Biases in estimating long-term recurrence intervals of extreme events due to regionalised sampling

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Abstract

Preparing for environmental risks requires estimating the frequencies of extreme events, often from data records that are too short to confirm them directly. This requires fitting a statistical distribution to the data. To improve precision, investigators often pool data from neighboring sites into single samples, referred to as "superstations," before fitting. We demonstrate that this technique can introduce unexpected biases in typical situations, using wind and rainfall extremes as case studies. When the combined locations have even small differences in the underlying statistics, the regionalization approach gives a fit that may tend toward the highest levels suggested by any of the individual sites. This bias may be large or small compared to the sampling error, for realistic record lengths, depending on the distribution of the quantity analysed. The results of this analysis indicate that previous analyses could potentially have overestimated the likelihood of extreme events arising from natural weather variability.

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Key Points:

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10	•	Grouping data of nearby locations into one larger sample or "superstation" can
11		induce biases at long recurrence intervals
12	•	The superstation fit tends to the highest levels suggested by any of the pooled lo-
13		cations
14	•	The bias may be large or small compared to random uncertainty, depending on
15		the distribution of the extreme event analysed

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16 Abstract

Preparing for environmental risks requires estimating the frequencies of extreme events, 17 often from data records that are too short to confirm them directly. This requires fit-18 ting a statistical distribution to the data. To improve precision, investigators often pool 19 data from neighboring sites into single samples, referred to as "superstations," before fit-20 ting. We demonstrate that this technique can introduce unexpected biases in typical sit-21 uations, using wind and rainfall extremes as case studies. When the combined locations 22 have even small differences in the underlying statistics, the regionalization approach gives 23 a fit that may tend toward the highest levels suggested by any of the individual sites. 24 This bias may be large or small compared to the sampling error, for realistic record lengths, 25 depending on the distribution of the quantity analysed. The results of this analysis in-26 dicate that previous analyses could potentially have overestimated the likelihood of ex-27 treme events arising from natural weather variability. 28

²⁹ Plain Language Summary

We report a previously unknown bias in a common method for estimating how of-30 ten extremely rare events such as extreme wind bursts or rain events will occur, when 31 return periods are longer than the available data record. The method analysed is one 32 where an investigator combines data from nearby locations to reduce sampling error. We 33 find by looking at new, high-resolution data that variations in behavior across sites can 34 sometimes produce biases much larger than the sampling error. The implication is that 35 some observed extreme events are even less likely to have occurred than previously thought, 36 assuming the underlying distribution hasn't changed over the period of observation. 37

38 1 Introduction

The statistical analysis of extreme-event frequencies and intensities is important 30 to many risk management problems. For example, estimating the appropriate design wind 40 speed requires the statistical analysis of historical wind data to estimate the strongest 41 wind that might occur over a long time interval (AS/NZS1170.2:2021, 2021; Holmes, 2002; 42 El Rafei et al., 2022). The design of offshore and coastal marine structures is governed 43 by statistical estimates of extreme waves (Gulev & Grigorieva, 2004; Meucci et al., 2020). 44 Similarly, statistical estimates of extreme rainfall values are essential for calculating flood 45 risk and designing stormwater infrastructure (Green et al., 2012; Johnson & Green, 2018). 46 To meet this need, extreme value theory has been widely used to estimate the proba-47 bility of events larger than any on record so far (Brabson & Palutikof, 2000; Coles, 2001; 48 Church et al., 2006; Wang et al., 2013). 49

The two usual approaches of extreme value theory are the generalized extreme value 50 distribution (GEV) and the generalized Pareto distribution (GPD). The GEV approach, 51 which combines three different statistical families (Weibull, Gumbel and Frechet), uses 52 block maxima in which the dataset is divided into blocks and the maximum over each 53 block is modelled (Gumbel, 1958; Palutikof et al., 1999; Coles, 2001). The GPD approach 54 is instead based on the peaks-over-threshold method for which a threshold value is spec-55 ified and all the values above this chosen threshold are used to fit the model (Pickands, 56 1975; Coles, 2001; Holmes & Moriarty, 1999). No matter which approach is used, the ex-57 trapolation to very rare events is subject to significant sampling errors when using short 58 data ranges and uncertainties are unavoidable as accurate observational records are com-59 monly short and/or geographically sparse. 60

To reduce statistical uncertainty, regionalization techniques have been used, whereby a larger sample is created by combining independent records of neighboring stations (J. Peterka, 1992; J. A. Peterka & Shahid, 1998; Holmes, 2002; Wang et al., 2013; Holmes, 2019) into what is sometimes called a "superstation". For example, regionalization has been

used for extreme wind assessment in the United States (J. A. Peterka & Shahid, 1998; 65 ANSI/AS CE 7-98, 1998; ASCE/SE I 7-16, 2016) and Australia (Holmes, 2002; AS/NZS1170.2:2021, 66 2021) to specify single design wind speed by compositing data from multiple stations; 67 for regional flood frequency estimates (Haddad & Rahman, 2012); and for Intensity-Frequency-68 Duration (IFD) rainfall relationships (Wallis et al., 2007; Norbiato et al., 2007; Green 69 et al., 2012; Johnson & Green, 2018). For IFD applications, nearby stations are pooled 70 together assuming they share a common distribution of rainfall and are independent. An-71 other place where this approach has been used is for so-called "regional frequency anal-72 ysis" of extreme wave heights, where data from sites with similar wave statistics are used 73 to estimate the distribution for a presumed homogeneous region (Van Gelder et al., 2001; 74 Bernardara et al., 2011; Lucas et al., 2017). 75

While regionalization allows the estimation of distribution parameters using a larger
dataset, the biases of this strategy are not explicitly quantified in the literature. Here
we report unexpected biases in estimating long-term recurrence intervals of extreme events
via regionalization, considering wind and rainfall extremes as case studies.

⁸⁰ 2 Data and Methods

2.1 Wind Data and Distribution Parameters

Wind data are obtained from the 1.5 km Bureau of Meteorology Atmospheric highresolution Regional Reanalysis for Australia Sydney region (BARRA-SY) and cover the period from 1996 to 2019. This fine-scale reanalysis may provide better representation of storm systems and mesoscale phenomena, which can allow better modelling and higher accuracy of the wind field compared to lower resolution models. The reader is referred to (Jakob et al., 2017; Su et al., 2019; Su et al., 2021) for more details.

The maximum hourly wind gust speed data was used to calculate the daily max-88 ima, which are then used to fit the GPD model. For computational tractability, we sub-89 sampled the BARRA-SY data to 5324 grid points covering New South Wales (NSW), 90 with a spacing of approximately 10 km. This is judged to be justified because gusts at 91 locations that are very close (i.e., 1.5 km apart) would capture the same gust events mul-92 tiple times hence not be independent. The domain is then decomposed into groups (su-93 perstations) of approximately 25 grid points (5×5 grid points neighborhood) such that 94 approximately 575 station years are obtained for each superstation. 95

The fully heterogeneous numerical experiment (Fig. 2B) considers threshold values between 15 and 23 m/s, shape factors between -0.2 to -0.1 and scale parameters ranging from 2 to 3. These ranges are representative of wind gust GPD distributions obtained using the BARRA-SY dataset. The partially heterogeneous case (Fig. 2B) considers a pre-fixed shape factor of -0.1 (similarly to what is used in the Australian standards) and a threshold value of 20 m/s. The range of scale factor values (between 3 and 4) was arbitrarily selected to represent a gradient of this parameter between the pseudo-stations.

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2.2 Rainfall Parameters

To provide parameter values for rainfall tests, 30 years (1992-2021) of daily rain-104 fall data were obtained from the Australian Bureau of Meteorology for five locations in 105 the Sydney area (Sydney Airport, Sans Suci, Randwick, Rose Bay and Peakhurst) pro-106 viding good data quality, using only low elevation stations (≤ 100 m). Incorrect values 107 (flagged as wrong via the quality checks) were removed, and the annual maxima were 108 used to fit a GEV model at each station. The GEV analysis suggests threshold values 109 ranging from 66.2 to 77.3, shape factors ranging from 0.08 to 0.27 and scale factors rang-110 ing from 22.7 to 25.3. These ranges were then used to generate synthetic random sam-111 ples for the fully heterogeneous scenario (Fig. 4B). The partially heterogeneous scenario 112

considers pre-fixed threshold and shape factor ($u_0 = 77.3$ and $\xi = 0.19$), which are selected from one of the stations distributions, and a range of scale factors from 17 to

25. The range of scale factors is randomly selected for this scenario to represent a gra dient of this parameter.

117 2.3 Extreme Value Theory

The GEV approach uses the block maxima technique where the maximum yearly value is considered for the fit. The GEV cumulative distribution function can be expressed as follows:

$$G(u_g) = \begin{cases} e^{-\left[1+\xi_g\left(\frac{u_g-u_{0g}}{\sigma_g}\right)\right]^{-1/\xi_g}}, & \text{for } \xi_g \neq 0\\ e^{-e^{(u_g-u_{0g})/\sigma_g}}, & \text{for } \xi_g = 0 \end{cases}$$

where u_{0g} , ξ_g and σ_g are the location, shape and scale parameters respectively. The return levels are given in terms of ARI as:

$$U_{Rg} = u_{0g} - \frac{\sigma_g}{\xi_g} [1 - (R)^{\xi_g}]$$
(1)

where R is the ARI.

The GPD model is based on the peaks-over-threshold approach (Palutikof et al., 1999; Coles, 2001) where all the data above a specified threshold are modelled. The GPD distribution can be expressed as:

$$H(u_p) = \begin{cases} 1 - \left(1 + \xi_p \frac{u_p - u_{0p}}{\sigma_p}\right)^{-1/\xi_p}, & \text{for } \xi_p \neq 0\\ 1 - e^{\frac{u_p - u_{0p}}{\sigma_p}}, & \text{for } \xi_p = 0 \end{cases}$$

¹¹⁹ where u_{0p} , ξ_p and σ_p are the threshold, shape and scale parameters respectively. ¹²⁰ This expression is defined on $\{u_p - u_{0p} > 0 \text{ and } (1 + \xi_p(u_p - u_{0p})/\tilde{\sigma_p}) > 0\}$ where ¹²¹ $\tilde{\sigma_p} = \sigma_p + \xi_p(u_p - u_{0p}).$

The return levels are given in terms of the average recurrence interval using:

$$U_{Rp} = u_{0p} - \frac{\sigma_p}{\xi_p} [1 - (\lambda R)^{\xi_p}]$$
(2)

where λ is the number of crossings of the threshold per year.

The rainfall estimates calculated using the GPD approach use the 99^{th} percentile 123 as a threshold value (Lazoglou & Anagnostopoulou, 2017). For the wind analysis, we used 124 a threshold selection algorithm that selects an optimal threshold value at each location. 125 The selection algorithm first classifies the gusts into convective (i.e., thunderstorms) and 126 synoptic (e.g., east coast lows, frontal systems) events. This is because gusts produced 127 by different mechanisms can have different statistical properties and distributions. A range 128 of thresholds is then tested to select the best fit. The reader is referred to (El Rafei et 129 al., 2022) for more details on the threshold selection algorithm and the storm classifi-130 cation technique. The current study only presents results of convective events, as the same 131 bias occur with either type using the superstation technique. 132

133 **3 Results**

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3.1 Extreme Wind Gust Example

135 3.1.1 Reanalysis Data Study

We begin by examining a high-resolution (1.5-km horizontal grid spacing) regional Australian reanalysis dataset (BARRA-SY) of 23 years length from Year 1996 to Year

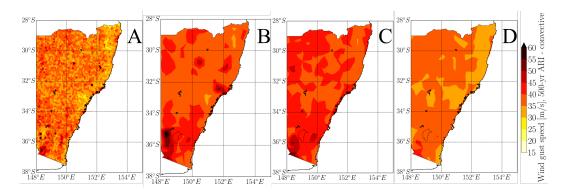


Figure 1: 500-year return wind gust speeds from convective events estimated several ways. Results are calculated using: A) the GPD algorithm applied at each grid point. B) the "superstation" approach to aggregate data from neighboring grid points. C) the nearby 90th percentile and D) mean of values shown in panel A, computed in the same neighborhood as the "superstations".

2019, using the GPD approach. The availability of this new high-resolution gridded dataset 138 has motivated this study, since wind data are available at sufficient density to directly 139 test the performance of the regionalization approach. The GPD method and reanaly-140 sis data are detailed in the Data and Methods section. Maps of 500-year convective (i.e., 141 thunderstorm) wind gust speeds calculated in several ways are shown in Fig. 1. We show 142 the 500-year ARI as this is used for the design criteria of many buildings in Australia 143 (Wang et al., 2013). Results of higher recurrence intervals (e.g. 1000- and 2000-vear ARIs, 144 not shown) present similar patterns. 145

The results obtained by GPD fitting and calculating return levels independently 146 at each grid cell (Fig. 1A) show speckling due to small differences in the estimated GPD 147 parameters. We expect this is mainly from sampling error due to the short record length, 148 as it looks random rather than geophysical. The regionalization approach (Fig. 1B), in 149 which data from 25 adjacent grid cells are combined before fitting, yields smoother re-150 sults but they show higher wind gust levels everywhere by about 10% compared to Fig. 1A. 151 They are also higher than when the grid-point results are instead spatially smoothed by 152 taking neighborhood means (2-D boxcar mean smoother of 5×5 grid points), as shown 153 in Fig. 1D. Indeed they are very close to the neighborhood 90th percentile results in Fig. 1C 154 (calculated as the 90th percentile wind gust speed of 25 adjacent grid points). This is 155 true even for regional maxima in neighborhood spread such as in the southeast corner 156 of the state. 157

This analysis suggests that the superstation technique gives a fit that is close to the highest levels suggested by any of the neighboring, noise-influenced, sites. We now examine two hypotheses for why this is happening: first, that we bias the return levels by using short data ranges; and second, that it is an effect of combining locations into a superstation.

3.1.2 Simulation Experiments

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A set of numerical experiments have been carried out to test these hypotheses, by generating synthetic records from pseudo-stations. Five synthetic "neighbor" data records are analytically generated from assumed GPD distributions with parameters based on wind gust distributions estimated using the BARRA-SY reanalysis, in two test scenarios. For the first, partially heterogeneous scenario (Fig. 2A,C), the five records are generated using the same threshold ($u_0 = 20 \text{ m/s}$) and shape factor ($\xi = -0.1$) but a range of scale factors from 3-4. Pre-fixing the threshold and shape factor is similar to to approaches sometimes used in structural design standards (AS/NZS1170.2:2021, 2021). For the second, fully heterogeneous scenario (Fig. 2B,D), all the distribution parameters vary among the records.

We first examine outcomes with long records (Figs. 2A and 2B), generating 1000 174 years of data, with a total of 5000 data points used by the GPD model, for each record 175 (i.e., five threshold exceedence events per year on average). Figs. 2A and B show that, 176 177 even when differences in the underlying distributions between the locations are small, the superstation fit is higher than the mean and the median of the individual fits and 178 tends to the highest levels suggested by the individual stations, which is consistent with 179 the wind data results in Fig. 1. Furthermore, in both simulations the bias increases at 180 longer recurrence intervals. The superstation bias in the fully heterogeneous scenario is 181 more significant, even at short recurrence intervals, because some locations are contribut-182 ing a lot more events than others to the superstation. This is not seen in the partially 183 homogeneous scenario where all station distributions have the same threshold and shape 184 factor and hence contribute similarly. 185

We next consider the effect of sample size by considering dataset lengths that range 186 from 30 to 100000 years. For each length we repeated the test 1000 times to yield a PDF, 187 the mean and 90th percentile of which is shown in Figs. 2C,D. The sampling uncertainty 188 monotonically reduces for longer datasets and gradually converges to the true supersta-189 tion bias. At 30 and 50 year record lengths (typical of real-world datasets), the bias may 190 be exceeded in magnitude by the random error as depicted by the large spread of the 191 PDF. Importantly however, all PDFs are centred on the true bias, showing that short 192 record lengths do not cause biases, only random sampling errors. The PDFs of error (Fig. 2C,D) 193 are narrow in both simulations when datasets are longer than 1000 years, implying a con-194 sistent bias site to site (or realisation to realisation), as implied by the geographic uni-195 formity of the difference between Fig. 1B vs. D. Moreover, as seen before, the simulated 196 superstation result is higher than the mean return level of the stations in the neighbor-197 hood. Hence we conclude from these tests that while the noise seen in the BARRA-SY 198 return-period map is from sampling errors due to the short record, the ubiquitous bias 199 toward high values is caused by regional pooling of data from nearby locations. 200

To understand what gives rise to this systematic bias, we compared the PDFs of 201 gust speed (Figs. 3A and B) corresponding to the exceedence curves shown in Fig. 2A, 202 B. In both scenarios, the slope of the superstation gust PDF at high gust speed tends 203 to be dictated by the stations that have the heaviest tail (i.e., where the most extremes 204 are recorded hence contributing most heavily to the superstation sample). Moreover, this 205 phenomenon increases as one goes farther out on the tail of the PDF; if for example a 206 very high threshold is used, nearly all data meeting the threshold come from one station 207 (open circles in Fig. 3B). This level-dependent bias imparts a shallower slope to the PDF 208 tail, which means that when calculating very long recurrence intervals, the extrapola-209 tion would tend to levels suggested by the locations that experience the most extreme 210 events, or possibly even higher. 211

212 **3.2 Ra**

3.2 Rainfall Example

We now repeat the above analysis for another type of natural hazard, extreme rainfall, to explore the generality of the result. In this case rather than GPD we use GEV, since this is how published rainfall estimates are typically calculated, but this choice does not substantially affect results (Data and Methods section). The synthetic data records are generated as before, considering the same two scenarios except for the parameter values. The range of rainfall scale factors varies from 17 to 25, based on distributions estimated from observed daily accumulated rainfall from weather stations in Sydney area.

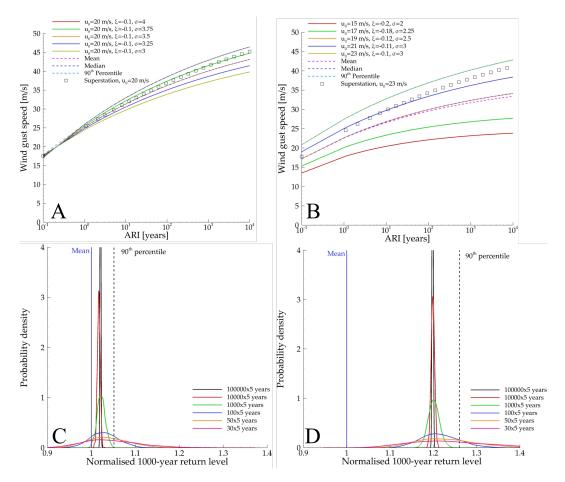


Figure 2: Simulated results for two different test cases: five pooled locations with (A,C) common threshold and shape factor values but varying scale factors, (B,D) all GPD distribution parameters varying among locations. Panels (A,B) show value vs. return period given 1000 years of data, and (C,D) show the PDF of the 1000-year return-period value given different data record lengths indicated in the legend (and assuming five threshold exceedences per year).

The return behavior of rainfall (Fig. 4A,B) is quite different from that of wind. Rain-220 fall is highly intermittent with a long tail on the PDF, and the fits are unbounded with 221 upward curvature to very high rain rates at extremely long return periods, due to the 222 positive shape factor of the distribution, unlike the case for wind which has a negative 223 shape factor and appears to asymptote toward a maximum possible value. Nonetheless 224 the rainfall superstation fit is again higher than the true mean and median, and tends 225 toward the 90th percentile of the neighborhood, consistent with the results based on wind 226 distributions. For dataset lengths less than 1000 years, however (Fig. 4C,D), this bias 227 is significantly outweighed by the sampling error such that the observed error in a sin-228 gle realization can be of either sign. This was not the case for gust distributions, where 229 the superstation bias stands out even with short records (Figs. 2C,D). 230

The large sampling errors observed in the rainfall case are independent of the distribution model, as shown in Fig. 5, where the results from Fig. 4D are compared between the GPD and GEV approaches. Both show a similar level of bias for all record lengths, although the biases are slightly smaller if GPD is used instead of the usual (for rainfall) GEV. Sampling uncertainties exceed the superstation bias regardless of the model,

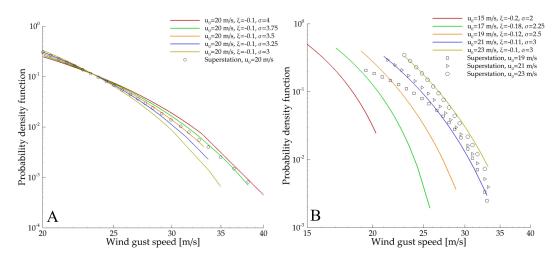


Figure 3: Probability density functions of the pseudo-station and superstation synthetic data shown in Figs. 2A, 2B. Open symbols in panel B show PDFs for three different threshold values.

for realistic dataset lengths. Thus, rainfall estimates are harder to constrain and the bias identified here is much less important compared to sampling error.

238 4 Conclusion

We demonstrate a previously unreported bias in estimating long-term recurrence 239 intervals of extreme events that results from the common practice of regionalization or 240 grouping data of nearby locations into one larger sample or "superstation". Wind gust 241 and rainfall extremes have been considered for this analysis, but the results are also likely 242 applicable to other types of weather extremes. Regionalisation assumes that all locations 243 grouped have the same underlying distribution. According to newly available, high-resolution 244 simulations of wind events in eastern Australia, differences in the underlying distribu-245 tion can be large enough to induce biases at long recurrence intervals that dominate sam-246 pling uncertainty. The superstation fit tends to the highest levels suggested by any of 247 the pooled locations and this bias increases with longer recurrence intervals. The tail of 248 the superstation distribution tends to get its slope from the locations that experienced 249 the most extremes. Moreover, the superstation PDF slope in our calculations is shallower 250 than the those of any of the contributing stations, such that extrapolation will result in 251 increasingly biased estimates at longer recurrence intervals. Our analysis suggests that 252 for highly intermittent processes with unbounded behavior at the extreme tail such as 253 rainfall with positive skewness for large values, the bias may be outweighed by random 254 uncertainty and so may not matter in some cases, but will become important for bounded 255 distributions. Since the importance of this bias depends on the distribution of the vari-256 able examined, we suggest that researchers should test for this bias before applying any 257 regionalization method. 258

259 Conflict of Interest

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The authors declare no conflicts of interest relevant to this study.

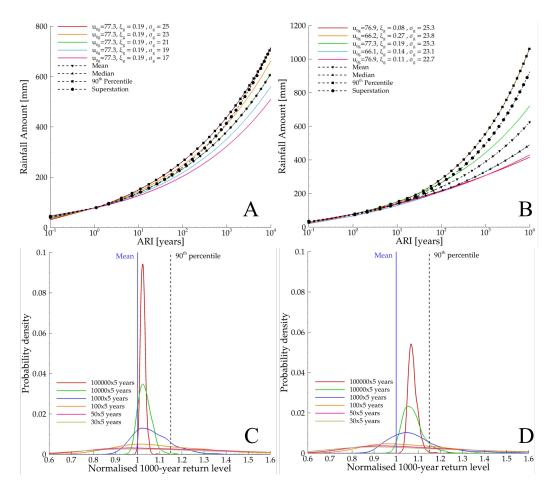


Figure 4: Same as Fig. 2 but for rainfall estimates.

²⁶¹ Data Availability Statement

BARRA-SY data are available in an-open repository on the NCI (http://dx.doi .org/10.4225/41/5993927b50f53).

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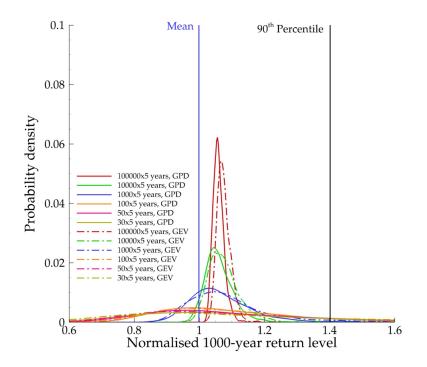


Figure 5: Same as Fig. 4D except comparing GPD and GEV models.

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Biases in estimating long-term recurrence intervals of extreme events due to regionalised sampling

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Key Points:

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10	•	Grouping data of nearby locations into one larger sample or "superstation" can
11		induce biases at long recurrence intervals
12	•	The superstation fit tends to the highest levels suggested by any of the pooled lo-
13		cations
14	•	The bias may be large or small compared to random uncertainty, depending on
15		the distribution of the extreme event analysed

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16 Abstract

Preparing for environmental risks requires estimating the frequencies of extreme events, 17 often from data records that are too short to confirm them directly. This requires fit-18 ting a statistical distribution to the data. To improve precision, investigators often pool 19 data from neighboring sites into single samples, referred to as "superstations," before fit-20 ting. We demonstrate that this technique can introduce unexpected biases in typical sit-21 uations, using wind and rainfall extremes as case studies. When the combined locations 22 have even small differences in the underlying statistics, the regionalization approach gives 23 a fit that may tend toward the highest levels suggested by any of the individual sites. 24 This bias may be large or small compared to the sampling error, for realistic record lengths, 25 depending on the distribution of the quantity analysed. The results of this analysis in-26 dicate that previous analyses could potentially have overestimated the likelihood of ex-27 treme events arising from natural weather variability. 28

²⁹ Plain Language Summary

We report a previously unknown bias in a common method for estimating how of-30 ten extremely rare events such as extreme wind bursts or rain events will occur, when 31 return periods are longer than the available data record. The method analysed is one 32 where an investigator combines data from nearby locations to reduce sampling error. We 33 find by looking at new, high-resolution data that variations in behavior across sites can 34 sometimes produce biases much larger than the sampling error. The implication is that 35 some observed extreme events are even less likely to have occurred than previously thought, 36 assuming the underlying distribution hasn't changed over the period of observation. 37

38 1 Introduction

The statistical analysis of extreme-event frequencies and intensities is important 30 to many risk management problems. For example, estimating the appropriate design wind 40 speed requires the statistical analysis of historical wind data to estimate the strongest 41 wind that might occur over a long time interval (AS/NZS1170.2:2021, 2021; Holmes, 2002; 42 El Rafei et al., 2022). The design of offshore and coastal marine structures is governed 43 by statistical estimates of extreme waves (Gulev & Grigorieva, 2004; Meucci et al., 2020). 44 Similarly, statistical estimates of extreme rainfall values are essential for calculating flood 45 risk and designing stormwater infrastructure (Green et al., 2012; Johnson & Green, 2018). 46 To meet this need, extreme value theory has been widely used to estimate the proba-47 bility of events larger than any on record so far (Brabson & Palutikof, 2000; Coles, 2001; 48 Church et al., 2006; Wang et al., 2013). 49

The two usual approaches of extreme value theory are the generalized extreme value 50 distribution (GEV) and the generalized Pareto distribution (GPD). The GEV approach, 51 which combines three different statistical families (Weibull, Gumbel and Frechet), uses 52 block maxima in which the dataset is divided into blocks and the maximum over each 53 block is modelled (Gumbel, 1958; Palutikof et al., 1999; Coles, 2001). The GPD approach 54 is instead based on the peaks-over-threshold method for which a threshold value is spec-55 ified and all the values above this chosen threshold are used to fit the model (Pickands, 56 1975; Coles, 2001; Holmes & Moriarty, 1999). No matter which approach is used, the ex-57 trapolation to very rare events is subject to significant sampling errors when using short 58 data ranges and uncertainties are unavoidable as accurate observational records are com-59 monly short and/or geographically sparse. 60

To reduce statistical uncertainty, regionalization techniques have been used, whereby a larger sample is created by combining independent records of neighboring stations (J. Peterka, 1992; J. A. Peterka & Shahid, 1998; Holmes, 2002; Wang et al., 2013; Holmes, 2019) into what is sometimes called a "superstation". For example, regionalization has been

used for extreme wind assessment in the United States (J. A. Peterka & Shahid, 1998; 65 ANSI/AS CE 7-98, 1998; ASCE/SE I 7-16, 2016) and Australia (Holmes, 2002; AS/NZS1170.2:2021, 66 2021) to specify single design wind speed by compositing data from multiple stations; 67 for regional flood frequency estimates (Haddad & Rahman, 2012); and for Intensity-Frequency-68 Duration (IFD) rainfall relationships (Wallis et al., 2007; Norbiato et al., 2007; Green 69 et al., 2012; Johnson & Green, 2018). For IFD applications, nearby stations are pooled 70 together assuming they share a common distribution of rainfall and are independent. An-71 other place where this approach has been used is for so-called "regional frequency anal-72 ysis" of extreme wave heights, where data from sites with similar wave statistics are used 73 to estimate the distribution for a presumed homogeneous region (Van Gelder et al., 2001; 74 Bernardara et al., 2011; Lucas et al., 2017). 75

While regionalization allows the estimation of distribution parameters using a larger
dataset, the biases of this strategy are not explicitly quantified in the literature. Here
we report unexpected biases in estimating long-term recurrence intervals of extreme events
via regionalization, considering wind and rainfall extremes as case studies.

⁸⁰ 2 Data and Methods

2.1 Wind Data and Distribution Parameters

Wind data are obtained from the 1.5 km Bureau of Meteorology Atmospheric highresolution Regional Reanalysis for Australia Sydney region (BARRA-SY) and cover the period from 1996 to 2019. This fine-scale reanalysis may provide better representation of storm systems and mesoscale phenomena, which can allow better modelling and higher accuracy of the wind field compared to lower resolution models. The reader is referred to (Jakob et al., 2017; Su et al., 2019; Su et al., 2021) for more details.

The maximum hourly wind gust speed data was used to calculate the daily max-88 ima, which are then used to fit the GPD model. For computational tractability, we sub-89 sampled the BARRA-SY data to 5324 grid points covering New South Wales (NSW), 90 with a spacing of approximately 10 km. This is judged to be justified because gusts at 91 locations that are very close (i.e., 1.5 km apart) would capture the same gust events mul-92 tiple times hence not be independent. The domain is then decomposed into groups (su-93 perstations) of approximately 25 grid points (5×5 grid points neighborhood) such that 94 approximately 575 station years are obtained for each superstation. 95

The fully heterogeneous numerical experiment (Fig. 2B) considers threshold values between 15 and 23 m/s, shape factors between -0.2 to -0.1 and scale parameters ranging from 2 to 3. These ranges are representative of wind gust GPD distributions obtained using the BARRA-SY dataset. The partially heterogeneous case (Fig. 2B) considers a pre-fixed shape factor of -0.1 (similarly to what is used in the Australian standards) and a threshold value of 20 m/s. The range of scale factor values (between 3 and 4) was arbitrarily selected to represent a gradient of this parameter between the pseudo-stations.

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2.2 Rainfall Parameters

To provide parameter values for rainfall tests, 30 years (1992-2021) of daily rain-104 fall data were obtained from the Australian Bureau of Meteorology for five locations in 105 the Sydney area (Sydney Airport, Sans Suci, Randwick, Rose Bay and Peakhurst) pro-106 viding good data quality, using only low elevation stations (≤ 100 m). Incorrect values 107 (flagged as wrong via the quality checks) were removed, and the annual maxima were 108 used to fit a GEV model at each station. The GEV analysis suggests threshold values 109 ranging from 66.2 to 77.3, shape factors ranging from 0.08 to 0.27 and scale factors rang-110 ing from 22.7 to 25.3. These ranges were then used to generate synthetic random sam-111 ples for the fully heterogeneous scenario (Fig. 4B). The partially heterogeneous scenario 112

considers pre-fixed threshold and shape factor ($u_0 = 77.3$ and $\xi = 0.19$), which are selected from one of the stations distributions, and a range of scale factors from 17 to

25. The range of scale factors is randomly selected for this scenario to represent a gra dient of this parameter.

117 2.3 Extreme Value Theory

The GEV approach uses the block maxima technique where the maximum yearly value is considered for the fit. The GEV cumulative distribution function can be expressed as follows:

$$G(u_g) = \begin{cases} e^{-\left[1+\xi_g\left(\frac{u_g-u_{0g}}{\sigma_g}\right)\right]^{-1/\xi_g}}, & \text{for } \xi_g \neq 0\\ e^{-e^{(u_g-u_{0g})/\sigma_g}}, & \text{for } \xi_g = 0 \end{cases}$$

where u_{0g} , ξ_g and σ_g are the location, shape and scale parameters respectively. The return levels are given in terms of ARI as:

$$U_{Rg} = u_{0g} - \frac{\sigma_g}{\xi_g} [1 - (R)^{\xi_g}]$$
(1)

where R is the ARI.

The GPD model is based on the peaks-over-threshold approach (Palutikof et al., 1999; Coles, 2001) where all the data above a specified threshold are modelled. The GPD distribution can be expressed as:

$$H(u_p) = \begin{cases} 1 - \left(1 + \xi_p \frac{u_p - u_{0p}}{\sigma_p}\right)^{-1/\xi_p}, & \text{for } \xi_p \neq 0\\ 1 - e^{\frac{u_p - u_{0p}}{\sigma_p}}, & \text{for } \xi_p = 0 \end{cases}$$

¹¹⁹ where u_{0p} , ξ_p and σ_p are the threshold, shape and scale parameters respectively. ¹²⁰ This expression is defined on $\{u_p - u_{0p} > 0 \text{ and } (1 + \xi_p(u_p - u_{0p})/\tilde{\sigma_p}) > 0\}$ where ¹²¹ $\tilde{\sigma_p} = \sigma_p + \xi_p(u_p - u_{0p}).$

The return levels are given in terms of the average recurrence interval using:

$$U_{Rp} = u_{0p} - \frac{\sigma_p}{\xi_p} [1 - (\lambda R)^{\xi_p}]$$
(2)

where λ is the number of crossings of the threshold per year.

The rainfall estimates calculated using the GPD approach use the 99^{th} percentile 123 as a threshold value (Lazoglou & Anagnostopoulou, 2017). For the wind analysis, we used 124 a threshold selection algorithm that selects an optimal threshold value at each location. 125 The selection algorithm first classifies the gusts into convective (i.e., thunderstorms) and 126 synoptic (e.g., east coast lows, frontal systems) events. This is because gusts produced 127 by different mechanisms can have different statistical properties and distributions. A range 128 of thresholds is then tested to select the best fit. The reader is referred to (El Rafei et 129 al., 2022) for more details on the threshold selection algorithm and the storm classifi-130 cation technique. The current study only presents results of convective events, as the same 131 bias occur with either type using the superstation technique. 132

133 **3 Results**

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3.1 Extreme Wind Gust Example

135 3.1.1 Reanalysis Data Study

We begin by examining a high-resolution (1.5-km horizontal grid spacing) regional Australian reanalysis dataset (BARRA-SY) of 23 years length from Year 1996 to Year

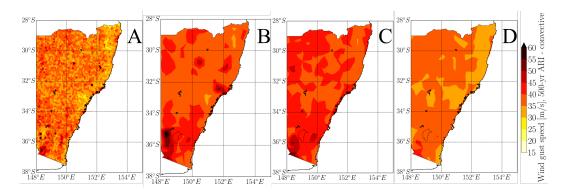


Figure 1: 500-year return wind gust speeds from convective events estimated several ways. Results are calculated using: A) the GPD algorithm applied at each grid point. B) the "superstation" approach to aggregate data from neighboring grid points. C) the nearby 90th percentile and D) mean of values shown in panel A, computed in the same neighborhood as the "superstations".

2019, using the GPD approach. The availability of this new high-resolution gridded dataset 138 has motivated this study, since wind data are available at sufficient density to directly 139 test the performance of the regionalization approach. The GPD method and reanaly-140 sis data are detailed in the Data and Methods section. Maps of 500-year convective (i.e., 141 thunderstorm) wind gust speeds calculated in several ways are shown in Fig. 1. We show 142 the 500-year ARI as this is used for the design criteria of many buildings in Australia 143 (Wang et al., 2013). Results of higher recurrence intervals (e.g. 1000- and 2000-vear ARIs, 144 not shown) present similar patterns. 145

The results obtained by GPD fitting and calculating return levels independently 146 at each grid cell (Fig. 1A) show speckling due to small differences in the estimated GPD 147 parameters. We expect this is mainly from sampling error due to the short record length, 148 as it looks random rather than geophysical. The regionalization approach (Fig. 1B), in 149 which data from 25 adjacent grid cells are combined before fitting, yields smoother re-150 sults but they show higher wind gust levels everywhere by about 10% compared to Fig. 1A. 151 They are also higher than when the grid-point results are instead spatially smoothed by 152 taking neighborhood means (2-D boxcar mean smoother of 5×5 grid points), as shown 153 in Fig. 1D. Indeed they are very close to the neighborhood 90th percentile results in Fig. 1C 154 (calculated as the 90th percentile wind gust speed of 25 adjacent grid points). This is 155 true even for regional maxima in neighborhood spread such as in the southeast corner 156 of the state. 157

This analysis suggests that the superstation technique gives a fit that is close to the highest levels suggested by any of the neighboring, noise-influenced, sites. We now examine two hypotheses for why this is happening: first, that we bias the return levels by using short data ranges; and second, that it is an effect of combining locations into a superstation.

3.1.2 Simulation Experiments

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A set of numerical experiments have been carried out to test these hypotheses, by generating synthetic records from pseudo-stations. Five synthetic "neighbor" data records are analytically generated from assumed GPD distributions with parameters based on wind gust distributions estimated using the BARRA-SY reanalysis, in two test scenarios. For the first, partially heterogeneous scenario (Fig. 2A,C), the five records are generated using the same threshold ($u_0 = 20 \text{ m/s}$) and shape factor ($\xi = -0.1$) but a range of scale factors from 3-4. Pre-fixing the threshold and shape factor is similar to to approaches sometimes used in structural design standards (AS/NZS1170.2:2021, 2021). For the second, fully heterogeneous scenario (Fig. 2B,D), all the distribution parameters vary among the records.

We first examine outcomes with long records (Figs. 2A and 2B), generating 1000 174 years of data, with a total of 5000 data points used by the GPD model, for each record 175 (i.e., five threshold exceedence events per year on average). Figs. 2A and B show that, 176 177 even when differences in the underlying distributions between the locations are small, the superstation fit is higher than the mean and the median of the individual fits and 178 tends to the highest levels suggested by the individual stations, which is consistent with 179 the wind data results in Fig. 1. Furthermore, in both simulations the bias increases at 180 longer recurrence intervals. The superstation bias in the fully heterogeneous scenario is 181 more significant, even at short recurrence intervals, because some locations are contribut-182 ing a lot more events than others to the superstation. This is not seen in the partially 183 homogeneous scenario where all station distributions have the same threshold and shape 184 factor and hence contribute similarly. 185

We next consider the effect of sample size by considering dataset lengths that range 186 from 30 to 100000 years. For each length we repeated the test 1000 times to yield a PDF, 187 the mean and 90th percentile of which is shown in Figs. 2C,D. The sampling uncertainty 188 monotonically reduces for longer datasets and gradually converges to the true supersta-189 tion bias. At 30 and 50 year record lengths (typical of real-world datasets), the bias may 190 be exceeded in magnitude by the random error as depicted by the large spread of the 191 PDF. Importantly however, all PDFs are centred on the true bias, showing that short 192 record lengths do not cause biases, only random sampling errors. The PDFs of error (Fig. 2C,D) 193 are narrow in both simulations when datasets are longer than 1000 years, implying a con-194 sistent bias site to site (or realisation to realisation), as implied by the geographic uni-195 formity of the difference between Fig. 1B vs. D. Moreover, as seen before, the simulated 196 superstation result is higher than the mean return level of the stations in the neighbor-197 hood. Hence we conclude from these tests that while the noise seen in the BARRA-SY 198 return-period map is from sampling errors due to the short record, the ubiquitous bias 199 toward high values is caused by regional pooling of data from nearby locations. 200

To understand what gives rise to this systematic bias, we compared the PDFs of 201 gust speed (Figs. 3A and B) corresponding to the exceedence curves shown in Fig. 2A, 202 B. In both scenarios, the slope of the superstation gust PDF at high gust speed tends 203 to be dictated by the stations that have the heaviest tail (i.e., where the most extremes 204 are recorded hence contributing most heavily to the superstation sample). Moreover, this 205 phenomenon increases as one goes farther out on the tail of the PDF; if for example a 206 very high threshold is used, nearly all data meeting the threshold come from one station 207 (open circles in Fig. 3B). This level-dependent bias imparts a shallower slope to the PDF 208 tail, which means that when calculating very long recurrence intervals, the extrapola-209 tion would tend to levels suggested by the locations that experience the most extreme 210 events, or possibly even higher. 211

212 **3.2 Ra**

3.2 Rainfall Example

We now repeat the above analysis for another type of natural hazard, extreme rainfall, to explore the generality of the result. In this case rather than GPD we use GEV, since this is how published rainfall estimates are typically calculated, but this choice does not substantially affect results (Data and Methods section). The synthetic data records are generated as before, considering the same two scenarios except for the parameter values. The range of rainfall scale factors varies from 17 to 25, based on distributions estimated from observed daily accumulated rainfall from weather stations in Sydney area.

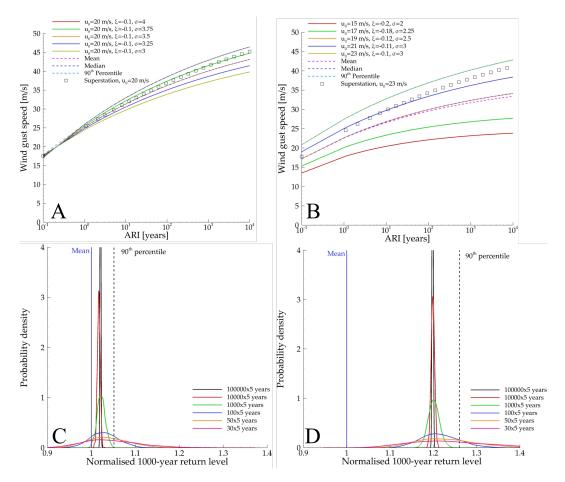


Figure 2: Simulated results for two different test cases: five pooled locations with (A,C) common threshold and shape factor values but varying scale factors, (B,D) all GPD distribution parameters varying among locations. Panels (A,B) show value vs. return period given 1000 years of data, and (C,D) show the PDF of the 1000-year return-period value given different data record lengths indicated in the legend (and assuming five threshold exceedences per year).

The return behavior of rainfall (Fig. 4A,B) is quite different from that of wind. Rain-220 fall is highly intermittent with a long tail on the PDF, and the fits are unbounded with 221 upward curvature to very high rain rates at extremely long return periods, due to the 222 positive shape factor of the distribution, unlike the case for wind which has a negative 223 shape factor and appears to asymptote toward a maximum possible value. Nonetheless 224 the rainfall superstation fit is again higher than the true mean and median, and tends 225 toward the 90th percentile of the neighborhood, consistent with the results based on wind 226 distributions. For dataset lengths less than 1000 years, however (Fig. 4C,D), this bias 227 is significantly outweighed by the sampling error such that the observed error in a sin-228 gle realization can be of either sign. This was not the case for gust distributions, where 229 the superstation bias stands out even with short records (Figs. 2C,D). 230

The large sampling errors observed in the rainfall case are independent of the distribution model, as shown in Fig. 5, where the results from Fig. 4D are compared between the GPD and GEV approaches. Both show a similar level of bias for all record lengths, although the biases are slightly smaller if GPD is used instead of the usual (for rainfall) GEV. Sampling uncertainties exceed the superstation bias regardless of the model,

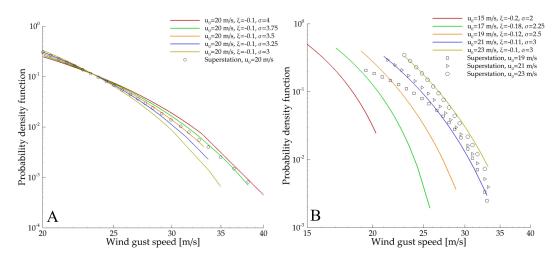


Figure 3: Probability density functions of the pseudo-station and superstation synthetic data shown in Figs. 2A, 2B. Open symbols in panel B show PDFs for three different threshold values.

for realistic dataset lengths. Thus, rainfall estimates are harder to constrain and the bias identified here is much less important compared to sampling error.

238 4 Conclusion

We demonstrate a previously unreported bias in estimating long-term recurrence 239 intervals of extreme events that results from the common practice of regionalization or 240 grouping data of nearby locations into one larger sample or "superstation". Wind gust 241 and rainfall extremes have been considered for this analysis, but the results are also likely 242 applicable to other types of weather extremes. Regionalisation assumes that all locations 243 grouped have the same underlying distribution. According to newly available, high-resolution 244 simulations of wind events in eastern Australia, differences in the underlying distribu-245 tion can be large enough to induce biases at long recurrence intervals that dominate sam-246 pling uncertainty. The superstation fit tends to the highest levels suggested by any of 247 the pooled locations and this bias increases with longer recurrence intervals. The tail of 248 the superstation distribution tends to get its slope from the locations that experienced 249 the most extremes. Moreover, the superstation PDF slope in our calculations is shallower 250 than the those of any of the contributing stations, such that extrapolation will result in 251 increasingly biased estimates at longer recurrence intervals. Our analysis suggests that 252 for highly intermittent processes with unbounded behavior at the extreme tail such as 253 rainfall with positive skewness for large values, the bias may be outweighed by random 254 uncertainty and so may not matter in some cases, but will become important for bounded 255 distributions. Since the importance of this bias depends on the distribution of the vari-256 able examined, we suggest that researchers should test for this bias before applying any 257 regionalization method. 258

259 Conflict of Interest

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The authors declare no conflicts of interest relevant to this study.

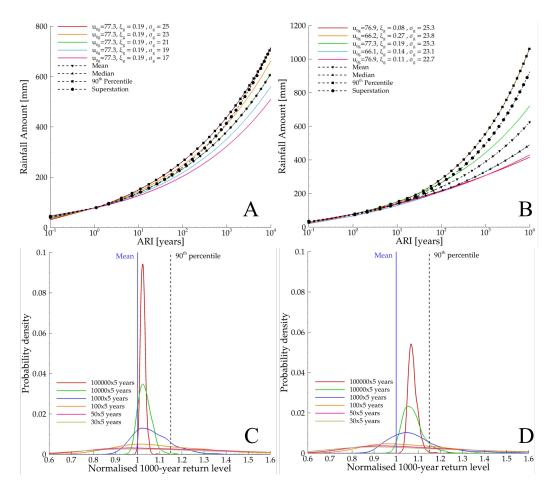


Figure 4: Same as Fig. 2 but for rainfall estimates.

²⁶¹ Data Availability Statement

BARRA-SY data are available in an-open repository on the NCI (http://dx.doi .org/10.4225/41/5993927b50f53).

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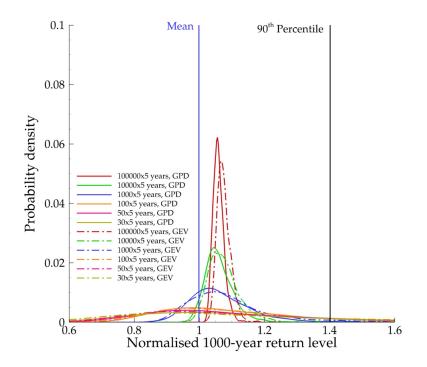


Figure 5: Same as Fig. 4D except comparing GPD and GEV models.

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