Uncertainty decomposition to understand the influence of water systems model error in climate vulnerability assessments

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Abstract

Climate vulnerability assessments rely on water infrastructure system models that imperfectly predict performance metrics under ensembles of future scenarios. There is a benefit to reduced complexity system representations to support these assessments, especially when large ensembles are used to better characterize future uncertainties. An important question is whether the total uncertainty in the output metrics is primarily attributable to the climate ensemble or to the systems model itself. Here we develop a method to address this question by combining time series error models of performance metrics with time-varying Sobol sensitivity analysis. The method is applied to a reduced complexity multi-reservoir systems model of the Sacramento-San Joaquin River Basin in California to demonstrate the decomposition of flood risk and water supply uncertainties under an ensemble of climate change scenarios. The results show that the contribution of systems model error to total uncertainty is small (~5-15%) relative to climate based uncertainties. This indicates that the reduced complexity systems model is sufficiently accurate for use in the context of the vulnerability assessment. We also observe that climate uncertainty is dominated by the choice of GCM and its interactive effects with the representative concentration pathway (RCP), rather than the RCP alone. This observation has implications for how climate vulnerabilities should be interpreted.

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23 Abstract

24 Climate vulnerability assessments rely on water infrastructure system models that imperfectly 25 predict performance metrics under ensembles of future scenarios. There is a benefit to reduced 26 complexity system representations to support these assessments, especially when large ensembles 27 are used to better characterize future uncertainties. An important question is whether the total 28 uncertainty in the output metrics is primarily attributable to the climate ensemble or to the systems 29 model itself. Here we develop a method to address this question by combining time series error 30 models of performance metrics with time-varying Sobol sensitivity analysis. The method is applied 31 to a reduced complexity multi-reservoir systems model of the Sacramento-San Joaquin River 32 Basin in California to demonstrate the decomposition of flood risk and water supply uncertainties 33 under an ensemble of climate change scenarios. The results show that the contribution of systems 34 model error to total uncertainty is small (~5-15%) relative to climate based uncertainties. This 35 indicates that the reduced complexity systems model is sufficiently accurate for use in the context 36 of the vulnerability assessment. We also observe that climate uncertainty is dominated by the 37 choice of GCM and its interactive effects with the representative concentration pathway (RCP), 38 rather than the RCP alone. This observation has implications for how climate vulnerabilities should 39 be interpreted.

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45 **1. Introduction**

46 Climate vulnerability assessments have become a common feature of water resources systems 47 planning studies (Arnell, 2011; Plummer et al., 2012; US Bureau of Reclamation, 2012; Weaver 48 et al., 2013). These assessments generally require ensemble simulations of future climate scenarios 49 that are passed through a combination of hydrologic models and water resources systems models 50 to measure the vulnerability of the water system to properties of future climate. Once these 51 vulnerabilities are identified, additional simulation or optimization experiments are used to 52 determine how well different adaptation actions mitigate these vulnerabilities (Herman et al., 2015; 53 Herman et al., 2020).

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55 The literature on the uncertainties that underlie future climate scenarios (Knutti et al., 2008; 56 Northrop and Chandler, 2014; Lehner et al 2020), associated hydrologic responses (Wilby and 57 Harris, 2006; Steinschneider et al., 2015a; Mendoza et al., 2016; Kundzewicz et al., 2018), and 58 water system vulnerability under climate change (Steinschneider et al., 2015b,c) is extensive. 59 However, water resources systems model uncertainties are usually neglected in these climate 60 impact assessments, presumably under the assumption that they are negligible in comparison to 61 other uncertainties. This assumption is likely valid for systems models underpinned by years to 62 decades of development. However, many of these high-fidelity models are computationally 63 expensive and ill-suited for ensemble experiments required by climate vulnerability and adaptation 64 assessments. More parsimonious system models and emulators of complex systems models have 65 become a popular means of reducing the computational cost of systems model simulation in 66 ensemble experiments (Haasnoot et al., 2014; Gijsbers et al., 2017; Basco-Carrera and Mendoza 67 2017; Voinov et al., 2018; Badham et al. 2019; Helgeson et al., 2021). These models provide faster

68 runtimes at the expense of some accuracy in system representation. Their use raises the question 69 of whether such reduced complexity models are suitable for use in climate vulnerability 70 assessments and how this should be assessed.

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72 Past work has considered the question of whether a systems model is fit-for-purpose (Haasnoot et 73 al. 2014; Hamilton et al., 2022). For example, Haasnoot et al. (2014) emphasize the ability of the 74 model to produce "credible outcomes with sufficient accuracy for the screening and ranking of 75 promising actions and pathways in order to support... strategic adaptive planning decisions." They 76 describe a simplified systems model as fit-for-purpose if it produces decisions that are consistent with a more complex model. In the context of climate vulnerability assessments, we investigate a 77 78 related concept: whether prediction errors arising from the systems model are negligible compared 79 to the uncertainty in forcing, particularly around key output metrics that are most relevant to 80 decision-making. This emphasizes the relative accuracy of the systems model against the 81 background of other exogenous uncertainties, and thus contributes a complementary viewpoint by 82 assessing if the model is fit-for-purpose in the context of planning under uncertain future 83 conditions. This viewpoint is consistent with recent recommendations to ensure greater 84 transparency and more robustness in climate change impact assessments (Wagener, 2022).

85

This technical note advances variance decomposition as an approach to assess the suitability of water systems models in climate vulnerability studies. Uncertainty decomposition is widely used to assess sources of uncertainty in climate models and their influence on key variables of interest (Hawkins and Sutton 2009, Lehner et al 2020). It is also used to identify factors that drive uncertainty in water systems performance metrics (Schlef et al. 2018; Greve et al 2018) and the 91 broader human-Earth system (Lamontagne et al., 2019). In this study we extend this technique to 92 assess whether a water systems model is sufficiently accurate for its intended purpose in a climate 93 vulnerability assessment, providing a diagnostic method to propagate and decompose systems 94 model error in the context of an ensemble of climate scenarios.

95

96 **2. Data and Methods**

97 2.1 Case Study & Simulation Model

The proposed method is demonstrated using a new, daily time step simulation model of the eight largest reservoirs in the Sacramento-San Joaquin River Basin (SSJRB), California, and the water supply pumping operations near the system outlet in the Sacramento-San Joaquin Delta (Figure 1). The model structure and data requirements were simplified from recently developed simulation models of this system (ORCA - Cohen et al. 2020, 2021; CALFEWS - Zeff et al. 2021) for the purpose of efficiently estimating water supply and flooding metrics in the Delta in large ensemble climate vulnerability assessments.

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Figure 1. Overview of the Sacramento-San Joaquin River Basin (SSJRB) simulation model.
(a) Reservoir operating policy relating releases *R* to storage *S*; (b) inflow and pumping
locations; (c) model accuracy (R²) for storage at the 8 reservoirs in the system and total delta
exports compared to the CALFEWS systems model (Zeff et al., 2021). The time periods for
the comparison are Oct 1997-Sept 2021 (SSJRB) and Oct 1996-Sept 2016 (CALFEWS). *The
Oroville result for the SSJRB model is impacted by the operational response to the Feb 2017
spillway failure over the subsequent year.

114

115 The SSJRB reduced complexity systems model (referenced hereafter as the SSJRB model) consists 116 of three components: reservoir release policies, gains, and Delta pumping. The model contains 43 117 parameters, including: five release policy parameters for each of the eight reservoirs, two 118 parameters for gains, and one parameter for Delta pumping. The model uses the historical observed 119 median operating pattern (storage and release for each day of the water year) over the period 1997-120 2021, and adjusts this pattern based on current hydrologic conditions. Initially, reservoir operations are described by 5-parameter $(x_0, \dots x_4)$ exponential water supply and linear flood hedging rules 121 122 (Figure 1a). The water supply rule is given by:

123
$$\frac{R_i(t)}{R_{i,m}(t)} = \left(\frac{S_i(t)}{S_{i,m}(t)}\right)^{x_0} \tag{1}$$

where $R_i(t)$ is the release for the *i*th reservoir, $S_i(t)$ is the storage, and $R_{i,m}(t)$ and $S_{i,m}(t)$ are the median release and storage for that day of the water year, respectively. The water supply release determined from Eq. 1 is then increased to model flood control operations. Specifically, if the day of the water year falls between $[x_1, x_2]$ and $S_i(t) > x_4 S_{i,m}(t)$, then $R_i(t)$ is increased by the amount $x_3(S_i(t) - x_4 S_{i,m}(t))$.

129

Next, the hydrologic gains into the Delta, G(t), are defined as the Delta inflow $D_{in}(t)$ minus the sum of reservoir outflows. This term represents the tributaries for which reservoirs are not modeled (see Figure 1b) as well as additional inflows downstream of the reservoirs. Gains can be either positive or negative. Positive gains represent winter inflows, while negative gains represent consumptive withdrawals in the summer. These gains are estimated from historical patterns using two parameters (x_5, x_6) :

136
$$G(t) = G_m(t) \left(\sum_i \frac{S_i(t)}{K_i} \right)^{x_5} + x_6 \sum_i Q_i(t)$$
(2)

where $G_m(t)$ are the median gains for that day of the water year, K_i is the storage capacity of reservoir $i \in [1,8]$, and $Q_i(t)$ is the inflow into reservoir *i*. The first term in (2) covers the broader seasonal patterns of withdrawals, while the second term represents additional Delta inflows that are assumed to be correlated with reservoir inflows included in the model.

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142 The Delta pumping policy is represented by the following equation with one parameter, x_7 :

143
$$P(t) = D_{in}(t) p_m(t) \left(\sum_i \frac{s_i(t)}{\kappa_i}\right)^{x_7}$$
(3)

144 where P(t) is the total pumping volume, $p_m(t)$ is the median pumping for that day of the water 145 year (percent of inflow), and the storage fraction term is the same as in Eq. 2. We impose an upper bound on pumping to approximate a combination of the infrastructure capacity and environmentalguidelines, though this amount can be exceeded if needed.

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149 The parameters for all of the model components are fit with differential evolution (Storn and Price, 150 1997). The historical data used to find the median daily patterns and to fit the parameters are taken 151 from the California Data Exchange Center (CDEC; cdec.water.ca.gov). These operating rules are 152 empirical simplifications based on the observed data, and do not exactly match those published in 153 water control manuals. However, they ensure that all reservoirs follow the same model structure, 154 and that the model is parsimonious enough to calibrate and modify. Figure 1c shows the ability of 155 the systems model to replicate historical storage for each of the eight reservoirs. Overall, the 156 systems model adequately represents the operations of these facilities, with R² values for daily 157 simulated and observed storage ranging between 0.66 and 0.91, with an average of 0.81. This 158 performance is slightly worse than that of a recently published, more detailed, state-of-the-art 159 simulation model of the California water system (CALFEWS; Zeff et al., 2021), which has an 160 average storage R² of 0.88 for the same reservoirs, and also contains a more detailed system 161 representation south of the Delta to describe deliveries to irrigation districts. However, the SSJRB 162 model is significantly faster (3-4 orders of magnitude) due to a combination of simplified structure 163 and Numba just-in-time compilation (Lam et al., 2015). In the analysis that follows, we investigate 164 whether the decrease in accuracy significantly influences our understanding of system 165 performance in the context of broader climate uncertainties.

166

167 **2.2 Error Models for Key System Model Outputs**

168 Two output variables from the SSJRB systems model are of interest in this study: a water supply 169 metric - delta pumping exports, P(t) - and a flood control metric - delta outflows, $D_{out}(t) =$ 170 $D_{in}(t) - P(t)$. We develop error models for these two metrics using the historical simulation from 171 the SSJRB systems model, which enables stochastic simulation of these errors under a wide range 172 of future scenarios at a daily time step.

173

We follow the general approach in McInerney et al. (2017) and define a residual term, ε_t , equal to the difference between the daily observations and simulations after transformation:

 $\varepsilon_t = f(O_t|\lambda) - f(M_t|\lambda)$

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177 Here, O_t is the observed data associated with the decision-relevant variable of interest (e.g., delta 178 outflows or delta pumping exports), M_t is the systems model simulation of that variable, t is a time 179 step (daily in this case) within the historical record of length T, and $f(\cdot | \lambda)$ is a transformation with 180 parameter λ . The transformation is used to simplify the probabilistic behavior of the observed and 181 simulated time series before calculating the residuals. Here, we employ the Box-Cox 182 transformation (Box and Cox, 1964), which becomes the identity transformation for $\lambda = 1$ and 183 approaches a logarithmic transformation as λ approaches 0. However, other transformations (e.g., 184 log-sinh) could also be applied.

185

186 To remove any systematic bias between the simulations and observations, the residuals are 187 regressed against the transformed simulation:

188

$$\varepsilon_t = \beta_0 + \beta_1 f(M_t | \lambda) + \epsilon_t \tag{5}$$

(4)

189 Two assumptions are made: 1) systems model bias is correlated with the magnitude of the 190 simulated response, e.g., the systems model tends to underestimate the observations when it 191 predicts large flows and overestimate the observations when it predicts low flows; 2) this bias is a 192 linear function of the magnitude of the simulation itself. The bias correction in Eq. 5 could be 193 made more general using a non-linear function (e.g., a local weighted regression or generalized 194 additive model), but initial analysis (not shown) suggested this was unnecessary for the case study 195 used in this work.

196

197 The bias corrected residuals ϵ_t are then decorrelated in time using an autoregressive (AR) model: 198

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$$\epsilon_t = \theta_0 + \theta_1 \epsilon_{t-1} + \xi_t \tag{6}$$

An AR(1) model was found to be sufficient to remove autocorrelation in the delta outflow and export residual time series. However, any higher-order autoregressive moving average (ARMA) model can be selected based on the behavior of the residual series ϵ_t .

203

Once the model above is fit to the historical series of data, stochastic traces $\widetilde{O_{1:T^*}}$ are simulated for new periods of interest ($t^* = 1, ..., T^*$, e.g., future decades under climate change). These simulations are produced with the following steps:

207 1) Bootstrap a value of ξ_t from the historical record.

208 2) For a new time t^* , estimate $\tilde{\epsilon_{t^*}}$ using Eq. 6, the resampled value of ξ_t , and the previous 209 value $\tilde{\epsilon_{t^*-1}}$.

210 3) Estimate
$$\tilde{\varepsilon_{t^*}}$$
 using Eq. 5, the value $\tilde{\epsilon_{t^*}}$, and the systems model simulation $\tilde{M_{t^*}}$.

211 4) Estimate
$$\widetilde{O_{t^*}} = f(\widetilde{\varepsilon_{t^*}} + f(\widetilde{M_{t^*}}|\lambda)|\lambda)^{-1}$$

212 Steps 1-4 are then repeated for all time steps $t^* = 1, ..., T^*$. $\tilde{\epsilon_0}$ is initialized as 0.

213

214 **2.3 Climate Scenarios**

215 In this study, we assess whether the SSJRB systems model error in key variables of interest (e.g., 216 delta outflows and exports) is sufficiently small in the broader context of climate uncertainty. This 217 is tested by forcing the systems model with an ensemble of projected flows between 2020-2099 218 for each of the 8 reservoir inflow points of the system (see Figure 1). The ensemble, developed in 219 Brekke et al. (2014), is derived from CMIP5 general circulation model (GCM) simulations (Taylor 220 et al., 2012), downscaled to a daily timescale and 1/8° spatial resolution using the updated Bias-221 Correction and Spatial Disaggregation (BCSD) technique (NCAR, 2014) followed by daily 222 disaggregation (Wood et al., 2004). The ensemble is composed of four different representative 223 concentration pathways (RCPs; 2.6, 4.5, 6.0, 8.5) and 31 different GCMs, with 97 scenarios 224 altogether (not every RCP is used with every GCM). The downscaled, daily climate data force the 225 Variable Infiltration Capacity (VIC) hydrologic model, previously calibrated for watersheds across 226 the US West (see Brekke et al. 2014).

227

228 To generate a balanced ensemble, we filter the full ensemble described above to include only those 229 GCMs with simulations under the RCP 4.5 and 8.5 emission scenarios (i.e., the most common 230 emission scenarios across all GCMs). This leads to 29 GCMs under these two RCPs, for a total of 231 58 scenarios. For each scenario and its associated trace of simulated daily delta outflows and delta 232 exports from the SSJRB systems model, we develop a stochastic ensemble of 100 traces of delta 233 outflow and exports using the error modeling procedure in Section 2.2. This results in a total of 234 5,800 80-year traces of daily delta outflows and exports, which are then analyzed using sensitivity 235 analysis (described next) to partition variance among the various uncertainty sources: GCMs, 236 RCPs, and systems model error. We also consider whether the variance partitioning changes

considerably if a smaller set of GCMs is used, for instance based on their ability to capture aspects
of regional climate (Gershunov et al., 2017; Pierce et al., 2018).

239

240 2.4 Sensitivity Analysis

We use Sobol sensitivity analysis to attribute variance within the ensemble of 5,800 traces of delta outflows and exports to the GCMs, the RCPs, systems model error, and interactive effects between these different sources of uncertainty. These three inputs are sampled as integer factors in the sensitivity analysis (i.e., the choice of GCM, RCP, and systems model error realization). Ultimately, the model is classified as fit-for-purpose for large ensemble experiments in a climate vulnerability assessment if the variance attributed to the systems model uncertainty and its interactive effects is small compared to the other sources of uncertainty.

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249 Sobol sensitivity analysis is described in detail elsewhere (Sobol, 2001; Saltelli et al., 2010; Pianosi 250 et al., 2016; Herman and Usher, 2017) and therefore only briefly reviewed here. Let Y_k be the 251 output metric of interest from the SSJRB systems model, which we define separately for each year of simulation (k=2020, ..., 2100). For delta outflows, we define Y_k as the annual maximum outflow 252 253 in year k; for delta exports, Y_k is defined as the annual sum of exports. While other metrics could 254 be chosen, these metrics are representative of annual flood and drought risk in the Sacramento-255 San Joaquin system. The Sobol method is used to attribute variance in Y_k to individual uncertainty 256 factors and their interactions, and can be written as follows:

257

$$D(Y_k) = \sum_i D_i + \sum_{i < j} D_{ij} + D_{12\dots h}$$
(7)

Here, $D(Y_k)$ represents the total variance in Y_k , D_i is the first-order variance contribution of the ith factor, D_{ij} is the second-order variance contribution of the interaction between the ith and jth factors, and $D_{12...h}$ represents the variance contribution of all higher-order interactions greater than secondorder. Sensitivity indices are then defined as the fraction of individual variance contribution terms to the total variance (e.g., $\frac{D_i}{D}$ and $\frac{D_{ij}}{D}$ represent the first-order sensitivity index for factor *i* and second-order sensitivity index for factors *i* and *j*, respectively). Similarly, the total-order sensitivity for a given factor $(1 - \frac{D_{\sim i}}{D})$ uses the variance associated with all factors besides factor *i* $(D_{\sim i})$ to define the variance attributed to all first-order and higher-order interactions associated with factor *i*. See Saltelli (2002) for additional detail on the numerical estimation of terms $(D_i, D_{ij}, D_{\sim i})$.

267

When making a determination of whether the systems model is sufficiently accurate for use in a climate impact analysis, we avoid setting a distinct threshold for the variance attributed to the systems model. This is ultimately a subjective choice based on user preference, and we believe that forwarding an (arbitrary) threshold here would discourage critical evaluation and collaborative decision-making on a case-by-case basis that is central to effective water resources planning. Similar logic supports recent efforts in the statistical literature to discourage the use of arbitrary pvalues when assessing the statistical significance of relationships (Wasserstein and Lazar, 2016).

275

3. Results and Discussion

277 **3.1 Systems Error Models**

Error accumulates throughout the system to influence the key metrics of interest: delta outflows and exports. Figures 2a and 2d show the observed and simulated time series of daily delta outflows and exports, respectively. The systems model captures the observed outflows well, with a Nash-Sutcliffe efficiency (NSE) over the entire period of 0.74 and a percent bias of 3.2%. Simulations of daily exports are less skillful, with an NSE of 0.32 and percent bias of 12.1%, although when aggregating to a weekly time step the simulations are better (NSE of 0.63). Notably, the systems
model tends to underestimate the largest delta outflows, albeit with one exception in 2017. The
SSJRB systems model imposes a cap on delta exports, but observed exports occasionally exceed
this limit. The systems model also overestimates the smallest delta outflows (Figure 2a),
particularly during dry years, and exhibits fewer and less extreme declines in daily exports (Figure 28).



Figure 2. a) Observed (blue) and simulated (red) daily delta outflows. The grey area shows 95% bounds of the stochastic delta outflow simulations. b) Target versus actual coverage probabilities (i.e., the probability that observations fall within the *p* percent bounds of the simulated ensemble), with *p* varied from 1% to 99%. c) The CDF of observed (blue) and simulated (red) annual maximum delta outflows, as well as the 95% bounds for the CDF of annual maxima from the stochastic traces. d-e) Same as (a-b) but for daily delta exports. f) Same as (c) but for the annual sum of delta exports.

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298 We fit error models for both the delta outflow and export series in order to generate stochastic

299 traces of these variables. A Box-Cox transformation with λ =0.3 is applied to the delta outflows,

while the delta exports are not transformed (i.e., $\lambda=1$). Also, delta outflow residuals ε_t did not vary 300 significantly with model simulations, so $\beta_1 = 0$ in Eq. 5. Figures 2a and 2d show 95% bounds 301 302 generated using the stochastic ensemble. These uncertainty bounds are developed by simulating 303 delta outflow and export residuals and adding them to the SSJRB simulation. Several important 304 features emerge from the ensemble. The delta outflows ensemble clearly captures some of the 305 lowest outflow values, but also captures several of the peak outflow events. Similarly, the 306 stochastic delta exports ensemble better captures many of the lowest observed exports and 307 observed exports that extend above the cap imposed in the SSJRB model.

308

This is clear in Figures 2b and 2e, which show coverage probabilities for the stochastic ensemble. Coverage probabilities quantify how often the observations fall between the $\frac{\alpha}{2}$ and $\left(1-\frac{\alpha}{2}\right)$ percentiles of the stochastic ensemble, with α ranging from 0.01 to 0.99 in Figures 2b,e. The target coverage probabilities (i.e., $\alpha \times 100\%$) are compared against the observed coverage probabilities to assess the reliability of the stochastic ensemble. The results show that the ensemble is very reliable, with the largest deviations between target and observed coverage probabilities only reaching $\sim 5.5\%$.

316

The assessments above focus on daily data, but we are often interested in specific aggregated metrics of model output. Figures 2c and 2f show the distribution of observed and simulated annual maximum delta outflows and annual total delta exports, respectively. These are the focus of the uncertainty decomposition shown in Section 3.2 below. Also shown in Figure 2c,f are the same distributions from the stochastic traces. For delta outflows, the simulated distribution tends to underestimate that of the observed, except for the very largest annual maxima. However, the 323 stochastic ensemble captures the observed distribution reasonably well, albeit with an 324 underestimation between the 40th and 75th percentiles of the distribution. The opposite issue is 325 apparent in the delta exports data, as the SSJRB model tends to overestimate the observed annual 326 totals. The stochastic ensemble corrects these discrepancies and adequately captures the observed 327 distribution.

328

329 3.2 Sensitivity Analysis

330 The SSJRB model was used to generate a 5,800-member ensemble of future simulations across 2 331 RCPs, 29 GCMs, and with 100 stochastic traces of model error (as described above). The Sobol 332 analysis resamples from these discrete factors to decompose the variance in the two output metrics 333 across the entire ensemble for each year, as shown in Figures 3a and 3b. Here we show the first-334 order sensitivity indices for RCP, GCM, and systems model error, as well as second-order 335 sensitivity indices that quantify the variance associated with interactive effects. Higher order 336 sensitivity is not shown, and so the cumulative variance shown in Figure 3 is always slightly less 337 than 100% of the total variance. All sensitivity indices are smoothed with a 10-year rolling average.

338

Several insights emerge from Figures 3a,b. First, the largest source of uncertainty in both metrics stems from the first-order effects of the GCMs. Averaged across all years, first-order GCM uncertainty accounts for 49% and 46% of the total variance in delta outflows and exports, respectively. This result demonstrates that the GCM used to simulate system inflows has the single largest impact on the variance in decision-relevant metrics of interest, regardless of emission scenario or other factors. We speculate that climate model uncertainty dominates the total uncertainty due to the range of regional precipitation responses (from both internal variability and
change) across the GCMs (Schlef et al., 2018), but we do not verify that here.

347

348 The second largest source of uncertainty stems from the interactive effect of GCMs with RCPs. 349 The magnitude of this uncertainty rivals that of the first-order GCM effect, and when averaged 350 across years, accounts for 41% and 35% of the variance in delta outflows and exports, respectively. 351 In contrast, the first-order effect of the RCPs is small, on average less than 2%. This result 352 highlights that neither annual maximum delta outflow nor annual total delta exports are explained 353 by consistent differences that emerge across emission scenarios. This is true even by the end of 354 the century when temperature differences between RCP 4.5 and RCP 8.5 are greatest. Instead, it 355 is the variable response of the GCMs to a given RCP that explains a large portion of variance in 356 the delta outflows and exports. That is, if one GCM projects drier conditions as emissions rise, but 357 another projects wetter conditions with greater emissions, only the combination of both the 358 emissions scenario *and* the specific GCM can explain the response of the water system.

359

360 Finally, we consider the variance explained by factors related to the systems model uncertainty. 361 Figures 3a,b show that for both delta outflows and delta pumping, uncertainties related to the 362 system model are small in comparison to the other sources. The uncertainty attributed to the first-363 order systems model effect is under 1% across time, as is the interactive effect between the system 364 model error and RCP. The interactive effect between systems model error and GCM is larger, with 365 time-averaged values of 4% for delta outflows and 7.9% for delta exports. However, the 366 cumulative uncertainty for all first and higher-order terms related to systems model error (i.e., 367 total-order sensitivity, not shown in Figure 3) is 8.2% for delta outflows and 16.5% for delta 368 exports, far below the uncertainties related to the GCMs and their interactive effects with emission 369 scenarios. From this perspective, we deem the systems model sufficiently accurate to support 370 climate change vulnerability studies of the system, given the set of 29 GCMs considered in the 371 ensemble.

372

373 It is reasonable to question if this result is consistent when a smaller set of GCMs with less 374 variability and a more consistent representation of the climate system is selected for analysis. To 375 answer this question, Figures 3c,d show the same results as in Figures 3a,b but based on a subset 376 of the original ensemble composed of simulations from the two RCPs and only four GCMs 377 (access 1-0, canesm2, cnrm-cm5, gfdl-cm3). These GCMs have been shown to capture atmospheric 378 river dynamics along the US West Coast that are important to water supply and flood risk in the 379 SSJRB system (Gershunov et al., 2017). Note that access1-3 is also highlighted as skillful in 380 Gershunov et al., 2017 but is not available in the ensemble of VIC simulations (see Section 2.3). 381

382 In the restricted set of GCMs, the uncertainty attributable to the first-order effect from the GCMs 383 declines, though it is still substantial. A similar result is seen for the interactive effect between the 384 GCMs and RCPs. Importantly, we only observe a small increase in the variance attributed to the 385 systems model. Instead, the variance contribution associated with the first-order effect of the RCPs 386 increases the most, from less than 2% in the full ensemble to 11% and 10% in the limited ensemble 387 for delta outflows and exports, respectively. This is consistent with the expectation that the effects 388 of emission scenario on hydrologic response are more apparent in GCMs which provide a more 389 consistent representation of regional climate. Still, the main result from the full ensemble persists: 390 systems model uncertainty is small compared to other uncertainties, therefore the reduced 391 complexity model is suitable for climate vulnerability analysis. We note that we repeated this same 392 experiment with 4 different GCMs that were shown to produce realistic simulations of 393 precipitation and temperature in California (Pierce et al., 2018), with very similar results (not 394 shown).

395





Figure 3. Variance contribution of GCMs, RCPs, and system model uncertainty to the total
variance of (a,c) annual maximum delta outflows and (b,d) annual total delta exports, shown
by year and for first-order effects and second-order interactions. Results are smoothed using
a lagged 10-year rolling average and shown when using (a,b) all GCMs and (c,d) only four

- 401 GCMs selected for their accurate representation of atmospheric rivers (ARs).
- 402

403 4. Conclusions

404 This technical note contributes a novel approach to determine whether a water resources systems

405 model is sufficiently accurate for ensemble experiments in climate vulnerability assessments. The

406 approach decomposes the variance in decision-relevant metrics (delta outflows and exports) to 407 compare systems model error to other uncertainty sources (e.g., future climate scenarios). Model 408 error in the output metrics is represented by a time series error model. To our knowledge, this is 409 the first time error models and variance decomposition have been used to assess model suitability 410 for climate impact experiments.

411

412 This technique is demonstrated with a new, computationally efficient, daily systems model of 413 reservoirs within the Sacramento – San Joaquin River Basin California. The analysis shows that 414 uncertainties from the systems model are small in comparison to those associated with future 415 climate scenarios, especially in the key metric of annual maximum delta outflow. Systems model 416 uncertainty contributes more to the total variance in delta exports, although the future climate 417 scenarios continue to dominate variance in the delta export metric. GCMs and their interactive 418 effects with RCPs contribute the most uncertainty in future delta outflows and exports. In 419 comparison, the first-order contributions from the RCPs themselves are small. The conclusions 420 above hold even if smaller ensembles of GCMs are selected based on their ability to represent 421 regional climate, although with a somewhat larger direct contribution from emission scenario.

422

The approach presented here is widely applicable and can be extended to consider other uncertainty sources not accounted for in this experiment. For instance, the effects of climate model error and natural climate variability, which were combined in the experiment presented here, could be separated using single-model initial condition large ensembles (Lehner et al., 2020). Similarly, hydrologic model uncertainty can significantly affect water resource impact assessments (Malek et al., 2022) and could be accounted for by using multiple model structures and behavioral 429 parameter sets. As the ensemble size grows, the computational costs of the approach (in particular 430 the Sobol method) may become infeasible. Therefore, future work is needed to consider alternative 431 methods for sensitivity analysis that are more efficient, such as ANOVA or the method of Morris 432 (Morris 1991; Herman et al., 2013). Further consideration of the tradeoff between model accuracy 433 and parsimony is also warranted.

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435 A few caveats of the approach deserve mention. The systems model uncertainty is quantified based 436 on historical errors between observed and modeled outcomes, but this approach will not capture 437 any nonstationarity in the error distribution that emerges under new boundary conditions. 438 Similarly, the current approach does not capture structural uncertainties related to changes in the 439 system, e.g., infrastructure adaptation or land-use change in response to exogenous forcing. 440 Despite these caveats, the combined use of time series error models and sensitivity analysis 441 provides an effective and straightforward screening method to assess the suitability of a systems 442 model in large ensemble climate vulnerability assessments.

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451 Data and Software Availability

452 The code, inputs, and final outputs for all analyses conducted in this study are available at

453 https://github.com/JohnRushKucharski/systemuncertainty and

- 454 https://zenodo.org/record/6331306#.Yin8uBDMLOQ.
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