Historic Trends and Sources of Year-over-Year Stability in Montana Winter Wheat Yields

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Abstract

Producers desire varieties that consistently perform with high yields and end-use qualities. Unlike easily recognized average yield improvements, yield stability over time is less examined, especially when considering the role of breeding relative to other factors like management and changing climatic conditions. Our study system was a 70-year historical dataset from which we estimated the year-over-year stability of Triticum aestivum, winter wheat varieties released by Montana's Agricultural Experimental Station. We examined yield stability within six locations representing diverse growing conditions across Montana and found evidence that breeding has improved stability at specific locations and not at others. Newer varieties showed improved year-over-year stability at locations that tended to have the lowest yields and more extreme weather conditions, reflecting that year-over-year stability has a genotype-by-environment component. We examined the role of climatic conditions, including temperature and rainfall to understand if reduced climatic variability was driving patterns of improved stability at these sites. However, the impact of breeding remained, or became evident when accounting for climatic variables. Together, these findings suggest that breeding's strong selective pressures improve second order traits.

Supplemental Information

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Release Year	Variety	Cumulative acres planted (thousands of acres), 1958 -2019*	Varieties g
1921	Karmont	2419	х
1926	Newturk	518	х
1932	Yogo	2128	х
1956	Itana	778	х
1956	Rego	111	х
1956	Westmont	165	х
1965	Sawmont	-	х
1967	Crest	67	х
1968	Froid	234	х
1971	Teton	30	
1979	Redwin	6447	х
1981	Rosebud	-	х
1981	Winridge	87	х
1983	Cree	514	х
1984	Norwin	143	х
1988	Tiber	4347	х
1989	Judith	1155	х
1994	NuWest	125	х

1995	Vanguard	1271	x
1996	Erhardt	67	х
1996	McGuire	50	х
1996	Rampart	4002	х
2001	BigSky	169	х
2001	NuSky	-	х
2003	Paul	-	х
2004	Genou	3840	х
2004	MT1159CL	-	
2005	Hyalite	-	
2005	Norris	-	
2005	Willow Creek	202	
2005	Yellowstone	4354	Х
2006	Bynum	-	
2010	Decade	835	х
2011	Bearpaw	232	х
2011	Judee	1515	х
2012	SY Clearstone 2CL	173	
2013	Colter	14	
2013	Warhorse	1210	х
2013	WB3768	-	
2015	Northern	-	
2016	Loma	72	
2016	Spur		
2018	FourOsix		
2018	MTF1435		
2018	Ray		
2019	Bobcat		
2019	Flathead		
2020	StandClear CLP		



Fig. S1 Yield residuals do not demonstrate increasing robustness over time

Residuals compared to fitted values from the linear regression fit of yield and year from data shown in Figure 1 at six locations are shown.



Fig. S2 Calculating a variety's yield robustness using CV is dependent on the number of years a variety was grown

a) The yield CVs are shown for each variety at each location grown and the corresponding number of years used to calculate the CV. The solid blue line is the fit of a quadratic plateau model. b) For each research centers, the yield CVs are shown for each variety and the corresponding number of years used to calculate the CV.



Fig. S3. Bootstrapping to calculate CV estimates and relate them to release year shows variability across locations

CV was calculated using sampling with replacement from five measurements for each variety that was measured at least five times. Linear regression was performed on CV values for release year. The p-value was determined for comparing that slope to 0 using a t-test. This was repeating 1000 times and the density of p-values is shown. The red vertical lines indicate p = 0.05. In a-c) results are shown across all locations. In a) the p-value distribution is shown. In b) the distribution of R^2 values is shown, and in c) the distribution of the slope estimates across the 1000 bootstraps is shown. In d-i) the p-value distributions are shown for the six locations.



Fig. S4 Annual and monthly weather patterns 1949-2019

a-d) Historic annual values for average a) mean, b) maximum, c) minimum temperatures, and d) total precipitation are shown for each location. The data was fit using linear regression to detect overall trends. e-h) The mean monthly values from 1949-2019 for e) mean, f) maximum, g) minimum temperatures, and h) total precipitation are shown for each research center.

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essoar.10511090.1.docx available at https://authorea.com/users/556406/articles/606433historic-trends-and-sources-of-year-over-year-stability-in-montana-winter-wheat-yields

2	Historic Trends and Sources of Year-over-Year Stability in Montana Winter Wheat Yields
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9	
10	ABSTRACT
11	Producers desire varieties that consistently perform with high yields and end-use qualities.
12	Unlike easily recognized average yield improvements, yield stability over time is less examined,
13	especially when considering the role of breeding relative to other factors like management and
14	changing climatic conditions. Our study system was a 70-year historical dataset from which we

15 estimated the year-over-year stability of *Triticum aestivum*, winter wheat varieties released by

16 Montana's Agricultural Experimental Station. We examined yield stability within six locations

17 representing diverse growing conditions across Montana and found evidence that breeding has

18 improved stability at specific locations and not at others. Newer varieties showed improved year-

19 over-year stability at locations that tended to have the lowest yields and more extreme weather

20 conditions, reflecting that year-over-year stability has a genotype-by-environment component.

We examined the role of climatic conditions, including temperature and rainfall to understand if reduced climatic variability was driving patterns of improved stability at these sites. However, the impact of breeding remained, or became evident when accounting for climatic variables. Together, these findings suggest that breeding's strong selective pressures improve second order traits.

26

INTRODUCTION

27 Volatile crop yields caused by environmental extremes threaten our global food supply (Lobell et 28 al., 2008). Further, producers lose economic value when crops grow variably across fields, due to increased difficulties in management and harvest and meeting end-use quality requirements. To 29 30 counteract these obstacles, multiple approaches are taken to optimize yield stability and quality, 31 including breeding and agronomic practices (Reckling et al., 2021). Crops grown at a specific geographic location experience not only varied weather and disease pressure year-over-year, but 32 also microenvironmental differences within a field and stochastic errors during development 33 (Lachowiec et al., 2015). Genotype greatly influences how plants withstand collective 34 perturbations (Hill & Mulder, 2010), herein referred to as robustness. Because genotypes 35 associated with differing degrees of robustness have been identified in plants (Hall et al., 2007; 36 Jimenez-Gomez et al., 2011; Sangster et al., 2008), including crops (Fisher & Zamir, 2021; 37 Makumburage & Stapleton, 2011; Ordas et al., 2008; Tollenaar & Wu, 1999), we hypothesize 38 39 that breeding has increased crop robustness to such perturbations.

Defining and measuring robustness is a complicated task. Not to be confused with the stability
(or conversely, the plasticity) of a trait across geographic location, robustness is a way to
describe the uniformity, repeatability, intra-genotypic variability (Bruijning et al., 2020), and

43	predictability of a genotype within a location. In many studies examining agronomic traits, the
44	coefficient of variation (CV) of a trait is calculated. CV traditionally is used to identify undesired
45	high levels of variability to determine experimental validity (Bowman, 2001) but also potentially
46	holds information about robustness (Fasoula & Fasoula, 2002). In agriculture, robustness also
47	has been captured using many statistics like Ve (Schou et al., 2020) and heterogeneity of
48	environmental variance (Hill & Mulder, 2010), among many others (Reckling et al., 2021),
49	revealing how different genotypes can influence robustness. Using these and similar robustness
50	statistics, the underlying genetic architecture and genetic loci controlling robustness have been
51	identified (Fisher & Zamir, 2021; Hall et al., 2007; Sangster et al., 2008).
52	Like most quantitative traits, robustness is affected not only by genotype, but also environment
53	and the interaction of genotype and environment (Falconer & Mackay, 1996). Most commonly,
54	how the environment affects robustness is measured within a single variety to assess the impact
55	of different agronomic practices. Varied planting density (Lu et al., 2020), application of
56	fertilizers, and irrigation (Kristensen et al., 2008; Kukal & Irmak, 2018) are examples of
57	agronomic practices examined for impacts on robustness. Long-term studies have produced
58	mixed findings regarding how robustness has changed over time in crops. An economic study of
59	global wheat and maize yields produced since the Green Revolution show an increase in
60	robustness over time (Gollin, 2006). In contrast, an agronomic study of a single, highly
61	controlled, long-term site found both barley and wheat yields decreased in robustness, dependent
62	on fertility management (Macholdt et al., 2021). A breeding study considering the impact of a
63	small number of specific maize hybrids released over five decades, found increased robustness is
64	associated with greater yields, due to improved stress tolerance (Tollenaar & Wu, 1999).

Additional information is needed to understand how breeding, location, and managementinfluence yield robustness.

To explore how breeding affects robustness, we examined historical data collected on winter 67 wheat grown in Montana. In Montana, winter wheat is one of the main cereal crops with a 68 planting area of over 627,000 hectares in 2020, accounting for nearly \$400 M in production 69 value (National Agricultural Statistics Service, 2021). Since 1949, Montana's winter wheat 70 breeding program grew winter wheat systematically at multiple research centers across the state, 71 72 collecting yield data. With these data, we examined yield year-to-year variability. We inspected 73 how yield robustness varies with geographic location. Given the important role of weather variability from year-to-year at a specific location, we examined the impacts of temperature and 74 75 precipitation on yields and its robustness. We propose the role of breeding in yield robustness 76 based on examining the most grown Montana Agricultural Experiment Station varieties 77 developed over the last 100 years. **MATERIALS AND METHODS** 78 79 **Data sets** 80 Yield data 81 Winter wheat yields were obtained from the Montana Winter Wheat Breeding Program, reported

as bushel per acre and converted to kg ha⁻¹. The 48 varieties included those that have been
planted across the state (Table S1, National Agricultural Statistics Service 1958-2019). The data
include the yields of released winter wheat varieties tested in Montana from 1949 to 2019 at six
research centers with the following soil types: Northern Agricultural Research Center with silt
loam (48° 30', 109° 48'; Havre, MT), Northwestern Agricultural Research Center with silt loam
(48° 10', 114° 15'; Kalispell/Creston, MT), Central Agricultural Research Center clay loam (47°

88 03', 109° 57'; Moccasin, MT), Southern Agricultural Research Center with silt loam (45° 55',

89 108° 15'; Huntley, MT), Eastern Agricultural Research Center with clay loam (47° 40', 104° 08';

90 Sidney, MT), and the Post Research Farm silt loam (45° 41′, 111° 00′; Bozeman, MT). Extreme

91 winterkill resulting in zero yields were excluded from all analyses as outliers (Reckling et al.,

92 2021). Only dryland results were included with no additional irrigation.

93

94 Weather data

Weather data was obtained using the web interface and search capabilities provided by the 95 National Oceanic and Atmospheric Administration's National Centers for Environmental 96 Information (https://www.ncdc.noaa.gov/cdo-web/search), accessed April 1, 2021. The datasets 97 chosen were "Global Summary of the Month" and "Global Summary of the Year" ranging from 98 January 1949 to December 2019. Weather stations were chosen for their proximity to the 99 research centers and data availability. Data were obtained from Kalispell Glacier Airport (Station 100 101 IDUSC00244558), Sidney (Station ID USC00247560), Huntley Experimental Station (Station ID USC00244345), and Moccasin Experiment Station (Station ID USC00245761). In Bozeman, 102 two stations were needed to obtain desired weather data. From January 1949 to October 1966, 103 104 data were obtained from Bozeman Montana State University (Station ID USC00241044), and from November 1966 to December 2019, data were obtained from Bozeman 6 W Experimental 105 106 Farm, (Station ID USC00241047). In Havre, two stations were used to obtain data spanning the 107 focus period: January 1949 to February 1961 Havre Weather Bureau (Station ID USW00024035) and March 1961 to December 2019 Havre Airport (Station ID USW00094). From each weather 108 109 station, four metrics of weather were obtained: monthly average (TAVG), monthly minimum 110 (TMIN), and monthly maximum (TMAX) air temperatures and monthly cumulative precipitation

111	(PRCP). The percentage data coverage across the weather metrics follows: Bozeman-98.5%,
112	Havre-100%, Huntley-96.4%, Kalispell-99.4%, Moccasin-95.8%, and Sidney-97.8%.
113	
114	Statistical analyses
115	All statistical analyses were performed using R version 4.1.0 (R Development Core Team,
116	2011).
117	Yield analyses
118	To examine the change in yields over time, we calculated the mean yield each year across
119	varieties at the six research centers. We used a linear model to evaluate the rates of change when
120	calculating the mean across locations per year and at each location separately.
121	
122	To examine the change in yield robustness over time, we utilized multiple approaches. First, we
123	visually examined the residuals of the linear model capturing the relationship between yield and
124	year. Second, we calculated the coefficient of variation (CV = $\frac{s}{\overline{Y}}$) of yield of a variety across the
125	years it was grown (Ray et al., 2015). Because the number of years a variety was tested was
126	positively correlated with the CV (Fig. S2), we employed two alternative methods to account for
127	this relationship. First, we observed a threshold in the impact of number of years grown on the
128	CV and used a quadratic plateau model to estimate the transition point, implemented with the
129	nlstools package in R (Baty et al., 2015). The 95% confidence interval for the transition point
130	was 6.57-12.14, with an estimate of 9.36. Therefore, we used a threshold of greater than 9 years
131	grown to estimate CV values for varieties. In a second approach, data were sampled with
132	replacement to eliminate effect of number of years sampled on the CV. Varieties were filtered to
133	those that included measurements from at least 5 years. For each bootstrap sample, five years of

yield data were selected for each variety. The linear relationship between the CV of the
bootstrapped samples and release year was computed. This process was repeated 1000 times, and
the distributions of *b*, R², and p-value for *b* for each bootstrapped sample was examined.
Bootstrap estimates were calculated pooling across locations as well as for each individual
location.

139

140 Weather analyses

To understand trends in weather metrics between 1949-2019, we examined annual mean values obtained for each location. The relationship between metrics and year were modeled with linear regression for each location. The change in each weather metric was calculated per decade using the slope estimate. Annual weather variability was examined by calculating a "moving CV" with a moving window of 30 years using the annual mean values for each weather metric.

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Seasonality in the weather metrics were also examined at each location to understand if certain months were highly variability in temperatures relative to others. The monthly means were calculated between the years of 1949-2019. Weather variability over time was examined by using the statistic CV after transforming the data to only include positive values by adding 30 to all values. The CV was then calculated for the monthly weather metrics obtained for each month.

152

153 Combined weather and yield analyses

Because weather data were available each year at each location, we examined the relationship
between the weather metrics and yields. Using principal components analysis, we examined how
the four weather metrics, release year (capturing the impact of breeding), and year grown

157 (capturing, at least, changes in agronomic practices) were acting in concert. We then calculated 158 the residual CV (CV_{res}) (Kukal & Irmak, 2018; Schou et al., 2020) by fitting a linear model with 159 yields as the response variable and using the mean annual temperature and cumulative 160 precipitation, which contributed most strongly to PC1 and PC2, and their interaction as the 161 explanatory variables. Using the residuals from this model after transformation to positive values 162 by adding the minimum residual, the CV_{res} was calculated for each variety. Next, the relationship 163 between CV_{res} and release year was examined using linear regression.

164

RESULTS

165 Montana wheat yields have increased with varied impact on yield robustness

To assess how winter wheat yields have changed historically, we focused on the most tested 166 167 varieties in Montana's winter wheat breeding program since the 1920s. We examined forty-eight 168 varieties ranked on acreage planted in Montana after their release (Table S1, Montana 169 Agriculture Statistics Service 1958-2019), assessing the results of yield trials across six research 170 centers representing different production environments in Montana. We first examined winter wheat's interannual improvement between 1949-2019 by calculating the mean yield each year. 171 On average across the six locations, yields increased 43.55 kg ha⁻¹ yr⁻¹based on fitting a linear 172 model (Fig. 1a). Depending on location, yields showed over a five-fold average improvement 173 over the last 70 years (Fig. 1b). Huntley and Kalispell had the largest year-over-years gains (Fig. 174 1c). These yield changes are not only dependent on breeding improvement but also on 175 176 improvement in agronomic practices (Lanning et al., 2010) and changing climatic conditions and atmospheric CO₂ (McGrath & Lobell, 2013). 177





179

Fig. 1 Winter wheat yield increases across Montana from 1949-2019

180a) The average yields from Montana's released varieties from breeding nurseries across six research centers181is shown from 1949-2019. The linear regression fit is indicated with the solid black line, with adjusted R²182and p-values for the slope shown. b) Yields from 1949-2019 is shown for each research station. The linear183regression fit is indicated with the solid black line, with adjusted R² and p-values for the slope displayed. c)184The linear regression coefficient for yield and year for each research center from panel b) is indicated with185the color scale for the corresponding county.

186

In addition to observing increased yields over time, we hypothesized increased year-over-year 187 yield stability. To distinguish stability year-over-year within a location from the more common 188 definition of stability across geographic locations (Becker & Leon, 1988; Finlay & Wilkinson, 189 190 1963), we will use the term robustness to refer to year-over-year stability. In other words, within a single geographic location we can describe how robust yields are to yearly fluctuations in 191 weather, plot location at a research center, biotic stresses, and more. We hypothesized that the 192 193 residuals of the linear regression models estimating the change in yield over time at each location would decrease with time. However, the residuals showed little change over time, except at 194 Kalispell, where they instead appeared to increase, contrary to our prediction (Fig. S1). These 195

data suggest that the collective impacts of changing agronomic practices, climate, breeding, etc.
are not resulting in greater yield robustness, with Kalispell decreasing in robustness, and other
locations unchanging.

199

200 The association of variety release year and yield robustness depends on location

201 Varieties were pooled together in the previous analyses, potentially obscuring the impact of breeding on robustness. Given that wheat is primarily released as inbred varieties, the selection 202 pressure to increase yields during breeding may also tend to select genotypes with more narrower 203 204 trait distributions (Gavrilets & Hastings, 1994; Wagner et al., 1997)—or more predictable genotypes. We first tested whether breeding could be a component affecting yield robustness by 205 comparing the release year of a variety to its robustness using the statistic CV. We calculated the 206 yield robustness for each variety across years planted at a single location; however, we observed 207 that robustness was sensitive to the number of years it was grown (Fig. S2). 208

209

To reduce or eliminate the impact of number of years grown on robustness, we used two approaches. First, based upon an estimated plateau in decreased robustness due to number of years grown (Fig. S2a), we set a conservative threshold of greater than 9 years planted for inclusion. This approach reduced the number of varieties included and eliminated those most recently publicly released from further analysis, because they have not existed long enough to be tested for 10 years. With these filtered data, we detected a very weak negative linear relationship between CV and release year among locations in this filtered dataset (Fig. 2a, Table 1).





218

Fig. 2 Yield CV for varieties decreases with release year

a) The yield CV was calculated for each variety at each research station for those grown at least 10 years.
The solid line indicates the fit of yield CV and release year of corresponding variety from linear regression.
b) The yield CV for each variety is shown. Black lines indicating the yield CV from each research center
for which the variety was grown at least 10 years. The red lines indicate the mean yield CV for each
variety, and the dotted black line shows the grand mean yield CV. The varieties are arranged in order of
release.

	Ų	<i>,</i>	j		-
Location	n	b	SE _b	R^2	
All	138	-0.001071***	0.0003359	+0.0651	
Havre	24	-0.002589***	0.0004373	+0.5969	
Sidney	17	-0.003069***	0.0009891	+0.4878	
Huntley	25	-0.001776***	0.0006205	+0.2306	
Moccasin	26	-0.000475	0.0003297	+0.04118	
Bozeman	27	-0.000340	0.0004239	-0.01395	
Kalispell	19	+0.003141*	0.001137	+0.2693	

TABLE 1 Linear regression of variety yield CV on variety release year

** Significant at the .01 probability level, *** Significant at the .001 probability level.

227 Given the yield of a variety is sensitive to geographic location in which it was grown, we

228 hypothesized that robustness of a variety also would depend on location. We examined the

229 location-specific yield CV values for varieties grown over nine years and observed that the CV

was not consistent across locations and could vary widely (Fig. 2b). Only two varieties, 230 Yellowstone and Judee, showed CV values across all locations that were less than the overall 231 mean. The impact of location on the relationship between release year and robustness suggests 232 multiple potential explanations, including that the environmental conditions experienced at each 233 locations drove the specific yield CV patterns observed over time. Indeed, when examining 234 235 robustness within specific locations, we found different tendencies due to location between the relationship of release year and yield CV. Havre and Sidney showed the strongest relationship 236 between release year and robustness (Fig. 3a, Table 1). Kalispell exhibited a pattern opposite to 237 238 the other locations with robustness decreasing in more recently released varieties (Fig. 3a).





Fig. 3 The relationship between yield CV and release year of a variety depends on location
a) The Yield CV was calculated for each variety at each research station for those grown at least 10 years.
The solid line indicates the fit of yield CV and release year of corresponding variety from linear regression.
b) The yield CV is shown for each variety grown at least ten years at each research station (black lines).
The mean yield CV for each research station is indicated with the red lines, and the grand mean is indicated with the dotted black line.

246

247 The threshold-based approach used above to determine the minimum number of years from248 which to calculate robustness may be biased. It is possible that lines that were planted for many

years may simply be more robust due to breeder selection. Therefore, we also performed a 249 second, similar analysis using bootstrapping to understand the relationship between robustness 250 251 and release year. In this approach, we calculated robustness using five randomly chosen yield values for each released line measured at least five years. Then we modelled the relationship 252 between release year and robustness. We performed 1000 bootstraps and examined the summary 253 254 statistics' distributions for evidence of a relationship between release year and robustness. Using the data across locations, we observed a negative relationship between CV and release year 255 (mean of 1000 bootstraps: b = -0.0125, P = 0.0197, $R^2 = 0.179$, Fig. S3a-c), a very similar 256 257 outcome as the threshold approach. Next, we performed this same bootstrap procedure for each location (Fig. S3d-i). Similar to the previous approach using a 10-year filter, we again found 258 support for a negative relationship between release year and robustness in Havre and Huntley. At 259 Kalispell and Sidney, the indication of relationships between robustness and release year was not 260 as evident; the data do not support a relationship. 261

262

To begin to understand why locations showed differences in the relationship between robustness 263 and release year, we first examined the range of CV values. We noted that the CV varied with 264 265 geographic location, with higher CV values at Huntley, Sidney, and Havre (Fig. 3b). We suspect that to observe a relationship between CV and release year, locations must have environments 266 267 that have weather and field perturbations to be overcome, and the absence of this variability 268 would prohibit observing a pattern. Thus, this analysis suggests that breeding has contributed to improved robustness in a manner that is dependent on location or can only be observed in certain 269 270 locations or environmental conditions.

271

272	We also found evidence that biotic perturbations influenced levels of robustness in a location-
273	specific manner. At Kalispell, Decade exhibited the lowest robustness in yield of the entire
274	dataset (yield $CV > 0.6$, Fig. 3b). This is due to very large range of yields produced over 12 years
275	of testing, ranging from 605.25 kg ha ⁻¹ to 9280.5 kg ha ⁻¹ . The low yields for Decade were
276	location specific and not observed at other sites. Examining the breeder's notes, we found that
277	second lowest value (1076 kg ha ⁻¹) measured was associated with stripe rust at Kalispell. The
278	lowest value (605.25 kg ha ⁻¹) was the following year. In addition to the abiotic, agronomic
279	practices, and breeding impacts considered, biotic stresses also can greatly impact the robustness
280	measured. In this case, however, we observe that the robustness is associated with genotype, and
281	conclude that Decade represents a genotype particularly susceptible relative to others tested at
282	the same location and year (Kertho, 2014; Riveland et al., 2011).
283	

284

Assessment of weather variability across locations

We hypothesized that the differences in yield robustness between locations in part reflects the 285 weather conditions experienced by a variety the years it was grown. Therefore, we first assessed 286 mean weather patterns at the six locations at multiple scales, focusing on temperature and 287 288 precipitation, both of which are critical input to crop yield (Kukal & Irmak, 2018). Between 1949-2019, the six locations varied in average annual temperature and precipitation, with 289 290 Huntley having the highest average annual temperatures and Havre having the lowest 291 precipitation (Fig. S4a-d). At the same time, all six locations increased in maximum annual temperatures between 1949-2019 (Table 2). Sidney had the largest average annual temperature 292 change between 1949 and 2019, increasing 0.41°C per decade. Sidney, Bozeman, and Moccasin 293 also had increasing minimum average annual temperatures during this time. Precipitation trends 294

295	also varied per location. Bozeman and Moccasin, which tended to have higher precipitation
296	among these locations, were unchanged or tending towards decreasing during this time period,
297	but Sidney had an annual average precipitation increasing 1.33 cm per decade between 1949-
298	2019. For a better resolved perspective of seasonal temperatures at each location, we calculated
299	the monthly means from 1949-2019 (Fig. S4e-g). Havre and Sidney experienced the most
300	extreme temperatures, not only facing high mean, minimum and maximum temperatures in the
301	summer months, but also the lowest temperatures in the winter months relative to the other four
302	locations. Monthly total precipitation means were also calculated for each month from 1949-
303	2019 (Fig. S4g). Kalispell has higher rainfall, especially during the winter and early spring,
304	relative to the other locations. Total monthly precipitation varied much more from month to
305	month and across locations relative to temperature (Fig. S4d).

	$b imes 10^{\circ}$				
Location	Temperature annual mean	Temperature annual maximum	Temperature annual minimum	Precipitation annual cumulative	
Havre	0.09212	0.17410*	0.01086	0.00453	
Sidney	0.41412***	0.39452***	0.43258***	1.33099*	
Huntley	0.15281*	0.19018*	0.11719	0.80292	
Bozeman	0.16543**	0.19713**	0.13363**	-0.15007	
Moccasin	0.22628**	0.19209*	0.24095***	-0.03251	
Kalispell	0.11965**	0.16146**	0.07672	0.19277	

306	TARLE 7 Linear	regression coef	fficients from rea	rression of weath	er metrics on ves	$r \text{ from } 1040_2010$
300		regression coel	melents nom reg	gression of weath	ier metries on yee	11011117772017.

307 * Significant at the .05 probability level, ** Significant at the .01 probability level, *** Significant at the .001
 308 probability level. ^aTo indicate trends per decade, yearly coefficients are shown multiplied by 10.
 309

310 Using these monthly and annual weather metrics, we examined weather variability over time in

the dataset. We calculated monthly CV statistics for the mean, minimum, and maximum

temperatures, and the precipitation at each location from 1949-2019 (Fig. 4a-d). For all three

summaries of temperature, we observed high CV in the fall and winter months, with a peak in 313 January. Across locations, Havre and Sidney ranked as having the highest CV values during 314 these winter months. In contrast, during the spring and summer months of winter wheat growth, 315 the temperature CV were indistinguishable among locations, indicating little year-over-year 316 variation in monthly temperatures. These patterns suggested that if year-over-year variability is 317 318 impacting yield robustness, the source may be variation in the fall and winter months. Because winter wheat is planted in the fall and vernalizes over winter, temperature throughout winter also 319 has direct impacts on the plants, including winter kill. Year-over-year precipitation showed a 320 321 different pattern from temperature, with more CV values ranging from 0.40 - 1.15, again with Sidney and Havre exhibiting a tendency for more variability. 322



323

324

Fig. 4 Climate metrics show location-specific patterns in variability

a-d) The CV of monthly values from 1949-2019 for a) mean, b) maximum, c) minimum temperatures, and
d) total precipitation are shown for each research center. e-h) CV values calculated from a 30-year moving
window are displayed for each location for the average e) mean, f) maximum, g) minimum temperatures,
and h) total precipitation.

329



331 (Whitlock et al., 2017). We used two approaches to examine historical trends in weather

variability at the six locations examined. First, we calculated a "moving CV", based upon the 332 annual temperature and precipitation values, using window sizes of 30 years, a standard in 333 calculating climate normal (U.S. Climate Normals, 2021). Contrary to our expectation that 334 weather variability is generally increasing, we observed that historic weather variability 335 depended on location (Fig. 4e-h). Fitting a linear model to the moving CV values, we found that 336 337 Bozeman weather in fact was converging for minimum temperature variability but showed no change for other weather metrics (Table 3). At Kalispell, all measures of temperature were 338 converging, with precipitation variability remaining steady, consistent with previous patterns 339 observed in Montana (Zhang et al., 2021). At Havre, Sidney, Huntley and Moccasin, the 340 different weather metrics showed mixed patterns of variability (Table 3). Sidney, for example 341 was diverging in all temperature metrics but converging for precipitation. Moreover, an abrupt 342 upward variability shift in all three temperature metrics occurred in Sidney in approximately 343 1986 (Fig. 4e-g). We confirmed this shift was not due to changes in equipment at the Sidney 344 weather station or missing weather data (only found in 2010s), but we have not identified sources 345 of this shift. Out of the 24 weather-location combinations, eight showed convergence in weather 346 variability, eight with divergence, and eight with no change. 347

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350	TABLE 3 Linear regression coefficients from 30-year moving window CV of multiple weather metrics
351	on year from 1954-2004.

	<i>b</i>								
Location	Temperature annual mean CV	Temperature annual maximum CV	Temperature annual minimum CV	Precipitation annual cumulative CV					
Havre	3.625×10 ⁻⁵	1.947×10 ⁻⁴ **	-2.460×10 ⁻⁴ **	1.6190×10 ⁻³ **					
Sidney	3.411×10 ⁻⁴ **	4.423×10 ⁻⁴ **	2.817×10 ⁻⁴ **	-1.2177×10 ⁻³ **					
Huntley	-2.430×10 ⁻⁵	-1.332×10 ⁻⁴ **	1.390×10 ⁻⁴ **	1.729×10 ⁻⁴					
Bozeman	-4.244×10 ⁻⁵	7.501×10 ⁻⁵	-2.238×10 ⁻⁴ **	1.507×10 ⁻⁴					
Moccasin	1.766×10 ⁻⁴ **	2.102×10 ⁻⁴ **	-5.304×10 ⁻⁵	-5.766×10 ⁻⁴ **					
Kalispell	-1.069×10 ⁻⁴ **	-7.245×10 ⁻⁵ *	-2.398×10 ⁻⁴ **	-1.416×10 ⁻⁴					

* Significant at the .05 probability level, ** Significant at the .01 probability level.

³⁵³

354	Accounting	for weather	demonstrates	remaining g	eographic influence	s on robustness
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Weather, breeding, and agronomic practice are expected to impact yield CV at all locations but 355 are not independent of one another. For example, while year grown reflects improvement in 356 357 agronomic practices, it also captures increasing temperatures and is strongly correlated to release year. Therefore, we first examined the relationships among release year (to reflect breeding), 358 year grown (to reflect agronomic practice at least), and the four weather metrics using principal 359 360 component analysis. The first dimension described 47.8% of the variation (Fig. 5a), primarily 361 capturing the three temperature metrics, along with a substantial contribution of year and release 362 year. The second dimension was driven by year and release year, representing an addition 27.4% 363 of the variation, which showed a close visual association with yield increase, as expected (Fig. 364 5b). The third dimension describing 17.4% of the variation primarily represented precipitation 365 differences (Fig. 5a). Because of the inter-relatedness of weather, agronomic practices, and breeding due to their correlations with time, we isolated their potential impacts on yields. A 366 partial correlation analysis showed that after accounting for the correlations between release 367

368 year, weather metrics, and year, a remaining correlation of r = 0.0583 was remaining between 369 yield and release year of varieties examined (Table 4) though the strongest correlation with yield



370 was with precipitation (r = 0.29).

371

Fig. 5 Non-independence of multiple metrics affecting robustness

a) Bar chart of percent of variance explained per principal component. b) Loadings of metrics to each

374 component (labeled Dim1, Dim2). c) Biplot for each variety tested, colored by yield.

375

			(°C)			
	Yield	Temperature annual mean	Temperature annual maximum	Temperature annual minimum	Precipitation annual cumulative (cm)	Year
Temperature annual mean	-0.0013					
Temperature annual maximum	-0.0024	0.9994***				
Temperature annual minimum	0.0112	0.9984***	-0.9965***			
Precipitation annual cumulative	0.2929***	0.0548*	-0.0682**	-0.0400		
Year	0.2026***	0.0257	-0.0238	-0.0225	-0.1124***	
Release year	0.0583*	-0.0220	0.0218	0.0202	0.0293	0.8551***

376	TABLE 4	Partial	correlations ^a	analysis	among	yield	and	potential	expl	anatory	factors,	all	locations
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377 * Significant at the .05 probability level, ** Significant at the .01 probability level, *** Significant at the .001
 378 probability level. ^aPearson correlation coefficients given.
 379

With this firmer basis of how yields are collectively affected by weather, we next aimed to parse 380 out the impact of weather for influencing yield robustness in these data. We revisited the filtered 381 dataset with years grown greater than nine, to avoid the impact of years grown. We obtained the 382 corresponding average annual temperature and annual average precipitation for each variety 383 during the specific years it was grown. Due to missing weather station data, we were unable to 384 include data for some years. We modeled how yields at each location were affected by average 385 386 temperature and precipitation, the largest contributors to principal components 1 and 3, as well as their interaction. Finally, we recalculated each varieties' yield robustness as the CV_{res} based on 387 the residuals of the models following the approach by Schou et al (2020). Overall, we observed 388 389 similar patterns as before. Havre continued to show a strong negative relationship between

release year and CV_{res} even after removing the impact of average temperature and precipitation (Fig. 6). Sidney and Huntley showed less strong relationships once removing the impact of weather. We detected no relationship between release year and CV_{res} in Bozeman, and in Kalispell, the opposite pattern was again found. In contrast, we noted a non-linear relationship in Moccasin, with a decrease in CV_{res} after the year 2000. Overall, even after some correction for weather differences, we continued to observe location specific impacts on robustness, pointing to many other factors involved in influencing patterns of robustness over time.



397

398Fig. 6 The relationship between yield CV_{res} and release year of a variety is dependent on location399After modeling the impact of average temperature and precipitation on yield, yield CV_{res} was calculated at400each research station for each variety grown at least 10 years and plotted against the variety release year.401Adjusted R² and p-values for the regression coefficient are displayed.

402

403

DISCUSSION

404 Historical data collected from specific varieties of winter wheat grown at defined locations in

405 Montana enable studies of the change in year-over-year stability, or robustness, of yields over

time. It is important to consider to which perturbations the trait is robust (Wagner, 2007). 406 Examining the robustness of repeated organs formed on an individual, such as tillers in wheat, 407 reveals robustness to stochastic errors of development. Examining individual plants within a 408 single plot reveals robustness to microenvironmental variation as well as developmental 409 differences. On a larger scale, measuring robustness across multiple plots captures robustness to 410 411 field variation, and finally, robustness has also been considered across years, revealing robustness to weather or variation in agronomic practices at the very least. In this experiment, we 412 measured aggregate robustness across these levels (Falconer & Mackay, 1996) and detected 413 tendencies for robustness to increase over time, dependent on breeding, but this was dependent 414 on geographic location. 415

416

A simple examination of the residuals of a linear fit of yield over time as a proxy for robustness 417 did not point to improved robustness over time. This approach cannot distinguish varieties, 418 agronomic practices, weather changes or other influences on yield. For example, improved 419 management and other non-genetic inputs at these research sites are known to affect yields. A 420 study of spring wheat at these six research centers from 1950-2007 found that the variety 421 "Thatcher" showed increased grain yields throughout this period, ranging from 14.8 kg ha⁻¹ yr⁻¹ 422 in Moccasin to 50.8 kg ha⁻¹ yr⁻¹ in Kalispell (Lanning et al., 2010). However, examining 423 424 robustness of each variety separately by calculating a standardized variance using the statistic 425 CV showed that more recently released varieties tend to be more robust (Fig. 2, 3). Importantly, estimating robustness requires a sufficiently large dataset. We found that estimates of robustness 426 427 using the statistic CV was dependent on the number of years a variety was grown. If a variety 428 was only grown for a few years, the CV was not representative of robustness. We detected a

plateau of approximately 10 years of data as sufficient to estimate robustness using CV (Fig.
S2a). We speculate that this plateau represents the number of years needed for a variety to
experience an extreme weather event on average.

432

Geographic location was a strong indicator of whether robustness increased with breeding. The 433 434 relationship between robustness and release year is observed at Havre, Sidney, and Huntley (Fig. 3a) and not at the other locations. We noted several potential shared characteristics at Havre, 435 Sidney and Huntley that may contribute to these locations demonstrating improved robustness. 436 437 First, Havre, Sidney, and Huntley had the lowest starting yields (Fig. 1c) in the 1950s. It is possible that these locations had the most to gain in both yield mean and robustness relative to 438 the other locations. Accordingly, the CV values across varieties tended to be higher and cover a 439 broad range at these locations (Fig. 3b), relative to Moccasin and Bozeman. Notably, Havre, 440 Sidney, and Huntley represent where most of the wheat is grown in the state by producers 441 (National Agricultural Statistics Service, 2021), possibly reflecting a focus of breeders to 442 optimize wheat's growth to those locations and/or that robustness improvement had the most 443 impact in locations with more extreme temperatures. Indeed, Havre, Sidney, and Huntley have 444 445 had the highest summer temperatures, with Sidney and Havre also having the lowest winter temperatures (Fig. S4e-g). The greater tolerance of perturbations at these highly variable 446 447 locations is consistent with observations in other crops, such as maize yields increasing due to 448 reduced plant-to-plant variability (Tollenaar & Wu, 1999). Kalispell stands out from the other locations that comprise the primary focus of the breeding program and represents a region of the 449 450 state more like eastern Washington. The selected lines are not well-adapted to this region, and 451 the decrease in robustness in Kalispell may reflect this. Pathogen and herbivore pressure also

vary over the years and were not considered in this study. Wheat stem sawfly and stripe rust have
devastating impacts on yields that can be location specific and breeding has focused on
producing resistant varieties that likely contribute to improved robustness. Finally, we also note
that management also did vary across research centers: during this period, Havre and Huntley
transitioned from tillage to chemical fallow, increasing soil moisture (Lenssen et al., 2007),
while other sites maintained tillage, including Sidney.

458

Weather conditions greatly impact yields of crops, evidenced by both the great impacts that 459 droughts and floods have on yields (Ray et al., 2015) within a region and how wheat yields are 460 sensitive to one-degree changes in temperatures in controlled conditions (Lanning et al., 2010). 461 We surmised that variability in weather conditions from year to year should be reflected in the 462 robustness of crops. Upon examining climate patterns of temperature and precipitation, we found 463 that temperatures are increasing at all research centers, and precipitation is unchanging or 464 increasing (Table 2). Second, the variabilities of weather metrics showed patterns of 465 convergence, divergence, and no change (Table 3). No general statements could be made to 466 relate change in robustness in yields as dependent on weather variability. After removing the 467 468 effect of weather metrics assessed, we continued to detect relationships between robustness and release year. Havre continued to show a strong negative relationship between release year and 469 CV_{res}. Sidney and Huntley showed less strong relationships once removing the impact of 470 471 weather. In the case of Sidney, this possibly reflects a tendency towards less variable precipitation (Fig. 4h, pink line) and at Huntley, this could reflect convergence in maximum 472 473 temperatures (Fig. 4f, orange line).

474

Overall, in this historical wheat dataset spanning 70 years, we can detect a signal that yield
robustness is dependent on breeding and location across many levels of perturbations, from
within plant to across years. Because of the extensive confounding of management and
environmental conditions, such as global CO₂ levels, the impact of breeding could best be tested
compared by growing historical varieties at the same site and time for multiple seasons to
estimate robustness. This work reveals how breeding influences robustness and how environment
influences its evolution—an inadvertent evolutionary selection experiment.

482

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AUTHOR CONTRIBUTIONS

JL: conceptualization, methodology, formal analysis, writing-original draft preparation,
writing-reviewing and editing, funding acquisition. JEB: data curation, writing-reviewing and
editing. ML: data curation, formal analysis, writing-reviewing and editing.

- 491 SUPPLEMENTAL MATERIAL
 492 Supplemental figures and tables are available as Supplemental Material.
- 493 CONFLICT OF INTEREST
- 494 The authors have no relevant financial or non-financial interests to disclose.

495	DATA AVAILABILITY
496	Code generated and data analyzed in this study are available at
497	https://github.com/Lachowiec-Lab/historical-stability.
498	REFERENCES
499 500 501	Baty, F., Ritz, C., Charles, S., Brutsche, M., Flandrois, JP., & Delignette-Muller, ML. (2015). A Toolbox for Nonlinear Regression in R: The Package nlstools. <i>Journal of Statistical Software</i> , <i>66</i> , 1–21. https://doi.org/10.18637/jss.v066.i05
502 503	Becker, H. C., & Leon, J. (1988). Stability Analysis in Plant Breeding. <i>Plant Breeding</i> , 101(1), 1–23. https://doi.org/10.1111/j.1439-0523.1988.tb00261.x
504 505	Bowman, D. T. (2001). Common use of the CV: a statistical aberration in crop performance trials (Contemporary Issue). <i>Journal of Cotton Science</i> , 5(2), 137–141.
506 507 508	Bruijning, M., Metcalf, C. J. E., Jongejans, E., & Ayroles, J. F. (2020). The Evolution of Variance Control. <i>Trends in Ecology & Evolution</i> , <i>35</i> (1), 22–33. https://doi.org/10.1016/j.tree.2019.08.005
509 510	Falconer, D. S., & Mackay, T. F. C. (1996). Introduction to quantitative genetics. Essex. UK: Longman Group.
511 512 513	Fasoula, V. A., & Fasoula, D. A. (2002). Principles underlying genetic improvement for high and stable crop yield potential. <i>Field Crops Research</i> , <i>75</i> (2–3), 191–209. http://dx.doi.org/10.1016/S0378-4290(02)00026-6
514 515 516	Finlay, K. W., & Wilkinson, G. N. (1963). The analysis of adaptation in a plant-breeding programme. <i>Australian Journal of Agricultural Research</i> , <i>14</i> (6), 742–754. https://doi.org/10.1071/ar9630742
517 518	Fisher, J., & Zamir, D. (2021). Genes for Yield Stability in Tomatoes. <i>Advanced Genetics</i> , 2(4), 2100049. https://doi.org/10.1002/ggn2.202100049
519 520 521	Gavrilets, S., & Hastings, A. (1994). A Quantitative-Genetic Model for Selection on Developmental Noise. <i>Evolution</i> , 48(5), 1478–1486. https://doi.org/10.1111/j.1558-5646.1994.tb02190.x
522 523	Gollin, D. (2006). Impacts of International Research on Intertemporal Yield Stability in Wheat and Maize: An Economic Assessment. CIMMYT.
524 525 526	Hall, M. C., Dworkin, I., Ungerer, M. C., & Purugganan, M. (2007). Genetics of microenvironmental canalization in Arabidopsis thaliana. <i>Proc Natl Acad Sci U S A</i> , 104(34), 13717–13722. https://doi.org/0701936104 [pii] 10.1073/pnas.0701936104
527 528	Hill, W. G., & Mulder, H. A. (2010). Genetic analysis of environmental variation. <i>Genetics Research</i> , 92(5–6), 381–395. https://doi.org/10.1017/S0016672310000546

- 529 Jimenez-Gomez, J. M., Corwin, J. A., Joseph, B., Maloof, J. N., & Kliebenstein, D. J. (2011).
- 530 Genomic analysis of QTLs and genes altering natural variation in stochastic noise. *PLoS Genet*,
- 531 7(9), e1002295. https://doi.org/10.1371/journal.pgen.1002295 PGENETICS-D-11-00547 [pii]
- 532 Kertho, A. O. (2014). *Evaluation of winter wheat germplasm for resistance to stripe rust and*
- 533 *leaf rust*. North Dakota State University.
- 534 Kristensen, L., Olsen, J., & Weiner, J. (2008). Crop Density, Sowing Pattern, and Nitrogen
- 535 Fertilization Effects on Weed Suppression and Yield In Spring Wheat. *Weed Science*, 56(1), 97–
- 536 102. https://doi.org/10.1614/WS-07-065.1
- 537 Kukal, M. S., & Irmak, S. (2018). Climate-Driven Crop Yield and Yield Variability and Climate
- 538 Change Impacts on the U.S. Great Plains Agricultural Production. *Scientific Reports*, 8(1), 3450.
- 539 https://doi.org/10.1038/s41598-018-21848-2
- Lachowiec, J., Queitsch, C., & Kliebenstein, D. J. (2015). Molecular mechanisms governing
- differential robustness of development and environmental responses in plants. *Annals of Botany*,
 117(5), 795–809. https://doi.org/10.1093/aob/mcv151
- Lanning, S. P., Kephart, K., Carlson, G. R., Eckhoff, J. E., Stougaard, R. N., Wichman, D. M.,
- Martin, J. M., & Talbert, L. E. (2010). Climatic Change and Agronomic Performance of Hard
- 545 Red Spring Wheat from 1950 to 2007. Crop Science, 50(3), 835–841.
- 546 https://doi.org/10.2135/cropsci2009.06.0314
- 547 Lenssen, A. W., Johnson, G. D., & Carlson, G. R. (2007). Cropping sequence and tillage system
- influences annual crop production and water use in semiarid Montana, USA. *Field Crops Research*, 100(1), 32–43. https://doi.org/10.1016/j.fcr.2006.05.004
- Lobell, D. B., Burke, M. B., Tebaldi, C., Mastrandrea, M. D., Falcon, W. P., & Naylor, R. L.
- 551 (2008). Prioritizing climate change adaptation needs for food security in 2030. *Science*
- 552 *319*(5863), 607–610. https://doi.org/10.1126/science.1152339
- 553 Lu, P., Jiang, B., & Weiner, J. (2020). Chapter Three—Crop spatial uniformity, yield and weed
- suppression. In D. L. Sparks (Ed.), *Advances in Agronomy* (Vol. 161, pp. 117–178). Academic
- 555 Press. https://doi.org/10.1016/bs.agron.2019.12.003
- 556 Macholdt, J., Hadasch, S., Piepho, H.-P., Reckling, M., Taghizadeh-Toosi, A., & Christensen, B.
- 557 T. (2021). Yield variability trends of winter wheat and spring barley grown during 1932–2019 in
- the Askov Long-term Experiment. *Field Crops Research*, *264*, 108083.
- 559 https://doi.org/10.1016/j.fcr.2021.108083
- 560 Makumburage, G. B., & Stapleton, A. E. (2011). Phenotype uniformity in combined-stress
- environments has a different genetic architecture than in single-stress treatments. *Frontiers in Plant Science*, 2, 12. https://doi.org/10.3389/fpls.2011.00012
- 563 McGrath, J. M., & Lobell, D. B. (2013). Regional disparities in the CO 2 fertilization effect and
- implications for crop yields. *Environmental Research Letters*, 8(1), 014054.
- 565 https://doi.org/10.1088/1748-9326/8/1/014054
- 566 National Agricultural Statistics Service. (2021). U.S. National Agricultural Statistics Service
- 567 NASS. [Web Archive].
- 568 https://www.nass.usda.gov/Quick_Stats/Ag_Overview/stateOverview.php?state=MONTANA

- 569 Ordas, B., Malvar, R. A., & Hill, W. G. (2008). Genetic variation and quantitative trait loci
- associated with developmental stability and the environmental correlation between traits in
- 571 maize. Genet Res (Camb), 90(5), 385–395. https://doi.org/10.1017/S0016672308009762
- 572 S0016672308009762 [pii]
- 573 R Development Core Team, R. (2011). R: A Language and Environment for Statistical
- 574 Computing. In R. D. C. Team (Ed.), *R Foundation for Statistical Computing* (Vol. 1, Issue
- 575 2.11.1, p. 409). R Foundation for Statistical Computing. https://doi.org/10.1007/978-3-540-
- 576 74686-7
- 577 Ray, D. K., Gerber, J. S., MacDonald, G. K., & West, P. C. (2015). Climate variation explains a
- third of global crop yield variability. *Nature Communications*, 6(1), 5989.
- 579 https://doi.org/10.1038/ncomms6989
- 580 Reckling, M., Ahrends, H., Chen, T.-W., Eugster, W., Hadasch, S., Knapp, S., Laidig, F.,
- Linstädter, A., Macholdt, J., Piepho, H.-P., Schiffers, K., & Döring, T. F. (2021). Methods of
- 582 yield stability analysis in long-term field experiments. A review. Agronomy for Sustainable
- 583 *Development*, 41(2), 27. https://doi.org/10.1007/s13593-021-00681-4
- 584 Riveland, N. R., Berg, J. E., Kephart, K. D., Wichman, D. M., Carlson, G. R., Kushnak, G. D.,
- 585 Stougaard, R. N., Eckhoff, J. L., Nash, D. L., Johnston, M., Grey, W. E., Jin, Y., Chen, X., &
- 586 Bruckner, P. L. (2011). Registration of 'Decade' Wheat. Journal of Plant Registrations, 5(3),
- 587 345–348. https://doi.org/10.3198/jpr2011.04.0191crc
- 588 Sangster, T., Salathia, N., Undurraga, Milo, R., Schellenberg, K., Lindquist, S. L., & Queitsch,
- 589 C. (2008). HSP90 affects the expression of genetic variation and developmental stability in
- quantitative traits. Proceedings of the National Academy of Sciences of the United States of
- 591 America, 105(8), 2963–2968. https://doi.org/10.1073/pnas.0712200105
- Schou, M. F., Kristensen, T. N., & Hoffmann, A. A. (2020). Patterns of environmental variance
 across environments and traits in domestic cattle. *Evolutionary Applications*, 13(5), 1090–1102.
 https://doi.org/10.1111/eva.12924
- 595 Tollenaar, M., & Wu, J. (1999). Yield Improvement in Temperate Maize is Attributable to
- 596 Greater Stress Tolerance. Crop Science, 39(6), 1597–1604.
- 597 https://doi.org/10.2135/cropsci1999.3961597x
- 598 *U.S. Climate Normals*. (2021, February 25). National Centers for Environmental Information 599 (NCEI). http://www.ncei.noaa.gov/products/land-based-station/us-climate-normals
- Wagner, A. (2007). *Robustness and evolvability in living systems*. Princeton University Press.
 https://press.princeton.edu/titles/8002.html
- Wagner, G. P., Booth, G., & Bagheri-Chaichian, H. (1997). A population genetic theory of
 canalization. *Evolution*, 329–347.
- 604 Whitlock, C., Cross, W., Maxwell, B., Silverman, N., & Wade, A. (2017). 2017 Montana
- 605 Climate Assessment (p. 318 p.). https://doi.org/doi:10.15788/m2ww8w
- Zhang, F., Biederman, J. A., Dannenberg, M. P., Yan, D., Reed, S. C., & Smith, W. K. (2021).
- 607 Five Decades of Observed Daily Precipitation Reveal Longer and More Variable Drought Events
- Across Much of the Western United States. *Geophysical Research Letters*, 48(7),
- 609 e2020GL092293. https://doi.org/10.1029/2020GL092293