

1 Identification of Time-varying Parameters of a 2 Monthly Budyko Function and its Implications

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Abstract: The Budyko framework, which can describe a simple but useful partitioning of precipitation under supply and energy limits, is used widely to estimate evapotranspiration (ET). Previous studies have investigated time-variant Budyko functions on annual or interannual scales but seldom on the intra-annual scale. This study used a monthly two-parameter (κ and y_0) Budyko function and three schemes that considered single observations (ET or streamflow (Q)) and dual observations (ET and Q) to assimilate the time-varying parameters using the ensemble Kalman filter method. The study considered the contiguous USA (CONUS) using the Model Parameter Estimation Experiment dataset. The time-varying parameters were explained on the basis of time series analysis and correlation with meteorological data. Three conclusions were as follows. (1) The identified time-varying parameters (κ and y_0) of the Budyko function could effectively simulate ET . (2) The assimilation using only ET observations could identify a plausible set for parameter κ but was inadequate for y_0 . (3) Most time-varying parameters exhibited a 12-month period, and the trend and change points detected for Midwest CONUS were related to anthropogenic influences such as extraction and use of groundwater. The findings show that changing environment can be detected by using the proposed time-varying parameters of the Budyko function.

Keywords: monthly Budyko function; time-varying parameters; ensemble Kalman filter; contiguous USA; MOPEX dataset;

1. Introduction

Hydrological models generalize the complex processes of the water cycle, which is of great importance in relation to research and experiments on the hydrological system. Parameters adopted in hydrological models play a crucial role in simulation performance and are often calibrated as constants. However, as with the effects of varying conditions such as afforestation, climate change, and human activities (Peel & Blöschl, 2011; Wu & Johnston, 2007), model parameters cannot be treated as time-invariant. Thus, it is more practical to analyze model parameters taking time-variant climatic factors and underlying surface conditions into consideration (Merz, Parajka,

44 & Blöschl, 2011; Patil & Stieglitz, 2015; Thirel et al., 2015).

45 The responses of time-varying parameters to changes in the underlying surface
46 and climatic conditions can be obtained using various methods. For example, the
47 functional form of parameters can be established with seasonal, annual, or interannual
48 trends (Jeremiah, Marshall, Sisson, & Sharma, 2013; Marshall, Sharma, & Nott, 2006;
49 Westra, Thyer, Leonard, Kavetski, & Lambert, 2014). However, additional judgement
50 is required to establish such functions and the parameters must always be adjusted for
51 different situations. An alternative method is to calibrate the model parameters by
52 dividing the available data into consecutive subsets using optimization algorithms (de
53 Vos, Rientjes, & Gupta, 2010; Seibert, McDonnell, & Woodsmith, 2010; Vaze et al.,
54 2010). Unfortunately, the selection of such subsets is subjective, and optimization
55 algorithms based on simple screening criteria (commonly, the Nash–Sutcliffe
56 efficiency (NSE) for runoff) cannot take full advantage of the information provided
57 by climatic data.

58 The data assimilation (DA) technique, the main concept of which is to compare
59 the errors between observations and model simulations to obtain optimal estimates of
60 the current system, is used widely to identify model parameters. Through the balance
61 of information (uncertainties) between a model and observations, the observations can
62 be used via DA to update model states and obtain model parameters simultaneously
63 (Abbaszadeh, Moradkhani, & Yan, 2018; Clark et al., 2008). In studies that focus on
64 identifying parameters, application of DA methods is generally effective (Chao Deng,
65 Liu, Guo, Li, & Wang, 2016; Feng et al., 2017; Smith, Beven, & Tawn, 2008; Vrugt,
66 ter Braak, Diks, & Schoups, 2013). Additionally, more than one set of observations
67 could be used to improve the efficiency of the DA method (M. S. Xiong et al., 2019).
68 However, similar to traditional optimization algorithms, DA methods have problems
69 regarding parameter equifinality when the hydrological model has a number of
70 parameters. To overcome this drawback, a simple model such as the Budyko
71 framework was used in this study.

72 The Budyko framework contains few parameters, and describes a simple but
73 useful partitioning of precipitation into evaporation and runoff under supply and
74 energy limits (Sankarasubramanian. et al., 2020). The Budyko hypothesis is that

75 multiyear average evaporation depends on available water capacity and atmospheric
76 water requirement. The framework based on the Budyko hypothesis has been used
77 extensively in the hydrological field to estimate evaporation (Bagrov, 1953; Fu, 1981;
78 Mezentsev, 1955; Milly, 1994; Pike, 1964; Porporato, Daly, & Rodriguez-Iturbe,
79 2004; Turc, 1954; Yang et al., 2007; L. Zhang, Dawes, & Walker, 2001; L. Zhang et
80 al., 2004).

81 Budyko-type equations have been applied widely on annual or interannual
82 scales, and time-varying investigations at such temporal scales have recently drawn
83 considerable attention. For example, Sinha, Jha, and Goyal (2019) estimated mean
84 annual-scale watershed parameter with watershed properties using the multiple linear
85 regression method and machine learning techniques. X. Zhang, Dong, Cheng, and Xia
86 (2019) improved the fitting capability for most catchments in the United States by
87 establishing a function between annual-scale time-varying parameters and time steps.
88 However, few studies have focused on the time-varying parameters of Budyko
89 functions at the intra-annual (e.g., monthly) scale. Model parameters on long temporal
90 scales represent an average reflection of the annual or interannual climatic and
91 underlying conditions, and their variation weakens at such scales. Consequently,
92 important change rules regarding both parameters and varying climatic conditions
93 might not be investigated adequately on annual or interannual scales. Moreover,
94 assumptions and applicability might lead to deviations when the Budyko framework
95 is applied on shorter temporal scales (Mianabadi, Davary, Pourreza-Bilondi, &
96 Coenders-Gerrits, 2020). For example, to investigate time-variant conditions using the
97 Budyko equation, soil water content must be changeable on the intra-annual scale,
98 which is markedly different to that on the annual scale.

99 Thus, this study investigated time-varying Budyko functions on the intra-annual
100 scale. The main objectives of this research were as follows: (1) to investigate an
101 effective approach for identification of intra-annual-scale time-varying parameters of
102 the Budyko function, (2) to determine whether evaporation observations alone or in
103 combination with streamflow observations could be used effectively to assimilate
104 parameters, and (3) to establish whether reasonable change rules regarding parameters
105 could be explained on the basis of the climatic or underlying conditions. The
106 remainder of this paper is organized as follows. The methods and experimental design

are introduced in Section 2. Section 3 describes the case study, including the study area and parameter setting. The results are presented and discussed in Section 4. Finally, our conclusions are stated in Section 5.

2. Methodology

An overview of the methods and analysis structure adopted for this study are shown in **Fig. 1**. Three schemes that considered single observations (ET or streamflow (Q)) and dual observations (ET and Q) were used to incorporate different observations into the assimilation. Subsequently, parameter sets were analyzed through time series analysis.

2.1 A two-parameter Budyko function for evaporation calculation

Greve, Gudmundsson, Orlowsky, and Seneviratne (2016) proposed a two-parameter Budyko function that can be used for the intra-annual dimension, e.g., the monthly scale. The parameters of this equation are κ and y_0 , which constitute the following equation:

$$\frac{E}{P} = F(\Phi, \kappa, y_0) = 1 + \Phi - (1 + (1 - y_0)^{\kappa-1} (\Phi)^\kappa)^{\frac{1}{\kappa}} \quad (1)$$

where E is the evaporation of the catchment, P represents precipitation, and Φ is the aridity index (AI), which is calculated as the ratio of potential evaporation and precipitation.

As in the traditional Budyko approach, the first parameter κ is a free model parameter, while the second parameter y_0 has physical significance because it might account for the maximum amount of additional water available to E (Greve et al., 2016). Specifically, Greve et al. (2016) interpret κ as the integrator of various factors other than the AI that influence the partitioning of precipitation. It can have wide application in relation to monthly data and is suitable for conditions when

evapotranspiration exceeds precipitation from the monthly to annual scale.

2.2 Calculation of monthly runoff

Monthly runoff can be calculated to assimilate runoff observations. Generally, runoff is connected with soil moisture on the monthly scale, and runoff can be assumed to have a hyperbolic tangent relationship with soil water content (L. H. Xiong & Guo, 1999), which can be expressed as follows:

$$Q_t = S_t \times \tanh(S_t / SC) \quad (2)$$

where Q_t is the monthly streamflow and S_t represents the soil water content, which is the specific parameter that can reflect the water storage capacity (SC) of the watershed.

Finally, the monthly soil water content can be expressed as follows:

$$S_t = S_{t-1} + P_t - ET_t - Q_t \quad (3)$$

Descriptions and ranges of the parameters introduced in Section 2.1 and 2.2, listed in **Table** , are based on previous research by M. S. Xiong et al. (2019) and Greve et al. (2016).

2.3 Ensemble Kalman filter

The ensemble Kalman filter (EnKF) is a type of DA technique that uses sequential data, in which an error covariance matrix is approximated using the Monte-Carlo method. An ensemble of state simulations is generated to regenerate the state variables and calibrate the model parameters, i.e., update is performed on the basis of a sequence of observed data (C. Deng, Liu, Guo, Wang, & Wang, 2015; Evensen, 1994). The steps necessary in application of the EnKF method are as follows.

(1) Initialize the background field. First, choose the state and observation variables to build the equations. Second, set the size of ensemble N and the length of the time step n for the DA, which is determined by the data length. Then, the state variables are created following a Gaussian distribution.

158 (2) Forecast the value. After initializing the background field at time step k ,
 159 forecasts of state variables and model parameters can be generated by disturbing the
 160 values updated from the previous time step. The state equation is as follows:

$$161 \quad \begin{pmatrix} \theta_{i,k+1}^f \\ x_{i,k+1}^f \end{pmatrix} = \begin{pmatrix} \theta_{i,k}^a \\ f(x_{i,k}^a, \theta_{i,k}^f) \end{pmatrix} + \begin{pmatrix} \delta_{i,k} \\ \varepsilon_{i,k} \end{pmatrix}, \quad \delta_{i,k} \sim N(0, R_k), \varepsilon_{i,k} \sim N(0, G_k) \quad (4)$$

162 where $\theta_{i,k+1}^f$ (parameter vector) and $x_{i,k+1}^f$ (state vector) mean the i -th ensemble
 163 forecast variable at time step $k+1$; $\theta_{i,k}^a$ and $x_{i,k}^a$ are analysis variables of the i -th
 164 ensemble member at time step k ; f can be described as a model operator; and $\delta_{i,k}$
 165 and $\varepsilon_{i,k}$ represent the model error (white noise) of the i -th ensemble member. It
 166 should be noted that the model errors obey a normal distribution with mean value of
 167 zero and variance (R_k, G_k) . All uncertainties regarding the model structure and data
 168 are reflected in $\varepsilon_{i,k}$.

169 (3) Observation perturbing. The universal observation equation is as follows:

$$170 \quad y_{i,k+1} = h(x_{i,k+1}, \theta_{i,k+1}) + \xi_{i,k+1}, \quad \xi_{i,k+1} \sim N(0, S_k) \quad (5)$$

171 where $y_{i,k+1}$ is the simulation value of the i -th ensemble member at time step $k+1$;
 172 y_{k+1} is the observation vector of time $k+1$; h is the observational operator that
 173 converts the model state variables to observations; and $\xi_{i,k+1}$ is the measurement error
 174 following the normal distribution.

175 (4) Update. The updating process is formalized as follows:

$$176 \quad \begin{pmatrix} \theta_{i,k+1}^a \\ x_{i,k+1}^a \end{pmatrix} = \begin{pmatrix} \theta_{i,k+1}^f \\ x_{i,k+1}^f \end{pmatrix} + \begin{pmatrix} K_{k+1}^\theta [y_{i,k+1} - h(x_{i,k+1}^f, \theta_{i,k+1}^f)] \\ K_{k+1}^x [y_{i,k+1} - h(x_{i,k+1}^f, \theta_{i,k+1}^f)] \end{pmatrix} \quad (6)$$

177 where K_{k+1} represents the Kalman gain matrix, the calculation method of which was
 178 described by Moradkhani, Sorooshian, Gupta, and Houser (2005) as follows:

$$179 \quad K_{k+1}^x = \sum_{k+1}^{xy} (\sum_{k+1}^{yy} + S_k)^{-1}, K_{k+1}^\theta = \sum_{k+1}^{\theta y} (\sum_{k+1}^{yy} + S_k)^{-1} \quad (7)$$

$$180 \quad \sum_{k+1}^{xy} = \frac{1}{N-1} \sum_{i=1}^N (x_{i,k+1}^f - \bar{x}_{k+1}^f)(\hat{y}_{i,k+1} - \hat{y}_{i,k+1}^m)^T \quad (8)$$

$$181 \quad \sum_{k+1}^{\theta y} = \frac{1}{N-1} \sum_{i=1}^N (\theta_{i,k+1}^f - \bar{\theta}_{k+1}^f)(\hat{y}_{i,k+1} - \hat{y}_{i,k+1}^m)^T \quad (9)$$

$$182 \quad \sum_{k+1}^{yy} = \frac{1}{N-1} \sum_{i=1}^N (\hat{y}_{i,k+1} - \hat{y}_{i,k+1}^m)(\hat{y}_{i,k+1} - \hat{y}_{i,k+1}^m)^T \quad (10)$$

183 where \sum_{k+1}^{xy} is the cross covariance of the simulated $\hat{y}_{i,k+1}$ and forecasted states $x_{i,k+1}^f$

184 (Snyder & Zhang, 2003); $\sum_{k+1}^{\theta y}$ is the cross covariance of the simulated $\hat{y}_{i,k+1}$ and

185 forecasted parameters; \sum_{k+1}^{yy} represents the error covariance of the simulated $\hat{y}_{i,k+1}$

186 (Mitchell & Houtekamer, 2000); \bar{x}_{k+1}^f and $\bar{\theta}_{k+1}^f$ are the means of the forecasted states

187 and parameter ensembles, respectively; and $\hat{y}_{i,k+1}^m$ is the mean of the simulated $\hat{y}_{i,k+1}$
 188 ensemble.

189 2.4 Experimental design

190 Using the evaporation and runoff calculation, the state and observation equations
 191 can be described as follows:

192 State equation:

$$\begin{pmatrix} \kappa_{i,k+1}^f \\ y_{0i,k+1}^f \end{pmatrix} = \begin{pmatrix} \kappa_{i,k}^a \\ y_{0i,k}^a \end{pmatrix} + \begin{pmatrix} \delta_{1,j,k} \\ \delta_{2,j,k} \end{pmatrix} \quad (11)$$

$$\begin{pmatrix} \kappa_{i,k+1}^f \\ y_{0i,k+1}^f \\ SC_{i,k+1}^f \\ S_{i,k+1}^f \end{pmatrix} = \begin{pmatrix} \kappa_{i,k}^a \\ y_{0i,k}^a \\ SC_{i,k}^a \\ f(S_{i,k}^a, \kappa_{i,k+1}^f, y_{0i,k+1}^f, SC_{i,k+1}^f) \end{pmatrix} + \begin{pmatrix} \delta_{1,j,k} \\ \delta_{2,i,k} \\ \delta_{3,j,k} \\ \varepsilon_{i,k} \end{pmatrix} \quad (12)$$

Observation equation:

$$ET_{k+1} = h(\kappa_{k+1}^f, y_{0k+1}^f) + \xi_{1,k+1} \quad (13)$$

$$Q_{k+1} = h(S_{k+1}^f, \kappa_{k+1}^f, y_{0k+1}^f, SC_{k+1}^f) + \xi_{2,k+1} \quad (14)$$

$$\begin{pmatrix} ET_{k+1} \\ Q_{k+1} \end{pmatrix} = \begin{pmatrix} h(\kappa_{k+1}^f, y_{0k+1}^f) \\ h(S_{k+1}^f, \kappa_{k+1}^f, y_{0k+1}^f, SC_{k+1}^f) \end{pmatrix} + \begin{pmatrix} \xi_{1,k+1} \\ \xi_{2,k+1} \end{pmatrix} \quad (15)$$

Three schemes were designed to obtain time-varying hydrological model parameters by the DA method using single data (ET or Q) and multivariate data (ET and Q). Experimental investigation was implemented to compare the three schemes in terms of model efficiency and parameter variation. Each of the three schemes is outlined below and brief descriptions are presented in Table 2. Descriptions of the assimilated time-varying parameters, calculation modules and observational data of the designed experimental schemes.

(1) Scheme 1 updates model parameters κ and y_0 by assimilating only ET data with the two-parameter Budyko function to investigate the performance of assimilation through the simple equation. Owing to the low complexity, there is no state variable in this scheme. Thus, Eqs. (11) and (13) are included in Scheme 1.

(2) Scheme 2 assimilates only Q data to update the model parameters. The difference from Scheme 1 is that the model combines the Budyko function with the monthly water balance model and the number of parameters is increased to three: κ , y_0 , and SC . The state variable S is updated solely by assimilating Q . Therefore,

214 Scheme 2 contains Eqs. (12) and (14).

215 (3) Scheme 3 uses an additional set of data. Multivariate data are used to update
 216 both model parameters κ , y_0 , and SC and state variable S . The model structure of
 217 Scheme 3 is similar to that of Scheme 2, in that both the Budyko function and the
 218 water balance model are used to obtain the parameters. Thus, Scheme 3 uses Eqs. (12)
 219 and (15).

220 2.5 Performance evaluation criteria

221 (1) The NSE (Nash & Sutcliffe, 1970) is an evaluation criterion applicable to
 222 evaluation of the effect of DA in this study:

$$223 \quad NSE = 1 - \frac{\sum_{i=1}^n (y_{sim,i} - y_{obs,i})^2}{\sum_{i=1}^n (y_{sim,i} - \bar{y}_{obs})^2} \quad (16)$$

224 where $y_{sim,i}$ and $y_{obs,i}$ are the simulated and observed variables for the i -th month,
 225 respectively, \bar{y}_{obs} is the average of the observed variable, and n is the total number of
 226 data time points. The NSE value ranges from $-\infty$ to 1, where the upper limit means
 227 perfect match between the simulation and observations.

228 (2) Kling–Gupta efficiency (KGE) is implemented through three components to
 229 avoid the NSE sensitivity to the data peak:

$$230 \quad KGE = 1 - \sqrt{(r-1)^2 + \left(\frac{\sigma_{sim}}{\sigma_{obs}} - 1\right)^2 + \left(\frac{\mu_{sim}}{\mu_{obs}} - 1\right)^2} \quad (17)$$

231 where r is the correlation coefficient of the simulated and observed data, and σ_{sim} ,
 232 σ_{obs} , μ_{sim} , and μ_{obs} are the standard deviations and mean values of the simulated and
 233 observed data, respectively.

234 (3) The Pearson correlation coefficient (R) was used to detect linear relationships
 235 between the assimilated parameters and between the parameters and the

236 meteorological factors:

$$R = \frac{\sum_{i=1}^n (a - \bar{a})(b - \bar{b})}{\sqrt{\sum_{i=1}^n (a - \bar{a})^2 (b - \bar{b})^2}} \quad (18)$$

238 where a and b represent the correlated variables, and \bar{a} and \bar{b} represent the mean
239 values of the variables.

240 (4) The root mean square error (RMSE) and volume error (VE), which were used
241 to evaluate the estimated parameter results in our study, reflect the differences
242 between simulated and observed variables:

$$RMSE = \sqrt{\frac{1}{n} \sum_{i=1}^n (y_{sim,i} - y_{obs,i})^2} \quad (19)$$

$$VE = \frac{\sum_{i=1}^n y_{sim,i} - \sum_{i=1}^n y_{obs,i}}{\sum_{i=1}^n y_{obs,i}} \quad (20)$$

245 where $y_{sim,i}$ and $y_{obs,i}$ are the simulated and observed variables, respectively.

246 (5) Diversity (defined just like the mean absolute relative error) is used to
247 measure differences between the absolute values of two sets of data:

$$D = \frac{1}{n} \sum_{i=1}^n \frac{|x_{exp,i} - x_{ref,i}|}{x_{ref,i}} \quad (21)$$

249 where $x_{exp,i}$ and $x_{ref,i}$ are the experimental and reference values, respectively.

250 2.6 Time series analytical methods

251 Three time series analytical methods were used to detect the trend, change
252 points, and periodicity of the parameters to elucidate potential variation rules.

(1) The Mann–Kendall trend test method (Kendall, 1975; Mann, 1945) has been used widely for detecting data trends and is recommended as a standard method by the World Meteorological Organization (Ziegler et al., 2003). If the standard normal statistical variables are positive, there is an upward trend in the time series data, and vice versa.

(2) The Pettitt test is a nonparametric approach that was first used to examine the change point of a dataset by Pettitt (1979). This method is simple and appropriate for identifying the change points of data and parameters. Thus, we applied this method to analyze the change points in this study.

(3) The wavelet transform analysis method, which can identify periodic changes by exploring the fluctuation signal law and distribution of multiple temporal scale hydrological sequences, has been used in general analysis of the laws of variation and periodicity of rainfall, runoff, and storage water (Shao Xiaomei, 2006). The function of the Morlet wavelet transform was used in this study to detect the periodicity of the parameters, because wavelet variance can reflect the principal period and other oscillation cycles of the parameters.

3. Case study

3.1 Study region and dataset

This study considered the contiguous USA (CONUS), which lies within the region 25°–49°N, 70°–130°W. The attributes of the studied catchments and relevant climatic data were obtained from the Model Parameter Estimation Experiment (MOPEX) (Duan et al., 2006). The hydrometeorological dataset extended between 1983 and 2003, in which precipitation and streamflow data were obtained from MOPEX, and potential evaporation and actual evaporation data were collected from a remote sensing dataset provided by the University of Montana (K. Zhang, Kimball, Nemani, & Running, 2010). In this dataset, actual evaporation is obtained using the Penman–Monteith equation (Monteith, 1972; Penman, 1948) and remote sensing results, and potential evaporation is computed using the Priestly–Taylor method (Priestley, 1972). To introduce data reflecting vegetation cover conditions, bimonthly

normalized difference vegetation index (*NDVI*) data were obtained from 8-km spatial resolution Advanced Very High Resolution Radiometer imagery (Tucker et al., 2005). All data were analyzed with consideration of first-year data as warmup.

Overall, 188 catchments were chosen according to the following criteria: (1) acceptable simulation performance, e.g., early screening NSE value of streamflow and evaporation larger than 0.6 (C. Deng, Liu, Wang, & Wang, 2018), (2) unabridged data record during the entire 21-year period, (3) relative water balance error within 10%, and (4) monthly *ET* no greater than potential *ET* (*PET*). The 188 selected catchments, which had area of 67–9806 km² and slope of 2.5%–25.6%, were classified into humid and (semi-)arid regions based on the *AI* (defined as the ratio of average annual *PET* to *P*) (Budyko, 1974). Catchments with *AI* > 1 were considered as (semi-)arid regions, while all others were considered as humid regions. The location and *AI* distribution of the selected catchments are shown in **Fig. 2**. The area and meteorological annual mean value of certain indices of the catchments (1983–2003) are shown in **Fig. 3**.

3.2 Parameter settings

Many previous studies have discussed the equilibrium between assimilation performance and computational demand, which is influenced by ensemble size (DeChant & Moradkhani, 2012; Evensen, 2004; Moradkhani et al., 2005). In our study, the ensemble size N was set as 1000, based on consideration of other research and the performance investigation of this study. The initial parameters and state variables were generated randomly as a Gaussian distribution, and the variance of the noise added to the parameters and the state and output variables was set on the basis of its magnitude (Leisenring & Moradkhani, 2012). The initial values of the proportionality factors were set as follows: 10% for κ and y_0 in Scheme 1; 10%, 10%, and 5% for κ , y_0 , and SC in both Scheme 2 and Scheme 3; and the observation error of streamflow and evaporation was 15%. Meanwhile, the parameter variance multipliers and uncertainty bounds were used to adjust the proportion of the error automatically in Scheme 1 and Scheme 2 (Leisenring & Moradkhani, 2012). Owing to the great difficulty in applying this method in multivariate assimilation, the multipliers and uncertainty bounds were not used in Scheme 3.

4. Results and discussion

4.1 Model performance

Results of several evaluation criteria (NSE , KGE , $RMSE$, VE) that reflect the performance of the evapotranspiration simulation in each of the three schemes are shown in **Fig. 4**. The values of NSE and KGE in Scheme 1 and Scheme 3 (hereafter, presented as S1 and S3, respectively) are high and close to 1, while the $RMSE$ value is reasonable compared with the magnitude of ET . Meanwhile, the values of VE in S1 and S3 are symmetrical on both sides of zero, and the median of which is approximate to zero. However, the relative magnitude in Scheme 2 (hereafter, presented as S2) is unacceptable. The above results demonstrate that the simulated ET of both S1 and S3 is plausible and that DA with the EnKF method is effective in terms of simulating ET well using assimilated time-varying parameters. However, the effect of S2 is unsatisfactory and its poor performance can be attributed to the lack of direct utilization of evaporation observations.

4.2 Relationships between parameters in schemes

Parameters κ and y_0 were used in all three schemes; thus, relationships between and comparisons of the two parameters among the three schemes need to be detected. Ternary diagrams reflecting the mean values of the two parameters in each of the three schemes are presented in **Fig. 5**. The mean values of κ and y_0 are nearly all situated on straight line l_2 , indicating that the mean values of the two parameters assimilated from the 188 catchments under S1 and S3 are of similar order of magnitude. However, the mean values of the parameters of the catchments in S2 differ markedly from those of the other two schemes.

The correlation coefficients and diversity of the two parameters were investigated among the three schemes two by two, the results of which are drawn in half-box plots in **Fig. 6a** and **6b**. Almost all the coefficients of κ between S1 and S3 are close to 1 and the median values of y_0 are near 0.5, whereas no correlation is evident among other compared combinations. In terms of diversity, the metrics of κ

and γ_0 are both <0.10 between S1 and S3, indicating similarity in the magnitudes of the parameters between these two schemes at each time step. Intuitively, it appears that there is a strong relationship and that κ has similar magnitude between the scheme solely using ET data and that using both ET and Q observations in the assimilation. Additionally, there are certain differences in the correlation coefficient in identifying γ_0 after the introduction of Q data (i.e., S3 compared with S1).

Mutual correlation coefficients between parameters κ and γ_0 in the same scheme were calculated and drawn as half-box plots in **Fig. 6c**. It can be seen that the median value of the coefficients in S1 and S2 is near zero. However, the median value of the coefficients in S3 lies near 0.6, representing clear correlation. The reason is the additional observation set in S3 (streamflow against S1; evaporation against S2) revises the model parameters better. It can activate the relationships among different climatic variables that constitute a more complete water cycle, which helps identify the relationship between two parameters.

The spatial distribution of the mean values of two parameters are drawn in Fig. S1. A similar distribution in central CONUS can be observed in relation to S1 and S3, especially for κ . There exists a type of *neighborhood effect* (i.e., similar values among neighboring catchments) in terms of the mean values of the parameters. Thus, it can be considered that the underlying surfaces and ecological attributes of these catchments are similar. In fact, these catchments are located in the same Bailey's ecoregion, which is a classification first proposed by Bailey (1994). Catchments in central CONUS belong to tropical continental and steppe regions, where similar types of natural vegetation grow. Interestingly, Carmona, Sivapalan, Yaeger, and Poveda (2014) obtained similar results using a single-parameter Budyko function. They found that catchments in the same area as identified in our study have almost the same value of parameter α .

4.3 Time series analysis of two parameters

In this section, analyses of the trend, change point, and periodicity of time-varying parameters are discussed. However, emphasis is placed on S1 and S3 owing

to the inadequate accuracy and poor performance of S2.

4.3.1 Trend analysis

As shown in **Fig. 7** and Fig. S2, there exist similar overall situations and spatial distributions of the trend of κ in S1 and S3. Decreasing and nonsignificant trends in y_0 are found in central CONUS, while increasing trends are evident in northwestern and eastern areas in S1. However, in S3, the decreasing and nonsignificant trends are retained, whereas the increasing trend in other catchments weakens. This is attributable to the introduction of streamflow data that can correct catchment attributes, especially water storage change. This aspect of this attribute is represented by y_0 (consistent with the definition by Greve et al. (2016)), which demonstrates that the method using only evaporation data cannot adequately identify the perfect trend of y_0 . It should be noted that there is no obvious difference for κ .

Additional analysis was conducted in relation to other parameters and state variable S in S3, which is a comparatively more reliable scheme. Several catchments situated in central CONUS exhibit the same decreasing trend, shown by the green color in **Fig. 8a**. Similar change points detected in these catchments are also discussed in Section 4.4. Carmona et al. (2014) concluded that these central catchments (shown by the red circle in **Fig. 8a**) have similar values of parameter α (in the classical Budyko function) with abnormal runoff coefficients and evaporation ratios. Human activities have substantial impact on the MOPEX catchments in this region. Dams, surface water utilization, groundwater extraction, and changes in agricultural land are the primary factors that cause the decreasing trends and change points in this area. Moreover, such grassland regions are mostly dominated by tile and agricultural drainage (especially in Iowa, shown by the brilliant blue circle in **Fig. 8a**), which results in changes in evaporation and groundwater level. Similar attribution can be found in both Wang and Hejazi (2011) and Ye, Yaeger, Coopersmith, Cheng, and Sivapalan (2012).

4.3.2 Change point analysis

Few catchments in S1 have change points. It is noteworthy that no catchment has a change point of κ , except for two catchments in northwestern CONUS in S3, which can illustrate a weakening of change points from S1 to S3. As for γ_0 , change points are identified in most catchments in S1, while the number declines to 63 in S3. These catchments are scattered over CONUS, although some are distributed in the middle of the continent; detailed information is presented in Fig. S3.

Results concerning detection of the change points of state variable *storage* are shown in **Fig. 8b**. Almost the same change points are identified in central CONUS, i.e., Iowa and surrounding states. Together with the results in Section 4.3.1, it can be considered that these catchments have the same trends of decrease and similar change points of storage water (state variable in our model). Furthermore, these discoveries reflect anthropogenic influences such as the extraction and usage of groundwater and other types of water resource. Hence, the parameters identified under the finer scale in our study can be considered as indicators of these changes.

4.3.3 Periodicity analysis

The spatial distribution of periodicity in the catchments is shown in Fig. S4. The colors from green to red represent a gradual increase in the number of identified period cycles. We suppose that a period of 12 or near 12 months can represent the characteristic annual hydrological cycle, which is reasonable for catchments whose weather characteristics cycle annually. In each catchment, κ is the general parameter that might have a period of one year. Intuitively, periodicity can be deemed unreasonable when the identified period cycle is relatively long (e.g., >60 months and even up to 200 months in our identification). First, the result of S2 is considered unreasonable because the periods of κ are discrepant and far from an exact value. A period of one year can be identified in most catchments in S1 and S3, and this phenomenon is more obvious in the latter than in the former. As shown in Fig. S4d and S4f, catchments in S1 do not have periodicity in γ_0 , while nearly half the catchments (92 catchments) have a period of 12 months in S3. Owing to the physical significance of water storage in γ_0 , this can prove that additional introduction of

streamflow data can help identify the physical mechanism and definition. Catchments for which a period cannot be identified are situated in central and southwestern CONUS.

Hydrologic regions defined by the US Geological Survey are drawn in **Fig. 9**. Those catchments without periodicity (red color) are in regions 3, 5, and 10, almost consistent with the boundaries of the hydrologic regions. Catchments in region 10 are near Midwest CONUS, where human activities, especially the pattern of use of water resources, are considered the main impacts, as mentioned in Section 4.2. Catchments in regions 3 and 5 suffered vegetation reduction during 1992–2001 (Wickham, Wade, & Riitters, 2011), which might account for the lack of periodicity.

4.4 Relationships between parameters and meteorological data

Correlation coefficients between meteorological data and the parameters of corresponding months were calculated for the three schemes. Moreover, the coefficients between the parameters of corresponding months and forcing from 1–6 months prior to 1–3 months following (marked as lag-1 to lag-6 or pre-1 to pre-3) were computed to detect whether meteorological factors could impact the catchment attributes by way of delay or advance. Details of the results are illustrated in Fig. S5.

The magnitudes and situations of hysteresis in S1 and S3 are almost identical. Parameter κ has positive relationships with ET , PET , and $NDVI$ and negative relationships with the difference in precipitation and evapotranspiration ($P-E$) and runoff. The absolute value of the correlation coefficient in S3 is larger than in S1, which proves that introducing streamflow data into the assimilation can slightly improve the correlation between κ and the data. Interestingly, there is certain correlation between κ and streamflow in S1, in which streamflow data were not used. The reason might be the consideration of a more complete water cycle process in calculating the other meteorological data.

Parameter y_0 has no correlation in S2 but there is weak correlation in S1

between y_0 and the meteorological data. It is noteworthy that there is positive correlation in S3 between y_0 and ET , PET , and $NDVI$, the medians of which are >0.5 . Meanwhile, parameter y_0 has negative correlation with $P-E$ and R . Introducing streamflow data can activate the relationship between the assimilated parameter and other meteorological factors. The definition of y_0 is related to water available in storage, and its consideration represents a more integrated water cycle. However, its impact is indirect, because there is no obvious relationship between y_0 and storage water in S1 or S3, even if the distribution of the scatter plot data is closer in S3.

5. Conclusions

To detect rules of the intra-annual-scale variation of parameters in a changing environment, this study implemented three schemes that differed in terms of the use of observational data. Time-varying parameters of the Budyko function (evaporation calculation module) were identified using the EnKF method, and assimilated model performance assessment, correlation analysis, and time series analysis were performed. The principal conclusions derived are as follows.

(1) Plausible time-varying parameters of the Budyko function can be identified using the EnKF approach. Faithful evapotranspiration simulations of Scheme 1 (ET observations only) and Scheme 3 (combined ET and Q observations) can be reflected by criteria such as NSE , KGE , $RMSE$, and VE .

(2) Using only evaporation data in the assimilation can help identify a plausible set for parameter K , but it is inadequate for parameter y_0 and relationships regarding meteorological data. Introducing streamflow data into the assimilation based on ET data can improve efficiency. More obvious periodicity of the parameters (especially y_0) can be detected. Additionally, the results obtained when using only streamflow data were unacceptable.

(3) There exist 12-month periodicities in parameters κ and γ_0 . Obvious trends and change points of the parameters were detected in certain regions, e.g., Midwest CONUS, as detected in previous related research. The distribution of the best estimates of parameter γ_0 was found consistent with the hydrologic regions defined by the US Geological Survey. It was proved that changing conditions could be identified using the time-varying parameters of the Budyko function because of the physical significance of the parameters.

The findings imply that evaporation observations are adequate when comparing simple catchment attributes. However, streamflow observations are also necessary when focusing on additional processes such as soil water content. Further research should focus on attributing physical mechanisms to the parameters and applying these relationships to predictions.

Acknowledgments

This study was supported by the Joint Funds of the National Natural Science Foundation of China (Grant No. U1865201), National Natural Science Foundation of China (Grant No. 51861125102), and Innovation Team in Key Field of the Ministry of Science and Technology (grant no. 2018RA4014).

Code/ Data availability

Precipitation and streamflow data were obtained from the Model Parameter Estimation Experiment dataset (Duan et al., 2006), and potential evaporation and actual evaporation data were collected from (K. Zhang et al., 2010). Other data, models, or code used in the study are available from the corresponding author upon reasonable request.

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Table 1. Descriptions and ranges of parameters set in this study.

Parameters	Description	Interval and unit
κ	Watershed parameter	1.0-4.0 (-)
y_0	Evaporation/storage parameter	0.1-0.7 (-)
SC	Water storage capacity	100-2000 (mm)

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Table 2. Descriptions of the assimilated time-varying parameters, calculation modules and observational data of the designed experimental schemes.

Schemes	Description
Scheme 1	Only use two-parameter Budyko model, P, PET, <i>ET</i> data for DA
Scheme 2	Use two-parameter Budyko model together with monthly runoff calculation, P, PET, <i>Q</i> data for DA
Scheme 3	Use two-parameter Budyko model together with monthly runoff calculation, P, PET, <i>Q and ET</i> data for DA

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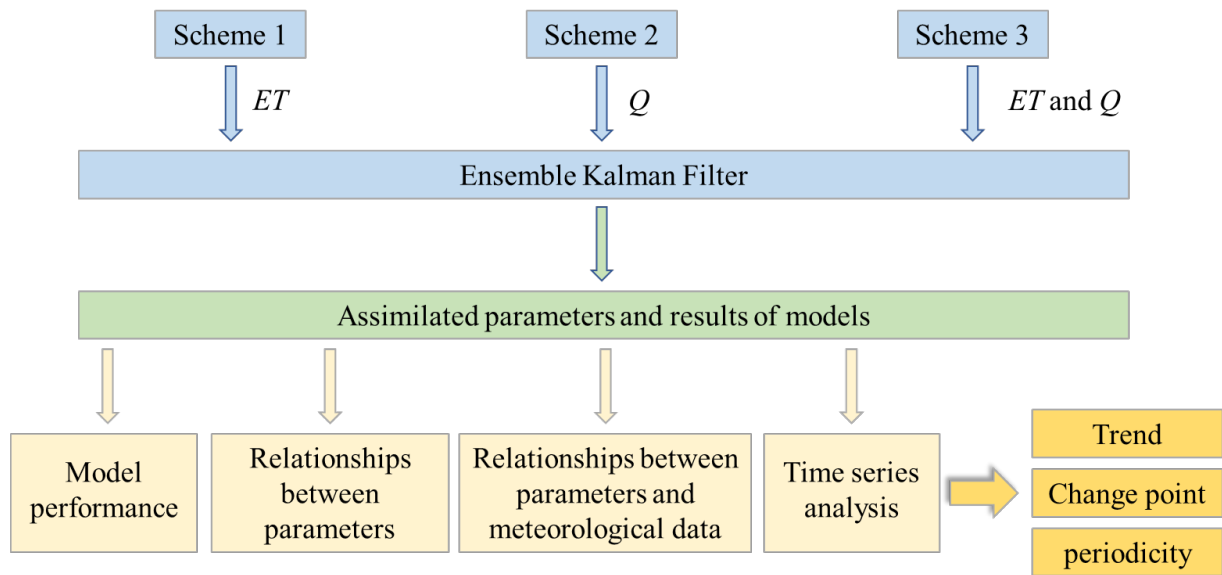


Fig. 1 Framework of the methodologies used and corresponding experimental design together with the structure of the analysis adopted in this paper. ET and Q next to the first row of arrows represent utilized observational data in Ensemble Kalman Filter (ET : evapotranspiration, Q : streamflow).

Miles

750 **Fig. 2** Locations of the 188 selected catchments and the corresponding aridity index
751 (*AI*) of humid catchments (blue) and (semi-)arid catchments (pale yellow).

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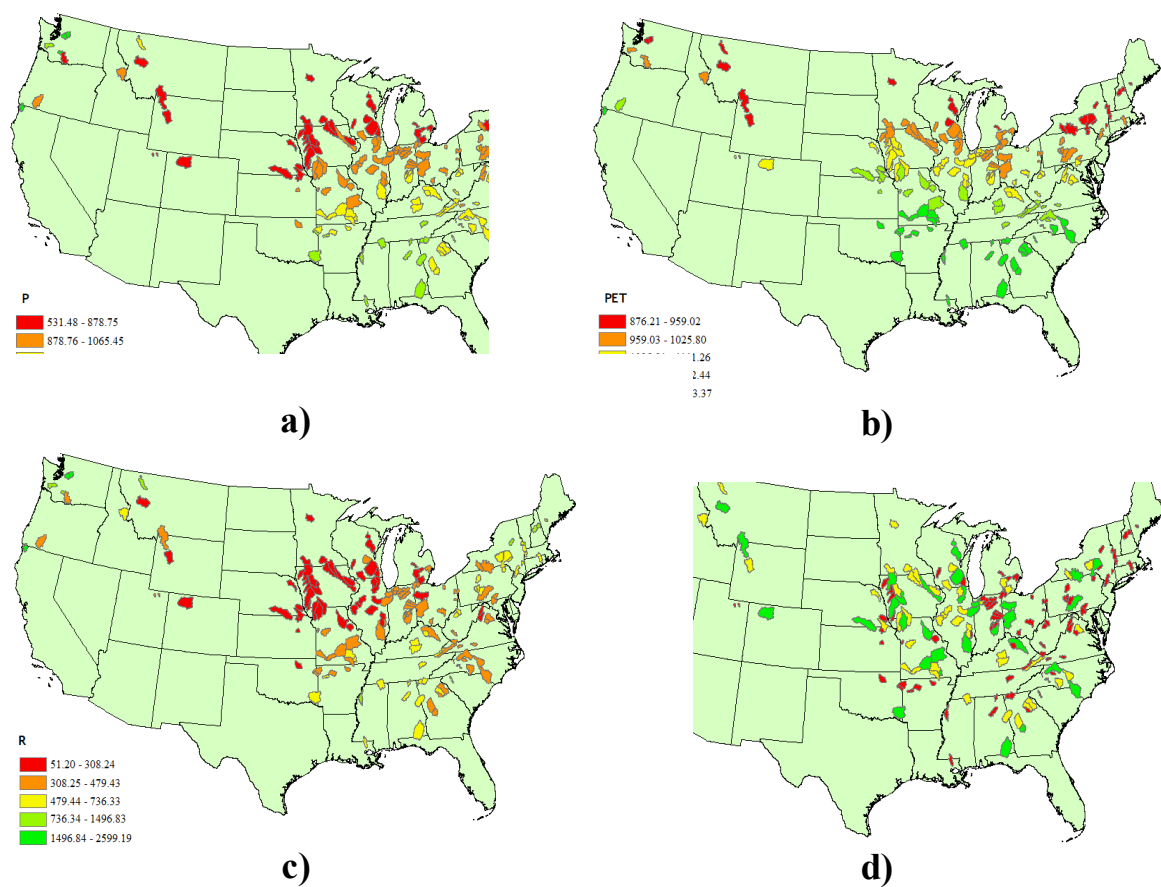


Fig. 3 Annual mean values of meteorological variables: **a)** precipitation, **b)** potential evapotranspiration, and **c)** runoff and **d)** the area of the 188 selected catchments.

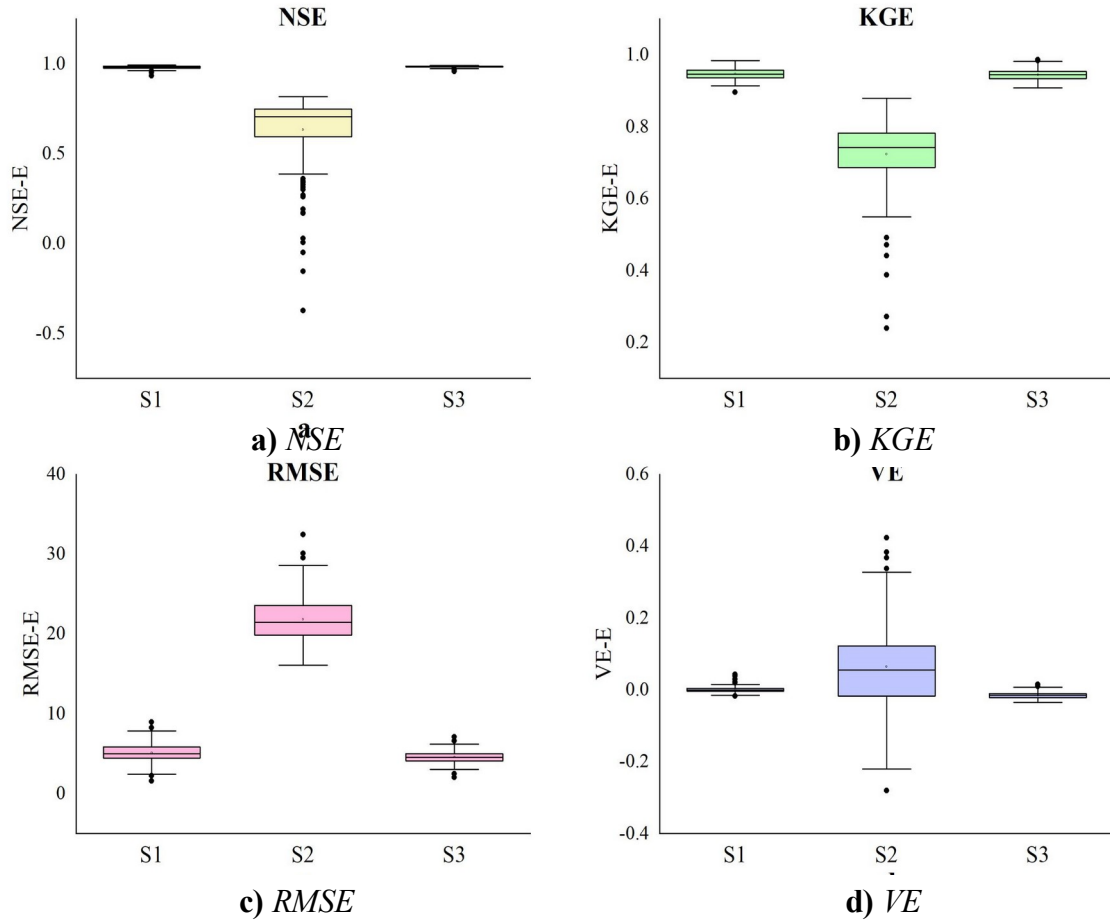


Fig. 4 Box figures of four evaluation criteria **a) NSE** **b) KGE** **c) RMSE** **d) VE** of *ET* simulations in the three schemes. In the plots, the whiskers represent the minimum and maximum values, the bottom and top of the boxes are the 25th and 75th percentiles, respectively, and the horizontal lines are the median values.

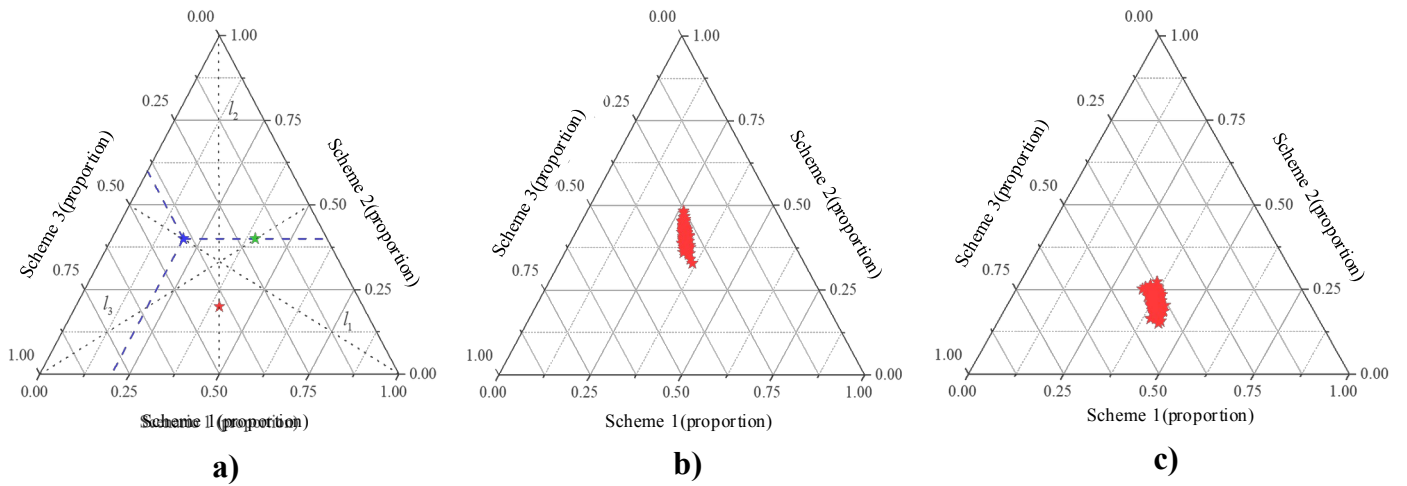


Fig. 5 Ternary diagrams of the mean values of the parameters in the three schemes. **a)**

Example in which the three stars represent three sets of data (each having three proportions of the parameter values). The proportions were calculated based on the ratio of the value in the corresponding scheme to the sum of the values in the three

schemes. Line l_1 indicates points at which the location of the blue star has the same proportion in S2 and S3, line l_2 is the same for S1 and S3, and line l_3 corresponds to S1 and S2. **b)** and **c)** the distribution of magnitude of κ and y_0 , respectively, where

the stars are situated on line l_2 .

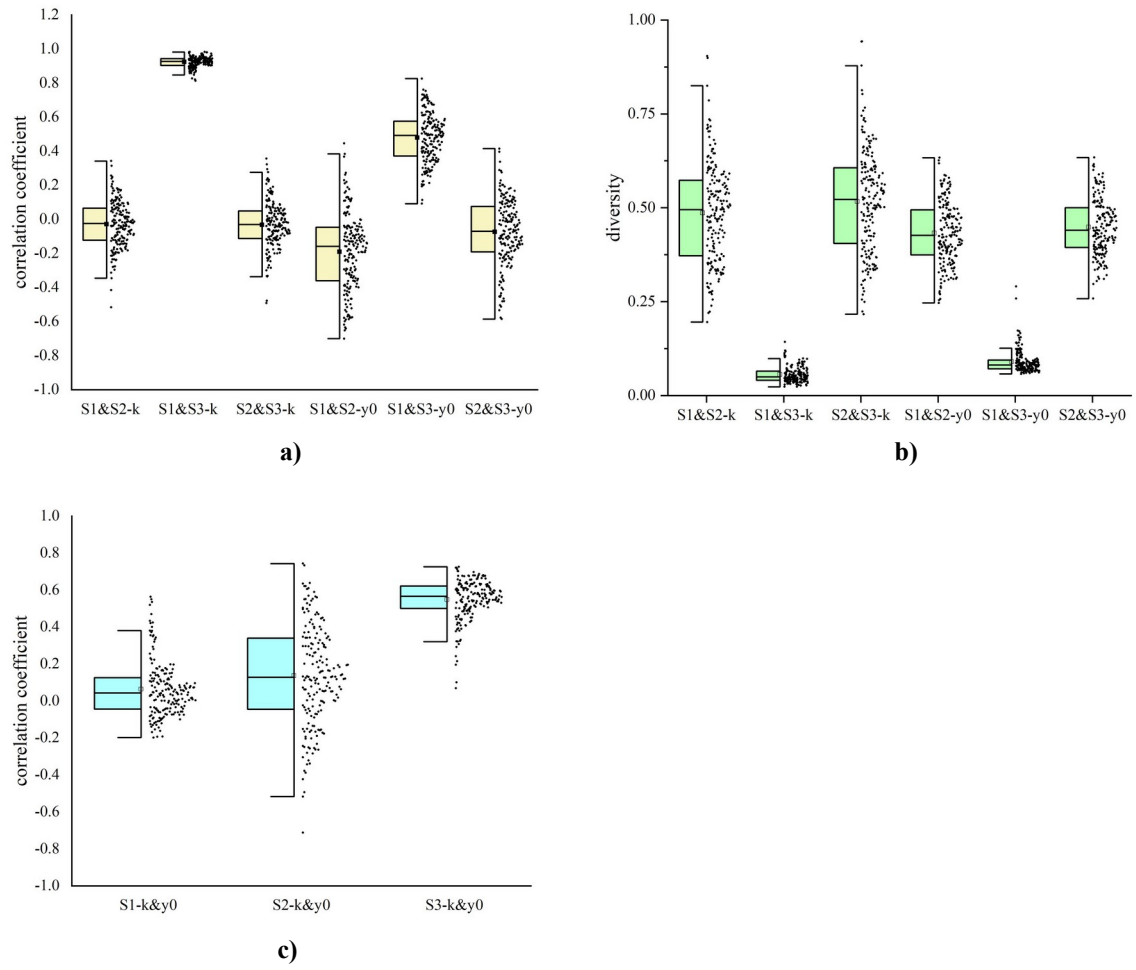


Fig. 6 Half-box plots of the correlation coefficients of **a)** κ and **b)** y_0 between the different schemes two by two, and **c)** half-box diagram of correlation between κ and y_0 in each scheme. In the plots, the whiskers represent the minimum and maximum values, the bottom and top of the boxes are the 25th and 75th percentiles, respectively, and the horizontal lines are the median values. The data scatter plot forms the right side of the half-box.

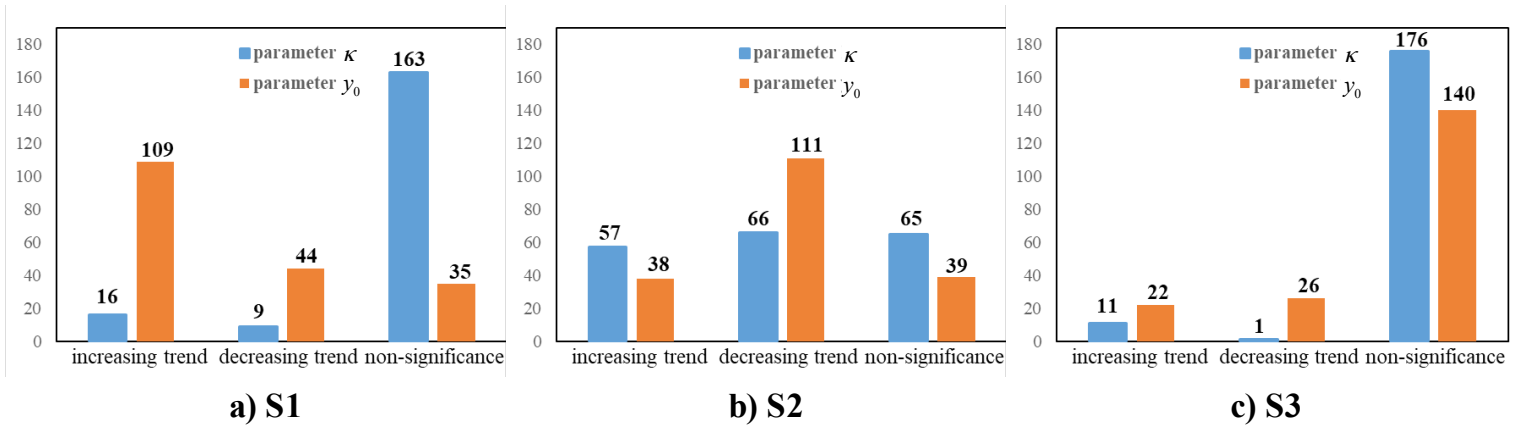


Fig. 7 Number of catchments with increasing, decreasing, and nonsignificant trends of the parameters (κ and γ_0) in **a) S1**, **b) S2**, and **c) S3**.

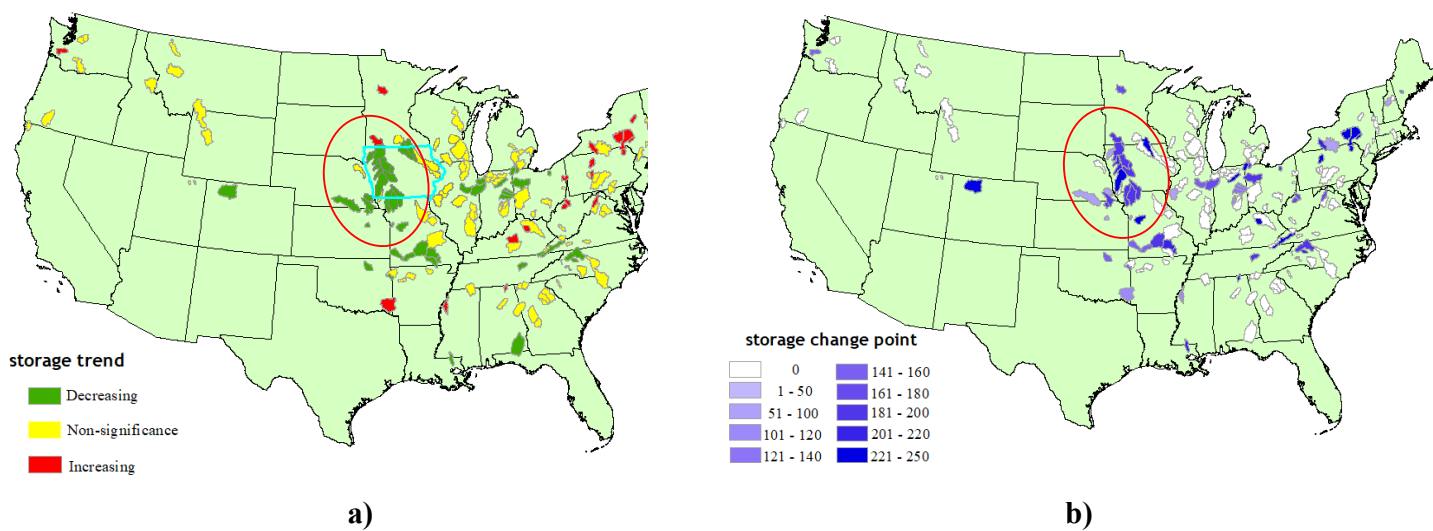
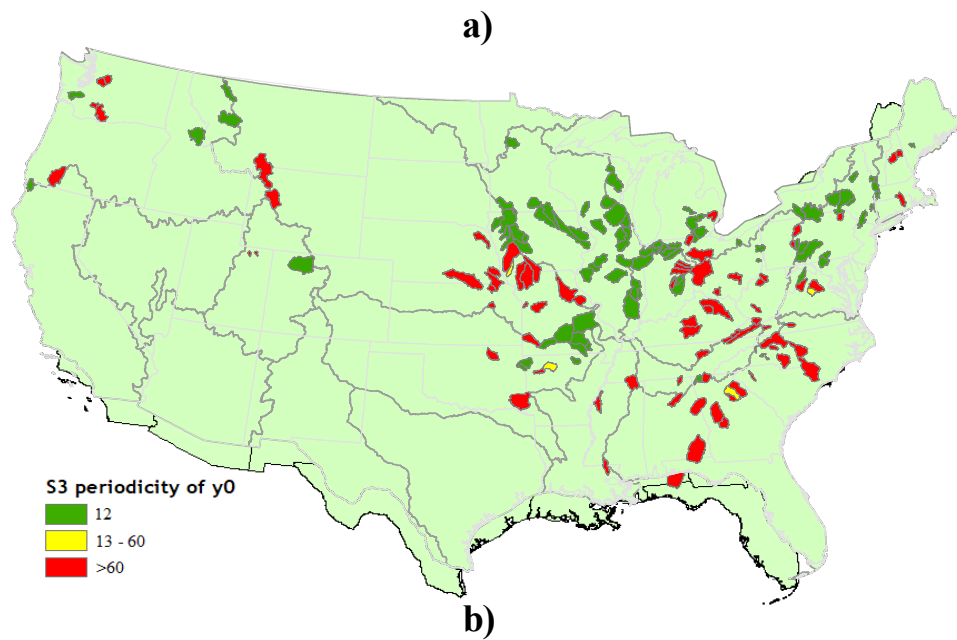


Fig. 8 Spatial distribution of **a)** trend and **b)** change point conditions of state variable *storage* in Scheme 3.



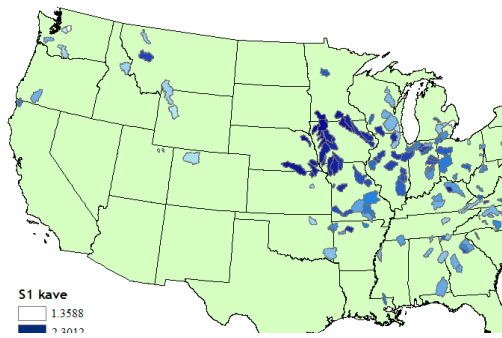
784

785 **Fig. 9 a)** US Geological Survey hydrologic regions with state boundaries where the786 number of each region is revealed. **b)** Periodicity of y_0 in Scheme 3, where

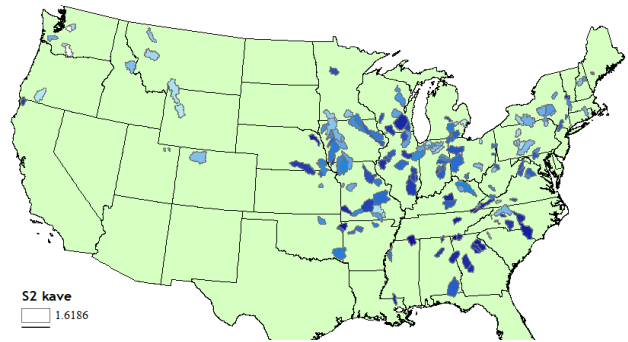
787 catchments colored green represent a period of 12 months, and the others generally

788 represent no period.

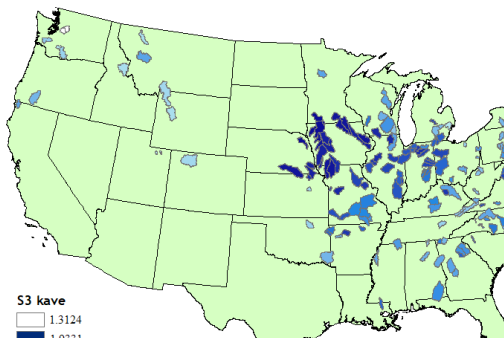
789



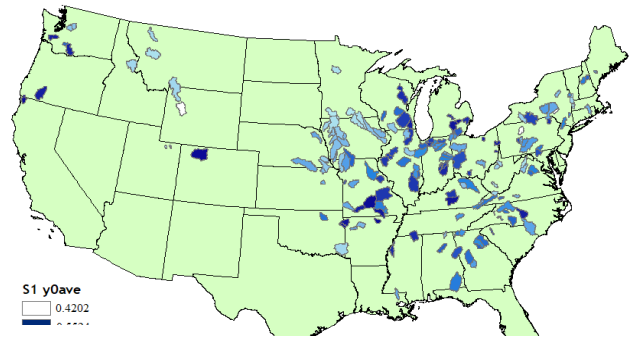
a)



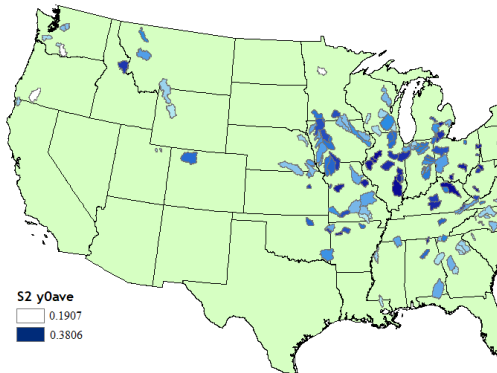
b)



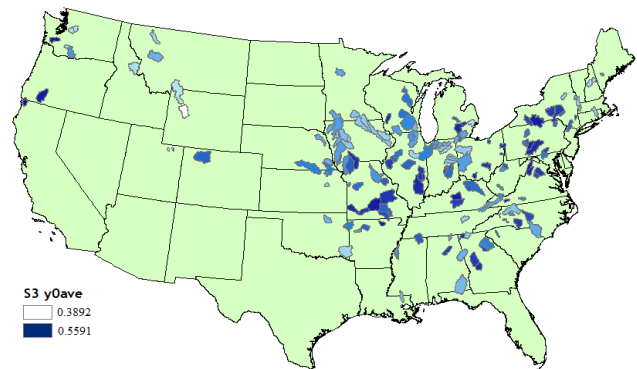
c)



d)



e)

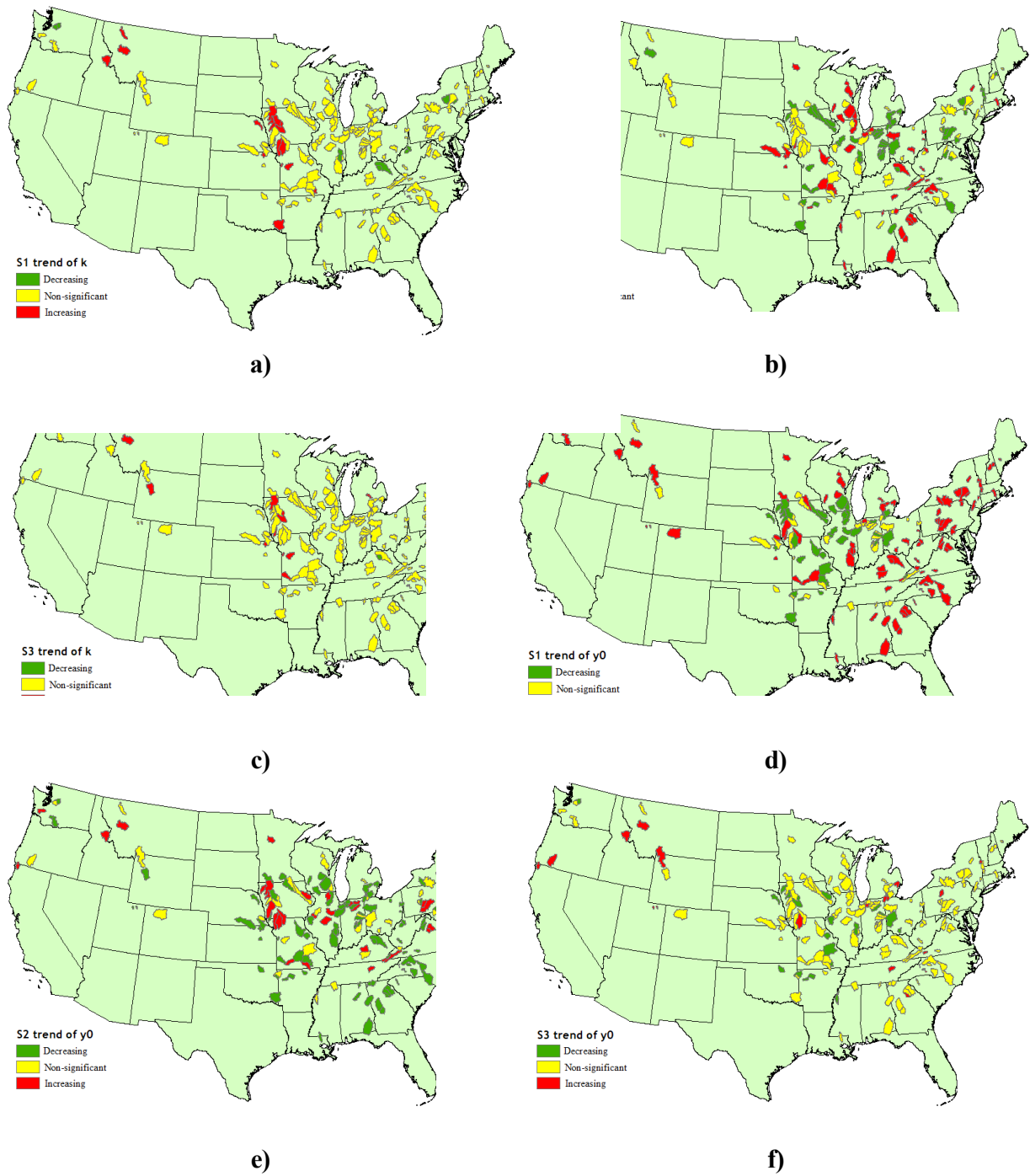


f)

790 **Fig. S1** Spatial distribution of mean values of parameters: **a) – c)** κ and **d) – f)** y_0 .
 791 **a)** and **d)** represent S1, **b)** and **e)** represent S2, and **c)** and **f)** represent S3. Color
 792 intensity indicates the magnitude of the parameter value. Accordingly, catchments
 793 with similar color have similar parameter values.

794

795



796 **Fig. S2** Spatial distribution of trends of parameters: **a) – c)** κ and **d) – f)** y_0 . **a)** and
 797 **d)** represent S1, **b)** and **e)** represent S2, and **c)** and **f)** represent S3. Red indicates an
 798 increasing trend, green means a decreasing trend, and yellow represents a
 799 nonsignificant result of the trend analysis.

800



Fig. S3 Overall number of change points of **a)** the two parameters, **b)** and **c)** κ , and

d) and **e)** y_0 . Color intensity represents the magnitude of the time step; a darker color indicates a greater number of change points with a large time step.

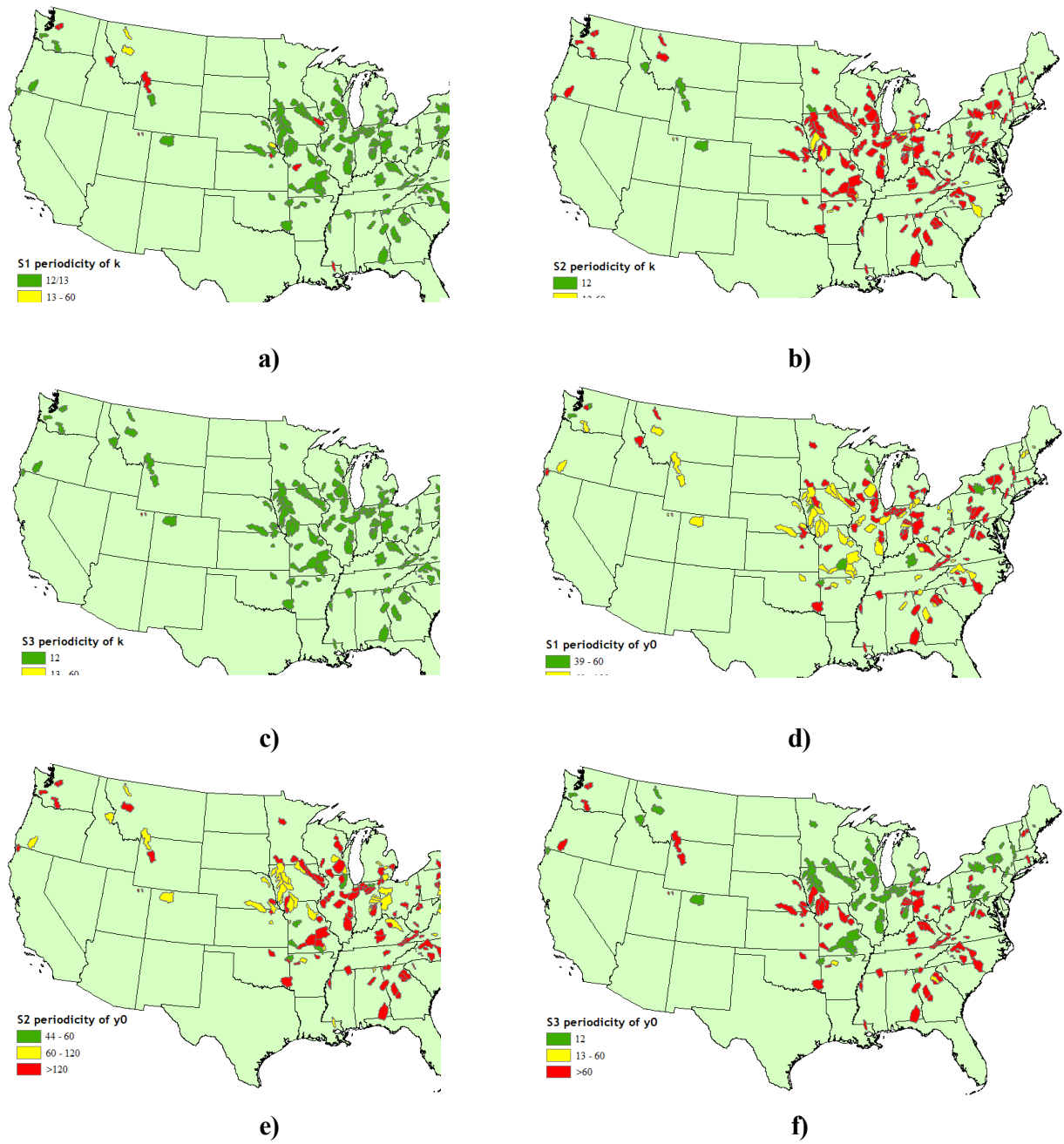


Fig. S4 Spatial distribution of the periodicity identified in the parameters: **a) – c)** k and **d) – f)** y_0 . **a)** and **d)** represent S1, **b)** and **e)** represent S2, and **c)** and **f)** represent S3. Colors of catchments indicate different periodicity corresponding to the legend at the bottom left of each panel.

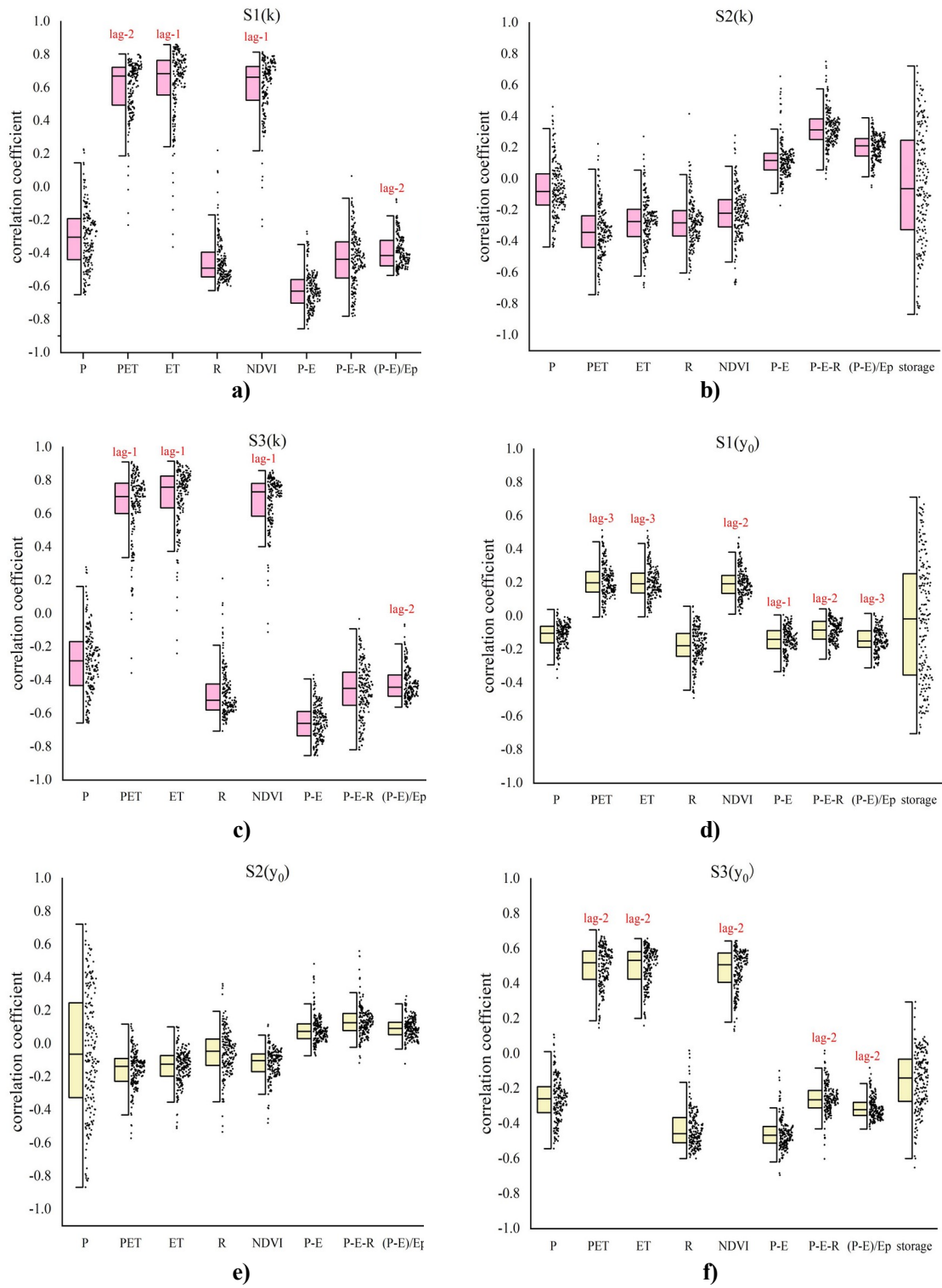


Fig. S5 Half-box plots of the correlation coefficient between meteorological data and parameters κ and y_0 in **a)** and **d)** S1, **b)** and **e)** S2, and **c)** and **f)** S3. Red labels represent the lag-1 to lag-6 and pre-1 to pre-3 mentioned in Section 4.4.