

# Climate change signal in Atlantic tropical cyclones today and near future

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## Abstract

This manuscript discusses the challenges in detecting and attributing recently observed trends in the Atlantic hurricanes and the epistemic uncertainty we face in assessing future hurricane risk. Data used here include synthetic storms downscaled from five CMIP5 models by the Columbia HAZard model (HAZ), and directly simulated storms from high-resolution climate models. We examine three aspects of recent hurricane activity: the upward trend and multi-decadal oscillation of the annual frequency, the increase in storm wind intensity, and the downward trend in the forward speed. Some datasets suggest that these trends and oscillation are forced while others suggest that they can be explained by natural variability. Future projections under warming climate scenarios also show a wide range of possibilities, especially for the annual frequencies, which increase or decrease depending on the choice of moisture variable used in the HAZ model and on the choice of climate model. The uncertainties in the annual frequency lead to epistemic uncertainties in the future hurricane risk assessment. Here, we investigate the reduction of epistemic uncertainties on annual frequency through a statistical practice – likelihood analysis. We find that historical observations are more consistent with the simulations with increasing frequency but we are not able to rule out other possibilities. We argue that the most rational way to treat epistemic uncertainty is to consider all outcomes contained in the results. In the context of hurricane risk assessment, since the results contain possible outcomes in which hurricane risk is increasing, this view implies that the risk is increasing.

1 **Climate change signal in Atlantic tropical cyclones**  
2 **today and near future**

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10 **Key Points:**

- 11 • Changes in the Atlantic hurricane risk are uncertain due to epistemic uncertainty  
12 in the projected annual frequency under global warming  
13 • Likelihood analysis shows that observations are more consistent with simulations  
14 with upward frequency projections than those without  
15 • Based on our results, it is more likely that the risk of hurricanes is increasing than  
16 that it is decreasing, though not by a large margin

**Abstract**

This manuscript discusses the challenges in detecting and attributing recently observed trends in the Atlantic hurricanes and the epistemic uncertainty we face in assessing future hurricane risk. Data used here include synthetic storms downscaled from five CMIP5 models by the Columbia HAZard model (CHAZ), and directly simulated storms from high-resolution climate models. We examine three aspects of recent hurricane activity: the upward trend and multi-decadal oscillation of the annual frequency, the increase in storm wind intensity, and the downward trend in the forward speed. Some datasets suggest that these trends and oscillation are forced while others suggest that they can be explained by natural variability. Future projections under warming climate scenarios also show a wide range of possibilities, especially for the annual frequencies, which increase or decrease depending on the choice of moisture variable used in the CHAZ model and on the choice of climate model. The uncertainties in the annual frequency lead to epistemic uncertainties in the future hurricane risk assessment. Here, we investigate the reduction of epistemic uncertainties on annual frequency through a statistical practice – likelihood analysis. We find that historical observations are more consistent with the simulations with increasing frequency but we are not able to rule out other possibilities. We argue that the most rational way to treat epistemic uncertainty is to consider all outcomes contained in the results. In the context of hurricane risk assessment, since the results contain possible outcomes in which hurricane risk is increasing, this view implies that the risk is increasing.

**Plain Language Summary**

We use a set of computer model simulations to study recent trends in Atlantic hurricanes. We looked at three aspects of these storms: the number of hurricanes each year, which has fluctuated up and down over time (but generally increased over the last several decades); the strength of their winds, which has been increasing; and the speed at which they move, which has been decreasing. These trends could be caused either by human-induced global warming or by natural variability; determining which cause is more important to overall hurricane risk requires us to understand how the number of hurricanes per year responds to warming. In our simulations, this number can either increase or decrease with warming, depending on which of two nearly identical versions of our model we use to simulate the storms. This uncertainty prevents us from reaching definitive conclusions about either present or future hurricane risk. Nonetheless, our analysis suggests that the risk of Atlantic hurricanes is more likely increasing than decreasing, and we argue that from a broader point of view, this is effectively equivalent to saying the risk is increasing.

**1 Introduction**

Rational measures to mitigate any risk must start from an assessment of that risk. Historical records can provide guidance, but in the case of atmospheric hazards such as hurricanes, we know that historical records are only a starting point for assessing current and future risk. This is both because the historical record is too short to fully sample the possibilities even in a stationary climate, and because the climate is changing (Schreck et al., 2014; Emanuel, 2021; D. Chan et al., 2022). Climate change makes the present different from the past, and requires us to consider whether the historical record alone, or catastrophe models that are built upon it, using purely statistical methods and assuming a stationary climate, are adequate, or need to be modified or supplemented to account for climate change.

Accounting for climate change is likely to require a greater use of physics than is historically typical in catastrophe models (Toumi & Restell, 2014; Emanuel, 2008). While one might instead try to assess the risk by using standard statistical methods but train-

67 ing only on the most recent observations (as opposed to the entire record), in the hope  
68 that those most recent observations represent the present and near-future climate ad-  
69 equately, this is likely to be challenging. Since hurricanes are rare, the number in the record  
70 over a period recent enough for this purpose is too small for risk assessment – especially  
71 when we also consider that low-frequency natural variability is present (i.e., Klotzbach  
72 & Gray, 2008; J. C. Chan, 2008; Wang et al., 2015), so that averaging times must be longer  
73 than might otherwise be necessary. To make the best possible assessment of present hur-  
74 ricane risk, then, we need to use our knowledge of the physics that connects hurricanes  
75 and climate (Emanuel, 2008).

76 The focus of this study is Atlantic tropical cyclones (TCs) risk in the present and  
77 near future. Future projections are useful for understanding how TCs may respond to  
78 climate changes of various sorts. Studies of historical observations, on the other hand, of-  
79 ten look for trends; but on their own, such studies do not establish the causes of the trends,  
80 nor whether they will persist. Establishing whether a trend is present (detection) is gen-  
81 erally viewed as a prerequisite to determining its cause (attribution) (Lloyd & Oreskes,  
82 2018). Detection can, in principle, be done with observations alone; attribution requires  
83 a model of some sort, in order to construct a counterfactual where the cause of interest  
84 is not present (Hegerl & Zwiers, 2011; Knutson, 2017). If a historical trend (or an os-  
85 cillatory signal) could be both detected and attributed to a specific cause, such as hu-  
86 man influence, or alternatively some specific natural processes, this would be of great  
87 scientific value, and would also allow us some insight into what to expect in the near fu-  
88 ture.

89 To develop such insight for Atlantic TCs, we will use recent observations and model  
90 simulations from historical (present), near future (up to 2040 or 2050), and pre-industrial  
91 control period. Simulations from pre-industrial control period contain no anthropogenic  
92 forcing signal and thus are used as a counterfactual. We use two types of model data.  
93 The first represents synthetic storms generated from a statistical-dynamical model, the  
94 Columbia (tropical cyclone) HAZard model (CHAZ), a model that encodes physical rela-  
95 tionships between tropical cyclones and their ambient large-scale environment (Lee et  
96 al., 2018). The second represents the directly simulated hurricanes from high-resolution  
97 global models, in which the above-mentioned relationships are simulated organically (Yoshida  
98 et al., 2017; Wehner et al., 2014; Roberts et al., 2020).

99 There are three objectives of this work. The first is to examine whether recently  
100 reported trends can be attributed to anthropogenic forcing. As summarized in Knutson  
101 et al. (2020a, 2020b), these trends are the recent variability of Atlantic annual TC fre-  
102 quency (Emanuel, 2007), an upward trend in the intensification rate (Bhatia et al., 2019)  
103 and lifetime maximum intensity (Kossin et al., 2013), and a slowing-down in the storm  
104 motion (Kossin, 2018). In particular, the cause of the recent increasing trend (since 1970)  
105 in Atlantic TC activity has been a subject of debate. On the one hand, reduced aerosols  
106 over the Atlantic since 1980s has been argued to be a dominant cause of the increasing  
107 TC activity in late 20<sup>th</sup> century (Mann & Emanuel, 2006; Sobel, Camargo, & Previdi,  
108 2019; Rousseau-Rizzi & Emanuel, 2020). On the other hand, several measures of Atlantic  
109 TC activity, including the major hurricane (TCs with LMI  $\geq$  93 kt) frequency (Goldenberg  
110 et al., 2001), are highly correlated to the the Atlantic Multi-decadal Oscillation (AMO)  
111 or Atlantic multidecadal variability (AMV), a low-frequency mode of variability iden-  
112 tified by the average sea surface temperature anomalies in the North Atlantic basin, typ-  
113 ically over 0-80°N (Ting et al., 2011). The recent AMO cycle, including both the upward  
114 trend from 1970 to 2005 and the downward trend from 2006 to 2018 have been associ-  
115 ated by some authors with natural variability (e.g., Yan et al., 2017, and others). How-  
116 ever, studies using CMIP5 historical runs simulated an ensemble-mean AMO that is sig-  
117 nificantly correlated with the observed AMO, suggesting that the recent historical vari-  
118 ability could be a consequence of radiative forcing (Clement et al., 2015; Bellomo et al.,  
119 2018). The future projections of TC frequency are subject to a similar degree of debate.

120 Many studies have suggested that the future should see a decline in the numbers of the  
 121 Atlantic TCs with warming (e.g., Knutson et al., 2010, and others), with a few excep-  
 122 tions (Emanuel, 2013; Bhatia et al., 2018; Vecchi et al., 2019).

123 The second objective is to compare historical simulations with observations to un-  
 124 derstand which modeling dataset is more consistent with the observations (Brunner et  
 125 al., 2020). Such analysis can provide guidance whether to favor one model over another,  
 126 which is especially useful for reducing uncertainty when the projections cover a wide range  
 127 even with an opposite sign, such as the projections of the divergent scenarios in the global  
 128 tropical cyclone genesis (i.e. Sobel et al., 2021). Lastly, we will assess hurricane risk over  
 129 a set of selected line gates in the present and future climates. Strictly speaking, risk in-  
 130 cludes severity of the hazard, exposure, and vulnerability of the properties of interest.  
 131 Only the hazard component is examined here.

## 132 2 Data, Experimental design and Method

### 133 2.1 Tropical cyclone datasets

#### 134 2.1.1 Observations

135 For reference, we use best-track data from National Hurricane Center obtained via  
 136 International Best Track Archive for Climate Stewardship v04r00 IBTrACS (Knapp et  
 137 al., 2010). We use 6-hourly storm positions (in longitude and latitude) and maximum  
 138 wind speeds (kt) from 1951 to 2020. Storm forward speed is derived from the position  
 139 data. We use only storms whose lifetime maximum intensity (LMI) reaches tropical storm  
 140 (TS) strength, 34 kt. Hurricanes are referred to storms with LMI of at least 64 kt.

#### 141 2.1.2 Synthetic events from the CHAZ model

142 The first set of model TCs used here consists of synthetic storm tracks from the  
 143 Columbia (tropical cyclone) Hazard (CHAZ) model (Lee et al., 2018). CHAZ is a statistical-  
 144 dynamical downscaling model that generates synthetic storms whose properties depend  
 145 on environmental conditions. The environmental conditions can come from an observation-  
 146 based reanalysis or a global climate model. There is no feedback of downscaled TC ac-  
 147 tivity to the global models. Three components in CHAZ describe storm formation and  
 148 subsequent evolution until shortly after landfall (or dissipation): the cyclone genesis in-  
 149 dex (TCGI; Tippet et al., 2011), the beta-advection track model (Emanuel, 2008), and  
 150 an auto-regressive intensity model (Lee et al., 2015, 2016). Details about CHAZ are re-  
 151 ported in Lee et al. (2018). The environmental variables required by the model are Po-  
 152 tential Intensity (Bister & Emanuel, 1997), deep-layer (850 to 250 hPa) vertical wind shear,  
 153 and one or more moisture variables: column integral relative humidity (CRH) and/or  
 154 column integral saturation deficit (SD), the absolute vorticity at 850 hPa, and the steer-  
 155 ing flow. The choice of moisture variables will prove particularly important in what fol-  
 156 lows. Both variables are calculated following Bretherton et al. (2004). The simulated trop-  
 157 ical cyclone activity in CHAZ, at global and basin scales, in both current and projected  
 158 future climates have been discussed in detail in Lee et al. (2018) and Lee et al. (2020),  
 159 respectively. The CHAZ model has been used for case studies in Texas (Hassanzadeh  
 160 et al., 2020), New York (Lee et al., 2022), Mumbai, India (Sobel, Lee, et al., 2019) and  
 161 the Philippines (Baldwin et al. 2022). Meiler et al. (2022) found that losses estimated  
 162 from CHAZ are comparable to those estimated using comparable academic tropical cy-  
 163 clone hazard models from Emanuel (2013) and Bloemendaal et al. (2020).

164 In this study, we use CHAZ to downscale five CMIP5 models (Taylor et al., 2012)  
 165 over the Atlantic basin. They are the National Center for Atmospheric Research (NCAR)  
 166 Community Climate System Model 4 (CCSM4) (Gent et al., 2011), the Geophysical Fluid  
 167 Dynamics Laboratory Climate Model version 3 (GFDL-CM3) (Donner et al., 2011), the

168 United Kingdom Meteorological Office Hadley Center Global Environment Model ver-  
 169 sion 2 Earth System (HadGEM2-ES) (Jones et al., 2011), the Max Planck Institute Earth  
 170 System Model Medium Resolution (MPI-ESM-MR) (Zanchettin et al., 2012), and the  
 171 Model for Interdisciplinary Research Climate Version 5 (MIROC5) (Watanabe et al., 2010)  
 172 from the University of Tokyo Center for Climate System Research, National Institute  
 173 for Environmental Studies, Japan, Japan Agency for Marine-Earth Science.

174 CHAZ’s projections of annual TC frequency, both in the Atlantic and globally, are  
 175 sensitive to whether CRH and SD are used in TCGI. Using TCGI with CRH leads to  
 176 a projected increase in global (and Atlantic) TC frequency, while SD leads to a projected  
 177 decrease (Lee et al., 2020). CRH and SD both measure the degree of the saturation of  
 178 the atmosphere with SD being the difference between the column integrated water vapor  
 179 and the same quantity at saturation, and CRH being their ratio. As saturated water  
 180 vapor increases with temperature in a warming climate, CRH remains close to con-  
 181 stant and SD decreases (Camargo et al., 2014). In the current climate, however, the be-  
 182 havior of these two variables are qualitatively similar, and the two TCGI formulations  
 183 yield similar results for the historical period, meaning that the historical evidence is in-  
 184 adequate to determine which of the two is more correct. Arguably, SD better reflects the  
 185 increase in the thermodynamic inhibition of TC formation in a warming climate (Emanuel,  
 186 1989, 2022), but the gaps in our understanding of the relationship between climate and  
 187 tropical cyclone frequency are so substantial that we do not view this argument as dis-  
 188 positive (Sobel et al., 2021). The diverging annual frequency projections from CHAZ thus,  
 189 in our view, reflects the broader state of the science, in that we have low confidence re-  
 190 garding whether one should expect more or fewer hurricanes as climate warms(i.e. Ca-  
 191 margo et al., 2020; Vecchi et al., 2019; Sugi et al., 2020). One reason for the low con-  
 192 fidence in TC frequency projection is the lack of theoretical understanding of tropical  
 193 cyclone genesis, and we refer the readers to a review article by Sobel et al. (2021) for a  
 194 detailed discussion.

195 Since total TC hazard and risk depend inextricably on TC frequency and we lack  
 196 a strong basis for choosing between SD and CRH, the sensitivity to the humidity vari-  
 197 able in our results causes a deep uncertainty in the projected risk. This uncertainty will  
 198 remain in the present study, in that we performed separate sets of simulations with ei-  
 199 ther CRH or SD as the humidity variable in the genesis module, referred to as CHAZ<sub>CRH</sub>  
 200 and CHAZ<sub>SD</sub>.

### 201 *2.1.3 Directly simulated hurricanes from General Circulation Models*

202 In addition to the CHAZ downscaling simulations described above, we use storms  
 203 tracked in a set of relatively high-resolution, i.e., tropical cyclone-permitting, global cli-  
 204 mate models. The first one is the 60-km MRI-AGCM3.2H large-ensemble simulation from  
 205 Mizuta et al. (2017) (MRI-LENS). Tropical cyclones in that model was discussed in Yoshida  
 206 et al. (2017). The second one is the 25-km High-Resolution Community Atmospheric Model  
 207 version 5 simulations, CAM5 (Wehner et al., 2014, 2015). Next, we use storms tracked  
 208 in the Coupled Model Intercomparison Project Phase 6 (CMIP6) (Eyring et al., 2016)  
 209 High Resolution Model Intercomparison Project (HighResMIP) (Haarsma et al., 2016).  
 210 Following Roberts et al. (2020) and Roberts et al. (2020), we use storms from CMCC-  
 211 CM2 (Cherchi et al., 2019), CNRM-CM6 (Voldoire et al., 2019), EC-Earth3P-HR (Haarsma  
 212 et al., 2020), HadGEM3-GC3.1 (Roberts et al., 2019), and MPI-ESM1.2 (Gutjahr et al.,  
 213 2019). There are two HighResMIP configurations, one is forced with prescribed SST while  
 214 the other is fully coupled. We only use the simulations from the fully coupled configu-  
 215 ration which allows natural variability to occur freely during the historical period. To  
 216 understand the sensitivity of model performance to the TC trackers, HighResMIP storms  
 217 are tracked by TRACK (Hodges et al., 2017) and TempestExtremes (Ullrich & Zarzy-  
 218 cki, 2017; Zarzycki & Ullrich, 2017; Ullrich et al., 2021), and both event sets are used

219 here. For convenience, we label modeled TCs from HighResMIP tracked with Tempest-  
 220 Extremes as Hi-TempExt and those tracked with TRACK as Hi-TRACK.

## 221 2.2 Experimental design

222 Except in MRI-LENS and CAM5, we use model TCs from the historical, near-term  
 223 future, and pre-industrial control (piC, no anthropogenic forcing) scenario simulations.  
 224 Note that the time range covered in each period varies by model. For the historical pe-  
 225 riod, they are 1951-2005 for CHAZ<sub>CRH</sub> and CHAZ<sub>SD</sub>, 1950-2010 for MRI-LENS, 1996-  
 226 2016 for CAM5, and 1951-2014 for the two HighResMIP datasets. In the future period,  
 227 CHAZ<sub>CRH</sub> and CHAZ<sub>SD</sub> contain storms from 2006-2040 under Representative Concentra-  
 228 tion Pathway 8.5 (rcp8.5) while HighResMIP storms are from 2015-2050 under Shared  
 229 Socioeconomic Pathways5-85 (ssp585). Both are high-emission scenarios with an addi-  
 230 tional radiative forcing of  $8.5 \text{ W m}^{-2}$  by the year 2100 (Riahi et al., 2017) in ssp585 which  
 231 considers a fossil-fueled development. Warming climate simulations for MRI-LENS and  
 232 CAM5 are under a  $4^\circ\text{C}$  (Yoshida et al., 2017) and  $1.5^\circ\text{C}$  warming (Wehner et al., 2018)  
 233 scenarios and thus are not used here. In piC, the labeling of year is arbitrarily in all datasets  
 234 as all years are equivalent. The MIR-LENS and CAM5 piC simulations are exceptions.  
 235 In MRI-LENS and CAM5, the observed SST information is given in both historical and  
 236 piC simulations as a lower boundary, but the long-term trend is removed in the piC sim-  
 237 ulations. In other words, MIR-LENS and CAM5 piC simulations still contain observed  
 238 variation. The piC simulations in MRI-LENS, called “no-warming” in Mizuta et al. (2017)  
 239 and those in CAM5, following “Nat-Hist” in Stone et al. (2019), are designed with an  
 240 underlying assumption that that only the linear trend is anthropogenic forced, not the  
 241 variability, which, as we will discussed in the next Section, is debatable.

242 In each period, the CHAZ model was used to generate 20 track ensemble members  
 243 per CMIP5 model and each track has 40 intensity ensembles (100 CMIP5 track ensemble  
 244 members and 4000 considering intensity ensemble), as is possible because the CHAZ  
 245 intensity module has a stochastic component. Hi-TRACK has 7 members (5 global cli-  
 246 mate models and two of them have 2 ensemble members) and Hi-TempExt has 6 (4 global  
 247 climate models and two of them have 2 ensemble members). MRI-LENS has 100 ensemble  
 248 members while CAM5 has 5. The data properties are listed in Table 1.

## 249 2.3 Frequency adjustment

There are biases in model TCs, because of biases in the models that generate them,  
 including the CHAZ model itself as well as the CMIP5 models from which CHAZ ob-  
 tains its environmental conditions, and the high-resolution global climate models used  
 here. In particular, all models have biases in TC frequency (Table 1), and directly-simulated  
 hurricanes from high-resolution global climate models have low-intensity biases, in gen-  
 eral, as the grid spacings of these models are too coarse to capture the full range of ob-  
 served hurricane strengths (e.g., Yoshida et al., 2017; Moon et al., 2022, and others). Here  
 we address only the frequency biases. Specifically, we derive an adjustment by compar-  
 ing the basin-wide annual TC frequency of models’ historical simulations to that of the  
 observations from the same period. The same adjustment will then be applied to both  
 historical and future simulations. Similarly, we compare the annual frequency of the piC  
 simulations to the observations to adjust piC’s annual frequency. In Lee et al. (2018) and  
 Lee et al. (2020), the basin-wide frequency adjustment is a multiplicative factor to en-  
 sure that the mean annual frequency over a basin in CHAZ is consistent to that in ob-  
 servations. However, some high-resolution global climate models used here, such MRI-  
 LENS, generate zero TCs in some years. A multiplicative factor would result in larger  
 variability but still have zeros in these years, which is unrealistic. Thus, here the basin-

wide frequency is adjusted as:

$$f_{\text{adj}} = \sigma_{\text{obs}} \times \frac{f_{\text{ori}} - \mu_{\text{model|ref}}}{\sigma_{\text{model|ref}}} + \mu_{\text{obs}}, \quad (1)$$

where  $f$  indicates annual frequency (each year) with the subscript indicating after ( $_{\text{adj}}$ ) or before ( $_{\text{ori}}$ ) frequency adjustment.  $\mu$  and  $\sigma$  are the mean and standard deviation of the frequency and the subscript indicates whether it is from simulations ( $_{\text{model}}$ ) or observations ( $_{\text{obs}}$ ). As we want to retain the climate change signal, reference  $\mu$  and  $\sigma$  ( $\mu_{\text{model|ref}}$  and  $\sigma_{\text{model|ref}}$ ) for adjusting frequencies in both historical and future simulations in each dataset are from its respective historical simulation. Observations are calculated from their respective historical periods. To adjust the annual frequencies of the piC simulations,  $\mu_{\text{model|ref}}$  and  $\sigma_{\text{model|ref}}$  are from piC. Biases in annual TC frequency of the piC simulations are different to those in the historical simulations. As we will discuss later, a basin-wide frequency adjustment may not correct regional biases, because model biases can have spatial dependence. When desired (in Section 5), we apply a multiplicative factor to ensure the annual frequency at storm with intensity greater than 40 kt in these data sets are consistent to observations, which is the same as the bias-correction approach used in (Lee et al., 2022).

An underlying assumption of our approach to bias correction, in common with many climate change studies, is that the bias of any given model remains the same in projected future climate periods as it is in the present, so that the influence of the projected climate change can still be captured when comparing simulations between rcp and hist periods. This assumption is analogous to that used to remove climatological biases in surface temperature and other quantities from the climate models themselves in global warming projections, for example those by the Intergovernmental Panel on Climate Change (Solomon et al., 2007). While this assumption of constant biases can be questioned, it is a simple assumption, and there is no empirical basis on which to base any more complex assumption one. Still, we will discuss the impacts of frequency adjustments on our findings.

## 2.4 Trend analysis

To calculate trends of TC activity, we fit second-order Legendre polynomials:

$$\hat{y} = a_0 + a_1x + \frac{a_2}{2}(3x^2 - 1), \quad x \in [-1, 1] \quad (2)$$

to the time series of the variables of interest from observations and model simulations. In Equation (2),  $x$  is years scaled to interval of  $[-1, 1]$ ,  $\hat{y}$  represents the fitted variables, the coefficient  $a_1$  shows linear trends and  $a_2$  shows quadratic trends. Considering quadratic trends allows the possibility that the observed multi-decadal variability is in fact forced (Clement et al., 2015; Bellomo et al., 2018). Here, we ask whether or not the observed trends lie within the ensemble spread from simulations. For example, if the observed trend is outside of the range of piC simulations but is within those from historical simulations, then the observed change (e.g., upward trend or increasing curvature) is unlikely to have occurred without anthropogenic forcing. When comparing the trends between observations and simulations,  $a_1$  and  $a_2$  are scaled back so that they have units of the variable's unit per year ( $yr^{-1}$ ) and per year square ( $yr^{-2}$ ), respectively.

## 3 Trend and multi-decadal variability

### 3.1 Atlantic TC frequency

We first examine the Atlantic TC frequency trends in the historical (present) climate and from historical to the warming future (i.e., using simulations from both historical and future periods). Figure 1a and b show the ensemble means of the time series of Atlantic hurricane frequency, i.e., the averaged total number of storms in the basin

293 each year whose maximum sustained winds exceed 34 kt from each dataset. The small  
 294 wiggles may be sampling variability. Figures 1c and d show the ensemble spread. By con-  
 295 struction, the time-mean annual frequency for each dataset over its respective histori-  
 296 cal period will be identical to observations after the frequency adjustment (Eq. (1)). The  
 297 original annual frequency of each dataset is shown in Table 1. Before 2000, the differ-  
 298 ent simulations are, by eye at least, indistinguishable in their overall envelopes, with none  
 299 showing any particular trend, and the observations (black thick line) lying well within  
 300 their spread (shown in Figure 1c). After 2000, the CHAZ<sub>SD</sub> (orange thick line) and CHAZ<sub>CRH</sub>  
 301 (blue thick line) results begin to diverge, with CHAZ<sub>SD</sub> showing a decreasing TC fre-  
 302 quency and CHAZ<sub>CRH</sub> showing an increasing TC frequency. It is possible that this is  
 303 related to the fact that the rcp8.5 scenario starts after 2005. The two HiResMIP datasets  
 304 show no considerable trend in the historical period but a sharp dip after 2030. The ssp585  
 305 scenario in HiResMIP starts after 2015, though. Hi-TRACK's annual TC frequency climbs  
 306 up by 2040. Roberts et al. (2020) reported that both Hi-TRACK and Hi-TempExt project  
 307 a reduction of ensemble mean annual frequency (less than 10%) from 1950-1980 to 2020-  
 308 2050, but the spread covers zero, indicating low confidence to the mean trend.

309 Figures 1b and 1d show analogous results for piC simulations. Note that the years  
 310 in the x-axis are not real; these labels are placed so we can compare the simulated trends  
 311 to the observed trend and those in Figures 1a and 1c. Two exceptions are MRI-LENS  
 312 and CAM5 simulations; both are uncoupled atmospheric models and forced with observed  
 313 SST with anthropogenic trend removed (See Section 2 for details). In the Figure 1b, CHAZ<sub>CRH</sub>  
 314 and CHAZ<sub>SD</sub> results do not diverge. There is no dip in the Hi-TRACK or Hi-TempExt.  
 315 Clearly, the separation between CHAZ<sub>CRH</sub> and CHAZ<sub>SD</sub> and the dip in the two High-  
 316 ResMIP datasets in Figure 1a represent forced responses.

317 Next we conduct the trend analyses of the annual TC frequency in Figure 1 using  
 318 second-order Legendre polynomials fits (Eq. (2)). As an example, Fig. 2a shows the anal-  
 319 ysis using the CHAZ<sub>CRH</sub> simulations and the observations. The observed fit (dashed black  
 320 line) has an upward trend of 0.085 storm year<sup>-2</sup> and a positive curvature of 0.005 storm year<sup>-2</sup>  
 321 (shown as the black line in Figs. 2b and 2c). The existence of a linear trend means that  
 322 there is an overall increasing trend in storm activity since 1951 while the quadratic terms  
 323 captures the multi-decadal variability, with high activity in the 1950s-1960s, low in the  
 324 1970s-80s, and high after that, which recent research suggests may be a forced signal rather  
 325 than natural variability (Clement et al., 2015; Bellomo et al., 2018). In Fig. 2a, the poly-  
 326 nomial fits of CHAZ<sub>CRH</sub> simulations from historical only (light blue dashed line) and from  
 327 historical to future (dark blue dashed line) both show an small upward curve while the  
 328 polynomial fit derived from the piC simulations (gray dashed line) is quite flat.

329 The ranges of the fit parameters from all ensemble members in each dataset are  
 330 also shown in Figures 2b-c. The observed linear trend are above most of the piC sim-  
 331 ulations except those from CAM5. However, CAM5 has only 10-years of simulations, which  
 332 is too short to be compared with 70-years of observations. The observed quadratic term  
 333 lies within the 25-75 percentile ensemble ranges of piC simulations from CHAZ<sub>CRH</sub>, CHAZ<sub>SD</sub>,  
 334 and MRI-LENS. It is outside of the ensemble ranges from two HighResMIP datasets which  
 335 have quadratic terms close to zero. The observed linear trend is at top 25 percentile (75-  
 336 100 percentile) of the hist simulations of CHAZ<sub>CRH</sub>, CHAZ<sub>SD</sub>, and is marginally included  
 337 in the simulations of MRI-LENS; the observed quadratic term is within the 25-75 per-  
 338 centile range the CHAZ<sub>CRH</sub> and MRI-LENS, and is at top 25 percentile in CHAZ<sub>SD</sub>. Only  
 339 the fit linear trend derived from historical + future simulations of the CHAZ<sub>CRH</sub> include  
 340 the observed value. For the quadratic trend, the fit parameter derived from CHAZ<sub>CRH</sub>  
 341 and CHAZ<sub>SD</sub> include the observed values but they are at top and bottom 25 percentile  
 342 range, respectively. (We do not use any warming simulations from CAM5 and MRI-LENS.)

343 Generally speaking, the polynomial fit analysis suggests that, first, CHAZ<sub>CRH</sub>, CHAZ<sub>SD</sub>  
 344 and MRI-LENS are better in capturing the observed trend and multi-decadal variabil-  
 345 ity as their historical spread covers the observed values. However, CAM5 has only 10 years

346 of data with 5 ensemble members and while Hi-TRACK and Hi-TempExt have only, re-  
 347 spectively, 7 and 6 ensemble members. These three datasets may be under-sampled. Sec-  
 348 ond, the observed linear trend is outside the spread of CHAZ<sub>CRH</sub>, CHAZ<sub>SD</sub> and MRI-  
 349 LENS' piC simulations but within the spread of these models' hist simulations, indicat-  
 350 ing that anthropogenic forcing is necessary to capture the upward trend in the past decades.  
 351 On the other hand, we can not rule out the possibility of the recent upward curvature  
 352 trend is within the range of natural variability. Although the MRI-LENS' piC simula-  
 353 tions is forced with the observed SST (with long-term trend removed) which results in  
 354 the upward curvature term right on top of observed values in Figure 2c. Simulations from  
 355 CHAZ<sub>CRH</sub> suggest that that anthropogenic forcing helps to capture the upward curva-  
 356 ture trend. Third, when considering the future period as well, the mean of CHAZ<sub>CRH</sub>  
 357 shows an upward trend, the mean of CHAZ<sub>SD</sub> shows a downward trend, while the mean  
 358 of the two HighResMIP simulations are close to zero. However, we have low confidence  
 359 in the projections as they include zero. Thus, we can not say for sure that the positive  
 360 linear and quadratic terms will continue into the future.

361 It should be noted that without the basin-wide frequency adjustment (not shown),  
 362 the observed linear and quadratic terms lie outside of the spread of MRI-LENS, Hi-TempExt  
 363 and Hi-TRACK in all three periods. They are within the spread of CHAZ<sub>CRH</sub> and CHAZ<sub>SD</sub>  
 364 simulations in piC and historical periods. With additional data from 2006 to 2040, only  
 365 CHAZ<sub>CRH</sub> shows such an upward trend will continue into the future.

### 366 3.2 Intensity and storm motion

367 Figure 2d shows the fit parameters of Atlantic TC lifetime maximum intensity (LMI).  
 368 Specifically, we look at the variability of the 95th percentile of LMI (LMI95), for which  
 369 an upward trend has been found in observations (Kossin et al., 2013). Here we focus on  
 370 the linear term only. There is an upward trend in the observations, meaning that the ex-  
 371 treme tail of observed intensity has increased with time, consistent with previous stud-  
 372 ies (e.g., Knutson et al., 2020a, and others). The positive linear trend is captured by the  
 373 ensemble spreads of two CHAZ datasets and those of MRI-LENS and CAM5 at both piC  
 374 and historical periods. It is outside of the ensemble spread of all simulations from from  
 375 Hi-TRACK and Hi-TempExt. Thus, at least from CHAZ<sub>CRH</sub>, CHAZ<sub>SD</sub>, MRI-LENS, and  
 376 CAM5, we can not rule out that the recent upward trend in the LMI95 is due to nat-  
 377 ural variability. When looking into the future, only the means of CHAZ<sub>CRH</sub> is positive  
 378 and the means of CHAZ<sub>SD</sub>, Hi-TempExt and Hi-TRACK are close to zero. Similar to  
 379 the results from TC frequency, the ensemble spread in Figure 2d include zero in the whole  
 380 historical + future periods, indicating, again, low-confidence in the projected changes.

381 Figure 2e shows the analysis for translation speed. Consistent with (Kossin, 2018),  
 382 the observations show a clear downward trend in the storm motion. This trend is within  
 383 ensemble spread in all periods, including piC, for all models, except the simulations from  
 384 Hi-TempExt. However, the mean and the 25-75 percentile ensemble spreads in these datasets  
 385 move toward different directions from piC to hist to hist +future periods. The Hi-Track  
 386 and MRL-LENS hist simulations show upward trends in the storm motion and this up-  
 387 ward trends continues in to the future. The differences in mean and 25-75 percentile en-  
 388 semble spreads from CHAZ<sub>CRH</sub> and CHAZ<sub>SD</sub> from these three period are small. The piC  
 389 and hist simulations from CAM5 shows that anthropogenic forcing may lead to a strong  
 390 downward trend in storm motion but again CAM5 simulations are shorter than do the  
 391 data from the other models. It seems unjustified, based on this set of models, to attribute  
 392 the observed slowing down to anthropogenic forcing. It also noteworthy that at a regional  
 393 scale, CHAZ projected an upward trend in storm motion speed for TCs affecting Texas  
 394 (Hassanzadeh et al., 2020) and an a downward trend for storms impacting New York (Lee  
 395 et al., 2022). Spatially inhomogeneous changes may dilute the basin-wide signal.

#### 4 Likelihood comparison

Figure 2 shows that the simulated trend in historical and historical + future vary from one dataset to another. This is especially true for the TC frequency projections between CHAZ<sub>CRH</sub> and CHAZ<sub>SD</sub>, but a qualitatively similar result, including both increasing and decreasing trends, holds for the rest of our ensemble of opportunity. It is natural to ask whether we can develop some criteria for determining which is correct. In climate science, multi-model ensemble mean is a common approach to obtain the consensus from multiple global climate models. However, such approach is only adequate when the ensemble spread represents variations that can be considered random, as might be the case with typical aleatoric uncertainties. The divergent scenarios in the frequency projections are a consequence of the epistemic uncertainty due to the lack of a satisfactory scientific understanding of tropical cyclone frequency (Sobel et al., 2021; Emanuel, 2022) and thus the multi-model mean may not be meaningful in this case. We can, however, use likelihood analysis, in which the probabilities that the observations occurs in the model simulated distribution were computed. Thus, we can determine which simulation the observation is more consistent with. This is similar to the Likelihood Skill Score used for evaluating weather and climate predictions (Barnston et al., 2010).

Specifically, we first assume that annual hurricane frequency is drawn from a Poisson distribution whose mean ( $\lambda_t$ ) has a trend in time ( $\lambda_t = at + b$ ). We then obtain  $a$  and  $b$  of each dataset by fitting the model annual TC frequency to a Poisson regression. We do so for all simulations with data throughout 2021 (up to 2005 for CAM5 and 2010 for MRI-LENS). Note that with  $a$  and  $b$ , we can derive  $\lambda_t$  even for years beyond the data coverage period, i.e., we can estimate  $f_{2020}$  with  $a$  and  $b$  derived from CAM5 data. The yearly likelihoods ( $L_t$ ) of the observed frequencies are assigned based on the Poisson distribution with a given  $\lambda_t$ . For example, the likelihood CHAZ<sub>CRH</sub> simulations will generate 29 TCs as observed in 2005 is 0.08%, which is based on a Poisson distribution with  $\lambda_{2005} = 15.7$ . The same calculation is applied to piC simulations, and the derived likelihood is denoted  $L_{piC,t}$ . For a given year, we then compare the log likelihood ratios  $L_t$  and  $L_{piC,t}$  (i.e.,  $\log(L_t/L_{piC,t}) = \log(L_t) - \log(L_{piC,t})$ ). If this ratio is larger than 0, the observations are more consistent with the simulations with anthropogenic forcing than with the piC simulations and vice versa.

We start by comparing the likelihoods of simulations with anthropogenic forcing to those with piC simulations (i.e.,  $\log(L_t/L_{piC,t})$  in Figure 3. The ratios of the likelihoods jointly up to 2020 (numbers on the upper-left in all panels) suggest that the observations are more consistent with the simulations with anthropogenic forcing than those without in CHAZ<sub>CRH</sub>, MRI-LENS, and Hi-TempExt. The annual likelihood ratios from these three datasets further show higher annual likelihood ( $L_t$ ) for the observed annual frequency values during the period of high TC activity in 1950-1970 and after 2000 while higher  $L_{piC,t}$  is found during 1970-2000. This is because there are upward trends in the simulated annual frequency in these three datasets when compared to in piC (Figure 2a). As  $\lambda_t$  increases with time, the distributions from these three datasets shift right with time and thus give greater likelihood to the high observed annual frequency when compared to those derived from piC simulations in which  $\lambda_t$  is close to constant in time. In contrast, CHAZ<sub>SD</sub> has a downward trend and its,  $\lambda_t$  shifts left in time and leads to lower likelihood when observed values are high. Consequently, we see a higher  $L_{piC,t}$  during high TC activity periods and higher  $L_t$  during the inactive TC seasons in CHAZ<sub>SD</sub>. The frequency slopes obtained from piC and hist in the Hi-TRACK data are similar and thus their log likelihood ratio is close to zero.

When we consider the likelihood over the whole observational period, we calculate the average of the likelihood, i.e., the roots of  $\prod_{2021}^{1950} L_t$  from all five datasets. Between CHAZ<sub>CRH</sub> and CHAZ<sub>SD</sub>, observations are more consistent with CHAZ<sub>CRH</sub>, which has an averaged likelihood of 5.24%, than to CHAZ<sub>SD</sub> which has the averaged likelihood of

448 3.46%. Among the five datasets, CHAZ<sub>CRH</sub> has highest likelihood, followed by Hi-TempExt  
 449 (5.13%), MRI-LENS (5.1%), Hi-TRACK (5.04%), and CAM5 (3.6%).

450 The basin-wide frequency adjustment (Eq. (1)) that we performed to correct model  
 451 biases is expected to affect the results of the likelihood analysis, because the frequency  
 452 adjustment both shifts the mean of the model's TC annual frequency distributions and  
 453 changes their shapes. The annual frequency distributions from historical and piC sim-  
 454 ulations are more distinct in the datasets without frequency adjustment, which indeed  
 455 leads to larger log likelihood ratios (not shown). Without the frequency adjustment, the  
 456 observed TC annual frequencies are more consistent with the historical simulations in  
 457 CHAZ<sub>CRH</sub>, CHAZ<sub>SD</sub>, and MRI-LENS than in their respective piC simulations due to the  
 458 large bias in these piC simulations. Without basin-wide TC frequency adjustment, Hi-  
 459 TRACK has the greatest averaged likelihood, followed by CAM5, CHAZ<sub>CRH</sub>, Hi-TempExt,  
 460 CHAZ<sub>SD</sub>, and MRI-LENS. MRI-LENS has the lowest likelihood because of its low bias  
 461 and zero storms in some years.

## 462 5 Climate change and regional hurricane risk at three line gates

463 Now we compute regional hurricane risk, from hazard perspective only, represented  
 464 by return periods of storms of given wind intensities passing through pre-defined gates,  
 465 derived using simulations from historical and future periods. We use simulations from  
 466 CHAZ<sub>CRH</sub>, CHAZ<sub>SD</sub>, Hi-TRACK, and Hi-TempExt. The three line gates used here (black  
 467 lines in Figures 4a-c) are the main development region (MDR) gate which can be thought  
 468 of as delineating Atlantic TC hazard in a general sense – how many storms form, and  
 469 at what intensity and move from the MDR toward the US and Caribbean Islands; the  
 470 GoM gate which records TC activity for those that enter the Gulf of Mexico; and the  
 471 NE gate which is parallel to a portion of the Northeastern US coast. As discussed ear-  
 472 lier (Section 2.3), to obtain more realistic return period curves for regional hurricane risk  
 473 assessment, we use a more localized frequency adjustment. As an example, Figures 4d-  
 474 4f show historical simulations from CHAZ<sub>CRH</sub> with basin-wide and regional frequency  
 475 adjustments (Eq. (1)). While the basin-wide frequency adjustment (dashed lines) yields  
 476 a TC frequency close to observations at the GoM gate, CHAZ<sub>CRH</sub> still overestimates storm  
 477 activity at the MDR gate and underestimates storm activity at the NE gate. The regional  
 478 frequency adjustment shifts the simulated return period curves (solid line, local adjust-  
 479 ment) by matching the return periods at 40 kt to the values derived from observations  
 480 (see Section 2.3 for details). In terms of the shape of the return period curve, as well as  
 481 the return periods at high intensities, CHAZ<sub>CRH</sub> performs better at the MDR gate than  
 482 at the GoM gate. It is difficult to directly compare the modeled curves to the observa-  
 483 tions at the NE gate, due to the significant underestimation of overall TC frequency at  
 484 the latter. However, even there, the shapes of the observed and modeled return period  
 485 curves are similar.

486 To show the changes in return periods between historical and future periods, Fig-  
 487 ures 4g-i show the return period curves derived from the four datasets that have rcp8.5/spp585  
 488 warming scenarios available. We use model storms from all ensemble members. Low-intensity  
 489 biases in the Hi-TRACK and Hi-TempExt lead to an underestimate of the TC risk. High-  
 490 ResMIP models barely simulate storms with major hurricane wind strength (Roberts et  
 491 al., 2020; Moon et al., 2022). The return period curves of CHAZ<sub>CRH</sub> and CHAZ<sub>SD</sub> hist  
 492 simulations are close to each other. The differences between simulations from historical  
 493 period and those from historical and future periods, i.e., the differences between the dashed  
 494 and solid lines, are small for the two CHAZ datasets in Figures 4g-i. Likewise the his-  
 495 torical and future period curves of GoM and NE gates for Hi-TRACK and Hi-TempExt  
 496 nearly indistinguishable. At the MDR gate, both Hi-TRACK and Hi-TempExt suggest  
 497 increases in the TC risk.

498 To make these differences more evident, we list the percentage changes in annual  
 499 TC frequency exceeding each Saffir-Simpson category on both sides of each panel in Fig-  
 500 ures 4g-i. As expected, there is an overall increase in the storm frequency at all thresh-  
 501 olds from historical to future periods for CHAZ<sub>CRH</sub> while there is an overall decrease for  
 502 CHAZ<sub>SD</sub>, consistent with the results in Figures 1 and 2a. The percentage changes are  
 503 larger at higher intensity thresholds in the CHAZ<sub>CRH</sub> but they are of similar or smaller  
 504 magnitude throughout the Saffir-Simpson categories in the CHAZ<sub>SD</sub>. This is probably  
 505 due to the increase in storm intensity as climate warms in CHAZ<sub>CRH</sub> and CHAZ<sub>SD</sub>.

506 The changes in the frequency of exceedance at the three gates from Hi-TRACK and  
 507 Hi-TempExt are not the same sign. Hi-TRACK shows a 67% decrease of Category 1+  
 508 ( $\leq 64$  kt) at the MDR gate but a 65 % increase at GoM gate. At the NE gate, Hi-TRACK  
 509 shows an 14 and 38% increase in the frequency of Category 1+ and 2+ storms, respec-  
 510 tively. Hi-TempExt shows a 68% decrease and 16% increase of Category 1+ storms at  
 511 the MDR and GoM gates, respectively. At the NE gate, it shows a 9% decrease and 92%  
 512 increase in the frequency of Category 1+ and 2+ storms. Storms from these two High-  
 513 ResMIP runs are undersampled and have low intensity biases (See Figure 7 in Roberts  
 514 et al. (2019)). The directly simulated storms are not suitable for risk assessment and these  
 515 numbers should be used with caution.

## 516 6 Discussion

517 The results of this study lead us to a view of Atlantic hurricane risk which requires  
 518 us to confront epistemic uncertainty. We have multiple sets of simulations which give dif-  
 519 ferent views of the risk, in particular more so as we look further into the future. TC fre-  
 520 quency increases in CHAZ<sub>CRH</sub> simulations and decreases in CHAZ<sub>SD</sub>, a difference that  
 521 hangs on a subtle modeling choice (saturation deficit vs. relative humidity as a predic-  
 522 tor of genesis). Changes in the high-resolution global climate model simulations are smaller,  
 523 but again their direction depends on which global climate models are considered.

524 The differences among these simulations are manifest not just in the future, but  
 525 also to some degree in the present, and the observations do not allow us to determine  
 526 with complete certainty which is correct. At present, no rigorous justification can be given  
 527 regarding which choice to make. Thus, all these outcomes — increasing, decreasing, and  
 528 no change in TC frequency in response to radiatively forced warming — have to be treated  
 529 as possible. One may favor a dataset over the others following the results of a statisti-  
 530 cal analysis, such as the likelihood analysis used here. Our calculations indicate that the  
 531 observations are somewhat more consistent with CHAZ<sub>CRH</sub>, followed by Hi-TempExt,  
 532 MRI-LENS, Hi-TRACK. However, the likelihood values among these four datasets are  
 533 close to each other, so it would not be justified to draw a definitive conclusion from this  
 534 analysis as to which model is most correct.

535 The epistemic uncertainty in CHAZ's projections on annual TC frequency comes  
 536 from our design of the CHAZ model, but the conclusion is that our results are consis-  
 537 tent with the level of broader understanding of TC frequency at present, including that  
 538 derived from the latest high-resolution models shown here as well as other downscaling  
 539 systems (Sobel et al., 2021). Other aspects of TC characteristics that could change with  
 540 anthropogenic climate change have been also evaluated here, namely the forward mo-  
 541 tion and LMI95, are less dramatically uncertain, although our analyses show that one  
 542 cannot rule out the role of natural variability. Still, the uncertainty regarding TC fre-  
 543 quency introduces a large uncertainty into any assessment of overall TC risk, since any  
 544 change of TC properties is only relevant to the extent that TCs actually occur.

545 The divergence between increasing and decreasing TC frequency scenarios becomes  
 546 most pronounced in the latter part of the 21st century, but has some impact on the present  
 547 and near future as well (Lee et al., 2020, 2022). In the situation when the change of fre-

548 quency is subtle, changes in other TC properties may lead to differences in regional TC  
549 risk assessment.

550 How one views the situation must ultimately be based on one's attitude towards  
551 risk and the consequences of being wrong in either direction. A priori, though, we ar-  
552 gue that the most rational way to treat epistemic uncertainty is to consider all outcomes  
553 contained in the results to be possible. In the present context, since the results contain  
554 possible outcomes in which TC risk — as estimated from a single model or subset of the  
555 entire multi-model ensemble — is increasing, that in itself means we should regard TC  
556 risk as increasing, at the highest level of understanding in which all available informa-  
557 tion is considered, even though there are other possible outcomes in which it is decreas-  
558 ing.

## 559 Open Research Section

560 CHAZ is an open-sourced model (<https://github.com/c13225/CHAZ>). IBTrACS  
561 data are available at (<https://www.ncdc.noaa.gov/ibtracs/>). Information for CMIP5  
562 data can be found at <https://pcmdi.llnl.gov/mips/cmip5/> and HighResMIP trop-  
563 ical cyclone information can be found at (<http://catalogue.ceda.ac.uk/uuid/e82a62d926d7448696a2b60c1925f8>).  
564 Underlying data for this publications are at ([https://github.com/c13225/Lee\\_etal\\_2023EarthsFuture](https://github.com/c13225/Lee_etal_2023EarthsFuture)).  
565

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**Table 1.** Data Characteristics

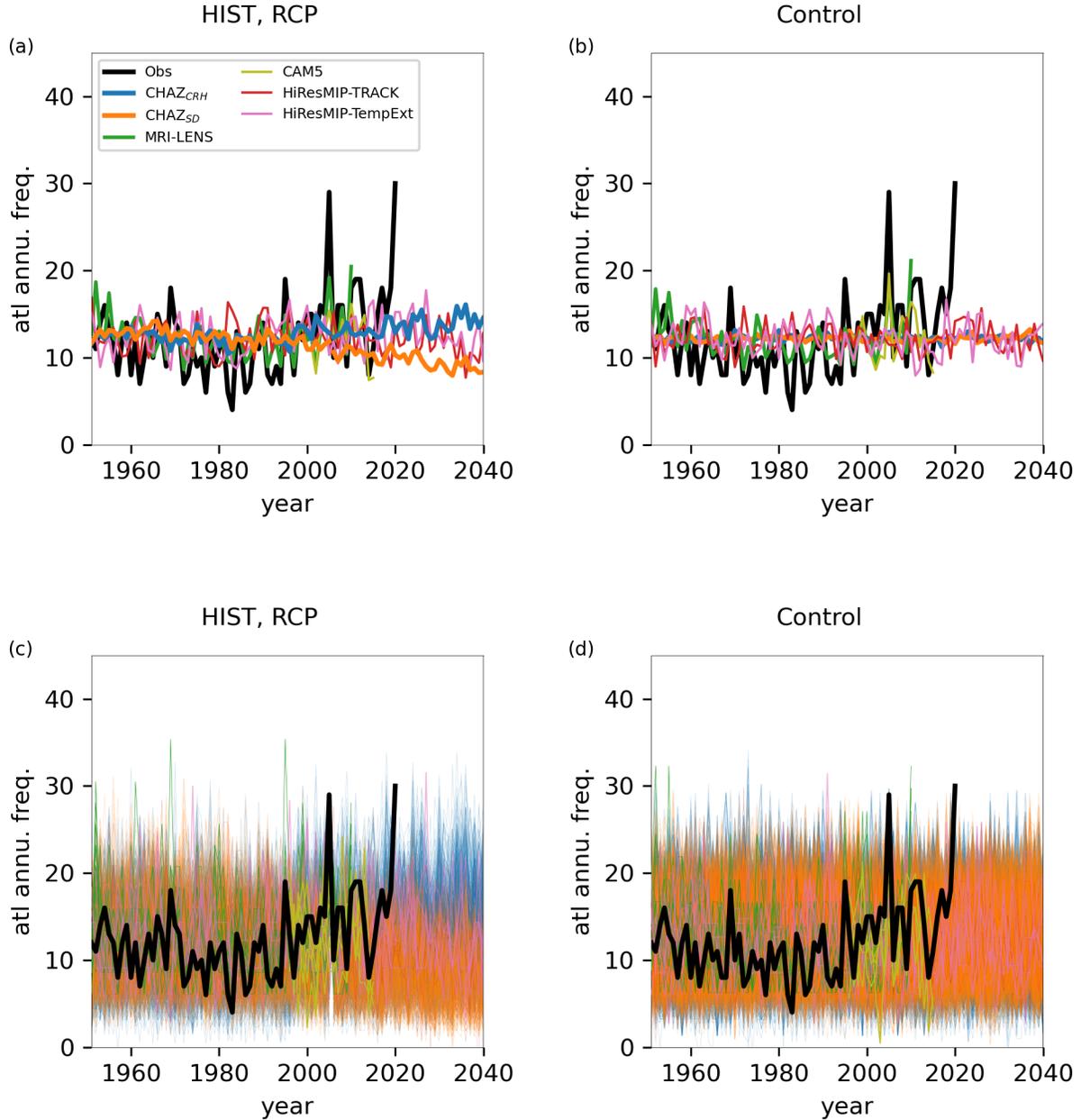
data	global climate models	resolution	ens	period	annual frequency
CHAZ <sub>CRH/SD</sub>	HadGEM2_ES	N/A	100	1951-2005;2006-2040; piC	8.8 /15.9
	CCSM4				11/16.1
	GFDL_CM3				16.5/19.1
	MPIESM_MR MIROC5				29.3/39.4 11.9/18.3
MRI-LENS	MRI-AGCM3.2H	60 km	100	1950-2010; piC	2.3
CAM5	CAM5	28km	5	1996-2005;piC	10.9
Hi-TRACK	CMCC-CM2-VHR4 (r1i1p1f1)	25 km	7	1951-2014;2015-2040;piC	5.0
	CNRM-CM6-1-HR(r1i1p1f2)	50 km			21.0
	EC-Earth3P-HR (r1i1p2f1)	50 km			6.8
	EC-Earth3P-HR (r2i1p2f1)	50 km			6.5
	HadGEM3-GC31-HH (r1i1p1f1)	50 km			21.5
	HadGEM3-GC31-HM (r1i1p1f1)	50 km			19.7
	MPI-ESM1-2-XR (r1i1p1f1)	50 km			4.5
Hi-TempExt	CNRM-CM6-1-HR(r1i1p1f2)	50 km	6	1951-2014;2015-2040;piC	13.4
	EC-Earth3P-HR (r1i1p2f1)	50 km			2
	EC-Earth3P-HR (r2i1p2f1)	50 km			2
	HadGEM3-GC31-HH (r1i1p1f1)	50 km			13.3
	HadGEM3-GC31-HM (r1i1p1f1)	50 km			12.4
MPI-ESM1-2-XR (r1i1p1f1)	50 km	0.63			

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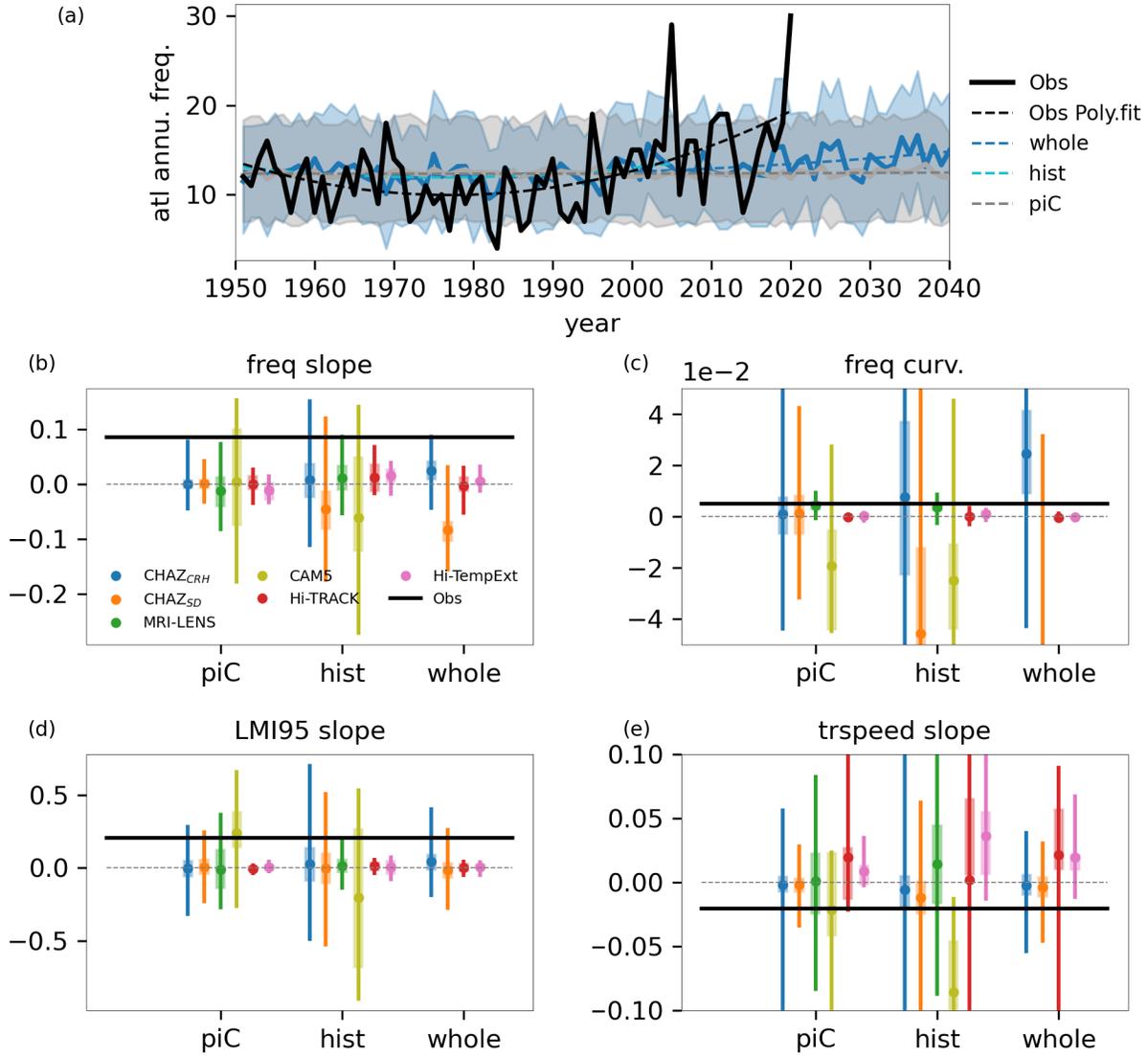
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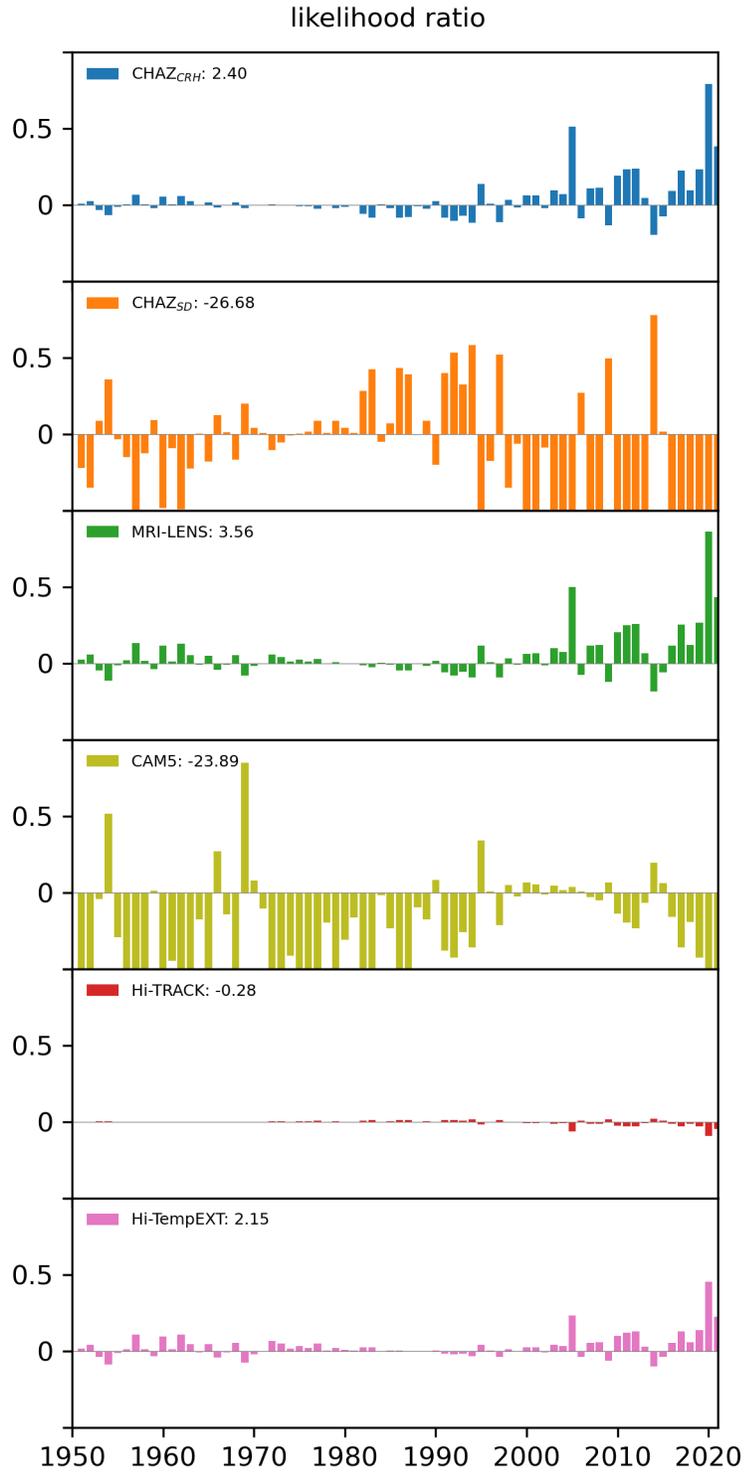
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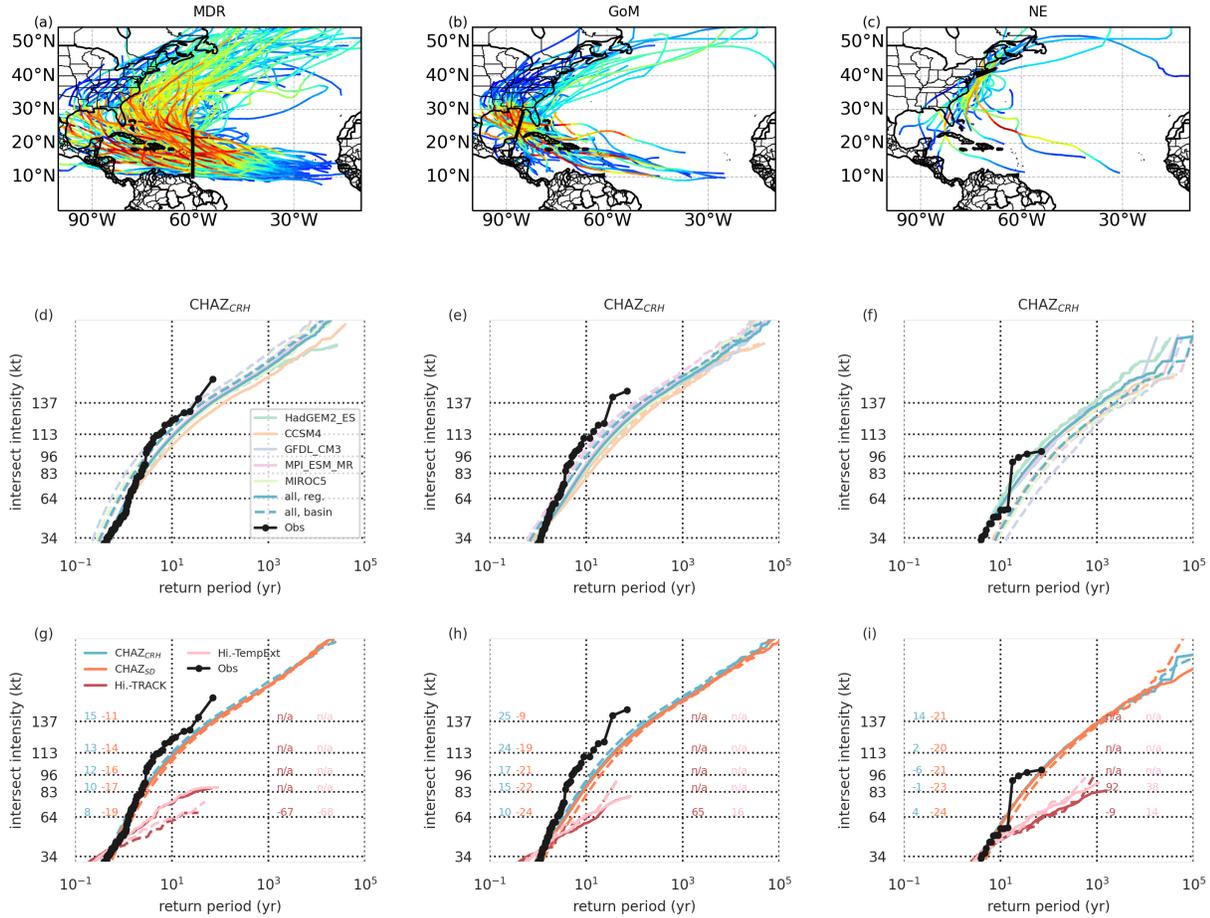
**Figure 1.** Annual frequency of Atlantic TCs exceeding 34 kt intensity threshold from 1951–2020 from best-track data (black), CMIP5 downscaling simulations using CHAZ<sub>CRH</sub> (blue) and CHAZ<sub>SD</sub> (pink), 25-km high-resolution CAM5 simulations (purple), 60km Japanese large-ensemble simulations (MRI-LENS, green), and HighResMIP simulations from (Roberts et al., 2020) and (Roberts et al., 2020). Storms from HighResMIP are tracked with TRACK (red) and TexmpExtreme (pink), respectively. In (a) and (c), simulations in their respective historical period are conducted with historical climate forcing while those in future period are with the rcp8.5 (for CHAZ) and ssp585 (for HighResMIP) warming scenarios. In (b) and (d), the simulations are under pre-industrial control climate (no anthropogenic forcing). (a) and (b) show the results from ensemble mean while (c) and (d) show the results from all ensemble members.



**Figure 2.** (a) Observed (black) and CHAZ<sub>CRH</sub> simulated mean annual hurricane frequency. The CHAZ simulations are from present (1951-2005) to future climate (2006-2040) periods (blue), and from those using pre-industrial control climate forcing (gray). Dashed lines show the polynomial fit. ‘hist’ shows the fit using synthetic storms from historical period only while ‘whole’ are from the historical and future periods. (b) Linear terms of the polynomial fit derived using synthetic storms’ annual frequency from all datasets. Datasets are indicated by color while the black line show the observed value. (c) Similar to (b) but for the quadratic terms. (d) and (e) are similar to (b) but for linear terms from the polynomial fit of LMI95 and storm forward motion speed. Units for (b), (c), (d) and (e) are, respectively, storm number year<sup>-1</sup>, storm number year<sup>-2</sup>, m s<sup>-1</sup> yr<sup>-1</sup>, and km hr<sup>-1</sup> yr<sup>-1</sup>.



**Figure 3.** Annual log-likelihood ratio in which  $\lambda_t$  is derived from historical (and future for the CHAZ and HighResMIP runs) simulations and the annual likelihood that is estimated based on piC simulations.



**Figure 4.** (a–c) Observed storm tracks from 1951–2020 at three line gates of interest. (d–f) Return period curves from 1951–2020 from best-track data (black lines), and CHAZ<sub>CRH</sub> historical simulations with basin-wide (dashed lines) and local (solid lines) basin corrections applied at the three gates. Global climate model forcings are indicated by colors and blue lines show the derived return period curves using all data. (g–i) Similar to (d–f) but for the four datasets. The solid lines show the return period curves using all historical simulations while dashed lines use all future simulations. Numbers at each Saffir-Simpson intensity threshold are the percentage changes of the frequency of the storms exceeding the threshold. Datasets are indicated by colors. Black curves show the empirical return curve using observations from 1951–2020.

1 **Climate change signal in Atlantic tropical cyclones**  
2 **today and near future**

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10 **Key Points:**

- 11 • Changes in the Atlantic hurricane risk are uncertain due to epistemic uncertainty  
12 in the projected annual frequency under global warming  
13 • Likelihood analysis shows that observations are more consistent with simulations  
14 with upward frequency projections than those without  
15 • Based on our results, it is more likely that the risk of hurricanes is increasing than  
16 that it is decreasing, though not by a large margin

**Abstract**

This manuscript discusses the challenges in detecting and attributing recently observed trends in the Atlantic hurricanes and the epistemic uncertainty we face in assessing future hurricane risk. Data used here include synthetic storms downscaled from five CMIP5 models by the Columbia HAZard model (CHAZ), and directly simulated storms from high-resolution climate models. We examine three aspects of recent hurricane activity: the upward trend and multi-decadal oscillation of the annual frequency, the increase in storm wind intensity, and the downward trend in the forward speed. Some datasets suggest that these trends and oscillation are forced while others suggest that they can be explained by natural variability. Future projections under warming climate scenarios also show a wide range of possibilities, especially for the annual frequencies, which increase or decrease depending on the choice of moisture variable used in the CHAZ model and on the choice of climate model. The uncertainties in the annual frequency lead to epistemic uncertainties in the future hurricane risk assessment. Here, we investigate the reduction of epistemic uncertainties on annual frequency through a statistical practice – likelihood analysis. We find that historical observations are more consistent with the simulations with increasing frequency but we are not able to rule out other possibilities. We argue that the most rational way to treat epistemic uncertainty is to consider all outcomes contained in the results. In the context of hurricane risk assessment, since the results contain possible outcomes in which hurricane risk is increasing, this view implies that the risk is increasing.

**Plain Language Summary**

We use a set of computer model simulations to study recent trends in Atlantic hurricanes. We looked at three aspects of these storms: the number of hurricanes each year, which has fluctuated up and down over time (but generally increased over the last several decades); the strength of their winds, which has been increasing; and the speed at which they move, which has been decreasing. These trends could be caused either by human-induced global warming or by natural variability; determining which cause is more important to overall hurricane risk requires us to understand how the number of hurricanes per year responds to warming. In our simulations, this number can either increase or decrease with warming, depending on which of two nearly identical versions of our model we use to simulate the storms. This uncertainty prevents us from reaching definitive conclusions about either present or future hurricane risk. Nonetheless, our analysis suggests that the risk of Atlantic hurricanes is more likely increasing than decreasing, and we argue that from a broader point of view, this is effectively equivalent to saying the risk is increasing.

**1 Introduction**

Rational measures to mitigate any risk must start from an assessment of that risk. Historical records can provide guidance, but in the case of atmospheric hazards such as hurricanes, we know that historical records are only a starting point for assessing current and future risk. This is both because the historical record is too short to fully sample the possibilities even in a stationary climate, and because the climate is changing (Schreck et al., 2014; Emanuel, 2021; D. Chan et al., 2022). Climate change makes the present different from the past, and requires us to consider whether the historical record alone, or catastrophe models that are built upon it, using purely statistical methods and assuming a stationary climate, are adequate, or need to be modified or supplemented to account for climate change.

Accounting for climate change is likely to require a greater use of physics than is historically typical in catastrophe models (Toumi & Restell, 2014; Emanuel, 2008). While one might instead try to assess the risk by using standard statistical methods but train-

67 ing only on the most recent observations (as opposed to the entire record), in the hope  
68 that those most recent observations represent the present and near-future climate ad-  
69 equately, this is likely to be challenging. Since hurricanes are rare, the number in the record  
70 over a period recent enough for this purpose is too small for risk assessment – especially  
71 when we also consider that low-frequency natural variability is present (i.e., Klotzbach  
72 & Gray, 2008; J. C. Chan, 2008; Wang et al., 2015), so that averaging times must be longer  
73 than might otherwise be necessary. To make the best possible assessment of present hur-  
74 ricane risk, then, we need to use our knowledge of the physics that connects hurricanes  
75 and climate (Emanuel, 2008).

76 The focus of this study is Atlantic tropical cyclones (TCs) risk in the present and  
77 near future. Future projections are useful for understanding how TCs may respond to  
78 climate changes of various sorts. Studies of historical observations, on the other hand, of-  
79 ten look for trends; but on their own, such studies do not establish the causes of the trends,  
80 nor whether they will persist. Establishing whether a trend is present (detection) is gen-  
81 erally viewed as a prerequisite to determining its cause (attribution) (Lloyd & Oreskes,  
82 2018). Detection can, in principle, be done with observations alone; attribution requires  
83 a model of some sort, in order to construct a counterfactual where the cause of interest  
84 is not present (Hegerl & Zwiers, 2011; Knutson, 2017). If a historical trend (or an os-  
85 cillatory signal) could be both detected and attributed to a specific cause, such as hu-  
86 man influence, or alternatively some specific natural processes, this would be of great  
87 scientific value, and would also allow us some insight into what to expect in the near fu-  
88 ture.

89 To develop such insight for Atlantic TCs, we will use recent observations and model  
90 simulations from historical (present), near future (up to 2040 or 2050), and pre-industrial  
91 control period. Simulations from pre-industrial control period contain no anthropogenic  
92 forcing signal and thus are used as a counterfactual. We use two types of model data.  
93 The first represents synthetic storms generated from a statistical-dynamical model, the  
94 Columbia (tropical cyclone) HAZard model (CHAZ), a model that encodes physical rela-  
95 tionships between tropical cyclones and their ambient large-scale environment (Lee et  
96 al., 2018). The second represents the directly simulated hurricanes from high-resolution  
97 global models, in which the above-mentioned relationships are simulated organically (Yoshida  
98 et al., 2017; Wehner et al., 2014; Roberts et al., 2020).

99 There are three objectives of this work. The first is to examine whether recently  
100 reported trends can be attributed to anthropogenic forcing. As summarized in Knutson  
101 et al. (2020a, 2020b), these trends are the recent variability of Atlantic annual TC fre-  
102 quency (Emanuel, 2007), an upward trend in the intensification rate (Bhatia et al., 2019)  
103 and lifetime maximum intensity (Kossin et al., 2013), and a slowing-down in the storm  
104 motion (Kossin, 2018). In particular, the cause of the recent increasing trend (since 1970)  
105 in Atlantic TC activity has been a subject of debate. On the one hand, reduced aerosols  
106 over the Atlantic since 1980s has been argued to be a dominant cause of the increasing  
107 TC activity in late 20<sup>th</sup> century (Mann & Emanuel, 2006; Sobel, Camargo, & Previdi,  
108 2019; Rousseau-Rizzi & Emanuel, 2020). On the other hand, several measures of Atlantic  
109 TC activity, including the major hurricane (TCs with LMI  $\geq$  93 kt) frequency (Goldenberg  
110 et al., 2001), are highly correlated to the the Atlantic Multi-decadal Oscillation (AMO)  
111 or Atlantic multidecadal variability (AMV), a low-frequency mode of variability iden-  
112 tified by the average sea surface temperature anomalies in the North Atlantic basin, typ-  
113 ically over 0-80°N (Ting et al., 2011). The recent AMO cycle, including both the upward  
114 trend from 1970 to 2005 and the downward trend from 2006 to 2018 have been associ-  
115 ated by some authors with natural variability (e.g., Yan et al., 2017, and others). How-  
116 ever, studies using CMIP5 historical runs simulated an ensemble-mean AMO that is sig-  
117 nificantly correlated with the observed AMO, suggesting that the recent historical vari-  
118 ability could be a consequence of radiative forcing (Clement et al., 2015; Bellomo et al.,  
119 2018). The future projections of TC frequency are subject to a similar degree of debate.

120 Many studies have suggested that the future should see a decline in the numbers of the  
 121 Atlantic TCs with warming (e.g., Knutson et al., 2010, and others), with a few excep-  
 122 tions (Emanuel, 2013; Bhatia et al., 2018; Vecchi et al., 2019).

123 The second objective is to compare historical simulations with observations to un-  
 124 derstand which modeling dataset is more consistent with the observations (Brunner et  
 125 al., 2020). Such analysis can provide guidance whether to favor one model over another,  
 126 which is especially useful for reducing uncertainty when the projections cover a wide range  
 127 even with an opposite sign, such as the projections of the divergent scenarios in the global  
 128 tropical cyclone genesis (i.e. Sobel et al., 2021). Lastly, we will assess hurricane risk over  
 129 a set of selected line gates in the present and future climates. Strictly speaking, risk in-  
 130 cludes severity of the hazard, exposure, and vulnerability of the properties of interest.  
 131 Only the hazard component is examined here.

## 132 2 Data, Experimental design and Method

### 133 2.1 Tropical cyclone datasets

#### 134 2.1.1 Observations

135 For reference, we use best-track data from National Hurricane Center obtained via  
 136 International Best Track Archive for Climate Stewardship v04r00 IBTrACS (Knapp et  
 137 al., 2010). We use 6-hourly storm positions (in longitude and latitude) and maximum  
 138 wind speeds (kt) from 1951 to 2020. Storm forward speed is derived from the position  
 139 data. We use only storms whose lifetime maximum intensity (LMI) reaches tropical storm  
 140 (TS) strength, 34 kt. Hurricanes are referred to storms with LMI of at least 64 kt.

#### 141 2.1.2 Synthetic events from the CHAZ model

142 The first set of model TCs used here consists of synthetic storm tracks from the  
 143 Columbia (tropical cyclone) Hazard (CHAZ) model (Lee et al., 2018). CHAZ is a statistical-  
 144 dynamical downscaling model that generates synthetic storms whose properties depend  
 145 on environmental conditions. The environmental conditions can come from an observation-  
 146 based reanalysis or a global climate model. There is no feedback of downscaled TC ac-  
 147 tivity to the global models. Three components in CHAZ describe storm formation and  
 148 subsequent evolution until shortly after landfall (or dissipation): the cyclone genesis in-  
 149 dex (TCGI; Tippett et al., 2011), the beta-advection track model (Emanuel, 2008), and  
 150 an auto-regressive intensity model (Lee et al., 2015, 2016). Details about CHAZ are re-  
 151 ported in Lee et al. (2018). The environmental variables required by the model are Po-  
 152 tential Intensity (Bister & Emanuel, 1997), deep-layer (850 to 250 hPa) vertical wind shear,  
 153 and one or more moisture variables: column integral relative humidity (CRH) and/or  
 154 column integral saturation deficit (SD), the absolute vorticity at 850 hPa, and the steer-  
 155 ing flow. The choice of moisture variables will prove particularly important in what fol-  
 156 lows. Both variables are calculated following Bretherton et al. (2004). The simulated trop-  
 157 ical cyclone activity in CHAZ, at global and basin scales, in both current and projected  
 158 future climates have been discussed in detail in Lee et al. (2018) and Lee et al. (2020),  
 159 respectively. The CHAZ model has been used for case studies in Texas (Hassanzadeh  
 160 et al., 2020), New York (Lee et al., 2022), Mumbai, India (Sobel, Lee, et al., 2019) and  
 161 the Philippines (Baldwin et al. 2022). Meiler et al. (2022) found that losses estimated  
 162 from CHAZ are comparable to those estimated using comparable academic tropical cy-  
 163 clone hazard models from Emanuel (2013) and Bloemendaal et al. (2020).

164 In this study, we use CHAZ to downscale five CMIP5 models (Taylor et al., 2012)  
 165 over the Atlantic basin. They are the National Center for Atmospheric Research (NCAR)  
 166 Community Climate System Model 4 (CCSM4) (Gent et al., 2011), the Geophysical Fluid  
 167 Dynamics Laboratory Climate Model version 3 (GFDL-CM3) (Donner et al., 2011), the

168 United Kingdom Meteorological Office Hadley Center Global Environment Model ver-  
 169 sion 2 Earth System (HadGEM2-ES) (Jones et al., 2011), the Max Planck Institute Earth  
 170 System Model Medium Resolution (MPI-ESM-MR) (Zanchettin et al., 2012), and the  
 171 Model for Interdisciplinary Research Climate Version 5 (MIROC5) (Watanabe et al., 2010)  
 172 from the University of Tokyo Center for Climate System Research, National Institute  
 173 for Environmental Studies, Japan, Japan Agency for Marine-Earth Science.

174 CHAZ’s projections of annual TC frequency, both in the Atlantic and globally, are  
 175 sensitive to whether CRH and SD are used in TCGI. Using TCGI with CRH leads to  
 176 a projected increase in global (and Atlantic) TC frequency, while SD leads to a projected  
 177 decrease (Lee et al., 2020). CRH and SD both measure the degree of the saturation of  
 178 the atmosphere with SD being the difference between the column integrated water vapor  
 179 and the same quantity at saturation, and CRH being their ratio. As saturated water  
 180 vapor increases with temperature in a warming climate, CRH remains close to con-  
 181 stant and SD decreases (Camargo et al., 2014). In the current climate, however, the be-  
 182 havior of these two variables are qualitatively similar, and the two TCGI formulations  
 183 yield similar results for the historical period, meaning that the historical evidence is in-  
 184 adequate to determine which of the two is more correct. Arguably, SD better reflects the  
 185 increase in the thermodynamic inhibition of TC formation in a warming climate (Emanuel,  
 186 1989, 2022), but the gaps in our understanding of the relationship between climate and  
 187 tropical cyclone frequency are so substantial that we do not view this argument as dis-  
 188 positive (Sobel et al., 2021). The diverging annual frequency projections from CHAZ thus,  
 189 in our view, reflects the broader state of the science, in that we have low confidence re-  
 190 garding whether one should expect more or fewer hurricanes as climate warms(i.e. Ca-  
 191 margo et al., 2020; Vecchi et al., 2019; Sugi et al., 2020). One reason for the low con-  
 192 fidence in TC frequency projection is the lack of theoretical understanding of tropical  
 193 cyclone genesis, and we refer the readers to a review article by Sobel et al. (2021) for a  
 194 detailed discussion.

195 Since total TC hazard and risk depend inextricably on TC frequency and we lack  
 196 a strong basis for choosing between SD and CRH, the sensitivity to the humidity vari-  
 197 able in our results causes a deep uncertainty in the projected risk. This uncertainty will  
 198 remain in the present study, in that we performed separate sets of simulations with ei-  
 199 ther CRH or SD as the humidity variable in the genesis module, referred to as CHAZ<sub>CRH</sub>  
 200 and CHAZ<sub>SD</sub>.

### 201 *2.1.3 Directly simulated hurricanes from General Circulation Models*

202 In addition to the CHAZ downscaling simulations described above, we use storms  
 203 tracked in a set of relatively high-resolution, i.e., tropical cyclone-permitting, global cli-  
 204 mate models. The first one is the 60-km MRI-AGCM3.2H large-ensemble simulation from  
 205 Mizuta et al. (2017) (MRI-LENS). Tropical cyclones in that model was discussed in Yoshida  
 206 et al. (2017). The second one is the 25-km High-Resolution Community Atmospheric Model  
 207 version 5 simulations, CAM5 (Wehner et al., 2014, 2015). Next, we use storms tracked  
 208 in the Coupled Model Intercomparison Project Phase 6 (CMIP6) (Eyring et al., 2016)  
 209 High Resolution Model Intercomparison Project (HighResMIP) (Haarsma et al., 2016).  
 210 Following Roberts et al. (2020) and Roberts et al. (2020), we use storms from CMCC-  
 211 CM2 (Cherchi et al., 2019), CNRM-CM6 (Volodire et al., 2019), EC-Earth3P-HR (Haarsma  
 212 et al., 2020), HadGEM3-GC3.1 (Roberts et al., 2019), and MPI-ESM1.2 (Gutjahr et al.,  
 213 2019). There are two HighResMIP configurations, one is forced with prescribed SST while  
 214 the other is fully coupled. We only use the simulations from the fully coupled configu-  
 215 ration which allows natural variability to occur freely during the historical period. To  
 216 understand the sensitivity of model performance to the TC trackers, HighResMIP storms  
 217 are tracked by TRACK (Hodges et al., 2017) and TempestExtremes (Ullrich & Zarzy-  
 218 cki, 2017; Zarzycki & Ullrich, 2017; Ullrich et al., 2021), and both event sets are used

219 here. For convenience, we label modeled TCs from HighResMIP tracked with Tempest-  
 220 Extremes as Hi-TempExt and those tracked with TRACK as Hi-TRACK.

## 221 2.2 Experimental design

222 Except in MRI-LENS and CAM5, we use model TCs from the historical, near-term  
 223 future, and pre-industrial control (piC, no anthropogenic forcing) scenario simulations.  
 224 Note that the time range covered in each period varies by model. For the historical pe-  
 225 riod, they are 1951-2005 for CHAZ<sub>CRH</sub> and CHAZ<sub>SD</sub>, 1950-2010 for MRI-LENS, 1996-  
 226 2016 for CAM5, and 1951-2014 for the two HighResMIP datasets. In the future period,  
 227 CHAZ<sub>CRH</sub> and CHAZ<sub>SD</sub> contain storms from 2006-2040 under Representative Concentra-  
 228 tion Pathway 8.5 (rcp8.5) while HighResMIP storms are from 2015-2050 under Shared  
 229 Socioeconomic Pathways5-85 (ssp585). Both are high-emission scenarios with an addi-  
 230 tional radiative forcing of  $8.5 \text{ W m}^{-2}$  by the year 2100 (Riahi et al., 2017) in ssp585 which  
 231 considers a fossil-fueled development. Warming climate simulations for MRI-LENS and  
 232 CAM5 are under a  $4^\circ\text{C}$  (Yoshida et al., 2017) and  $1.5^\circ\text{C}$  warming (Wehner et al., 2018)  
 233 scenarios and thus are not used here. In piC, the labeling of year is arbitrarily in all datasets  
 234 as all years are equivalent. The MIR-LENS and CAM5 piC simulations are exceptions.  
 235 In MRI-LENS and CAM5, the observed SST information is given in both historical and  
 236 piC simulations as a lower boundary, but the long-term trend is removed in the piC sim-  
 237 ulations. In other words, MIR-LENS and CAM5 piC simulations still contain observed  
 238 variation. The piC simulations in MRI-LENS, called “no-warming” in Mizuta et al. (2017)  
 239 and those in CAM5, following “Nat-Hist” in Stone et al. (2019), are designed with an  
 240 underlying assumption that that only the linear trend is anthropogenic forced, not the  
 241 variability, which, as we will discussed in the next Section, is debatable.

242 In each period, the CHAZ model was used to generate 20 track ensemble members  
 243 per CMIP5 model and each track has 40 intensity ensembles (100 CMIP5 track ensemble  
 244 members and 4000 considering intensity ensemble), as is possible because the CHAZ  
 245 intensity module has a stochastic component. Hi-TRACK has 7 members (5 global cli-  
 246 mate models and two of them have 2 ensemble members) and Hi-TempExt has 6 (4 global  
 247 climate models and two of them have 2 ensemble members). MRI-LENS has 100 ensemble  
 248 members while CAM5 has 5. The data properties are listed in Table 1.

## 249 2.3 Frequency adjustment

There are biases in model TCs, because of biases in the models that generate them,  
 including the CHAZ model itself as well as the CMIP5 models from which CHAZ ob-  
 tains its environmental conditions, and the high-resolution global climate models used  
 here. In particular, all models have biases in TC frequency (Table 1), and directly-simulated  
 hurricanes from high-resolution global climate models have low-intensity biases, in gen-  
 eral, as the grid spacings of these models are too coarse to capture the full range of ob-  
 served hurricane strengths (e.g., Yoshida et al., 2017; Moon et al., 2022, and others). Here  
 we address only the frequency biases. Specifically, we derive an adjustment by compar-  
 ing the basin-wide annual TC frequency of models’ historical simulations to that of the  
 observations from the same period. The same adjustment will then be applied to both  
 historical and future simulations. Similarly, we compare the annual frequency of the piC  
 simulations to the observations to adjust piC’s annual frequency. In Lee et al. (2018) and  
 Lee et al. (2020), the basin-wide frequency adjustment is a multiplicative factor to en-  
 sure that the mean annual frequency over a basin in CHAZ is consistent to that in ob-  
 servations. However, some high-resolution global climate models used here, such MRI-  
 LENS, generate zero TCs in some years. A multiplicative factor would result in larger  
 variability but still have zeros in these years, which is unrealistic. Thus, here the basin-

wide frequency is adjusted as:

$$f_{\text{adj}} = \sigma_{\text{obs}} \times \frac{f_{\text{ori}} - \mu_{\text{model|ref}}}{\sigma_{\text{model|ref}}} + \mu_{\text{obs}}, \quad (1)$$

where  $f$  indicates annual frequency (each year) with the subscript indicating after (*adj*) or before (*ori*) frequency adjustment.  $\mu$  and  $\sigma$  are the mean and standard deviation of the frequency and the subscript indicates whether it is from simulations (*model*) or observations (*obs*). As we want to retain the climate change signal, reference  $\mu$  and  $\sigma$  ( $\mu_{\text{model|ref}}$  and  $\sigma_{\text{model|ref}}$ ) for adjusting frequencies in both historical and future simulations in each dataset are from its respective historical simulation. Observations are calculated from their respective historical periods. To adjust the annual frequencies of the piC simulations,  $\mu_{\text{model|ref}}$  and  $\sigma_{\text{model|ref}}$  are from piC. Biases in annual TC frequency of the piC simulations are different to those in the historical simulations. As we will discuss later, a basin-wide frequency adjustment may not correct regional biases, because model biases can have spatial dependence. When desired (in Section 5), we apply a multiplicative factor to ensure the annual frequency at storm with intensity greater than 40 kt in these data sets are consistent to observations, which is the same as the bias-correction approach used in (Lee et al., 2022).

An underlying assumption of our approach to bias correction, in common with many climate change studies, is that the bias of any given model remains the same in projected future climate periods as it is in the present, so that the influence of the projected climate change can still be captured when comparing simulations between rcp and hist periods. This assumption is analogous to that used to remove climatological biases in surface temperature and other quantities from the climate models themselves in global warming projections, for example those by the Intergovernmental Panel on Climate Change (Solomon et al., 2007). While this assumption of constant biases can be questioned, it is a simple assumption, and there is no empirical basis on which to base any more complex assumption one. Still, we will discuss the impacts of frequency adjustments on our findings.

## 2.4 Trend analysis

To calculate trends of TC activity, we fit second-order Legendre polynomials:

$$\hat{y} = a_0 + a_1x + \frac{a_2}{2}(3x^2 - 1), \quad x \in [-1, 1] \quad (2)$$

to the time series of the variables of interest from observations and model simulations. In Equation (2),  $x$  is years scaled to interval of  $[-1, 1]$ ,  $\hat{y}$  represents the fitted variables, the coefficient  $a_1$  shows linear trends and  $a_2$  shows quadratic trends. Considering quadratic trends allows the possibility that the observed multi-decadal variability is in fact forced (Clement et al., 2015; Bellomo et al., 2018). Here, we ask whether or not the observed trends lie within the ensemble spread from simulations. For example, if the observed trend is outside of the range of piC simulations but is within those from historical simulations, then the observed change (e.g., upward trend or increasing curvature) is unlikely to have occurred without anthropogenic forcing. When comparing the trends between observations and simulations,  $a_1$  and  $a_2$  are scaled back so that they have units of the variable's unit per year ( $yr^{-1}$ ) and per year square ( $yr^{-2}$ ), respectively.

## 3 Trend and multi-decadal variability

### 3.1 Atlantic TC frequency

We first examine the Atlantic TC frequency trends in the historical (present) climate and from historical to the warming future (i.e., using simulations from both historical and future periods). Figure 1a and b show the ensemble means of the time series of Atlantic hurricane frequency, i.e., the averaged total number of storms in the basin

293 each year whose maximum sustained winds exceed 34 kt from each dataset. The small  
 294 wiggles may be sampling variability. Figures 1c and d show the ensemble spread. By con-  
 295 struction, the time-mean annual frequency for each dataset over its respective histori-  
 296 cal period will be identical to observations after the frequency adjustment (Eq. (1)). The  
 297 original annual frequency of each dataset is shown in Table 1. Before 2000, the differ-  
 298 ent simulations are, by eye at least, indistinguishable in their overall envelopes, with none  
 299 showing any particular trend, and the observations (black thick line) lying well within  
 300 their spread (shown in Figure 1c). After 2000, the CHAZ<sub>SD</sub> (orange thick line) and CHAZ<sub>CRH</sub>  
 301 (blue thick line) results begin to diverge, with CHAZ<sub>SD</sub> showing a decreasing TC fre-  
 302 quency and CHAZ<sub>CRH</sub> showing an increasing TC frequency. It is possible that this is  
 303 related to the fact that the rcp8.5 scenario starts after 2005. The two HiResMIP datasets  
 304 show no considerable trend in the historical period but a sharp dip after 2030. The ssp585  
 305 scenario in HiResMIP starts after 2015, though. Hi-TRACK's annual TC frequency climbs  
 306 up by 2040. Roberts et al. (2020) reported that both Hi-TRACK and Hi-TempExt project  
 307 a reduction of ensemble mean annual frequency (less than 10%) from 1950-1980 to 2020-  
 308 2050, but the spread covers zero, indicating low confidence to the mean trend.

309 Figures 1b and 1d show analogous results for piC simulations. Note that the years  
 310 in the x-axis are not real; these labels are placed so we can compare the simulated trends  
 311 to the observed trend and those in Figures 1a and 1c. Two exceptions are MRI-LENS  
 312 and CAM5 simulations; both are uncoupled atmospheric models and forced with observed  
 313 SST with anthropogenic trend removed (See Section 2 for details). In the Figure 1b, CHAZ<sub>CRH</sub>  
 314 and CHAZ<sub>SD</sub> results do not diverge. There is no dip in the Hi-TRACK or Hi-TempExt.  
 315 Clearly, the separation between CHAZ<sub>CRH</sub> and CHAZ<sub>SD</sub> and the dip in the two High-  
 316 ResMIP datasets in Figure 1a represent forced responses.

317 Next we conduct the trend analyses of the annual TC frequency in Figure 1 using  
 318 second-order Legendre polynomials fits (Eq. (2)). As an example, Fig. 2a shows the anal-  
 319 ysis using the CHAZ<sub>CRH</sub> simulations and the observations. The observed fit (dashed black  
 320 line) has an upward trend of 0.085 storm year<sup>-2</sup> and a positive curvature of 0.005 storm year<sup>-2</sup>  
 321 (shown as the black line in Figs. 2b and 2c). The existence of a linear trend means that  
 322 there is an overall increasing trend in storm activity since 1951 while the quadratic terms  
 323 captures the multi-decadal variability, with high activity in the 1950s-1960s, low in the  
 324 1970s-80s, and high after that, which recent research suggests may be a forced signal rather  
 325 than natural variability (Clement et al., 2015; Bellomo et al., 2018). In Fig. 2a, the poly-  
 326 nomial fits of CHAZ<sub>CRH</sub> simulations from historical only (light blue dashed line) and from  
 327 historical to future (dark blue dashed line) both show an small upward curve while the  
 328 polynomial fit derived from the piC simulations (gray dashed line) is quite flat.

329 The ranges of the fit parameters from all ensemble members in each dataset are  
 330 also shown in Figures 2b-c. The observed linear trend are above most of the piC sim-  
 331 ulations except those from CAM5. However, CAM5 has only 10-years of simulations, which  
 332 is too short to be compared with 70-years of observations. The observed quadratic term  
 333 lies within the 25-75 percentile ensemble ranges of piC simulations from CHAZ<sub>CRH</sub>, CHAZ<sub>SD</sub>,  
 334 and MRI-LENS. It is outside of the ensemble ranges from two HighResMIP datasets which  
 335 have quadratic terms close to zero. The observed linear trend is at top 25 percentile (75-  
 336 100 percentile) of the hist simulations of CHAZ<sub>CRH</sub>, CHAZ<sub>SD</sub>, and is marginally included  
 337 in the simulations of MRI-LENS; the observed quadratic term is within the 25-75 per-  
 338 centile range the CHAZ<sub>CRH</sub> and MRI-LENS, and is at top 25 percentile in CHAZ<sub>SD</sub>. Only  
 339 the fit linear trend derived from historical + future simulations of the CHAZ<sub>CRH</sub> include  
 340 the observed value. For the quadratic trend, the fit parameter derived from CHAZ<sub>CRH</sub>  
 341 and CHAZ<sub>SD</sub> include the observed values but they are at top and bottom 25 percentile  
 342 range, respectively. (We do not use any warming simulations from CAM5 and MRI-LENS.)

343 Generally speaking, the polynomial fit analysis suggests that, first, CHAZ<sub>CRH</sub>, CHAZ<sub>SD</sub>  
 344 and MRI-LENS are better in capturing the observed trend and multi-decadal variabil-  
 345 ity as their historical spread covers the observed values. However, CAM5 has only 10 years

346 of data with 5 ensemble members and while Hi-TRACK and Hi-TempExt have only, re-  
 347 spectively, 7 and 6 ensemble members. These three datasets may be under-sampled. Sec-  
 348 ond, the observed linear trend is outside the spread of CHAZ<sub>CRH</sub>, CHAZ<sub>SD</sub> and MRI-  
 349 LENS' piC simulations but within the spread of these models' hist simulations, indicat-  
 350 ing that anthropogenic forcing is necessary to capture the upward trend in the past decades.  
 351 On the other hand, we can not rule out the possibility of the recent upward curvature  
 352 trend is within the range of natural variability. Although the MRI-LENS' piC simula-  
 353 tions is forced with the observed SST (with long-term trend removed) which results in  
 354 the upward curvature term right on top of observed values in Figure 2c. Simulations from  
 355 CHAZ<sub>CRH</sub> suggest that that anthropogenic forcing helps to capture the upward curva-  
 356 ture trend. Third, when considering the future period as well, the mean of CHAZ<sub>CRH</sub>  
 357 shows an upward trend, the mean of CHAZ<sub>SD</sub> shows a downward trend, while the mean  
 358 of the two HighResMIP simulations are close to zero. However, we have low confidence  
 359 in the projections as they include zero. Thus, we can not say for sure that the positive  
 360 linear and quadratic terms will continue into the future.

361 It should be noted that without the basin-wide frequency adjustment (not shown),  
 362 the observed linear and quadratic terms lie outside of the spread of MRI-LENS, Hi-TempExt  
 363 and Hi-TRACK in all three periods. They are within the spread of CHAZ<sub>CRH</sub> and CHAZ<sub>SD</sub>  
 364 simulations in piC and historical periods. With additional data from 2006 to 2040, only  
 365 CHAZ<sub>CRH</sub> shows such an upward trend will continue into the future.

### 366 3.2 Intensity and storm motion

367 Figure 2d shows the fit parameters of Atlantic TC lifetime maximum intensity (LMI).  
 368 Specifically, we look at the variability of the 95th percentile of LMI (LMI95), for which  
 369 an upward trend has been found in observations (Kossin et al., 2013). Here we focus on  
 370 the linear term only. There is an upward trend in the observations, meaning that the ex-  
 371 treme tail of observed intensity has increased with time, consistent with previous stud-  
 372 ies (e.g., Knutson et al., 2020a, and others). The positive linear trend is captured by the  
 373 ensemble spreads of two CHAZ datasets and those of MRI-LENS and CAM5 at both piC  
 374 and historical periods. It is outside of the ensemble spread of all simulations from from  
 375 Hi-TRACK and Hi-TempExt. Thus, at least from CHAZ<sub>CRH</sub>, CHAZ<sub>SD</sub>, MRI-LENS, and  
 376 CAM5, we can not rule out that the recent upward trend in the LMI95 is due to nat-  
 377 ural variability. When looking into the future, only the means of CHAZ<sub>CRH</sub> is positive  
 378 and the means of CHAZ<sub>SD</sub>, Hi-TempExt and Hi-TRACK are close to zero. Similar to  
 379 the results from TC frequency, the ensemble spread in Figure 2d include zero in the whole  
 380 historical + future periods, indicating, again, low-confidence in the projected changes.

381 Figure 2e shows the analysis for translation speed. Consistent with (Kossin, 2018),  
 382 the observations show a clear downward trend in the storm motion. This trend is within  
 383 ensemble spread in all periods, including piC, for all models, except the simulations from  
 384 Hi-TempExt. However, the mean and the 25-75 percentile ensemble spreads in these datasets  
 385 move toward different directions from piC to hist to hist +future periods. The Hi-Track  
 386 and MRL-LENS hist simulations show upward trends in the storm motion and this up-  
 387 ward trends continues in to the future. The differences in mean and 25-75 percentile en-  
 388 semble spreads from CHAZ<sub>CRH</sub> and CHAZ<sub>SD</sub> from these three period are small. The piC  
 389 and hist simulations from CAM5 shows that anthropogenic forcing may lead to a strong  
 390 downward trend in storm motion but again CAM5 simulations are shorter than do the  
 391 data from the other models. It seems unjustified, based on this set of models, to attribute  
 392 the observed slowing down to anthropogenic forcing. It also noteworthy that at a regional  
 393 scale, CHAZ projected an upward trend in storm motion speed for TCs affecting Texas  
 394 (Hassanzadeh et al., 2020) and an a downward trend for storms impacting New York (Lee  
 395 et al., 2022). Spatially inhomogeneous changes may dilute the basin-wide signal.

#### 4 Likelihood comparison

Figure 2 shows that the simulated trend in historical and historical + future vary from one dataset to another. This is especially true for the TC frequency projections between CHAZ<sub>CRH</sub> and CHAZ<sub>SD</sub>, but a qualitatively similar result, including both increasing and decreasing trends, holds for the rest of our ensemble of opportunity. It is natural to ask whether we can develop some criteria for determining which is correct. In climate science, multi-model ensemble mean is a common approach to obtain the consensus from multiple global climate models. However, such approach is only adequate when the ensemble spread represents variations that can be considered random, as might be the case with typical aleatoric uncertainties. The divergent scenarios in the frequency projections are a consequence of the epistemic uncertainty due to the lack of a satisfactory scientific understanding of tropical cyclone frequency (Sobel et al., 2021; Emanuel, 2022) and thus the multi-model mean may not be meaningful in this case. We can, however, use likelihood analysis, in which the probabilities that the observations occurs in the model simulated distribution were computed. Thus, we can determine which simulation the observation is more consistent with. This is similar to the Likelihood Skill Score used for evaluating weather and climate predictions (Barnston et al., 2010).

Specifically, we first assume that annual hurricane frequency is drawn from a Poisson distribution whose mean ( $\lambda_t$ ) has a trend in time ( $\lambda_t = at + b$ ). We then obtain  $a$  and  $b$  of each dataset by fitting the model annual TC frequency to a Poisson regression. We do so for all simulations with data throughout 2021 (up to 2005 for CAM5 and 2010 for MRI-LENS). Note that with  $a$  and  $b$ , we can derive  $\lambda_t$  even for years beyond the data coverage period, i.e., we can estimate  $f_{2020}$  with  $a$  and  $b$  derived from CAM5 data. The yearly likelihoods ( $L_t$ ) of the observed frequencies are assigned based on the Poisson distribution with a given  $\lambda_t$ . For example, the likelihood CHAZ<sub>CRH</sub> simulations will generate 29 TCs as observed in 2005 is 0.08%, which is based on a Poisson distribution with  $\lambda_{2005} = 15.7$ . The same calculation is applied to piC simulations, and the derived likelihood is denoted  $L_{piC,t}$ . For a given year, we then compare the log likelihood ratios  $L_t$  and  $L_{piC,t}$  (i.e.,  $\log(L_t/L_{piC,t}) = \log(L_t) - \log(L_{piC,t})$ ). If this ratio is larger than 0, the observations are more consistent with the simulations with anthropogenic forcing than with the piC simulations and vice versa.

We start by comparing the likelihoods of simulations with anthropogenic forcing to those with piC simulations (i.e.,  $\log(L_t/L_{piC,t})$  in Figure 3. The ratios of the likelihoods jointly up to 2020 (numbers on the upper-left in all panels) suggest that the observations are more consistent with the simulations with anthropogenic forcing than those without in CHAZ<sub>CRH</sub>, MRI-LENS, and Hi-TempExt. The annual likelihood ratios from these three datasets further show higher annual likelihood ( $L_t$ ) for the observed annual frequency values during the period of high TC activity in 1950-1970 and after 2000 while higher  $L_{piC,t}$  is found during 1970-2000. This is because there are upward trends in the simulated annual frequency in these three datasets when compared to in piC (Figure 2a). As  $\lambda_t$  increases with time, the distributions from these three datasets shift right with time and thus give greater likelihood to the high observed annual frequency when compared to those derived from piC simulations in which  $\lambda_t$  is close to constant in time. In contrast, CHAZ<sub>SD</sub> has a downward trend and its,  $\lambda_t$  shifts left in time and leads to lower likelihood when observed values are high. Consequently, we see a higher  $L_{piC,t}$  during high TC activity periods and higher  $L_t$  during the inactive TC seasons in CHAZ<sub>SD</sub>. The frequency slopes obtained from piC and hist in the Hi-TRACK data are similar and thus their log likelihood ratio is close to zero.

When we consider the likelihood over the whole observational period, we calculate the average of the likelihood, i.e., the roots of  $\prod_{2021}^{1950} L_t$  from all five datasets. Between CHAZ<sub>CRH</sub> and CHAZ<sub>SD</sub>, observations are more consistent with CHAZ<sub>CRH</sub>, which has an averaged likelihood of 5.24%, than to CHAZ<sub>SD</sub> which has the averaged likelihood of

448 3.46%. Among the five datasets, CHAZ<sub>CRH</sub> has highest likelihood, followed by Hi-TempExt  
 449 (5.13%), MRI-LENS (5.1%), Hi-TRACK (5.04%), and CAM5 (3.6%).

450 The basin-wide frequency adjustment (Eq. (1)) that we performed to correct model  
 451 biases is expected to affect the results of the likelihood analysis, because the frequency  
 452 adjustment both shifts the mean of the model's TC annual frequency distributions and  
 453 changes their shapes. The annual frequency distributions from historical and piC sim-  
 454 ulations are more distinct in the datasets without frequency adjustment, which indeed  
 455 leads to larger log likelihood ratios (not shown). Without the frequency adjustment, the  
 456 observed TC annual frequencies are more consistent with the historical simulations in  
 457 CHAZ<sub>CRH</sub>, CHAZ<sub>SD</sub>, and MRI-LENS than in their respective piC simulations due to the  
 458 large bias in these piC simulations. Without basin-wide TC frequency adjustment, Hi-  
 459 TRACK has the greatest averaged likelihood, followed by CAM5, CHAZ<sub>CRH</sub>, Hi-TempExt,  
 460 CHAZ<sub>SD</sub>, and MRI-LENS. MRI-LENS has the lowest likelihood because of its low bias  
 461 and zero storms in some years.

## 462 5 Climate change and regional hurricane risk at three line gates

463 Now we compute regional hurricane risk, from hazard perspective only, represented  
 464 by return periods of storms of given wind intensities passing through pre-defined gates,  
 465 derived using simulations from historical and future periods. We use simulations from  
 466 CHAZ<sub>CRH</sub>, CHAZ<sub>SD</sub>, Hi-TRACK, and Hi-TempExt. The three line gates used here (black  
 467 lines in Figures 4a-c) are the main development region (MDR) gate which can be thought  
 468 of as delineating Atlantic TC hazard in a general sense – how many storms form, and  
 469 at what intensity and move from the MDR toward the US and Caribbean Islands; the  
 470 GoM gate which records TC activity for those that enter the Gulf of Mexico; and the  
 471 NE gate which is parallel to a portion of the Northeastern US coast. As discussed ear-  
 472 lier (Section 2.3), to obtain more realistic return period curves for regional hurricane risk  
 473 assessment, we use a more localized frequency adjustment. As an example, Figures 4d-  
 474 4f show historical simulations from CHAZ<sub>CRH</sub> with basin-wide and regional frequency  
 475 adjustments (Eq. (1)). While the basin-wide frequency adjustment (dashed lines) yields  
 476 a TC frequency close to observations at the GoM gate, CHAZ<sub>CRH</sub> still overestimates storm  
 477 activity at the MDR gate and underestimates storm activity at the NE gate. The regional  
 478 frequency adjustment shifts the simulated return period curves (solid line, local adjust-  
 479 ment) by matching the return periods at 40 kt to the values derived from observations  
 480 (see Section 2.3 for details). In terms of the shape of the return period curve, as well as  
 481 the return periods at high intensities, CHAZ<sub>CRH</sub> performs better at the MDR gate than  
 482 at the GoM gate. It is difficult to directly compare the modeled curves to the observa-  
 483 tions at the NE gate, due to the significant underestimation of overall TC frequency at  
 484 the latter. However, even there, the shapes of the observed and modeled return period  
 485 curves are similar.

486 To show the changes in return periods between historical and future periods, Fig-  
 487 ures 4g-i show the return period curves derived from the four datasets that have rcp8.5/spp585  
 488 warming scenarios available. We use model storms from all ensemble members. Low-intensity  
 489 biases in the Hi-TRACK and Hi-TempExt lead to an underestimate of the TC risk. High-  
 490 ResMIP models barely simulate storms with major hurricane wind strength (Roberts et  
 491 al., 2020; Moon et al., 2022). The return period curves of CHAZ<sub>CRH</sub> and CHAZ<sub>SD</sub> hist  
 492 simulations are close to each other. The differences between simulations from historical  
 493 period and those from historical and future periods, i.e., the differences between the dashed  
 494 and solid lines, are small for the two CHAZ datasets in Figures 4g-i. Likewise the his-  
 495 torical and future period curves of GoM and NE gates for Hi-TRACK and Hi-TempExt  
 496 nearly indistinguishable. At the MDR gate, both Hi-TRACK and Hi-TempExt suggest  
 497 increases in the TC risk.

498 To make these differences more evident, we list the percentage changes in annual  
 499 TC frequency exceeding each Saffir-Simpson category on both sides of each panel in Fig-  
 500 ures 4g-i. As expected, there is an overall increase in the storm frequency at all thresh-  
 501 olds from historical to future periods for CHAZ<sub>CRH</sub> while there is an overall decrease for  
 502 CHAZ<sub>SD</sub>, consistent with the results in Figures 1 and 2a. The percentage changes are  
 503 larger at higher intensity thresholds in the CHAZ<sub>CRH</sub> but they are of similar or smaller  
 504 magnitude throughout the Saffir-Simpson categories in the CHAZ<sub>SD</sub>. This is probably  
 505 due to the increase in storm intensity as climate warms in CHAZ<sub>CRH</sub> and CHAZ<sub>SD</sub>.

506 The changes in the frequency of exceedance at the three gates from Hi-TRACK and  
 507 Hi-TempExt are not the same sign. Hi-TRACK shows a 67% decrease of Category 1+  
 508 ( $\leq 64$  kt) at the MDR gate but a 65 % increase at GoM gate. At the NE gate, Hi-TRACK  
 509 shows an 14 and 38% increase in the frequency of Category 1+ and 2+ storms, respec-  
 510 tively. Hi-TempExt shows a 68% decrease and 16% increase of Category 1+ storms at  
 511 the MDR and GoM gates, respectively. At the NE gate, it shows a 9% decrease and 92%  
 512 increase in the frequency of Category 1+ and 2+ storms. Storms from these two High-  
 513 ResMIP runs are undersampled and have low intensity biases (See Figure 7 in Roberts  
 514 et al. (2019)). The directly simulated storms are not suitable for risk assessment and these  
 515 numbers should be used with caution.

## 516 6 Discussion

517 The results of this study lead us to a view of Atlantic hurricane risk which requires  
 518 us to confront epistemic uncertainty. We have multiple sets of simulations which give dif-  
 519 ferent views of the risk, in particular more so as we look further into the future. TC fre-  
 520 quency increases in CHAZ<sub>CRH</sub> simulations and decreases in CHAZ<sub>SD</sub>, a difference that  
 521 hangs on a subtle modeling choice (saturation deficit vs. relative humidity as a predic-  
 522 tor of genesis). Changes in the high-resolution global climate model simulations are smaller,  
 523 but again their direction depends on which global climate models are considered.

524 The differences among these simulations are manifest not just in the future, but  
 525 also to some degree in the present, and the observations do not allow us to determine  
 526 with complete certainty which is correct. At present, no rigorous justification can be given  
 527 regarding which choice to make. Thus, all these outcomes — increasing, decreasing, and  
 528 no change in TC frequency in response to radiatively forced warming — have to be treated  
 529 as possible. One may favor a dataset over the others following the results of a statisti-  
 530 cal analysis, such as the likelihood analysis used here. Our calculations indicate that the  
 531 observations are somewhat more consistent with CHAZ<sub>CRH</sub>, followed by Hi-TempExt,  
 532 MRI-LENS, Hi-TRACK. However, the likelihood values among these four datasets are  
 533 close to each other, so it would not be justified to draw a definitive conclusion from this  
 534 analysis as to which model is most correct.

535 The epistemic uncertainty in CHAZ's projections on annual TC frequency comes  
 536 from our design of the CHAZ model, but the conclusion is that our results are consis-  
 537 tent with the level of broader understanding of TC frequency at present, including that  
 538 derived from the latest high-resolution models shown here as well as other downscaling  
 539 systems (Sobel et al., 2021). Other aspects of TC characteristics that could change with  
 540 anthropogenic climate change have been also evaluated here, namely the forward mo-  
 541 tion and LMI95, are less dramatically uncertain, although our analyses show that one  
 542 cannot rule out the role of natural variability. Still, the uncertainty regarding TC fre-  
 543 quency introduces a large uncertainty into any assessment of overall TC risk, since any  
 544 change of TC properties is only relevant to the extent that TCs actually occur.

545 The divergence between increasing and decreasing TC frequency scenarios becomes  
 546 most pronounced in the latter part of the 21st century, but has some impact on the present  
 547 and near future as well (Lee et al., 2020, 2022). In the situation when the change of fre-

548 quency is subtle, changes in other TC properties may lead to differences in regional TC  
549 risk assessment.

550 How one views the situation must ultimately be based on one's attitude towards  
551 risk and the consequences of being wrong in either direction. A priori, though, we ar-  
552 gue that the most rational way to treat epistemic uncertainty is to consider all outcomes  
553 contained in the results to be possible. In the present context, since the results contain  
554 possible outcomes in which TC risk — as estimated from a single model or subset of the  
555 entire multi-model ensemble — is increasing, that in itself means we should regard TC  
556 risk as increasing, at the highest level of understanding in which all available informa-  
557 tion is considered, even though there are other possible outcomes in which it is decreas-  
558 ing.

## 559 Open Research Section

560 CHAZ is an open-sourced model (<https://github.com/c13225/CHAZ>). IBTrACS  
561 data are available at (<https://www.ncdc.noaa.gov/ibtracs/>). Information for CMIP5  
562 data can be found at <https://pcmdi.llnl.gov/mips/cmip5/> and HighResMIP trop-  
563 ical cyclone information can be found at (<http://catalogue.ceda.ac.uk/uuid/e82a62d926d7448696a2b60c1925f8>).  
564 Underlying data for this publications are at ([https://github.com/c13225/Lee\\_etal\\_2023EarthsFuture](https://github.com/c13225/Lee_etal_2023EarthsFuture)).  
565

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**Table 1.** Data Characteristics

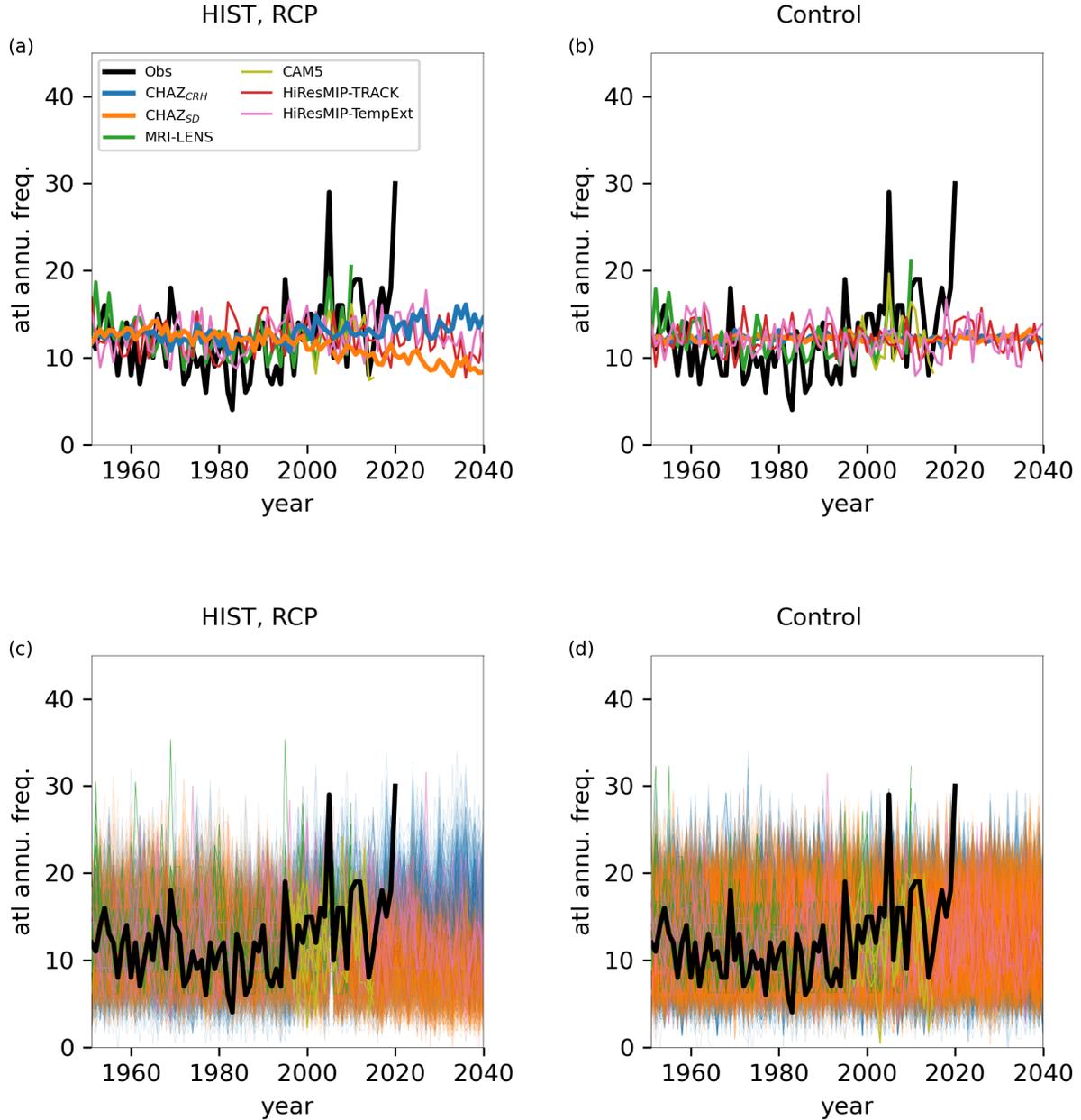
data	global climate models	resolution	ens	period	annual frequency
CHAZ <sub>CRH/SD</sub>	HadGEM2_ES	N/A	100	1951-2005;2006-2040; piC	8.8 /15.9
	CCSM4				11/16.1
	GFDL_CM3				16.5/19.1
	MPIESM_MR MIROC5				29.3/39.4 11.9/18.3
MRI-LENS	MRI-AGCM3.2H	60 km	100	1950-2010; piC	2.3
CAM5	CAM5	28km	5	1996-2005;piC	10.9
Hi-TRACK	CMCC-CM2-VHR4 (r1i1p1f1)	25 km	7	1951-2014;2015-2040;piC	5.0
	CNRM-CM6-1-HR(r1i1p1f2)	50 km			21.0
	EC-Earth3P-HR (r1i1p2f1)	50 km			6.8
	EC-Earth3P-HR (r2i1p2f1)	50 km			6.5
	HadGEM3-GC31-HH (r1i1p1f1)	50 km			21.5
	HadGEM3-GC31-HM (r1i1p1f1)	50 km			19.7
	MPI-ESM1-2-XR (r1i1p1f1)	50 km			4.5
Hi-TempExt	CNRM-CM6-1-HR(r1i1p1f2)	50 km	6	1951-2014;2015-2040;piC	13.4
	EC-Earth3P-HR (r1i1p2f1)	50 km			2
	EC-Earth3P-HR (r2i1p2f1)	50 km			2
	HadGEM3-GC31-HH (r1i1p1f1)	50 km			13.3
	HadGEM3-GC31-HM (r1i1p1f1)	50 km			12.4
MPI-ESM1-2-XR (r1i1p1f1)	50 km	0.63			

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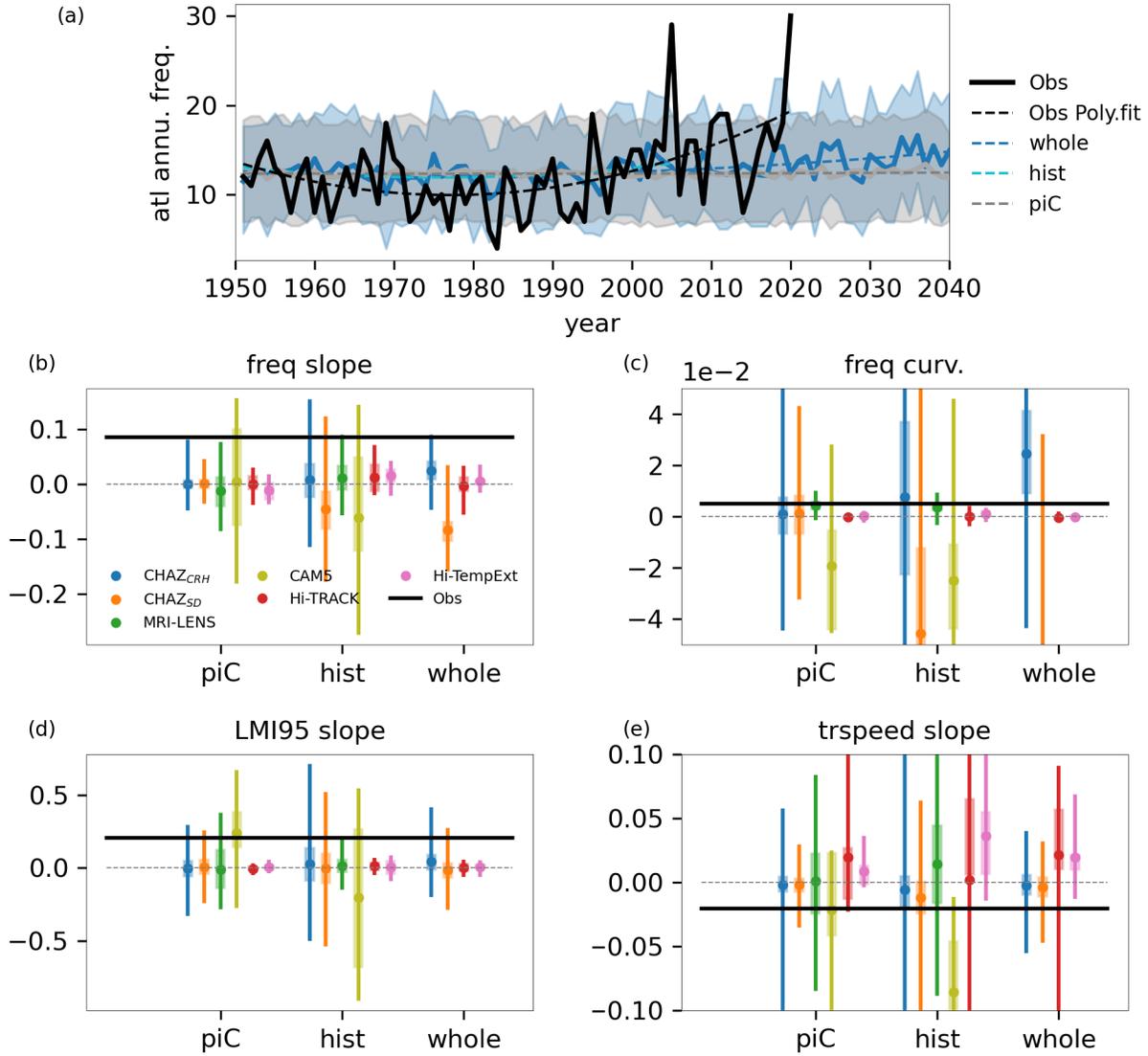
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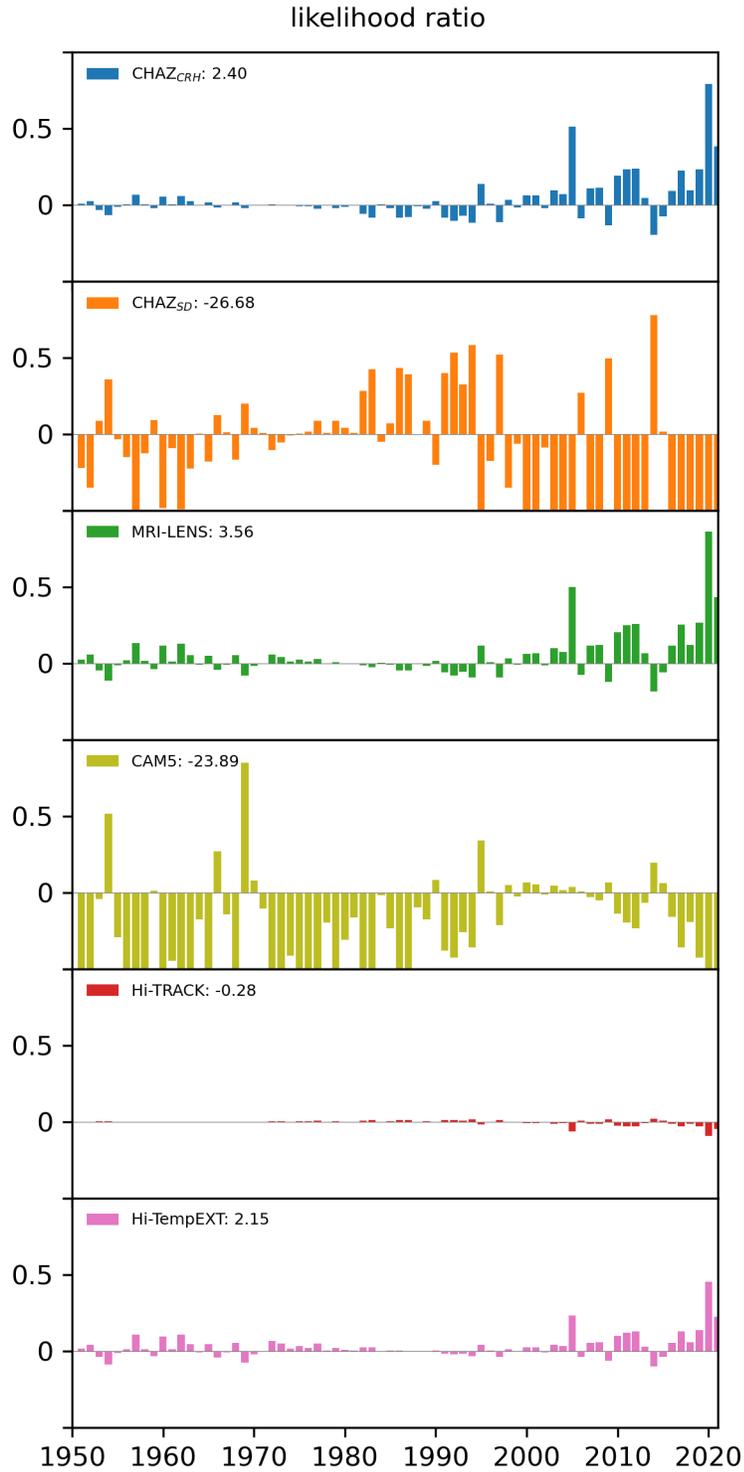
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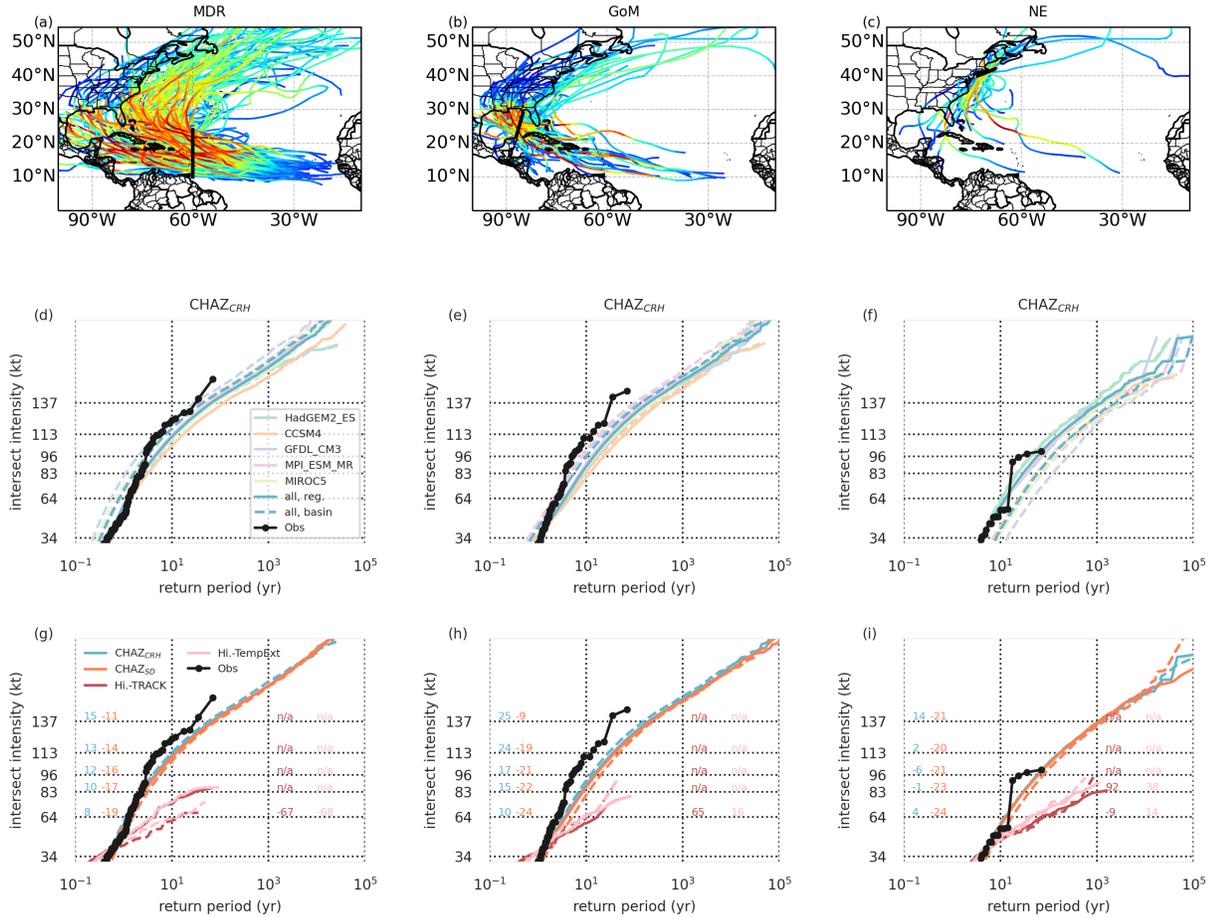
**Figure 1.** Annual frequency of Atlantic TCs exceeding 34 kt intensity threshold from 1951–2020 from best-track data (black), CMIP5 downscaling simulations using CHAZ<sub>CRH</sub> (blue) and CHAZ<sub>SD</sub> (pink), 25-km high-resolution CAM5 simulations (purple), 60km Japanese large-ensemble simulations (MRI-LENS, green), and HighResMIP simulations from (Roberts et al., 2020) and (Roberts et al., 2020). Storms from HighResMIP are tracked with TRACK (red) and TexmpExtreme (pink), respectively. In (a) and (c), simulations in their respective historical period are conducted with historical climate forcing while those in future period are with the rcp8.5 (for CHAZ) and ssp585 (for HighResMIP) warming scenarios. In (b) and (d), the simulations are under pre-industrial control climate (no anthropogenic forcing). (a) and (b) show the results from ensemble mean while (c) and (d) show the results from all ensemble members.



**Figure 2.** (a) Observed (black) and CHAZ<sub>CRH</sub> simulated mean annual hurricane frequency. The CHAZ simulations are from present (1951-2005) to future climate (2006-2040) periods (blue), and from those using pre-industrial control climate forcing (gray). Dashed lines show the polynomial fit. ‘hist’ shows the fit using synthetic storms from historical period only while ‘whole’ are from the historical and future periods. (b) Linear terms of the polynomial fit derived using synthetic storms’ annual frequency from all datasets. Datasets are indicated by color while the black line show the observed value. (c) Similar to (b) but for the quadratic terms. (d) and (e) are similar to (b) but for linear terms from the polynomial fit of LMI95 and storm forward motion speed. Units for (b), (c), (d) and (e) are, respectively, storm number year<sup>-1</sup>, storm number year<sup>-2</sup>, m s<sup>-1</sup> yr<sup>-1</sup>, and km hr<sup>-1</sup> yr<sup>-1</sup>.



**Figure 3.** Annual log-likelihood ratio in which  $\lambda_t$  is derived from historical (and future for the CHAZ and HighResMIP runs) simulations and the annual likelihood that is estimated based on piC simulations.



**Figure 4.** (a–c) Observed storm tracks from 1951–2020 at three line gates of interest. (d–f) Return period curves from 1951–2020 from best-track data (black lines), and CHAZ<sub>CRH</sub> historical simulations with basin-wide (dashed lines) and local (solid lines) basin corrections applied at the three gates. Global climate model forcings are indicated by colors and blue lines show the derived return period curves using all data. (g–i) Similar to (d–f) but for the four datasets. The solid lines show the return period curves using all historical simulations while dashed lines use all future simulations. Numbers at each Saffir-Simpson intensity threshold are the percentage changes of the frequency of the storms exceeding the threshold. Datasets are indicated by colors. Black curves show the empirical return curve using observations from 1951–2020.