al.* [2013]  
 616 and Poulter *et al.* [2010a]. The mismatch between EC and model responses to temperature  
 617 is striking (Fig. 5). Not one is as welcoming of warmer temperature as EC GPP. Dispropor-  
 618 tionately strong responses to high temperature (Fig. 6) mean that differences between model  
 619 predictions will increase as the climate warms.

620 Independent indicators of tropical plant response to higher temperatures are mixed.  
 621 At La Selva, Costa Rica, the net of GPP and respiration fell strongly with increases over  
 622 12 years in daily minimum temperatures [Clark *et al.*, 2013]. Temperature appears to be a  
 623 positive and stronger driver of net ecosystem exchange globally, and precipitation a weaker  
 624 driver, than is represented in most dynamic global vegetation models [Wang *et al.*, 2013].  
 625 For some models, EBF's optimal temperature parameter is influential. MacBean *et al.* [2018]  
 626 found that one MsTMIP model's temperature flex point (Fig. 6) seems too low. Optimal tem-  
 627 perature for EBF is difficult to specify well including because the real value may vary among  
 628 the Amazon's thousands of tree species.

629 Mismatches in GPP seasonal timing (Fig. 7, Text S5), consistent with *Poulter et al.*  
630 [2010b], suggest that during the dry season plants experience less water stress than modeled.  
631 An exploratory assessment of cumulative rain as a proxy for soil moisture (Text S6) shows  
632 that while at individual sites cumulative rain is an important predictor, no lag duration works  
633 well everywhere. GPP's site-specific dependence peaks at three months at CAX and eight at  
634 BAN. When the only lag durations whose cumulative rain predictor is statistically significant  
635 at all sites, six or seven months, is used at all sites, the regression has a worse fit than if only  
636 current month's rain is included. Sites' differing optimal lag periods cancel each other when  
637 generalized and blur the importance for GPP of seasonal drought.

### 638 **6.3 A water- versus light-limitation dichotomy poorly characterizes the models.**

639 GPP increases with light in mild models, and falls with light in lively models (Fig. 5).  
640 The lively models also respond more to light strongly. It is tempting to associate each group  
641 with either strong response to light at the expense of temperature sensitivity or vice versa.  
642 If more rain causes GPP to rise enough, and tropical rain clouds reduce radiation, then GPP  
643 necessarily would fall with increasing light.

644 Clouds' opposing consequences for photosynthesis of more rain but less light [*Huete*  
645 *et al.*, 2006; *Nemani et al.*, 2003] has been labeled as light-limited versus water-limited  
646 [*Arias et al.*, 2011; *Baker et al.*, 2013, 2019; *Myneni et al.*, 2007]. If a site is water-limited,  
647 GPP falls during sunny periods due to soil dryness. At light-limited sites water is more  
648 plentiful and GPP rises during sunny periods [*Graham et al.*, 2003]. Some observational  
649 evidence contradicts the hypothesis that light limitations and water limitations represent a  
650 trade-off in the tropics, however, and instead that GPP may have little response to variations  
651 in light [*Restrepo-Coupe et al.*, 2013]. For four of the nine mild models, GPP is lowest  
652 during a dark month for at least one site, but for at least one other site GPP is lowest during a  
653 dry month when light is likely to be stronger (Fig. S3).

654 The dichotomy requires that light and rain be anti-correlated, with a tendency for in-  
655 creases in tropical clouds both to reduce light and to increase rain. In the MsTMIP driver  
656 data rain explains only 3% of the variation in light, and the correlation of EC GPP and light  
657 is negative (-0.14). Neither in the MsTMIP driver data does low rainfall bring more sensible  
658 heat in presumed Bowen ratio response to drought stress; Rain explains similarly little (4%)  
659 of the variation in temperature. While the lively models do respond strongly to precipitation,

660 as would be expected for modeling approaches that focus mostly on water limitations, there  
661 are few indications the mild models are strongly light-limited.

#### 662 **6.4 Hindcast GPP's weak responses to CO<sub>2</sub> do not reveal predicted responses to** 663 **CO<sub>2</sub>.**

664 CO<sub>2</sub> was a significant predictor of GPP only for two especially mild process models  
665 (Fig. S11). The CO<sub>2</sub> slope for five models was statistically indistinguishable from zero Much  
666 more than for the other weather drivers, however, the effect of CO<sub>2</sub> is likely to differ in com-  
667 ing decades from either real or modeled responses for 2000-2010.

668 Over time CO<sub>2</sub> can overwhelm trends in other environmental drivers due to its persis-  
669 tence and much larger relative change [Fisher *et al.*, 2013]. And GPP responses to CO<sub>2</sub> are  
670 unlikely to be linear, mainly because multiple and sometimes conflicting components shape  
671 a net trade-off between CO<sub>2</sub> fertilization and increased water use efficiency [Swann *et al.*,  
672 2016]. Land surface models differ substantially in how strongly they amplify atmospheric  
673 CO<sub>2</sub> increases [Piao *et al.*, 2013]. As an example, this study's model D, Tem6, simulates  
674 logarithmic increase [Jain *et al.*, 1994]. Other methods than comparison with this EC dataset  
675 will be needed to assess the accuracy of model sensitivity to CO<sub>2</sub> in rainforest vegetation.

#### 676 **6.5 Fluxcom GPP's low variance logically reflects flattening.**

677 Statistical models of GPP derive from a limited set of core time series data: satellites,  
678 flux towers, and ground-based weather observations. The statistical models included in this  
679 study use all three. There is no fourth independent source data with which the statistical  
680 models can be benchmarked. Some tentative conclusions about each statistical model's ac-  
681 curacy for the Amazon still can be drawn from the comparisons made above to flux towers,  
682 however.

683 There is conspicuous absence of a Fluxcom (Model G) assessment that features rain-  
684 forests even though rainforest is the most productive PFT on the planet and for climate per-  
685 haps the most important. One comparison of GPP from 53 eddy covariance towers to an  
686 early version of Fluxcom included EBF sites only in Australia and Italy [Joiner *et al.*, 2014].

687 The defining source for Fluxcom is EC data. Reassuringly, Fluxcom's gross fit with EC  
688 GPP is among the closest for this study's models. Only model B exceeds Fluxcom's overall

689 correlation with EC GPP ( $r = 0.37$ , Fig. S8). While Fluxcom's mean GPP in site cells is, like  
 690 all but one model, outside EC credible bounds, it is among the half-dozen models closest to  
 691 the EC mean. Fluxcom's rain responsiveness slope is one of the closest to EC estimates (Fig.  
 692 5). Its scaled temperature slope, 0.006, is much shallower than the ECs' slope of 0.460. But  
 693 so to some degree is every other model's. The sign of Fluxcom's slope for light matches that  
 694 of the EC towers although its response is stronger by almost half. Fluxcom's month of peak  
 695 GPP is within two months of EC estimates for all sites except RJA (Fig. S4), again among  
 696 the best matching of models.

697 Fluxcom's complex algorithms resemble linear regression in ways that make flattening  
 698 applicable. For example, in model tree ensembles, one of Fluxcom's options, machine learn-  
 699 ing stratifies spatially and temporally defined outcomes into bins but each simulated value  
 700 is the prediction of a particular bin's regression. Predictions from regressions are systemati-  
 701 cally flattened, with lower variance than the underlying data.

702 Flattening could help explain *Piao et al.* [2013]'s finding that Fluxcom's GPP is less  
 703 variable than any of ten DGVMs'. In our study, Fluxcom's GPP variance (1.4 v. ECs' 3.3  
 704  $\text{gCm}^{-2}\text{d}^{-1}$ ) and seasonal amplitude for site cells (2.2 v. 3.4  $\text{gCm}^{-2}\text{d}^{-1}$ ) are about half as vari-  
 705 able. Fluxcom and Wecann are less accurate for evergreen broadleaf forests (EBF) than their  
 706 global average [*Alemohammad et al.*, 2017; *Badgley et al.*, 2019]. Reasons for the poor per-  
 707 formance include the dearth of both eddy covariance towers and clear satellite retrievals in  
 708 the tropics [*Tramontana et al.*, 2016; *Jung et al.*, 2020]. Fluxcom would thus likely have es-  
 709 pecially well-smoothed tropical predictions.

710 Fluxcom's globally low GPP variability has been called "undersampled" [*Piao et al.*,  
 711 2013], "poorly captured" [*Tramontana et al.*, 2016], underestimated for reasons that are not  
 712 fully clear [*Jung et al.*, 2020], and, on the product's website, "too small" ([https://www.bgc-  
 713 jena.mpg.de/geodb/projects/Data.php](https://www.bgc-jena.mpg.de/geodb/projects/Data.php)). The tendency bears consideration when using Flux-  
 714 com. But in light of flattening, we disagree that Fluxcom's mildness is necessarily a weak-  
 715 ness. The low variance appears instead to be an inherent consequence of omitted drivers and  
 716 flattening, and suggests theoretically good potential for relatively accurate driver responsive-  
 717 ness globally.

718 Wecann's (Model F's) temporal span prevents reasonable direct comparisons to the EC  
 719 towers used in this study. In each comparison across the Amazon basin, Wecann resembles  
 720 Fluxcom (Text S3 and Figs. 6, S2, S10, S11, and S12). GPP differs at the warmest decile

721 (Fig. 6). Wecann is slightly less sensitive to light than is Fluxcom and more so to CO<sub>2</sub> (Fig.  
722 S11). Current weather explains less of the variance. But the models' residual standard errors  
723 (Fig. S12), which tend to indicate the extent of mismatch in a cycle's phase [Taylor, 2001],  
724 are similar. The detailed nature of these differences reinforce Wecann and Fluxcom's overall  
725 similarity.

726 While source data for Fluxcom and Wecann feature eddy covariance towers, VPM  
727 (Model ) emphasizes satellite sources. It uses especially tight coupling and straightforward  
728 equations to merge remotely sensed products. VPM's unusual spatial pattern of rain respon-  
729 siveness in the Amazon (Fig. 1, Model I) corresponds to likely seasonal trends in cloud con-  
730 tamination of satellite data retrievals. Cloud cover and therefore scarcity of acceptable satel-  
731 lite retrievals globally tends to peak in the tropics and during the wettest months. Gap-filled  
732 or missing data therefore overlap heavily with periods of high greenness and potentially of  
733 peak rainforest GPP. Photosynthetically active radiation from the NCEP II weather reanaly-  
734 sis is the model's multiplicative component that cloudiness is most likely to skew. Radiation  
735 from weather reanalyses was specifically omitted as an input to another statistical model due  
736 to the product's high uncertainty [Gentine et al., 2019; Jung et al., 2011].

737 VPM is more strongly anticorrelated with EC GPP than any other model ( $r = -0.19$ ).  
738 It is the only model for which rain's linear regression slope is negative. GPP falls with in-  
739 creasing precipitation for all but the driest decile (Fig. S2). VPM's response to light also is  
740 an outlier, increasing at every decile with no inflection point (Fig. 6). In terms of the month  
741 with lowest GPP at each site, while other models' average timing difference from ECs ranges  
742 from 1.5 to 3.3 months, VPM averages 4.5 months (Fig. S3). The phase of VPM's seasonal  
743 cycle is nearly opposite that of EC GPP at most sites. At no site does VPM simulate min-  
744 imum GPP during a dry season month. While VPM's GPP estimates are unrepresentative  
745 of the best reference data available for the Amazon, a logically underlying reason of cloud  
746 contamination applies most strongly to the tropics and could have little effect on VPM's ac-  
747 curacy elsewhere.

748 In summary, Fluxcom's match with EC GPP is weak. Their correlation is 0.37. But the  
749 fit is better than for almost all process models. Wecann compares similarly. The fidelity does  
750 not establish Fluxcom's or Wecann's veracity, merely their anticipated conformity with EC  
751 GPP in all its imperfections. Of this study's models, Fluxcom and Wecann appear to be the  
752 best currently available wall-to-wall estimates of mean Amazon GPP.

## 6.6 Flattening has practical implications for ESMs.

Flattening certainly does not mean that ESM outputs necessarily have low variability. GPP in the Amazon is a case in point. For many ESM subprocesses, other sources of model uncertainty and imperfection remain large enough to obscure and/or overwhelm flattening due to omitted random variables.

Flattening is likely to be a dominant problem for models privileged to have high accuracy (modest RSE) combined with low precision (low  $r^2$ ). These models suffer little from uncertainty about included predictors' responsiveness but lack some key drivers. Currently it is flattening's side effects that are most pertinent to ESMs. As described below, they affect studies of climate outliers; intermediate model calculations; trade-offs in the development, assessment, and use of models; and consequences of increasing model complexity.

There is enormous practical value in understanding change in the frequencies of droughts, wildland fires, heat waves, floods, and tropical cyclones [Katz and Brown, 1992]. Studies of climate outliers' consequences often base predictions on driver change or z-scores [Abatzoglou and Williams, 2016, for example] rather than on the variability of predictions directly, sidestepping any bias in variance of predictions that results from flattening. A conceptually similar option is to build model predictor metrics from simulated history [Camargo, 2013] rather than from observed driver history [Westerling *et al.*, 2011].

Feedbacks mean that variability internal to a model can affect mean outcomes. Precipitation intensity in the Amazon rainforest is an illustration. In typical ESM runs, rainfall depth is spread uniformly across a grid cell and can simulate overabundant frequency and duration of light mist [Baker *et al.*, 2019]. Rain reaching the soil is a step function that breaks when rain exceeds the depth that leaf surfaces can store. Modeled rainforest foliage remains wet far longer than it actually does after the region's legend cloudbursts. So much water evaporates from leaves that soil recharge is weak, in turn reducing GPP. Cloud superparameterization distributes rain among and within virtual sub-cells rather than spreading it uniformly, improving representation of soil moisture. Greater variability in rainfall intensity within each cell was required for simulating accurate mean rates of plant processes. But the superparameterization adjustment creates new inaccuracies in modeled global precipitation patterns [Phillips, 2019].

783           When variance of intermediate outputs needs to be similar to actual variability, fea-  
784           sible workarounds may be scarce. Theoretically they include explicit addition of statistical  
785           noise in the form of deterministic or random draws from exogenous distributions [*Pelletier,*  
786           1997; *Khodaparast et al., 2008*], finer temporal or spatial scaling, and ensemble modeling  
787           of probabilistic outcomes. An example of the first is LPJmL's semi-stochastic distribution of  
788           monthly rainfall to individual days [*Poulter et al., 2010a*]. Fuzzy parameters can be used  
789           to address not uncertainty in parameter estimates [*Hoffman and Miller, 1983; Ersoy and*  
790           *Yünsel, 2006*] but as a proxy for statistical noise due to omitted variables.

791           Driver sensitivity affects both trends and variability of outcomes [*Nijssen et al., 2019*].  
792           As long as models lack some influential real drivers, optimizing the predicted variance of  
793           predictions can compromise the accuracy of responsiveness to drivers and vice versa. Ide-  
794           ally the trade-offs are deliberate and publicly documented. The short list of diagnostics to  
795           which the Max Planck Institute's ESM was finely-tuned for CMIP5 included both means and  
796           variabilities [*Mauritsen et al., 2012*]. Especially when omitted variables are known to be  
797           highly influential, conversations about the relative importance of sensitivity versus variance  
798           for specific equations, models or applications may be fruitful. For example, flattening may  
799           affect benchmark selection; Fluxcom appears likely to be reasonable if noisy reference data  
800           for GPP's driver sensitivity.

801           Weather sensitivity describes, by definition, the consequences of a changing climate.  
802           It is important to focus at least as much on the accuracy of sensitivity as on predictions'  
803           variance despite possibly less certain benchmark data. Accuracy of outcome variances al-  
804           ready is integral to ILAMB *Collier et al. [2018]* and to many model intercomparison projects  
805           [*Houghton et al., 2001; Jupp et al., 2010; Li et al., 2019*]. Responsiveness to drivers tends to  
806           be more difficult to benchmark than the variance of outputs. But inadvertently prioritizing  
807           spread over responsiveness in accuracy assessments can be counterproductive.

808           As ESMs become increasingly accurate for the processes and drivers they include,  
809           and incorporate more of the drivers that cause real variability, simulated variance will tend  
810           to rise. Processes and drivers added over time to the ESMs in IPCC's Comprehensive As-  
811           sessment Reports have done little to reduce uncertainty around mean temperature trends but  
812           higher complexity has increased the accuracy of simulated variability [*Dahan, 2010*]. There  
813           are practical limits to adding global drivers to climate models, however.

## 814 **7 Conclusions**

815 We compared 15 process models and three statistical models to GPP estimates from  
 816 six eddy covariance towers in the Amazon rainforest. Models split almost equally between  
 817 weaker and stronger responsiveness than EC data to the environmental drivers of current  
 818 month's rain, temperature, and light. Most striking about the wide spreads across modeled  
 819 GPP's means, responsiveness to drivers, and amplitude of the annual cycle is how little virtu-  
 820 ally every model resembles EC estimates. Similarity to Amazon flux towers is one of many  
 821 important ESM accuracy metrics. Models that poorly replicate driver responsiveness in  
 822 tropical rainforest may do very well by other criteria. The implication, provided EC GPP  
 823 is somewhat realistic, is that lively models overreact to rain and have opposite signs of re-  
 824 sponse for light and temperature. Since temperature is likely to rise and rain to become more  
 825 variable, the liveliest models may substantially exaggerate the Amazon's future change and  
 826 peril.

827 As this article's title asserts, accurate deterministic simulation of both sensitivity to  
 828 drivers and variability for Amazonian GPP is unattainable. The reason is that weather ex-  
 829 plains so little of EC GPP variability that flattening is a strong influence. Of wider relevance  
 830 is the role of omitted processes and uncertainty due to other sources of modeling error in re-  
 831 ducing the variability of model predictions. It is generically appropriate to be more skeptical  
 832 of too much variability than of too little. In the interest of accurate sensitivity, low variance  
 833 of predictions relative to a benchmark may sometimes deserve acclaim.

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 836 ers in the Amazon, developed the models assessed, and participated in MsTMIP.  
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 838 of this study. GPP for MsTMIP [Huntzinger *et al.*, 2014] was downloaded from  
 839 [https://daac.ornl.gov/cgi-bin/dsviewer.pl?ds\\_id=1225](https://daac.ornl.gov/cgi-bin/dsviewer.pl?ds_id=1225), and driver data [Wei *et al.*, 2014] from  
 840 [https://daac.ornl.gov/NACP/guides/NACP\\_MsTMIP\\_Model\\_Driver.html](https://daac.ornl.gov/NACP/guides/NACP_MsTMIP_Model_Driver.html). For statistical  
 841 model GPP, Fluxcom [Jung *et al.*, 2019; Tramontana *et al.*, 2016] was downloaded from  
 842 <https://www.bgc-jena.mpg.de/geodb/projects/Data.php>, Wecann [Alemohammad *et al.*, 2017]  
 843 from <https://avdc.gsfc.nasa.gov/pub/data/project/WECANN/>, and VPM [Xiao *et al.*, 2005;  
 844 Zhang *et al.*, 2017] from [https://figshare.com/articles/Monthly\\_GPP\\_at\\_0\\_5\\_degree/5048011](https://figshare.com/articles/Monthly_GPP_at_0_5_degree/5048011).

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848 is underway. Url will go here.] SiB4 GPP derives from coauthors' prior work. The authors  
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