Estimating the CO2 fertilization effect on extratropical forest productivity from Flux-tower observations

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Abstract

The land sink of anthropogenic carbon emissions, a crucial component of mitigating climate change, is primarily attributed to the CO2 fertilization effect on global gross primary productivity (GPP). However, direct observational evidence of this effect remains scarce, hampered by challenges in disentangling the CO2 fertilization effect from other long-term drivers, particularly climatic changes. Here, we introduce a novel statistical approach to separate the CO2 fertilization effect on GPP and daily maximum net ecosystem production (NEPmax) using eddy covariance records across 38 extratropical forest sites. We find the median stimulation rate of GPP and NEPmax to be $16.4 \pm 4\%$ and $17.2 \pm 4\%$ per 100 ppm increase in atmospheric CO2 across these sites, respectively. To validate the robustness of our findings, we test our statistical method using factorial simulations of an ensemble of process-based land surface models. We acknowledge that additional factors, including nitrogen deposition and land management, may impact plant productivity, potentially confounding the attribution to the CO2 fertilization effect. Assuming these site-specific effects offset to some extent across sites as random factors, the estimated median value still reflects the strength of the CO2 fertilization effect. However, disentanglement of these long-term effects, often inseparable by timescale, requires further causal research. Our study provides direct evidence that the photosynthetic stimulation is maintained under long-term CO2 fertilization across multiple eddy covariance sites. Such observation-based quantification is key to constraining the long-standing uncertainties in the land carbon cycle under rising CO2 concentrations.

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Estimating the CO₂ fertilization effect on extratropical forest productivity from Flux tower observations

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11 Key Points:

- We present a novel statistical method to disentangle the variability of photosynthetic rates related to climate and non-climate drivers
 The analysis from 38 eddy covariance sites shows a 16.4 ± 4% increase in photosynthetic carbon uptake for a 100 ppm rise in atmospheric CO₂
 Our statistical method is successfully validated against idealized model simulations with
- 17 and without increasing CO₂

18 Abstract

The land sink of anthropogenic carbon emissions, a crucial component of mitigating climate 19 change, is primarily attributed to the CO₂ fertilization effect on global gross primary productivity 20 (GPP). However, direct observational evidence of this effect remains scarce, hampered by 21 challenges in disentangling the CO₂ fertilization effect from other long-term drivers, particularly 22 climatic changes. Here, we introduce a novel statistical approach to separate the CO₂ fertilization 23 effect on GPP and daily maximum net ecosystem production (NEP_{max}) using eddy covariance 24 records across 38 extratropical forest sites. We find the median stimulation rate of GPP and 25 26 NEP_{max} to be $16.4 \pm 4\%$ and $17.2 \pm 4\%$ per 100 ppm increase in atmospheric CO₂ across these sites, respectively. To validate the robustness of our findings, we test our statistical method using 27 factorial simulations of an ensemble of process-based land surface models. We acknowledge that 28 additional factors, including nitrogen deposition and land management, may impact plant 29 30 productivity, potentially confounding the attribution to the CO₂ fertilization effect. Assuming these site-specific effects offset to some extent across sites as random factors, the estimated 31 32 median value still reflects the strength of the CO₂ fertilization effect. However, disentanglement of these long-term effects, often inseparable by timescale, requires further causal research. Our 33 study provides direct evidence that the photosynthetic stimulation is maintained under long-term 34 CO₂ fertilization across multiple eddy covariance sites. Such observation-based quantification is 35 key to constraining the long-standing uncertainties in the land carbon cycle under rising CO₂ 36 concentrations. 37

38 Plain Language Summary

Through photosynthesis, plants convert CO₂ and water into sugars and oxygen using solar 39 energy, one of the most important chemical reactions on Earth. Human-made carbon emissions 40 are increasing increasing atmospheric CO_2 levels, impacting global photosynthesis. The 41 additional carbon is believed to have a fertilizing effect on photosynthesis, causing vegetation to 42 absorb a significant portion of the emitted CO₂. However, the strength of this CO₂ fertilization 43 effect on photosynthesis is uncertain, but is a crucial factor in determining the future trajectory of 44 atmospheric CO₂ concentrations. In this study, we introduce a new statistical method to quantify 45 46 the increase in photosynthetic carbon uptake, stimulated by rising CO₂, based on measurements 47 from 38 forest sites. Our results reveal that a 100 ppm increase in CO₂ enhances photosynthesis by approximately 16%. Validation of the statistical method with artificial model experiments 48

supports the robustness of our findings. Our study improves the understanding of the impacts of
 human-made CO₂ emissions on the global carbon cycle.

51 **1 Introduction**

Forests play a crucial role in the global carbon cycle, acting as a significant sink for 52 53 anthropogenic carbon emissions. Approximately 25% of annual carbon emissions are estimated to be sequestered and stored by forests via photosynthesis, with boreal and temperate forests 54 making substantial contributions (Pan et al., 2011). The physiological effects of increasing 55 atmospheric carbon dioxide (CO₂) on plant productivity, known as the CO₂ fertilization effect, 56 57 are expected to stimulate photosynthesis and drive the enhanced carbon uptake. However, obtaining observational evidence for these effects in natural ecosystems and understanding how 58 this process has changed historically remains a key knowledge gap. 59 Multiple lines of evidence support an enhancement in photosynthesis (or gross primary 60 production; GPP) in response to an increase in CO₂: CO₂ enrichment experiments (Norby et al., 61 2010; Walker et al., 2021), ecosystem monitoring (Keenan et al., 2013; Fernández-Martínez et 62 al., 2017; Mastrotheodoros et al., 2017) and indirect proxies based on long-term atmospheric 63 carbonyl sulfide records (Campbell et al., 2017) or isotopomer signal (Ehlers et al., 2015). 64 Process-based land-surface models, which simulate the physiological responses of vegetation to 65 environmental changes, also predict a stimulation of photosynthesis with increasing CO₂ levels 66 (Sitch et al., 2015). However, multi-model projections of the CO₂ effect on long-term GPP 67 diverge considerably due to uncertainties in process parameterizations and feedback 68 69 mechanisms, particularly in response to meteorological extremes and climatic changes associated to rising CO₂ (Zaehle et al., 2005; De Kauwe et al., 2013; Rogers et al., 2014). Constraining the 70 71 CO₂ fertilization effect in models through direct observational evidence is a long-called-for necessity to advance our understanding of carbon cycling and essential for more reliable future 72 73 projections of carbon sequestration. 74 The global network of eddy covariance (EC) flux towers observing the exchange of CO₂ at the

cosystem scale provides a valuable resource, as it has accumulated sufficiently long time series

data to potentially provide direct evidence of the CO₂ fertilization effect (Knauer et al., 2017;

⁷⁷ Baldocchi, 2020; Zhan et al., 2022). Previous studies have attempted to attribute the CO₂

fertilization effect on GPP by utilizing EC records. However, these studies have not adequately

accounted for confounded drivers of GPP, potentially leading to a misattribution, or only relied 79 on an indirect use of measurements for attribution, such as using EC data to calibrate a 80 predefined model. For instance, (Chen et al., 2022) used an eco-evolutionary optimality 81 framework to reproduce EC-inferred GPP and subsequently attribute the CO₂ fertilization effect 82 on GPP. Their analysis estimated a global GPP enhancement of 63 g C m⁻² yr⁻¹ from 2001-2014 83 due to rising CO₂. Similarly, (Ueyama et al., 2020) utilized a model constrained with data from 84 104 EC towers and estimated a 12.4 g C m⁻² yr⁻¹ increase in GPP. (Fernández-Martínez et al., 85 2017), employing generalized mixed linear models, attributed an increase of 11.2 ± 2.5 g C m⁻² 86 yr^{-1} in the GPP from 1995-2011 at 23 forest sites to CO₂. These variations emphasize the 87 importance of disentangling the CO₂ effect on GPP directly, *i.e.*, not using predefined model 88 structures, and by carefully considering confounding drivers in leveraging the continuously 89

90 growing EC records.

91 In this study, we aim to disentangle the CO₂ fertilization effect on photosynthetic uptake directly and exclusively from long-term multi-site flux measurements and accompanying meteorological 92 93 data. Several factors, such as CO₂, climate changes, land-use and land-cover changes, affect 94 ecosystem productivity and are correlated and confounded on the multi-decadal time-scale. We introduce the so-called GPP residual method that statistically captures the sensitivity of GPP to 95 CO₂ and climate variables at different time scales to account for co-linearities among the drivers. 96 97 First, we detrend the GPP time series to separate the long-term variability of GPP (the trend) primarily driven by CO₂ and climate, from the shorter-term variability (anomalies) primarily 98 driven by climate variabilities. The method estimates the sensitivity of GPP to climate, referred 99 to as the γ effect with temperature change as the proxy for climatic change, based on these 100 anomalies. Next, we can quantify the γ effect based on long-term changes assuming the 101 sensitivity remains consistent over the time scales of a few decades. The difference between the 102 original observed GPP trend and the inferred GPP trend due to climate changes yields the 103 unexplained GPP residual, which can be attributed to the long-term CO₂ effect on GPP, here 104 referred to as the β effect. Specifically, we define the β factor as the relative change in GPP (%) 105 per 100 ppm increase in CO₂, and the γ factor as the relative change in GPP (%) per Kelvin 106 increase in surface air temperature. It's important to note that the GPP residual may also include 107 other long-term effects specific to individual sites, such as signals related to nitrogen deposition 108 or land management. First, we validate the robustness of the GPP residual method using site-109

110 level simulations with the QUINCY model (QUantifying Interactions between terrestrial

- 111 Nutrient CYcles and the climate system; Thum et al., 2019). We use the GPP residual method to
- 112 further estimate the β and γ effects in both GPP and daily maximum net ecosystem production
- 113 (NEP_{max}) using long-term EC records. At last, we compare the results from EC records and

114 factorial simulations from a set of land-surface models (TRENDY version 9; Sitch et al., 2015).

- Additionally, we discuss the relevance of other potential long-term impacts on GPP in more
- 116 detail.

117 2 Materials and Methods

118 2.1 Eddy covariance data

119 This study comprises 38 forested eddy covariance (EC) sites (Table S1), where CO_2 flux data and meteorological data are collected by flux towers from Integrated Carbon Observation System 120 (ICOS; Rebmann et al., 2018), and AmeriFlux (Novick et al., 2018). We focus on tree-dominated 121 ecosystems as they exhibit less sensitivity to short-term climate variability compared to grass-122 dominated ecosystems. We focus on tree-dominated ecosystems due to the heightened sensitivity 123 of grass-dominated ecosystems to short-term climate variability, adding complexity to the 124 disentanglement of their response to CO₂ (Hovenden et al., 2014; Reich et al., 2018). The sites 125 span geographically across Europe and North America. The forest types can be broadly classified 126 into: deciduous broadleaf forest (DBF; 12 sites), evergreen need-leaved forest (ENF; 20 sites), 127 and mixed deciduous-coniferous forest (MF; 6 sites). 128

129 We obtain long-term recorded (≥ 10 years) eddy covariance data at daily scale and net ecosystem

production (NEE) at half-hourly scale from 1994 to 2022 (Table S2). GPP is estimated from the

131 nighttime partitioning algorithm (Reichstein et al., 2005). Meteorological variables used in this

132 study include temperature, incoming shortwave radiation and vapour pressure deficit (VPD).

133 Due to the limited depth of soil moisture measurements (Yu et al., 2022), we calculate a water

134 availability index (WAI; Tramontana et al., 2016) following a bucket model approach. The

135 maximum cumulative water deficit (Aragão et al., 2007) represents the available water content

136 (awc). WAI is calculated as the balance of precipitation recharge and observed

137 evapotranspiration as follows:

$$input(t) = \min(Precipitation(t), awc - WAI(t))$$
(1)

139
$$WAI(t+1) = \max\left(\left(WAI(t) + input(t) - Evapotranspiration(t)\right), 0\right)$$
(2)

140 Where t represents the timestep t. We exclusively include data with daily quality control 141 indicators for NEE and meteorological variables surpassing 0.6, denoting a 60% or higher 142 percentage of measured and high-quality gap-filled data. In summary, EC sites are selected for 143 this study based on three criterias: (a) only sites dominated by tree-ecosystems are selected; (b) 144 there has to be a long-term record of EC observations (≥ 10 years) after the quality control when 145 (c) at least 60% data per day is measured or gap-filled with good quality. In total we estimate β 146 and γ for 228 site-months.

Maximum leaf area index (LAI_{max}, m² m⁻²), forest age (years), site-specific disturbance or
management information are collected from the literature (Musavi et al., 2017; Besnard et al.,
2018; Flechard et al., 2020; Migliavacca et al., 2021), the BADM product, and/or site principal

150 investigators.

138

151 Due to the systematic biases of the atmospheric CO_2 measurements in the eddy covariance data,

152 for consistency, we replace the CO_2 measurement with the CO_2 product CAMS CF-1.6

153 (Chevallier et al., 2005, 2010) from the nearest pixel to each EC site. The CO_2 reanalysis data

154 spans from 1994 to 2022 with daily resolution, thus sufficient to match the time period of the 155 eddy covariance data.

156 2.2 Estimating β and γ using the *GPP residual method*

We develop the *GPP residual method* (Figure 1) to isolate the CO₂ fertilization effect (β) from 157 other confounding factors (e.g., climate). β is inferred for each site and each month-of-year 158 159 separately, using the median values of GPP and hydro-meteorological data across five-day intervals within the considered months to filter out synoptic weather variability and its impact on 160 GPP dynamics. The calculation of β consists of three steps: (1) Data preparation (Figure 1b). 161 The growing season when plant photosynthesis is active is defined based on the mean seasonal 162 cycle of GPP (averages by day-of-year) across the time series. A month is considered within the 163 growing season, if there are more than twenty days when GPP is greater than 25% of the 164

maximum of GPP as inferred from the mean seasonal cycle. Within each month, the median 165 values of all variables are retrieved for every five-day interval. We then calculate anomalies 166 using the median values by subtracting the long-term trend of the linear-fit for each month-of-167 year (e.g., July in 1999, ..., 2020). We rescale the anomalies of all variables by adding the 168 average value across the considered time period. The rescaling allows the random forest model 169 in step (2) is trained and applied at an identical magnitude to extrapolate; (2) Model training and 170 predicting climatic effects on GPP (Figure 1c). We train a random forest regression model for 171 GPP anomalies using anomalies of hydro-meteorological variables (*i.e.*, *temperature (Temp*), 172 incoming shortwave radiation (SW_{in}), vapor pressure deficit (VPD), water availability index 173 (WAI), growing degree days (GDD)). We use the model to predict GPP using the actual hydro-174 meteorological data (including both trends and anomalies) at each month-of-year. The predicted 175 GPP (GPP_{climatic}) thus only reflects the effect of climate. Next to the random forest regression 176 model, we call a multivariate linear regression model to test the robustness of the results from the 177 178 random forest regression model (Figure S1); (3) Isolating non-climatic effects on GPP (Figure 1d). The non-climatic effects on GPP (GPP_{residual}) are derived by removing the GPP_{climatic} from 179 the actual GPP time series ($GPP_{residual} = GPP - GPP_{climatic}$). $GPP_{baseline}$ is calculated as the mean 180 GPP over the first two years in the considered time series. The relative change of GPP in 181 response to CO_2 (*i.e.*, β) is derived as the trend of the linear-fit between CO_2 concentration and 182 GPP_{residual} in relative to GPP_{baseline}. Similarly, the sensitivity of GPP to temperature (*i.e.*, γ) is 183 184 derived as the trend of linear-fit between temperature and GPP_{climatic} in relative to GPP_{baseline}:

$$\beta = \frac{\frac{\Delta GPP_{residual}}{GPP_{baseline}}}{\Delta CO_2} \times 100\%$$
(3)

.

186
$$\gamma = \frac{\frac{\Delta GPP_{climatic}}{GPP_{baseline}}}{\Delta Temp} \times 100\%$$
(4)

The advantage of the *GPP residual method* is that we separate the confounding factors at different time scales, thus, it can solve the issue of multicollinearity to some extent, when the independent variables are highly correlated to one another. To show the different results yield from the *GPP residual method* and a multivariate regression method, we adopt a simple multivariate regression model as the following:

$$\frac{GPP}{GPP_{baseline}} * 100\% = \beta \times CO_2 + \gamma \times Temp + \rho_{VPD} \times VPD + \rho_{WAI} \times WAI + \rho_{GDD} \times GDD + \rho_{SW_{in}} \times SW_{in}$$
(5)

Where β is the sensitivity of GPP to CO₂, and γ is the sensitivity of GPP to temperature. After obtaining β values for each site-month, we calculate the median β . This approach helps to mitigate the influence of outliers. We further estimate the uncertainty of the median β using the bootstrap method. By repeatedly sampling from the considered β distribution, we create multiple bootstrap samples. Each sample is then used to calculate the median β . The standard deviation across these bootstrap estimates provides an estimate of the uncertainty associated with median β . We calculate the median γ and its uncertainty in the same way.

To consider the seasonal and spatial variation of GPP, we further calculate annual β and γ by aggregating monthly β and γ weighted by monthly

202 GPP_{baseline}:

203
$$Annual \beta = \sum_{i=1}^{n} \beta_{month_i} \times \frac{GPP_{baseline_i}}{GPP_{gs}}$$
(6)

Where i represents a specific month, and n is the total number of months. GPP_{gs} is the sum of baseline GPP across the considered growing season. Similarly, the annual β and γ at each site can be further aggregated across space:

207
$$mean \beta = \sum_{i=1}^{n} Annual \beta_i \times \frac{GPP_{gs_i}}{GPP_{sum}}$$
(8)

208
$$mean \gamma = \sum_{i=1}^{n} Annual \gamma_i \times \frac{GPP_{gs_i}}{GPP_{sum}}$$
(9)

209 Where i represents a specific site, and n is the total number of sites. GPP_{sum} is the sum of

baseline GPP across all sites. To assess the robustness of the median β or γ values and determine

211 if they are influenced by site selection, we compare the mean β or γ calculated across all sites,

weighted by baseline GPP, with the median β or γ derived from the distribution of monthly β or

 γ values. If the median value remains relatively stable and comparable to the mean value across

all sites, it suggests that the selection of sites does not significantly impact the robustness of the

215 median β or γ estimations.



216

Figure 1. Schematic of the statistical GPP residual method to isolate the CO₂ fertilization effect 217 and climatic effect in observational data of GPP. (a) Hypothesis. The overall goal of the GPP 218 residual method is to isolate the CO₂ fertilization effect on GPP by removing long-term climate 219 effects inferred from short-term variability. (b) Data preparation. All the time series of climate 220 variables and GPP are detrended and individually rescaled to the long-term mean of each 221 variable. The black lines denote the actual time series for each variable, and the red line denotes 222 the detrended time series. (c) A random forest model or multivariate linear regression model is 223 trained to learn the sensitivity of GPP to the climate variables based on the detrended time series, 224 i.e., interannual variability. (d) The trained model predicts the long-term changes in GPP caused 225 by climate changes using the original time series of climate predictors, including the long-term 226 trend. The non-climate-induced effect on GPP is therefore estimated from the residual of 227 absolute GPP minus climate-induced GPP, shown as the red line. 228

229 The EC technique allows for direct measurement of NEE, which is the difference between GPP

and ecosystem respiration (RECO):

$$NEE = RECO - GPP \tag{10}$$

231

- 232 The maximum net ecosystem production (NEP_{max}) is defined as the negative sign of the
- 233 minimum NEE during a day from half-hourly measurement:

$$NEP_{max} = -NEE_{min} \tag{11}$$

In addition to GPP, we further estimate β or γ for NEP_{max} following the same method.

236 2.3 Validate the *GPP residual method* with a land surface model

We use the terrestrial biosphere model QUINCY (QUantifying Interactions between terrestrial 237 Nutrient CYcles and the climate system; Thum et al., 2019), which has been evaluated against a 238 subset of FLUXNET sites across large geographical ranges and different ecosystem types, to 239 validate the GPP residual method. We perform two simulations with identical climate but 240 varying CO₂ concentrations (transient CO₂ as observed, and constant CO₂ at levels of the year 241 1988) at 166 forested sites distributed across the globe. The model setup and model simulations 242 are identical with the "freeze-CO₂ experiment" in Zhan et al. (2022). For better comparison with 243 eddy covariance records, we take the last twenty years (1999-2018) in the simulations as the time 244 period of the validation. 245

246 The advantage of this method is that we can compare the β estimated by the *GPP residual*

247 *method* ($\beta_{\text{estimated}}$) with β modeled by QUINCY (β_{QUINCY}), which is regarded as a surrogate truth

that represents theory of photosynthetic responses to CO₂, climate and water availability.

249 β_{QUINCY} is calculated as the sensitivity of the CO₂-induced change in GPP to CO₂ concentration,

250 in which GPP is calculated as the difference between simulations forced with transient CO_2 and

251 constant CO₂ during the considered time period. β is calculated for each site and each month-of-

252 year. The selection of months follows the same rule as the data preparation in the previous

253 section. $\beta_{\text{estimated}}$ is calculated using the *GPP residual method* elaborated in the previous section

from the simulation forced with transient CO₂.

255 We evaluate the agreement of $\beta_{\text{estimated}}$ and β_{QUINCY} for each forested site (166 sites in total) in

the model. We use the root-mean-square error (RMSE) to measure the difference between

257 $\beta_{\text{estimated}}$ and β_{QUINCY} across the growing season. The RMSE of β estimation per site is calculated

258 as:

259
$$RMSE(\beta) = \sqrt{\frac{\sum_{i=1}^{n} (\beta_{estimated_i} - \beta_{QUINCY_i})^2}{n}}$$
(10)

where n is the number of months when β is estimated; $\beta_{estimated_i}$ and β_{QUINCY_i} is the β estimated and modeled at month i, respectively. In this study, we use the validated *GPP residual method* to estimate β in tree-dominated ecosystems based on measured meteorological data and the CO₂ atmospheric inversion product.

264 2.4 β and γ determined from the TRENDY v9 ensemble

265 We use simulations from twelve process-based global dynamic vegetation models (DGVM) within the TRENDY projects (Sitch et al., 2015; Le Quéré et al., 2018) to derive the modeled β 266 and γ . We use four simulations (called S0, S1, S2, S3 in the TRENDY v9 protocol; see Table S3) 267 with and without land use and land cover changes (LULCCs) under both transient (historically 268 observed) and pre-industrial (constant) environmental conditions. CO₂ effect on GPP modeled by 269 TRENDY (β_{S1-S0}) is calculated as the difference between output from S1 and S0, to avoid the 270 effect from climate recycling. To test the robustness of the GPP residual method and the 271 potential LULCCs effect, we apply the same statistical method (*i.e.*, the GPP residual method) 272 on simulations in S2 and S3, respectively. We derive γ from S2 simulations by calculating the 273 sensitivity of GPP in S2 to temperature. We select grid-cells containing the 38 considered eddy 274 covariance sites in all models. β and γ is calculated for the same site-months as data analyzed in 275 276 EC records.

277 **3 Results and Discussion**

278 3.1 Evaluating the *GPP residual method* with a land surface model

279 We develop and test the *GPP residual method* with QUINCY model simulations (Methods) from

which we cannot only infer β with our method, but also directly compare with the modeled β .

281 Overall, we find that our method can satisfactorily estimate β , and can capture the seasonal

variations of β across biomes. We find β in tropical forest is overall well reproduced by our

statistical method, supported by a mean root-mean-square error (RMSE) of 1.2 gC/m²/year/ppm.

However, the performance in the cold northern high latitude regions, where part of the boreal

needle leaf forests and temperate forests are located, is slightly diminished, with a mean RMSE 285 of 1.5 gC/m²/year/ppm (Figure 2a, b). The *GPP residual method* with implemented random 286 forests exhibits better estimation in temperate broadleaf summergreen trees (TeBS), compared 287 with both multivariate regression methods that estimate negative β in TeBS during the beginning 288 and the late growing season (Figure S1, S2). Additionally, we find a slight overestimation 289 accompanied with higher RMSE, during summer months in boreal needleleaf evergreen (BNE) 290 forested sites and boreal needleleaf summergreen (BNS) forested sites. The discrepancy between 291 292 estimated β and modeled β can be attributed to the limitations associated with constructing a statistical model to estimate the sensitivity of GPP to climate variables relying on interannual 293 variabilities. This means this statistical model does not account for vegetation acclimation on 294 climatic variability in the long-term, such as phenological changes, which cannot be learned 295 296 from interannual variabilities. Thus, the statistical method exhibits decreased accuracy, particularly in ecosystems where seasonality exerts strong control. 297

In addition to the limitation of capturing vegetation phenology, we individually consider the 298 effect of rising CO₂ and the effect of changing climatic conditions. Thus, the synergetic effect of 299 rising CO₂ and temperature (Drake et al., 1997) is not considered in our approach, where *e.g.*, 300 increasing CO₂ can modify plants' response to temperature. This simplification could result in 301 the overestimation of the CO₂ fertilization effect on GPP. On the other hand, the anomalies 302 303 associated with extreme events can be theoretically reproduced by the statistical method. However, given that only a few instances of extreme events are in the training dataset, we 304 acknowledge that the non-linear relationship between climate and GPP during extreme 305 conditions can induce errors in the estimation of β . Overall, we find encouraging consistency 306 between the β estimated by the *GPP residual method* and β modeled by QUINCY. 307



308

Figure 2. Validation of the *GPP residual method* with QUINCY model simulations. (a) Seasonal variation of β across vegetation types estimated by the *GPP residual method* with a random forest model in red and QUINCY in black (166 sites with 1220 site-months). The red and black shaded area depicts one standard deviation around the mean value of β across multiple sitemonths (solid lines). (b) The map shows the root mean square error (RMSE) between estimated β and modeled β in the growing season for each site in the QUINCY model. Brighter color indicates lower bias and thus a better performance of the *GPP residual method*.

316 3.2 CO₂ fertilization effect in forested ecosystems inferred from eddy covariance records

317 Using the *GPP residual method*, we estimate the strength of the CO₂ fertilization effect on

- 318 photosynthetic carbon uptake as recorded in eddy covariance (EC) time series at 38 forested sites
- (Figure 3). We assess the sensitivity of GPP to CO_2 , denoted β (Methods), separately for each
- 320 individual month across the years of the time series to account for seasonal variations. The
- 321 median β value across all sites and months is 16.4 ± 4.1 % increase in GPP per 100 ppm rise in
- 322 atmospheric CO₂. While β displays considerable variability across sites and months, positive β
- values are consistently observed in 61 % of sites for at least two months in the record. The
- 324 strongest enhancement of β occurs during the boreal summer months, although a selection of
- sites exhibits stronger effects in spring (*e.g.*, *CA-Ca3*, *IT-Ren*) or autumn (*e.g.*, *US-GLE*, *FR- FBn*).
- Among the analyzed sites, the top seven sites listed in Figure 3a (DE-Hzd, CA-LP1, CA-Ca3,
- 328 CA-Cbo, US-GLE, IT-Ren, CA-TP4) exhibit the most pronounced GPP enhancement, with their
- 329 site-specific annual mean β values surpassing the top 20% of all the sites. In addition to the
- median β value across all sites and months, we aggregate monthly β based on the monthly
- baseline GPP, to represent the mean β across selected sites. The aggregated mean β is 14.9 %
- per 100 ppm, indicating the median β is representative and not biased by the site selection,
- 333 considering the variation in GPP across sites.
- The presented approach lacks the ability to isolate additional hidden long-term effects stemming 334 from human activities, and these effects may manifest in an over- or under-estimation of β 335 derived at individual sites. Notably, certain sites (e.g., DE-Obe, CH-Dav, FI-Let, DE-Hai) 336 exhibit negative β values consistently throughout the growing months. The negative β identified 337 338 at the "CH-Dav" site may be associated with a disturbance event, specifically a harvest conducted in the year 2006. Similarly, the thinning activity at the "FI-Let" site in 2016 induced a 339 declining of GPP trend, leading to a negative β estimate throughout the year. Forest in the "DE-340 Hai" site is recoving from a sever drought. Conversely, other drivers such as nitrogen deposition 341 at nitrogen-limited sites (De Vries et al., 2006; Sutton et al., 2008; de Vries et al., 2014) or 342 forests undergoing succession (Pugh et al., 2019) can induce a long-term increase of GPP, 343 potentially resulting in an overestimation of β . However, these rather site-specific trends in 344 opposing directions may offset each other within a well-distributed and sufficiently large sample 345 size, enabling the median β across sites to predominantly capture the widespread CO₂ 346 347 fertilization effect. This notion is supported by excluding known disturbance sites (e.g., forest

thinning) from the analysis, resulting in a median β (11.9 ± 4.1 % per 100 ppm) that does not significantly differ from the β estimated using all sites (Figure 3b).



350

Figure 3. Estimation of β from eddy covariance data using the *GPP residual method* with a 351 random forest model. (a) Plot showing the estimated β for each eddy covariance site across 352 months in the growing season. The size of circles represents the magnitude of monthly baseline 353 GPP. Sites are shown in descending order of the annual mean β (Methods). Site-codes marked 354 by a star are presented separately at the end of the list, indicating that disturbances have been 355 recorded at those specific sites. Site-codes shown in blue and black color locate in Europe (c) 356 357 and North America (d). (b) The histogram of β values shown in panel (a). The grey (yellow) vertical dashed lines denote the median β (Methods) estimated from all sites and months 358 (excluding the disturbed sites). The grey (yellow) shaded area indicates the bootstrap estimates 359 for the uncertainty of median β from all sites (excluding the disturbed sites). Maps (c) and (d) 360 display the annual mean β values at each site. 361

- 362 The daily maximum net ecosystem production (NEP_{max}) provides insights into the peak
- 363 photosynthetic activity of the ecosystem during optimal conditions within a day. It is valuable for
- understanding the ecosystem's contribution to carbon sequestration. In addition to GPP, we
- identify the CO₂ fertilization effect on NEP_{max} as 17.2 ± 3.6 % per 100 ppm (Figure 4). The
- temporal and spatial variation of β in NEP_{max} is consistent with β in GPP (Figure 3), adding
- additional observational evidence of the CO₂ fertilization effect for better understanding of the
- 368 global carbon cycle dynamics.



Figure 4. Estimation of β from eddy covariance data using the *GPP residual method* with a random forest model. Analogous to Figure 3 but β is estimated for NEP_{max}.

Next, we assess the robustness of our findings by testing multiple regression methods in 372 estimating the GPP sensitivity to climatic changes, and evaluate their statistical performance. 373 374 Testing a multivariate linear regression instead of a random forest regressor, we find that the median β yields a slightly different estimate of 15.1 ± 5.0 % per 100 ppm (Figure S3). If we 375 376 however apply the multivariate regression model without accounting for confounding drivers of rising CO₂ and climatic changes (Methods), the median β is notably lower and amounts to 11.4 ± 377 5.6 % per 100 ppm (Figure. S4). We utilize the "Out-of-Bag" (OOB) score to estimate the 378 performance of the random forest regressor on unseen data without the need for a separate 379 validation set (Methods). Although there are instances of relatively low OOB score at certain 380 sites and months, no clear relationship emerges between estimated β values and model 381 performances (Figure S5). 382

383 3.3 Exploring the spatial variation of the CO₂ fertilization effect

369

We further explore the spatial variability in estimated β . Thereby, we assess the roles of plant

functional types (PFTs), forest age, temperature, VPD and maximum leaf area index (LAI). Past

386 studies have found stronger stomatal responsiveness to changes in CO₂ in deciduous trees versus

conifers (Saxe et al., 1998; Medlyn et al., 2001; Brodribb et al., 2009; Klein and Ramon, 2019),

although variability exists when assessing their different responses in photosynthesis (Saxe et al., 388 1998). Overall, we find a greater enhancement in GPP in evergreen needle-leaved forest (ENF) 389 390 in response to increasing CO₂ (Figure S6), compared with deciduous broadleaf forest (DBF). The difference in GPP responses to increasing CO₂ across PFTs may vary with scales, or complex 391 environmental conditions (e.g., under stress or not). Future work may focus on this difference in 392 more detail. An open question is whether mature forests, which may be approaching a quasi-393 equilibrium state, are responding to CO₂ and climate in the same fashion as younger stands (Kira 394 and Shidei, 1967; Odum, 1969; Luyssaert et al., 2008). We find no significant relationship 395 between β and forest age, but we show a tendency of GPP enhancement to decline towards older 396 stands (Figure S7a). Theoretically, the enhancement in GPP would relate to differences in 397 growing season temperature and VPD, with greater enhancement at warmer growth temperatures 398 399 (suppression of photorespiration; Baig et al., 2015). However, we find no significant but slight negative trend in the relationship between β and temperature as well as VPD (Figure S7b, c). 400 This trend may be attributed to the combined impact of temperature and VPD. We find a positive 401 tendency of β with increasing site maximum LAI (Figure S7d), which could be an interaction 402 between rising CO₂ and exponential growth phase (*i.e.*, *regrowth*). 403

3.4 Comparing the CO₂ fertilization effect inferred from eddy covariance sites and TRENDY
 ensemble

406 We compare our EC based β estimates with an ensemble of twelve process-based global

407 dynamic vegetation models (DGVM) following the TRENDY simulation protocol (Sitch et al.,

408 2015; Le Quéré et al., 2018). The TRENDY ensemble consists of four experiments (Table S3):

409 the pre-industrial control run (S0), the run considering only CO₂ changes (S1), the run

410 considering CO₂ and climate change forcings (S2), and the latter with additional prescribed land-

411 use and land-cover changes (S3). To conduct the comparison, we extract model time series from

412 the individual grid-cells containing the 38 considered eddy-covariance sites. The modeled CO₂

413 fertilization effect inferred by calculating the difference β_{S1-S0} , exhibits a large spread among the

414 TRENDY models. The median β_{S1-S0} across the grid-cells and models is 17.7 % per 100 ppm,

415 which is remarkably close to the median β obtained through the *GPP residual method* using EC

416 records. Seven models (VISIT, CLM5.0, JULES-ES-1p0, CLASSIC, ISAM, ORCHIDEE and

417 LPJ-GUESS) fall within the bootstrapped uncertainty range of the median β estimated from

418 observations (Figure 5; Methods). We acknowledge the limitation of this comparison, as it

- 419 involves contrasting site-level estimates with grid-level results, which is influenced by the
- 420 heterogeneity within each grid-cell. Nevertheless, we argue that the median values across the
- sites and grid-cells provide a more aggregated perspective that helps mitigate the influence of
- 422 sub-grid heterogeneity.



423

Figure 5. Comparing β estimated from eddy covariance data and the TRENDY model ensemble. The medians and interquartile ranges of β are shown for each model and for the ensemble mean, as horizontal lines within the boxes, and the upper and bottom lines of the box, respectively. Each box includes grid-cells containing the 38 considered eddy covariance sites. Box plots for individual models are in an ascending order based to the median β . The dotted red line represents the median β derived from eddy covariance records (as shown also in Figure 3), with the uncertainty showing in shaded area.

431 Other factors, such as nitrogen deposition, disturbances, and particularly land management, can

432 influence ecosystem productivity as recorded in EC data. These factors potentially influence the

- estimation of β using the *GPP residual method*. To assess the effect of land-use and land-cover
- 434 changes (LULCCs) on β estimation, we compare β derived from the TRENDY S2 and S3
- simulations using the GPP residual method. With the exception of the LPJ-GUESS and JULES-
- 436 ES-1p0 models, the ensemble mean of β_{S2} closely aligns with β_{S3} (Figure 5), suggesting that the
- 437 neglected effects of LULCCs do not substantially affect the β estimation. Furthermore, in line

- 438 with the validation using QUINCY simulations, the *GPP residual method* tends to slightly
- 439 overestimate β when comparing β_{S2} with β_{S1-S0} derived from TRENDY. This further emphasizes
- that the method cannot account for long-term vegetation acclimation and phenological changes;
- 441 however, these effects are minor within the considered time period.
- 442 3.5 Influence of climatic changes on productivity throughout the season
- Conventionally, the γ factor is defined as the sensitivity of land carbon storage to climate 443 variations using temperature change as the proxy (Friedlingstein et al., 2006; Gregory et al., 444 2009; Arora et al., 2020). Analogously, we define γ as the relative change in the climate-driven 445 GPP component over temperature change, which is already obtained in the GPP residual method 446 (Methods). The median γ in GPP estimated from the EC dataset is 1.8 ± 0.2 % per Kelvin 447 (Figure 5). Comparing this to the sensitivity of GPP to CO₂, assuming a 100 pm increase in 448 atmospheric CO₂ concentration is roughly equivalent to 1 Kelvin temperature increase in the 449 historical period, we find that γ is considerably lower than β , in line with previous studies 450 (Fernández-Martínez et al., 2017; Chen et al., 2022). The median γ in NEP_{max} (-0.2 ± 0.2 % per 451 Kelvin) is much lower than γ in GPP, reflecting a negative response of NEP_{max} to temperature 452 variations, particularly at the peak of the growing season (Figure S8). The median γ estimated 453 from the TRENDY ensemble (S2 simulations) is 0.8 ± 0.2 % per Kelvin (Figure S9). Also, γ 454 exhibits a large spread among models compared to γ from EC, suggesting a more pronounced 455 uncertainty in the process representation in estimating ecosystem responses to climate changes 456 among the land-surface models (Figure S10). A clear seasonality of γ emerges in both 457 observations and models (Figure 5, Figure S8, Figure S9). While γ is higher at the beginning and 458 the end of the growing season for most of the sites, most sites show negative γ in at least one 459 460 month of the growing season (26 out of 38 in the EC estimated GPP; 36 out of 38 in the EC measured NEP_{max}; 33 out of 38 in the TRENDY ensemble mean). Our results indicate that 461 warming may have a positive effect on vegetation productivity at colder conditions and a 462 potential negative effect in warm climate. A high temperature is usually accompanied by a high 463 VPD, which limits the stomatal conductance and evapotranspiration (Park Williams et al., 2013; 464 Novick et al., 2016). On the other hand, plant productivity response to temperature is associated 465 with water availability; a positive effect of temperature can occur when water is not limiting the 466 467 ecosystem functioning (Fernández-Martínez et al., 2019). Additionally, the negative climate-

carbon feedback may contribute to, and potentially partly mask or enhance the CO₂ fertilization
effect.



Figure 6. Estimation of γ from eddy covariance dataset using the *GPP residual method* with a 471 random forest model. (a) Plot showing the estimated γ for each eddy covariance site across 472 months in the growing season. The size of circles represents the magnitude of monthly baseline 473 GPP. Sites are shown in descending order of the annual mean γ (Methods). Site-codes marked by 474 a star are presented separately at the end of the list, indicating that disturbances have been 475 recorded at those specific sites. Site-codes shown in blue and black color locate in Europe (c) 476 and North America (d). (b) The histogram of γ values shown in panel (a). The grey (yellow) 477 vertical dashed lines denote the median γ (Methods) estimated from all sites and months 478 (excluding the disturbed sites). The grey (yellow) shaded area indicates the bootstrap estimates 479 for the uncertainty of median γ from all sites (excluding the disturbed sites). Maps (c) and (d) 480 display the annual mean γ values at each site. 481

470

Overall, we recognize the inherent limitations in EC-based data acquisition, the assumptions of 482 the GPP residual method, and the potential influence of other long-term factors such as human 483 activities, which can introduce biases in the estimation of β and γ from observations. However, 484 despite these challenges, the β and γ values estimated from the EC records align with the β and γ 485 values simulated by the TRENDY model, both in terms of the median values across multiple 486 sites and the seasonality of γ . Furthermore, utilizing the TRENDY simulations, we demonstrate 487 that the discrepancy in β estimation, influenced by land use and land cover changes, remains 488 489 within an acceptable range.

490 **4 Conclusions**

Our study isolates a robust, multi-decadal enhancement in vegetation productivity ($\beta_{GPP} = 16.4$ 491 ± 4 % per 100 ppm, $\beta_{NEP_{max}} = 17.2 \pm 4$ % per 100 ppm) across Northern Hemisphere forests in 492 response to the rising atmospheric CO₂ concentration. We further diagnose the median value of 493 GPP sensitivity to temperature (γ) of 1.8 ± 0.2 % per Kelvin, and find evidence of a negative 494 495 effect of temperature on photosynthesis at the peak of the growing season. Assuming a 100 pm increase in CO₂ concentration is equivalent to 1 Kelvin temperature increase, the negative 496 temperature effect potentially masks the positive increasing CO₂ effect on GPP. While the 497 TRENDY ensemble captures the median β and γ inferred from eddy covariance records, there is 498 a notable variation in β and γ exhibited between individual models. To improve the 499 500 representation of the temperature effect on GPP in TRENDY models, further studies should 501 focus on reducing uncertainties associated with physiological processes and incorporating multimodel constraints. This study paves the way for future investigations into long-term drivers of 502 change and ecosystem functioning. We anticipate that our approach could be readily applied to 503 other ecosystems (e.g., drylands), other datasets (long-term satellite records of change, i.e., 504 vegetation greenness, etc.), and other variables that describe ecosystem function (e.g., 505 evapotranspiration). Together with a wide range of evidence of global vegetation productivity in 506 507 response to rising atmospheric CO₂ concentration and climate change, a better estimation of the remaining carbon budget to achieve the climate goals in the Paris Agreement is possible. 508

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