Challenges in scaling up greenhouse gas fluxes: experience from the UK Greenhouse Gas Emissions and Feedbacks Programme

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Abstract

The role of greenhouse gases (GHGs) in global climate change is now well recognised and there is a clear need to measure emissions and verify the efficacy of mitigation measures. To this end, reliable estimates are needed of the GHG balance at national scale and over long time periods, but these estimates are difficult to make accurately.

Because measurement techniques are generally restricted to relatively small spatial and temporal scales, there is a fundamental problem in translating these into long-term estimates on a regional scale.

The key challenge lies in spatial and temporal upscaling of short-term, point observations to estimate large-scale annual totals, and quantifying the uncertainty associated with this upscaling.

Here, we review some approaches to this problem, and synthesise the work in the recent UK Greenhouse Gas Emissions and Feedbacks Programme, which was designed to identify and address these challenges.

Approaches to the scaling problem included:

instrumentation developments which mean that near-continuous data sets can be produced with larger spatial coverage;

geostatistical methods which address the problem of extrapolating to larger domains, using spatial information in the data; more rigorous statistical methods which characterise the uncertainty in extrapolating to longer time scales;

analytical approaches to estimating model aggregation error; enhanced estimates of C flux measurement error;

and novel uses of remote sensing data to calibrate process models for generating probabilistic regional C flux estimates.

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Key Points:

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10	•	We reviewed some of the challenges in accurately estimating fluxes of GHGs at
11		national scale.
12	•	Uncertainty arises from imperfectly-known models, parameters and inputs used
13		in the extrapolation.
14	•	Bayesian principles allow us to quantify this whilst combining information sources
15		in a coherent way.

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16 Abstract

The role of greenhouse gases (GHGs) in global climate change is now well recognised and 17 there is a clear need to measure emissions and verify the efficacy of mitigation measures. 18 To this end, reliable estimates are needed of the GHG balance at national scale and over 19 long time periods, but these estimates are difficult to make accurately. Because measure-20 ment techniques are generally restricted to relatively small spatial and temporal scales, 21 there is a fundamental problem in translating these into long-term estimates on a regional 22 scale. The key challenge lies in spatial and temporal upscaling of short-term, point ob-23 servations to estimate large-scale annual totals, and quantifying the uncertainty asso-24 ciated with this upscaling. Here, we review some approaches to this problem, and syn-25 thesise the work in the recent UK Greenhouse Gas Emissions and Feedbacks Programme, 26 which was designed to identify and address these challenges. 27

Approaches to the scaling problem included: instrumentation developments which 28 mean that near-continuous data sets can be produced with larger spatial coverage; geo-29 statistical methods which address the problem of extrapolating to larger domains, us-30 ing spatial information in the data; more rigorous statistical methods which characterise 31 the uncertainty in extrapolating to longer time scales; analytical approaches to estimat-32 ing model aggregation error; enhanced estimates of C flux measurement error; and novel 33 uses of remote sensing data to calibrate process models for generating probabilistic re-34 gional C flux estimates. 35

³⁶ Plain Language Summary

Greenhouse gases cause climate change, and we need to know how much is emit-37 ted each year across the globe. As well as coming from burning fossil fuels, plants and 38 soil also take up and emit these gases, and we need to be able to quantify this in order 39 to understand how best to tackle climate change. However, we can only measure these 40 emissions over very small areas, at only a few locations, and for relatively short periods 41 of time. Extrapolating from these measurements to a whole country introduces several 42 uncertainties which are often largely ignored. Here, we examine progress in tackling this 43 problem, and focus on better statistical methods to properly identify and account for the 44 errors that are introduced by the large change in scale. Another is the development of 45 instrumentation which can measure the gas emissions over larger scales and run contin-46 uously. Earth observation from satellites provides a promising source of data for the fu-47 ture, but cannot yet provide direct measurements of gas emissions. The Bayesian approach 48 to modelling provides us with a coherent method for combining data from different sources, 49 accounting for their uncertainties, and propagating this through to the uncertainties as-50 sociated with predictions of national scale fluxes. 51

52 **1** Introduction

The role of greenhouse gases (GHGs) in causing global climate change is now well 53 recognised (IPCC, 2013). Emissions of GHGs from terrestrial ecosystems play an im-54 portant part in this, and the potential for feedbacks within the climate system which am-55 plify the emissions of GHGs from natural ecosystems is substantial. Accurate estimates 56 are therefore needed of the GHG balance of the land surface at regional and national scales, 57 and over long time periods, if we are to understand the key driver of global change. Be-58 cause of the large scale involved, in relation to the scale at which we can make obser-59 vations, this presents a major challenge which spans the domains of biogeochemistry, ecol-60 ogy, remote sensing, and atmospheric science. 61

For directly measuring GHG fluxes, we have two approaches available, based on either enclosing a small area within a chamber and monitoring the change in GHG concentration, or based on micrometeorological measurements of GHG concentration and

turbulence in the near the surface (see section below). However, both of these operate 65 at scales much smaller than the spatial scale of interest - that of a region, nation or the 66 whole globe. This means that we need to use a model to predict the large-scale flux. The 67 fundamental upscaling issue is that we are forced to rely on predictions from a model 68 which cannot be parameterised or tested at the true scale of interest. If we introduce some 69 generic notation, we can consider this as three inter-related problems¹. We need to pre-70 dict y, the large-scale GHG flux, based on parameters θ derived at a small scale and in-71 put variables x estimated over the large-scale domain: 72

$$y = f(\theta, x). \tag{1}$$

Firstly the parameters θ are only inferred from a very small subset of the condi-73 tions prevailing over the whole domain. Because of basic sampling error, there is uncer-74 tainty in the parameter estimates. Secondly, there is also usually considerable uncertainty 75 in the values of the inputs x over the large-scale domain. GHG fluxes from an ecosys-76 tem depend upon such things as incident radiation, leaf area index, soil aerobic status, 77 soil microbial populations, and time elapsed since disturbance events. None of these is 78 easily measured over a wide region, and we inevitably rely on some proxy or modelled 79 estimate, and this introduces uncertainty in the true value of x over the whole domain. 80 Thirdly, the model f is commonly non-linear, which complicates the upscaling procedure. 81 The goal for science in this field is to quantify and reduce the uncertainty associated with 82 parameters θ , input variables x, and model f in making the jump between the small scale 83 of measurement and the large scale of prediction, which we illustrate in Figure 1. Progress 84 here is necessary if we are to estimate the large-scale GHG balance accurately (Leip et 85 al., 2018), and to demonstrate the efficacy of mitigation policies (Gifford, 1994; Smith 86 & Smith, 2004; Smith et al., 2008). 87

In this paper, we review the challenges in upscaling small-scale GHG flux measure-88 ments to produce national-scale estimates. The UK recently developed a novel, multi-89 disciplinary programme to identify and tackle some of these challenges, and our exam-90 ples come from this programme. The overall aim was to improve the quantification of 91 uncertainty where it arises in the upscaling process, and to reduce this uncertainty by 92 improvements to instrumentation, measurement methods or modelling procedures. Specif-93 ically we focused on five challenges which focus on components of the problem, illustrated 94 in Figure 1: 95

1. Quantifying uncertainty in spatial upscaling of chamber fluxes to field scale. Chamber measurements sample only a very small area, even in relation to a single agricultural field. The challenge is to quantify the mean and uncertainty in the estimate of the field-scale mean flux.

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- 2. Quantifying uncertainty in temporal upscaling of chamber fluxes to annual scale. Similarly, chamber measurements typically sample only during a very few hours, in relation to the total flux over a year. The challenge here is to quantify the annual cumulative emission and its uncertainty, based on a sparse and spatially variable sample set.
- 3. Reducing uncertainty in spatial and temporal upscaling of chamber fluxes via improved instrumentation. An alternative approach to both of the above

¹ In addition to these, new phenomena may arise at the larger scale because of feedbacks in the system, which are not apparent at the small scale. For example, evapotranspiration from an individual leaf is strongly controlled by the stomatal conductance. However, because of the effect of regional-scale evapotranspiration on the vapour pressure deficit of the air in the boundary layer, regional-scale evapotranspiration is more strongly controlled by radiation input. This is a serious issue with water vapour fluxes, but less so for GHG fluxes themselves, because the magnitude of such feedbacks is much smaller.

is to developed new measurement systems which can provide better spatial and temporal coverage.

- 4. Quantifying uncertainty in eddy covariance measurements of field scale
 fluxes. Eddy covariance systems are expensive and complex to run, so are rarely
 operated with any replication. It is therefore usually very difficulty to estimate
 the systematic and random errors associated with these measurements. Here we
 evaluate five co-located eddy flux systems to determine the measurement error on
 net exchanges of CO₂ at both instantaneous times and daily scales.
- 5. Quantifying aggregation error in spatial upscaling. When non-linear models are parameterised at a small scale, but applied at a larger scale, the results will generally be in error wherever small-scale heterogeneity is not accounted for. A further challenge is to estimate and account for this kind of error. We evaluated this in the context of national-scale GHG flux estimates in the UK.

After outlining the basic approach to measuring GHG fluxes, we address each of these challenges in turn, with reference to specific analyses for a range of GHGs. For clarity, each section provides its own methods, results and discussion. We then conclude with a synthesis of the findings from the individual studies. Our synthesis allows us to assess advantages and limitations of current research, and to make suggestions for the development of new studies and approaches necessary to make better inferences about GHG fluxes at regional, national and global scales.

¹²⁷ 2 Measurement Methods for GHG fluxes

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For directly observing GHG fluxes, we have two broad techniques available: chamberbased and micrometeorological. In the former, part of the plant or soil surface is enclosed in a gas-tight chamber, and the flux is inferred from measurements of the mixing ratio. In the case of static (non-steady-state) chambers the mixing ratio is measured on a sequence of gas samples extracted from the chamber over a short time period. From mass balance, the mixing ratio within the chamber is predicted to follow:

$$\chi = \chi_0 + \frac{F}{h\rho} \mathrm{d}t \tag{2}$$

where χ_0 is the initial mixing ratio of a GHG, *h* is the height of the chamber, ρ is the molar density of dry air, and dt is the time increment since enclosure. We thus have an inverse problem, which can be rearranged to estimate the flux as:

$$F = \frac{\mathrm{d}\chi}{\mathrm{d}t_0} h\rho \tag{3}$$

where $d\chi/dt_0$ is the initial rate of change in the mixing ratio. As an approximation, we 137 can assume linearity in $d\chi/dt$, and solve for $d\chi/dt_0$ using linear regression. If we account 138 for the non-linearity of diffusion into the chamber, we have to apply non-linear regres-139 sion, optimisation methods, and potentially, complex 2-D diffusion models (Livingston 140 et al., 2006; Pedersen et al., 2010; Sahoo & Mayya, 2010; Levy et al., 2011). Because part 141 of the ecosystem has to be physically enclosed, the spatial scale of these measurements 142 is necessarily restricted, typically to 0.1 m^2 and rarely more than 1 m^2 . Similarly, the 143 temporal scale of measurements is restricted because the physical enclosure changes the 144 environment within - the emitted gas concentrations build up, and the effect of wind and 145 rain is removed. 146

¹⁴⁷ Micrometeorological techniques make use of measurements in the atmosphere near ¹⁴⁸ the surface. Historically, these were based on measuring the gradients in wind and GHG ¹⁴⁹ mixing ratios and making some assumptions about the turbulent transport. With the ¹⁵⁰ advent of fast-response infra-red analysers for CO_2 , and more recently for CH_4 and N_2O ¹⁵¹ based on QCL or CRD laser absorption spectroscopy, the eddy covariance method has ¹⁵² become the default approach (Kroon et al., 2010; Mammarella et al., 2010; Haszpra et ¹⁵³ al., 2018). If we can assume stationarity and horizontal homogeneity, it follows from mass

balance that we can equate the surface flux to the eddy covariance term i.e.:

$$F = \overline{w'\chi'}\overline{\rho} \tag{4}$$

where w' and χ' represent the instantaneous deviations from the means. To measure this 155 term accurately, we need high frequency (10-20 Hz) measurements of the vertical wind-156 speed w and χ at an appropriate height above the surface. Various corrections are re-157 quired to account for the frequency response of the measurement system, non-zero ver-158 tical windspeed, deviations from stationarity, and density fluctuations (Lee et al., 2006; 159 Aubinet et al., 2012). The advantage of the approach is that it measures the integrated 160 surface flux over an area much larger than a chamber, typically several hundred square 161 metres (of the order of a small agricultural field), and can run near-continuously. 162

¹⁶³ 3 Spatial upscaling of chamber fluxes to field scale

As described above, chambers used to measure gas fluxes typically have small di-164 mensions $(< 1 \text{ m}^2)$, several orders of magnitude smaller than the domains, such as agri-165 cultural fields, that we want to make inferences about, so the potential for sampling er-166 ror is large. That is, the naïve sample mean of the chambers may deviate substantially 167 from the true mean for the field. We want to improve this estimate, and quantify the 168 associated uncertainty in extrapolating the field mean. This is a common problem in the 169 area of geostatistics, where observations are only available at point locations, but pre-170 dictions are required over a larger spatial domain. The classical geostatistical approach 171 to this problem is to represent the spatial domain as a grid of discrete cells (a "raster"), 172 and to use kriging to predict the values at all the unobserved locations in this grid. Krig-173 ing and its terminology originated in the mining industry, but is now a widespread and 174 generally applicable technique for extrapolation problems. In essence, it is a form of weighted 175 local averaging, where the estimates of values at unrecorded places are weighted aver-176 ages of the observations. The kriging weights are calculated on the basis of the semivar-177 iogram, which quantifies the form of the increasing variance between pairs of points as 178 the distance between them increases. Graphically, this shows the scale at which values 179 are highly correlated, and how this changes with spatial scale. Prediction at a new lo-180 cation is based on all the observations, each weighted according to the degree of corre-181 lation at that distance predicted by the semivariogram. Kriging has been shown to be 182 optimal in the sense that it provides estimates with minimum variance and without bias 183 (in the long-term statistical sense). It is also often described as providing estimates of 184 known variance, but this is only true if the form of the semivariogram is known with cer-185 tainty; in real-world applications, this is never the case. Here, we use kriging to extrap-186 olate chamber fluxes to the field scale, but for the purposes of characterising the uncer-187 tainty correctly, we apply it in a Bayesian framework. In brief, this means we account 188 for the uncertainty in the variogram model, and represent each of the parameters as a 189 probability distribution. Rather than assuming the variance is known, we calculate the 190 posterior distribution of the parameters, given the observed data, and sample many re-191 alisations of these to represent the uncertainty. 192

We can attempt to test the success of this upscaling method because we can also 193 measure at the field scale using eddy covariance. However, eddy covariance also does not 194 directly give the field-scale domain mean, as its spatial sampling characteristics are af-195 fected by wind speed, wind direction, sensible heat flux and friction velocity: the so-called 196 "flux footprint" (Schuepp et al., 1990, Schmid and Oke (1990), Leclerc and Foken (2014)). 197 The footprint defines the relative contribution of each element of the surface area to the 198 measured vertical flux, according to the advection-diffusion equation. This acts as a weight-199 ing function, such that some areas contribute strongly to the measured flux, and oth-200 ers not at all. If the mean flux F of a scalar over a landscape represented by a discre-201 tised gridded domain with dimensions n_x by n_y at time t is given by: 202

$$\bar{F}_t = \sum_{x=1}^{n_x} \sum_{y=1}^{n_y} F_{xyt} \frac{1}{n_x n_y}$$
(5)

where the overbar denotes spatial averaging, eddy covariance effectively measures a weighted mean, where the footprint provides the set of weights, ϕ , to give:

$$\widehat{F}_{t} = \sum_{x=1}^{n_{x}} \sum_{y=1}^{n_{y}} F_{xyt} \phi_{xyt}$$
(6)

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where we use the hat-symbol to indicate that this is an estimator, not \overline{F}_t itself.

So the appropriate way to upscale chamber measurements is to use Bayesian kriging to estimate F_{xyt} over the whole grid, and thereby the domain mean \bar{F}_t . However, to compared with eddy covariance, we need to apply the footprint weighting ϕ to estimate the flux that we would expect eddy covariance to measure, \hat{F}_t .

Here, we applied the method to data from an arable field of oilseed rape in Lin-210 colnshire, U.K., where fluxes of N_2O were measured by chambers and eddy covariance 211 in the same field (Keane et al., 2017). The Bayesian kriging method provides a way of 212 scaling chamber flux to the field scale, incorporating spatial pattern and its associated 213 uncertainty. This prediction is weighted by the flux footprint to produce our expecta-214 tion of the flux measured by eddy covariance. This allowed us to compare one method 215 with another, accounting for the difference in spatial sampling characteristics inherent 216 in the two methods as best we can, and account for the associated uncertainty properly. 217

Figure 2 shows the chambers fluxes of N_2O upscaled by Bayesian kriging and weighted 218 by the footprint probabilities ϕ , to give a valid comparison against the eddy covariance 219 data. The results show that the upscaled values can deviate substantially from the naïve 220 sample mean of the chambers. More importantly, the uncertainty in the mean estimated 221 by Bayesian kriging is substantially larger than the conventional 95 % CI in most cases. 222 This is a source of uncertainty that is typically ignored, and this demonstrates the im-223 portance of representing spatial upscaling effects explicitly. This also indicates that we 224 need to be careful in drawing conclusions from such measurements when interpreting the 225 difference in means among experimental treatment plots. The magnitude of error will 226 depend on the spatial pattern in the surface flux; in extreme situations, the sample mean 227 could be quite poorly representative of the large-scale mean. In these cases, spatial up-228 scaling clearly needs to be considered explicitly, and the Bayesian kriging described here 229 provides a rigorous method to do this. 230

When comparing chambers and eddy covariance measurements, the difference in 231 232 spatial sampling is usually ignored, and the arithmetic mean of chamber flux samples is compared with the value from eddy covariance for the corresponding time period. This 233 ignore the fact that the flux footprint acts as moving spatial filter, whereby the location 234 and extent of the area that influence the measured flux changes each half-hour, accord-235 ing to wind speed, direction etc. The method described here also provides a rigorous way 236 to compare chamber measurements so that they represent the same area that is sampled 237 by eddy covariance. 238

²³⁹ 4 Temporal upscaling of chamber fluxes to annual scale

There are considerable challenges in interpolating and extrapolating cumulative N₂O fluxes, based on relatively sparse and variable measurements from only a few time points. This leads to substantial uncertainty in the "emission factor" (EF or Ω , the total N₂O released as a percentage of the fertiliser nitrogen added) which is used in the national

inventory. This is analogous to the problem of estimating the spatial mean from mea-244 surements at a limited number of locations, but in the time domain. The method most 245 commonly used in the literature to calculate cumulative N_2O fluxes is to interpolate and 246 integrate using trapezoidal rule integration. However, this method is very sensitive to 247 noise in the data, there is no straightforward way to quantify the uncertainty introduced 248 or to extrapolate beyond the sample data, and it does not account for the typically log-249 normal spatial distribution of fluxes. Here, we used two approaches to examine spatial 250 and temporal upscaling: firstly using a process-based model directly, and secondly us-251 ing Bayesian emulation of this model. 252

4.1 DNDC model

²⁵⁴ DNDC is a process-based biogeochemical model, widely used to estimate agricul-²⁵⁵ tural soil N₂O emissions. We parameterised the model using chamber measurements of ²⁵⁶ N₂O fluxes, along with data on harvest crop grain N content, soil mineral N and soil mois-²⁵⁷ ture, from four experimental sites between 2010 and 2012, accounting for time lags be-²⁵⁸ tween measured and simulated time-series (Myrgiotis et al., 2016).

The model was applied across a 3800 km² area of Scotland where >90 of its croplands are located. Spatial data on soil properties, crop coverage and weather, and UKspecific crop calendars and fertiliser-use recommendations were used to create model inputs for 2011-2013 at a 1 km² resolution (Myrgiotis et al., 2018). The distribution of Ω estimated from the regional simulations was compared against the site data of measurementsbased Ω to evaluate the effect of the upscaling process.

Measured and simulated distributions of Ω have similar range and shape (Figure 265 3), but there were important differences in their distributions, e.g. their inter-quartile 266 range (upper/lower dotted lines in Figure 3). The wider inter-quartile range of the mea-267 sured Ω is a result of greater variability in weather across the UK field sites and mea-268 surement period compared to the conditions over the Scottish arable region for the up-269 scaling period. There were higher temperatures and precipitation at some of the field 270 sites responsible for peaks in Ω during experimental measurements. The mean of the sim-271 ulated Ω is 0.47%, 14% less the mean measured Ω of 0.55%. The inter-quartile range 272 of the simulated Ω was around half that found in the measurements. We conclude that 273 upscaling causes substantive changes in Ω linked to the different range of conditions en-274 countered across the wider region. 275

4.2 Meta-model of N₂O fluxes

To address this issue, we developed an emulator of the DNDC model with which we could apply Bayesian calibration to characterise the uncertainties in the spatial and temporal distribution of emissions. Following a fertilisation event, the time course of N_2O flux is expected to rise to a peak, then decay exponentially. This pattern in time is reproduced by DNDC and similar models, and is well described very simply by the lognormal equation:

$$\mu_t = \frac{1}{\sqrt{2\pi}kt} e^{-(\log(t) - \Delta)^2 / 2k^2} N_{\rm in} \Omega$$
(7)

where μ_t is the spatial mean of the N₂O flux at time t, Δ and k are analogues for the location and scale parameters, N_{in} is the nitrogen input, and Ω is the fraction of this nitrogen which is released as N₂O. Because the lognormal function integrates to unity at $t = \infty$, Ω is implicitly based on the total cumulative emission, rather than at an arbitrarily defined time. The symbol Δ can be interpreted as the natural logarithm of the delay between fertiliser application and peak flux; k is a decay rate term. This equation provides a simple meta-model which can be used to emulate the behaviour of DNDC (and
 similar models).

Because μ_t typically has a very skewed spatial distribution, there is a high prob-291 ability of the sample means underestimating the true value; the problem increases as vari-292 ance increases and sample size decreases. Several approaches have been proposed as more 293 efficient estimators of the location and scale of lognormal distributions, but none of these 294 entirely solve the problem when σ is large and n is small, as is generally the case with 295 flux measurements. In this study, we applied a Bayesian approach, using the Markov Chain 296 297 Monte Carlo (MCMC) method with Gibbs sampling (Gelman et al., 2013). In this way, we estimated the parameters of the underlying distribution. 298

So, at time t following fertilisation, the mean flux is given by Equation 7, at which time the N_2O flux has a distribution

$$F \sim \ln \mathcal{N}(\mu_{\log,t}, \sigma_{\log}^2)$$

$$\mu_{\log,t} = \log(\mu_t) - 0.5\sigma_{\log}^2$$
(8)

To obtain the cumulative flux at time t, we use the standard lognormal cumulative distribution function $(\ln t - \Lambda)$

$$F_{\text{cum},t} = \Phi\left(\frac{\ln t - \Delta}{k}\right) N_{\text{in}}\Omega \tag{9}$$

where Φ is the cumulative distribution function of the standard Normal distribution. The model was encoded in the JAGS language, and fitted to a number of data sets from across the UK.

Priors for Δ and k were specified as Normal distributions based on the temporal patterns produced by the DNDC model (see below). A Normal distribution was also assigned to σ_{\log} , based on earlier data from various sites in the UK, mainly from Cowan et al. (2014, 2016). The prior distribution for σ_{\log} was truncated at zero to exclude negative values. Ω was given a lognormal distribution, fitted to the data collation of Stehfest and Bouwman (2006) which included data on emission factors from all over the world.

Figure 4 shows the posterior distribution of cumulative fluxes calculated from the 312 UK data sets, expressed as the emission factor, Ω . This distribution is very narrowly de-313 fined in some cases (e.g. Dum 2012-10-16, EBS 2009-03-17), and very wide in other cases 314 (e.g. EBS 2007-05-16, EBS 2008-06-18), meaning that uncertainty in the emission fac-315 tor can be very small or very large. Ω is generally in the range zero to 5 %, but some 316 events have substantially higher emission factors. The value estimated by the trapezoidal 317 method is generally within the posterior distribution of the lognormal model, but the 318 values are usually lower. The emission factor is generally rather poorly constrained by 319 flux chamber measurements, because of the difficulties of accurately estimating the mean 320 of a lognormal distribution with large variance when n is small. The standard approach 321 fails to capture this uncertainty. Our new approach performs well in that it appropri-322 ately quantifies the uncertainty, and removes some of the bias by accounting explicitly 323 for the lognormal distribution. 324

5 Reducing uncertainty in large-scale fluxes via improved instrumentation

The direct measurement of GHG fluxes has developed in tandem with the instrument technologies that allow GHG mixing ratios to be measured. Compared with CO_2 , N_2O is less amenable to measurement by infra-red absorption: it is present at lower background concentrations; the typical fluxes are smaller relative to the background concentrations; and the infra-red absorption bands are narrower, making the technicalities of

measurement more difficult. Until recently, N₂O was only accurately measurable by gas 332 chromatography. Hence, for the most part, observations of N_2O fluxes are only available 333 using static chamber methods (Hutchinson & Mosier, 1981; Matson & Harriss, 2009), 334 which necessarily sample small areas (typically $< 0.1 \text{ m}^2$) over short time periods, based 335 on few (2-4) points. However, N₂O fluxes show a wide variability, ranging over orders 336 of magnitudes on small spatial scales, which is not predictable. This is attributed to two 337 main causes: the variety of unobserved microbial controls on gas production, including 338 their physiological activity and population dynamics; and the sensitivity of physical trans-339 port in the soil, which interacts with the measurement process (enclosure) in a complex 340 way (Xu et al., 2006; Sahoo & Mayya, 2010). Fast-response sensors for N₂O have recently 341 become available, making high-precision chamber measurements (Cowan et al., 2014) and 342 the micrometeorological eddy covariance method (Kroon et al., 2010) feasible. The chal-343 lenge here is to use these new sensors to develop continuous measurement systems for 344 fluxes of N_2O . With continuous measurements, we overcome the need to interpolate and 345 extrapolate in space and time, and thereby remove the large uncertainties this introduces. 346 In the GHGEF programme, we developed and applied two systems which provide near-347 continuous measurements of N_2O flux: a robotic auto-chamber system using a cavity ring-348 down spectroscopic instrument ("SkyLine"), and an eddy covariance system based on 349 quantum cascade laser (QCL) spectroscopy. 350

351 5.0.1 SkyLine

A detailed description of the system is available in Keane et al. (2018). Briefly, the 352 SkyLine2D automated chamber system used a single, cylindrical chamber (internal di-353 ameter 40 cm, height 62 cm), suspended from a motorized trolley mounted on parallel 354 horizontal ropes held above the crop by 2.5-m tall aluminium trellis arches (Figure 5). 355 The trolley repeatedly traversed a transect across the crop of up to 40 m, enabling mea-356 surements at high spatial resolution (< 1 m) across replicated manipulations or under-357 lying variation in the landscape. At designated locations where collars were placed in 358 the soil, the chamber automatically lowered and sealed on the collar to conduct a flux 359 measurement. The base of the chamber was fitted with a rubber gasket which formed 360 a gas-tight seal when dropped on the flange of the landing base. Guides around the cham-361 ber bases ensured the chamber landed accurately. A vent was fitted into the chamber 362 to minimize pressure differences between the chamber and the external atmosphere (af-363 ter Xu et al. 2006). The chamber operated as a non-steady state dynamic system, with headspace gas being circulated between the chamber and a cavity ring-down spectroscopic 365 (CRDS) analyser for N₂O (LGR isotopic N₂O analyser, Los Gatos Research, CA, USA) 366 housed in an enclosed shed at one end of the SkyLine2D apparatus (Figure 5). The CRDS 367 analyser operated at 1 Hz, giving a precise measurement of the rise in mixing ratio within 368 the chamber, allowing a relatively short chamber closure (ca. 5 min), thus minimising 369 the time during which the underlying plants and soil are isolated from ambient condi-370 tions. 371

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5.0.2 QCL eddy covariance system

The system used a continuous wave quantum cascade laser (QCL) absorption spec-373 trometer (CW-QC-TILDAS-76-CS, Aerodyne Research Inc., Billerica, MA, USA), with 374 an ultra-sonic anemometer (WindMaster Pro 3-axis, Gill, Lymington, UK) to measure 375 fluctuations in 3-D wind components at a frequency of 20 Hz. The QCL was fitted with 376 a laser capable of measuring N_2O with a precision of 0.3 ppb, together with H2O and 377 either CO_2 or CO, using absorption features at 12 and 22 m⁻¹. Internal software fits the 378 observed spectra to a template of known spectral line profiles from the HITRAN (High-379 resolution TRANsmission) molecular spectroscopic database. Absolute gas concentra-380 tions can then be calculated from the strength of the absorption line measured, the tem-381 perature, pressure and path length. A vacuum pump (Triscroll 600, Agilent Technolo-382

gies, US) was used to draw air through the inlet and instrument with a flow rate of approximately 14 l min⁻¹. Data from the sonic anemometer and QCL was logged in tandem using a custom program written in LabView (National Instruments, TX, USA).

Fluxes were calculated over 30 minute intervals using the EddyPro software (Ver-386 sion 6.2.1, Li-COR, Lincoln, NE, U.S.A.), based on the covariance between gas concen-387 tration and vertical wind speed. For flux data taken with a low signal-to-noise ratio, time-388 lag identification by maximisation of the cross covariance introduces systematic biases 389 (Langford et al., 2015). Here, we investigated methods for estimating the timelag, based 390 391 on either maximising the timelag over a longer time window, or using the timelag established for CO_2 . Both CO_2 and N_2O share the same path in the sample line to the mea-392 surement cell, and would be expected to travel at the same rate. Fluxes of CO_2 are an 393 order of magnitude larger, so the timelag which gives the maximum covariance is usu-394 ally clearly defined within each half-hour, and should be equally applicable to N_2O , and 395 not subject to the systematic error described by Langford et al. (2015). Depending which 396 laser was installed in the QCL, measurements of CO_2 were not always available, so the 397 method using N_2O only was also used. After investigating options, we decided on a six-398 hour window in which we found the timelag which maximised covariance. This timelag 399 was then fixed for all data within the six-hour window, and fluxes were calculated on a 400 30-minute basis. Standard corrections were made in the flux calculation, following Moncrieff 401 et al. (1997), Ibrom et al. (2007) and Burba (2013). Random uncertainty was estimated 402 by the method of Finkelstein and Sims (2001). 403

Figure 6 shows a near-continuous time series of measurements of N_2O flux over a 404 whole month by both automated SkyLine chamber and eddy covariance methods. Con-405 ventionally, only infrequent static chamber measurements (blue symbols) would be avail-406 able. Given the variability in the data - irregular peaks in time and spatial variability 407 between chambers - the near-continuous measurements provide a much more accurate 408 estimate of the cumulative emission over the period following fertiliser application, and 409 thereby the emission factor. Measurement techniques such as these, with better data cov-410 erage, are clearly needed to enable appropriate temporal upscaling, so that longer-term 411 means and cumulative emissions can be estimated accurately from the observations. 412

6 Quantifying uncertainty in eddy covariance measurements of field scale fluxes

As described above, eddy covariance (EC) is a one of the key techniques for mea-415 suring GHG fluxes, but subject to instrument noise, uncertainties in the processing steps 416 and micrometeorological conditions required to meet the underlying assumptions. The 417 measurement error for eddy covariance is thus challenging to quantify. Due to their ex-418 pense, EC systems are usually deployed singly, so there is a lack of replication – the most 419 obvious means by which this can be estimated. By using two EC systems 800 m apart 420 (Hollinger & Richardson, 2005) were able to quantify the measurement difference (er-421 ror) in flux calculations. However, the tower separation was large enough so that foot-422 print regions did not overlap, so the comparison is confounded. Hollinger and Richard-423 son proposed a time-for-space substitution to allow error calculations from a single EC 424 system, the so-called successive days method, and showed it has utility. 425

Here for the first time we present a comparison of NEE flux estimation from five 426 co-located flux systems, all sampling the same pasture in southwest Scotland (Crichton 427 Research Farm, Dumfries). The low stature of the vegetation allowed the EC systems 428 to be set up within 10 m of each other, each sampling at 5 m above ground level. By 429 having five systems the variance between these can be directly determined for the same 430 conditions to assess instrumental and processing error. All systems used a Gill R3 sonic 431 anemometer; 4 used a Vaisala GMP343 CO₂ sensor and a Honeywell HIH-4000 relative 432 humidity sensor, while one used a LI-COR LI-7500 (Hill et al., 2017). Fluxes were cal-433 culated for 30-min periods using EdiRe (version 1.5.0.50). There were 13 days during 434

June 2015 with almost continuous and simultaneous measurements by all sensor systems. The pasture was cut on 13 May, and LAI was tracked using an LAI-2000 (LiCor Inc), growing from 3 to 4 during the study period from 5-17 June. The deviations of the 30 minute time series, and the daily time series from aggregated data were calculated, and their distributions tested for normality using a Shapiro-Wilk test (R software).

There was broadly a good agreement between all the flux systems in their estimates 440 of NEE at 30 minute and daily time steps (7). The median standard deviation of the 30441 minute data was 2 μ mol $m^{-2}s^{-1}$. The 30-minute error distribution deviated significantly 442 from normal (P < 0.001), being clearly log-normal with a high tails (7). For the daily ag-443 gregate NEE estimates, the median value was 0.45 gC $m^{-2}d^{-1}$, and the error distribu-444 tion was not significantly different from normal (Shapiro-Wilk test p=0.93). These re-445 sults are close to the results of space-for-time error estimates at Harvard Forest, which 446 reported a daily standard deviation of 0.58 gC $m^{-2}d^{-1}$ (Hill et al., 2012) for NEE. 447

The value of EC error estimates like these are clear. Errors determine the capac-448 ity of EC to detect C sources and sinks, which are driven by small differences between 449 large input and output fluxes. Error data are also important to set the weighting of EC 450 observations in the calibration of C process models. Bayesian calibration methods weigh 451 the importance of observational data constraints in model fitting on the basis of mea-452 surement error. Bayesian calibrations propagate measurement uncertainty into model 453 predictions, for example by finding an ensemble of model parameters that produce es-454 timates of NEE consistent with observations and their uncertainty. Eddy covariance data 455 have been used successfully across the UK to produce more robust calibration of C cy-456 cle models. The calibrated models then propagate the flux uncertainty from localised 457 and incomplete data set at a few sites into complete and regional assessments of C cy-458 cling (Myrgiotis et al., 2020; Smallman et al., 2017; Revill et al., 2016). 459

⁴⁶⁰ 7 Quantifying aggregation error in spatial upscaling

7.1 Analytical approach

461

Typically, we derive a model of the GHG flux based on small-scale measurements 462 as a function of input variables x, measured at the corresponding scale. We want to ap-463 ply this to a larger scale, usually using the mean value of x, averaged over a grid cell, 464 region, or longer time period. However, the output of a model f at the larger scale, with spatially varying inputs x should be calculated as $\int_{-\infty}^{+\infty} f(x) p(x) dx$, where p(x) is the probability distribution of input values in the region. We can denote this integral more 465 466 467 briefly as E[f(x)], using the expectation operator E[]. Whenever models are run with 468 the averaged inputs, what is being calculated is f(E[x]) rather than E[f(x)]. If the mod-469 els are nonlinear, this results in an upscaling error δ defined as: 470

$$\Delta = f(E[x]) - E[f(x)] \tag{10}$$

When models are applied to large regions, the domain is typically subdivided using a spatial grid, with each grid cell covering an area of tens or hundreds of square kilometres. In contrast, the models themselves tend to be based on observations made at much finer scales such as flux chambers, eddy covariance footprints, individual crop fields, or forest stands. The challenge here is in using a small-scale model to estimate GHG fluxes in a large grid cell, accounting for within-cell spatial heterogeneity, and the error this produces.

In the GHGEF programme, we derived a method for characterising this error, based on a multivariate Taylor-expansion approach, a development of earlier work in crop and forest modelling (Bresler & Dagan, 1988; Band et al., 1991; Rastetter et al., 1992). The approach is to estimate Δ and correct for it. For functions of one variable, we can use the formula derived by applying the expectation operator to the second-order Taylor expansion of f(x):

$$E[f(x)] \approx f(E[x]) + \frac{1}{2} Var[x] f^{(2)}(E[x]), \qquad (11)$$

where $f^{(2)}(E[x])$ is the second derivative of f evaluated at the mean of x, and Var[x]is the variance of x within the region. Combining this formula with Eq. 10 gives us an approximate formula for the upscaling error of models with one input variable:

$$\widehat{\Delta} = -\frac{1}{2} Var[x] f^{(2)}(E[x]), \qquad (12)$$

where we use the $\hat{}$ -symbol to indicate that the formula provides an estimator for Δ , not Δ itself. For functions of multiple input variables where x is a vector, multivariate Taylor expansion yields:

$$\widehat{\Delta} = -\frac{1}{2} \operatorname{tr}(S H), \tag{13}$$

where S is the variance-covariance matrix of x, H is the Hessian matrix of second order partial derivatives of f(x), and tr denotes the trace (sum of diagonal elements) of the matrix product SH. For simple models, analysis shows that the formula is exact (i.e. $\hat{\Delta} = \Delta$), allowing full correction of model upscaling errors. In other cases, the formula provides an approximation.

We demonstrated the application of this approach to example models (Van Oijen 495 et al., 2017) of methane flux (Levy et al., 2012), ammonia and nitrous oxide flux (Flechard 496 et al., 2007). Using high-resolution data, we could calculate the true model output E[f(x)]. 497 After averaging the input data in 32-km grid cells, we re-evaluated the model using this, 498 corresponding to f(E[x]), and by difference, we calculated the aggregation error Δ . Us-499 ing the Taylor-expansion approach described above, we then calculated our approxima-500 tion of this error Δ , for verification against the known values of Δ . Depending on the 501 model, aggregation error could be substantial, ranging from -3 to - 48 %. The error var-502 ied spatially (Figure 8), depending on the range of the input variables, their sub-grid vari-503 ance, and the non-linearity of the model. The Δ formula gave reasonable approximations 504 to the true error (Figure 8, right-hand panel), correcting model output to within 2 to 505 -9 % of the true values. 506

For nonlinear models, the effects of spatial upscaling need to be accounted for, and the $\hat{\Delta}$ formula described here is a generally applicable means to do this. The approach can be applied to more complex, high-dimensional process-based models, but exploration of the accuracy of the approximation is needed.

511

7.2 An approach using earth observation and model-data fusion

An alternative approach to quantify aggregation error is to use earth observation 512 data (and other relevant mapped products) to provide inputs at increasingly fine scales. 513 We conducted four model-data fusion analyses at differing spatial resolutions of Great 514 Britain's (GB) terrestrial carbon cycle at a monthly time step for an 18 year period (2001-515 2018) using the CARbon DAta Model framework (CARDAMOM, Bloom et al., 2016). 516 CARDAMOM uses a Bayesian approach within an Adaptive Proposal - Markov Chain 517 Monte Carlo (AP-MHMCMC, Haario & Tamminen, 2001; Roberts & Rosenthal, 2009) 518 to estimate location (i.e. pixel) specific ensembles of parameters for an intermediate com-519 plexity model of the terrestrial C-cycle (DALEC Bloom & Williams, 2015). The param-520 eter ensembles are consistent with observational constraints, their uncertainties, model 521 structure, meteorology and disturbance (fire & forest loss). From these parameter en-522 sembles we are able directly estimate at pixel level the uncertainty of DALECs C-cycle 523 simulation of terrestrial fluxes and stocks. CARDAMOM analyses were conducted at 4 524

spatial resolutions, 111 x 111 km, 56 x 56 km, 28 x 28 km and 5 x 5 km, all at monthly 525 temporal resolution. Observational constraints assimilated by CARDAMOM are monthly 526 time series information on leaf area index (LAI), a single estimate of above ground biomass 527 (AGB) and soil C stocks. LAI is extracted from the 1 x 1 km, 8 day product from the 528 Copernicus Service Information (2020). A single per-pixel estimate of AGB for 2017 and 529 its uncertainty is drawn from ESA's 1 x 1 km CCI Biomass product (Santoro, 2021). From 530 this AGB estimate we derive the total woody biomass (which corresponds with DALECs 531 model structure) following Saatchi et al. (2011). A single per-pixel estimate of soil car-532 bon stock is extracted from the SoilGrids database (Hengl et al., 2017). Meteorological 533 drivers are drawn from the ERA5 reanalysis (Hersbach et al., 2020). 534

Our analyses all estimate that GB's terrestrial ecosystems were a net sink of car-535 bon with a net biome exchange (NBE) of -6.7 to -10 TgC yr⁻¹ (-0.32 to -0.41 MgC ha 536 yr^{-1}) between 2001 and 2017 (Table 1). Spatial resolution has a substantial impact on 537 the magnitude and spatial patterning estimated by the CARDAMOM analyses (Table 538 1, Figure 9). For example, at the coarsest resolution of the model grid (111 x 111 km) 539 NBE is near neutral across much of GB except the far north east and south west. In con-540 trast the highest resolution (5 x 5 km) shows fine scale variation across the whole of GB. 541 Moreover, the range of C flux magnitudes estimates is smaller in the coarser resolution 542 analyses due to aggregation of sub-grid variability. The distributions of pixel-level mean 543 annual fluxes progressively converge towards that estimated by our finest spatial reso-544 lution analysis (Figures 9, 10). 545

The estimates of the gross biological fluxes (i.e. GPP and Reco) are relatively in-546 sensitive to spatial resolution between the 5 km, 28 km and 56 km analyses. The mean 547 annual flux estimate of GPP and Reco in the 5 km, 28 km and 56 km resolution anal-548 yses vary by less than 0.2 MgC ha yr⁻¹ (< 2%; Table 1). Estimation of emissions due 549 to disturbance, both fire and forest cover loss, are progressively underestimated at coarser 550 spatial resolutions (Table 1) and show varied time series dynamics (Figure 10). The mean 551 pixel-level forest loss estimates plateaus between 5 km and 28 km (Table 1). However, 552 there remains disagreement in the temporal interannual variability and magnitude of for-553 est loss estimates particularly after 2010 (Figure 10). Emissions due to fire have not con-554 verged in terms of mean annual emissions (Table 1), or temporal dynamics and magni-555 tude (Figure 10). Overall, these results suggest that there remains substantial sub-grid 556 scale disturbance information missing from these analyses, even at 5 km spatial resolu-557 tion. 558

559 8 Conclusions

We reviewed some of the major challenges in accurately estimating net fluxes of 560 GHGs at national scale. These revolve around the difficulties of extrapolating small-scale, 561 short-term observations and models. Uncertainty necessarily arises if these are extrap-562 olated to larger scales because of the imperfectly known models, parameters and inputs 563 used in the extrapolation. Where spatial pattern exists in the small-scale observations, 564 this can be accounted for using a geostatistical approach such as kriging; applying this 565 in the Bayesian framework allows the uncertainty to be propagated correctly. To tackle 566 this same problem in the time dimension, process-based models can be used, which bring 567 in prior knowledge of how we expect the flux to change over time. Temporal upscaling 568 of N_2O fluxes has been difficult because of the lack of high frequency measurements over 569 whole growing seasons. Substantial random error and systematic bias can be introduced 570 when extrapolating in time to estimate cumulative fluxes. Again, the Bayesian approach 571 can be used to characterise the uncertainty in their application. Modelling the spatial 572 and temporal pattern distribution in the observations in a Bayesian framework allows 573 this uncertainty to be accounted for appropriately. Furthermore, new measurement tech-574 niques are now allowing near-continuous, long-term observations to be made of these GHG 575 fluxes, which reduce the need for extrapolation. Models applied at large scales need to 576

account for sub-grid-scale heterogeneity; significant upscaling error can occur otherwise.

This aggregation error can be approximated closely by a Taylor-expansion formulation,

based only on the model Hessian matrix and the estimated variance-covariance matrixof the inputs.

With advances in computing power, running models of GHG fluxes over large do-581 mains at relatively high resolution is increasingly feasible. For example, we can now rou-582 tinely run models over the UK at 1-km resolution (Robinson et al., 2016), requiring 240,000 583 grid cells. In moving from the 1 km^2 scale to national scale, the issues of model upscal-584 ing do not apply in the same way, so long as the model has been correctly scaled from 585 the original observations to the 1-km² scale. This is because the variance over the whole 586 domain is represented, albeit as 1-km² means, i.e. we evaluate E[f(x)], so $\Delta = 0$, so the 587 total flux is simply the sum of all the grid cells. In moving between these scales, the main 588 problem encountered is in estimating the input variables accurately over this domain. 589 Because of the large size of the domain, our information is often incomplete or based on 590 proxies (e.g. NDVI) Disney et al. (2016). This is an area where data availability is rapidly 591 improving, with new and improved satellite products becoming available, as well as other 592 new technologies - improved laser-based spectroscopic methods, and unmanned airborne 593 vehicles. 594

A key point is simply to recognise that the scaling problem exists, and that some 595 attempt to explicitly deal with it, however inadequate, is better than none, and gives a 596 basis for improvement upon. The field of geostatistics has dealt with similar problems 597 over several decades, but the techniques are not yet widely applied in biogeochemistry 598 or process-based modelling. A common theme throughout is the adoption of Bayesian 599 principles for combining multiple information sources in a coherent way, which allows 600 us to make better use of the available data. We expect this trend to continue and ex-601 pand into the machine learning domain. We can foresee further application of data as-602 similation methods which combine bottom-up models with atmospheric observations, that 603 are thereby better constrained, potentially more powerful, and easier to interpret in terms 604 of the national-scale inventories. 605

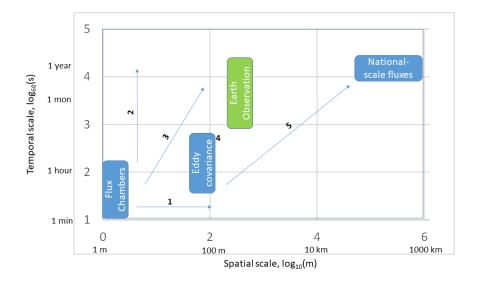
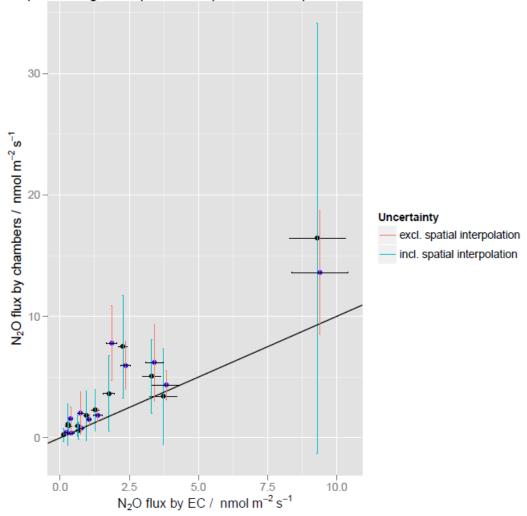


Figure 1. Diagram illustrating the spatial and temporal scales of the upscaling problem, from measurements using chambers and eddy covariance which cover small scales, to national-scale annual fluxes. The numbers respond to the challenges which we identify in tackling this problem. Earth observation represents a rapidly expanding source of data with wide coverage and increasingly fine resolution.



Footprint-weighted spatial interpolation with prediction errors

Figure 2. Comparison of fluxes of N_2O measured by chambers and eddy covariance for 12 sampling occasions. For each occasion, symbols show (i) the naïve sample mean of the chamber N_2O fluxes, with the conventional 95 % confidence intervals representing uncertainty (blue symbols, red vertical bars), and (ii) the maximum a posteriori mean flux of the chambers N_2O fluxes produced by Bayesian kriging, with the 2.5 and 97.5 percentiles of the posterior distribution representing uncertainty (black symbols, blue vertical bars). The latter are offset fractionally for visibility. The solid black line shows the 1:1 relationship.

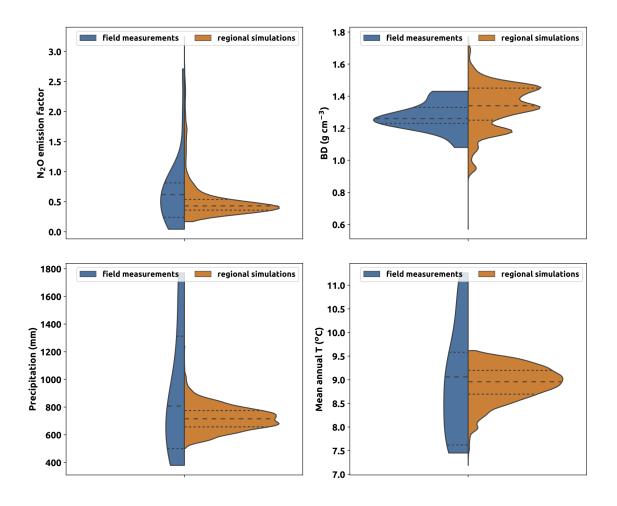


Figure 3. A comparison of distributions of emission factor Ω for N₂O generated from (i) field measurements collected across range of different management approaches, sites and seasons at locations in the UK and (ii) from upscaling estimates based on a model calibrated using the site level data, applied across eastern Scotland over several years. Also shown are the distributions of soil conditions (bulk density, BD) and weather conditions during measurement periods for all sites, and the conditions relevant to the modelled upscaling. The crops grown in the measured and simulated sites were winter wheat, winter barley, spring barley and winter oilseed rape.

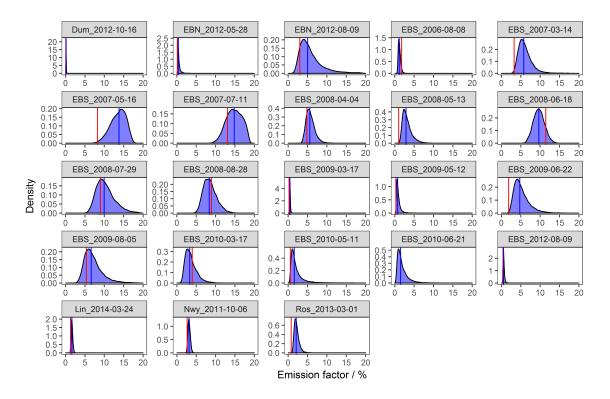


Figure 4. Posterior distribution of emission factor for each fertilizer application obtained by Bayesian estimation of the lognormal model. The blue vertical line shows the median of this distribution and the red vertical line shows the value calculated by the trapezoidal method. A lognormal distribution was used to provide the prior distribution of the emission factor.

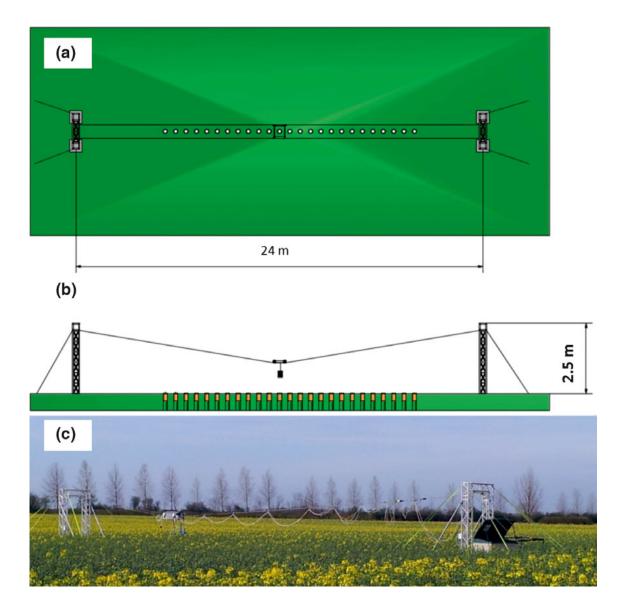


Figure 5. Aerial and side profile schematics of the SkyLine2D system showing (a), the trellis arch supports at either end, supporting the Kevlar ropes between. The motorized trolley is depicted at the mid-point of the two supports (b). Cross section of the in situ system at the OSR field site and (c) the N_2O and CH_4 Los Gatos cavity ring-down analysers were housed in the green garden box by the right-hand trellis support.

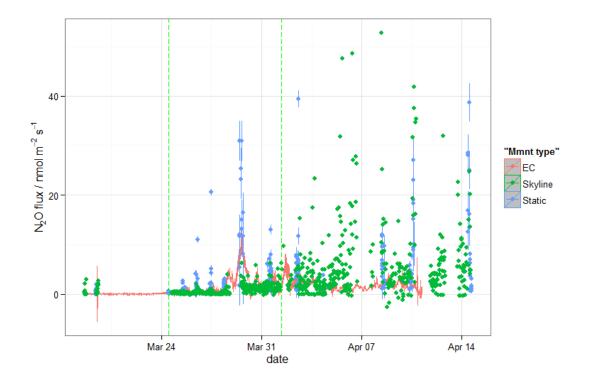


Figure 6. Time course of N_2O fluxes at the Lincolnshire field site, following fertilisation events (dotted vertical lines) as measured by eddy covariance (EC), static chambers, or the Sky-Line2D system.

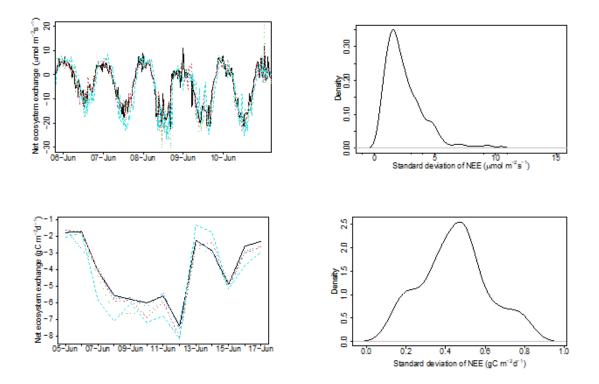


Figure 7. Comparison of five simultaneous eddy flux measurement systems (multiple time series during 2015 in left panels) and their variability (density of standard deviation across all systems for each time interval, right panels) over 30 minute (top panels) and daily (bottom panels) time integrals. The five eddy systems were all installed in the same field, sown with grass, and all within 10 m of each other. For 13 days there were continuous 30-minute time samples available for comparison (except 9-June, only 43 samples available for comparison) or for summing to produce a daily estimate. Only five days are shown in the top left panel for clarity.

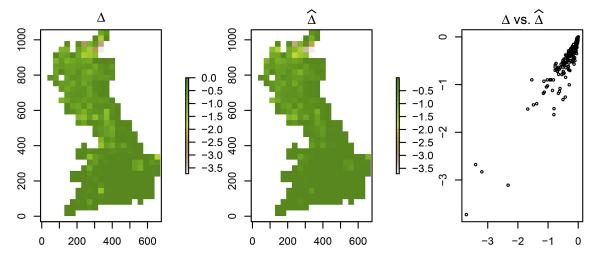


Figure 8. Left panel: The upscaling error Δ for the methane model is calculated by subtracting the correctly upscaled model results (using high-resolution input data) from the incorrect upscaling (using aggregated (mean) input data), and is mapped. Middle panel: The upscaling errors that were predicted by the $\hat{\Delta}$ -formula applied to the inputs and their (co)variances. Right panel: The quality of the error-prediction can be evaluated from the scatterplot of $\hat{\Delta}$ vs. Δ .

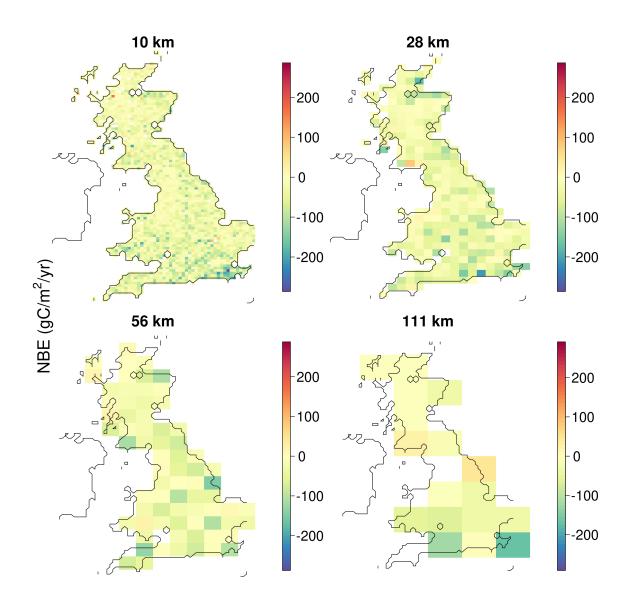


Figure 9. Net Biome Exchange (NBE = -GPP + Reco + Fire) estimated by 4 CARDAMOM analyses at a range of spatial resolutions. A negative value indicates a net uptake of carbon.

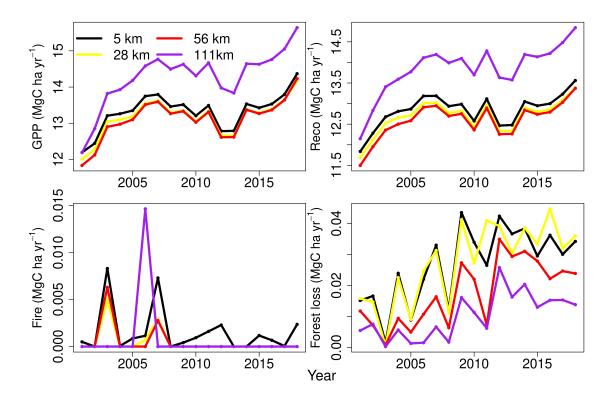


Figure 10. Time series of GB wide mean Gross Primary Productivity (GPP), ecosystem respiration (Reco), fire and forest loss estimated by 4 CARDAMOM analyses at a range of spatial resolutions.

Table 1. Mean annual C-budgets for Great Britain for each target resolution. Budget terms are presented as the mean annual flux in MgC ha⁻¹ yr⁻¹. Budget terms presented are gross primary productivity (GPP), ecosystem respiration (Reco), net ecosystem exchange of carbon (NEE = Reco - GPP) C loss due to forest removal (Forest loss) and net biome exchange of carbon (NBE = NEE + Fire loss). The 95 % confidence interval of each term is presented in parenthesis. Fire emissions were low and so are not included in the summary.

Resolution	GPP	Reco	NEE	Forest loss	NBE
$5 \mathrm{km}$	13.3 (10.1 / 16.0)		-0.39 (-4.1 / 4.4)	$0.03\ (0.01\ /\ 0.06)$	-0.39 (-4.1 / 4.4)
$28 \mathrm{km}$	$13.2\ (10.0\ /\ 15.9)$	$12.7 \ (8.5 \ / \ 18.2)$	-0.37 (-4.1 / 4.5)	$0.03\ (0.01\ /\ 0.06)$	-0.37 (-4.1 / 4.5)
56 km	$13.1 \ (9.9 \ / \ 16.1)$	$12.6 \ (8.5 \ / \ 18.2)$	-0.41 (-4.1 / 4.3)	$0.02 \ (0.01 \ / \ 0.04)$	-0.41 (-4.1 / 4.3)
111 km	$14.3\ (11.4\ /\ 16.6)$	$13.8 \ (9.9 \ / \ 19)$	-0.32 (-3.7 / 4.2)	$0.01 \ (0.005 \ / \ 0.02)$	-0.32 (-3.7 / 4.2)

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- and the British Atmospheric Data Centre (https://catalogue.ceda.ac.uk/uuid/c117a2e6f451405393cac1c6fbf8f7a3).

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