No Internal Connections Detected Between Leading Decadal to Multidecadal Climate Modes in North Atlantic and North Pacific Basins

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November 21, 2022

Abstract

Previous studies have attempted to unravel possible connections between leading decadal to multidecadal climate modes in the North Atlantic and North Pacific ocean basins, the Atlantic Multidecadal Variability in the North Atlantic, and the Pacific Decadal Oscillation and Victoria Mode in the North Pacific. We use newly available climate model data and apply improvements to existing methods to rexamine relationships among the different modes. Our main tool is the Multi-Model Large Ensemble Archive, which includes 270 ensemble members and allows for isolation of the forced and internal components of climate variability. Our results suggest that any internal connections between these modes are indistinguishable from random noise. Further, external forcing is shown to affect each region in similar ways, suggesting that climate change could be an indirect link between the two basins, and can confound the interpretation of he relationship between the basins.

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Supporting Information for

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Additional Supporting Information (Files uploaded separately)

None

| Modeling Center | CCCma | CESM | CSIRO | GFDL | GFDL | MPI |
|-----------------|---------|------------|-------|------|-------|------------|
| Model Version | CanESM2 | CESM1-CAM5 | MK3.6 | CM3 | ESM2M | MPI-ESM-LF |

| Modeling Center | CCCma | CESM | CSIRO | GFDL | GFDL | MPI |
|------------------------------------|------------------------|----------------------|-----------------------|---------------------|--------------------|-----------------------|
| Initialization Method | Macro/Micro | Micro | Macro | Micro | Macro | Macro |
| # of Ensemble Members | 50 | 39 | 30 | 20 | 30 | 100 |
| Latitude Resolution (grid points) | $2.75^{\circ}(64)$ | $<1^{\circ}(192)$ | 1.85° (96) | 2° (90) | 2° (90) | 1.875° (96) |
| Longitude Resolution (grid points) | 2.8125° (128) | 1.25° (288) | 1.875° (192) | 2.5° (144) | $2.5^{\circ}(144)$ | 1.875° (192) |
| Start Year | 1950 | 1920 | 1850 | 1920 | 1950 | 1850 |
| PI Length | 296 years | 319 years | 500 years | 200 years | 200 years | 281 years |

Table S1. Details of the six large ensembles from the MMLEA and observations used in this study, adapted from Deser et al. (2020) and from the MMLEA data webpage.

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image2.emf available at https://authorea.com/users/529866/articles/597210-no-internalconnections-detected-between-leading-decadal-to-multidecadal-climate-modes-in-northatlantic-and-north-pacific-basins

Supplemental Figure 1. Cross-correlation matrix of HADISST observed linearly-detrended filtered climate pattern time series. All possible cross-correlations between each of the three EOF modes in each basin are shown. From left to right: NA-EOF1, NA-EOF2, and NA-EOF3. From top to bottom: NP-EOF1, NP-EOF2, and NP-EOF3. Black lines represent statistical significance thresholds generated using the peak test, while green lines represent the cross-correlations for the corresponding modes. Lags correspond to the NA (NP) mode leading for positive (negative) lags.

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image3.emf available at https://authorea.com/users/529866/articles/597210-no-internalconnections-detected-between-leading-decadal-to-multidecadal-climate-modes-in-northatlantic-and-north-pacific-basins

Supplemental Figure 2. The effect linear detrending has on observed EOFs in the NP. For observations where no detrending occurs, A) and C) show the first EOF spatial pattern and time series respectively. B) and D) show the same but for the second EOF. For linearly detrended observations, E) and G) show the first EOF spatial pattern and time series respectively, while F) and H) show the same for the second EOF. I) shows the time series from C) (linearly detrended) and H) together to show their similarity (93% correlation, significant at the 99% threshold). The red curves are the same time series, while the blue curve in I) is equivalent to the time series in C) but with the linear slope removed.

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Supplemental Figure 3. EOF spatial patterns. From top to bottom: HADISST (observations), CANESM2, CESM, CSIRO MK36, GFDL CM3, GFDL ESM2M, and MPI (six MMLEA models). From left to right: NA-EOF1, NP-EOF1, NP-EOF2, and NP-EOF3. MMLEA patterns are internal (full minus ensemble mean) composite means (EOF analysis is performed on each member, and then the resulting EOFs are averaged across all members). Color scale is dimensionless, following standard EOF output.



Effect of Linear Detrending on Observed North Pacific EOFs



HADISST Observed (Linear Detrending) North Atlantic vs. North Pacific Cross Correlations





| T. Fenske¹ and A. Clement¹ ¹Rosenstiel School of Marine and Atmospheric Science, University of Miami, M Corresponding author: Tyler Fenske (tyler.fenske@rsmas.miami.edu) Key Points: Novel methods are applied to decadal to multidecadal climate modes to analy relationships between modes in different basins. An analysis of relationships between leading North Atlantic and North Pacifidos not reveal any internal connections, challenging previous results. External forcing such as global warming is shown to be a possible confounding climate relationships. | asins |
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30 Abstract

Previous studies have attempted to unravel possible connections between leading decadal 31 to multidecadal climate modes in the North Atlantic and North Pacific ocean basins, the Atlantic 32 Multidecadal Variability in the North Atlantic, and the Pacific Decadal Oscillation and Victoria 33 Mode in the North Pacific. We use newly available climate model data and apply improvements 34 to existing methods to rexamine relationships among the different modes. Our main tool is the 35 Multi-Model Large Ensemble Archive, which includes 270 ensemble members and allows for 36 37 isolation of the forced and internal components of climate variability. Our results suggest that any internal connections between these modes are indistinguishable from random noise. Further, 38 external forcing is shown to affect each region in similar ways, suggesting that climate change 39 could be an indirect link between the two basins, and can confound the interpretation of he 40 relationship between the basins. 41

42 Plain Language Summary

We examine possible connections between climate patterns in the North Atlantic and North Pacific oceans. New climate model data and improved statistical methods allow us to build on previous research of these links. In contrast to previous studies, no natural connections are detected. However, global warming is shown to affect each region in similar ways, suggesting that climate change could be an indirect link between the two basins.

48

49 Introduction

Climate modes are considered to be the leading source of internal climate variability, 50 affecting weather and climate patterns across the globe. These long-distance effects are 51 sometimes referred to as teleconnections and are driven by atmospheric bridges (Alexander et al. 52 53 2002, Liu and Alexander 2007). Sea-surface temperature (SST) variability associated with a particular climate mode is coupled to the atmosphere, allowing the mode to change the overlying 54 atmospheric circulation. This signal is then transported through the atmosphere to other regions, 55 where the variability influences the ocean in a distant location, potentially imprinting on or even 56 57 exciting a different climate mode there (e.g. Liu and Alexander 2007, Dommenget and Latif 2008). The magnitude of control a climate mode has on another region can also vary in time, 58 59 adding another dimension to potential interactions and making them more difficult to identify (Raible et al. 2014). The possibility of climate mode interactions must be considered to fully 60 understand the sources of internal climate variability. 61

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Here, we focus on potential interactions between decadal to multidecadal climate modes 63 in the North Atlantic (NA) and North Pacific (NP) ocean basins. In the NA, low-frequency 64 variability is captured via the Atlantic Multidecadal Oscillation (AMO) or Atlantic Multidecadal 65 Variability (AMV) (Enfield et al. 2001). In the NP, two modes are commonly used to capture the 66 low-frequency variability. The Pacific Decadal Oscillation (Mantua et al. 1997) and the Victoria 67 Mode (VM) (Bond et al. 2003) or North Pacific Gyre Oscillation (NPGO) (Di Lorenzo et al. 68 2008) are the two leading modes of decadal and multidecadal variability respectively. Many 69 other methods of capturing variability in these basins have since been developed (Eden and Jung 70 2001, Salinger et al. 2001, Martin et al. 2019, Nigam et al. 2020, etc.), although convention has 71

maintained scientific usage of the AMV and PDO as the dominant low-frequency modes.

73 Despite their extensive usage, these modes, especially the AMV, may not adequately isolate and

- capture a single source of internal variability (Marini and Frankignoul 2014; Wills et al. 2018;
- 75 O'Reilly et al. 2019, etc.). Representing multidecadal variability with relatively short observed
- periods (order of 100 years) is also challenging, especially when low-pass filtering is applied
- (Cane et al. 2017). Assuming oscillatory variability, only one or two full cycles may be observed
- (Mann et al. 2021). This reduces the effective degrees of freedom and subsequently requires care
 to be taken during statistical analysis, especially regarding significance testing.
- 79 80

Several previous studies have worked on the NA-NP relationship, all of which suggest 81 that the two basins have some statistical relationship with each other. Both d'Orgeville and 82 83 Peltier (2007) and Zhang and Delworth (2007) relate the first two Empirical Orthogonal Functions or Principal Components (EOFs or PCs, hereafter referred to as EOFs) of NP SSTs to 84 a metric for the AMV (d'Orgeville and Peltier use the first EOF of NA SSTs, while Zhang and 85 86 Delworth use NA area mean SSTs), although they utilize those EOFs in different ways. d'Orgeville and Peltier combine them and then isolate the 20 year (analogous to the PDO) and 60 87 year (analogous to the VM) period wavelets and conclude that there is a singular source driving 88 variability in each region, while Zhang and Delworth use the two EOFs and conclude that the 89 Atlantic Meridional Overturning Current (AMOC) drives the AMV, which in turn drives the 90 PDO/VM through atmospheric teleconnections. Wu et al. (2011) use the first two EOFs of each 91 92 basin and finds a statistically significant link. Marini and Frankignoul (2014) use several methods attempting to deconstruct the origin of the AMV, such as dynamical filtering and 93 removing trends in various manners. Their analysis includes a comparison of the AMV and 94 PDO, where they come to a similar conclusion as previous studies. Nigam et al. (2020) uses 95 global rotated empirical orthogonal functions (EOFs) to represent all major global modes, and 96 their modes most similar to the AMV and PDO also support a relationship existing. An et al. 97 (2021) use ensemble pacemaker experiments to suggest that multidecadal Pacific variability is 98 generated by AMO forcing and local air-sea interactions. These studies use the student's t-test for 99 significance thresholds, with Wu et al. (2011) also using customized bootstrap methods. All of 100 these studies are in agreement that a modest but statistically significant correlation exists with the 101 AMV leading the PDO by 12-14 years. 102

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Here we will build upon those previous examples using a novel dataset, the Multi-Model Large Ensemble Archive (MMLEA), and improve on existing methods, such as significance testing and mode definitions. Using these tools, we reanalyze the potential relationships between low-frequency climate variability in the NA and NP ocean basins. Our results challenge previous findings, suggesting that a relationship is not statistically significant. We also focus on the role that external forcing plays in the inter-basin relationship and find that it may be a confounding factor.

111

112 **Data**

For SST observational data, we use the UK Met Office's Hadley Centre Sea Ice and SST data set (HADISST) (Rayner et al. 2003). Data is linearly detrended at each grid point, and we use 151 years of monthly data from 1870-2020. We repeat our analysis with other common SST datasets and find no qualitative differences.

Our primary tool is the new Multi-Model Large Ensemble Archive (MMLEA) (Deser et 118 al. 2020). Necessary output is currently available for six ensembles included in the MMLEA 119 (each hereafter as CANESM2, CESM, CSIRO-MK36, GFDL-CM3, GFDL-ESM2M, and MPI). 120 Each ensemble contains at least 20 members for a total of 269 members (50, 39, 30, 20, 30, and 121 100 respetively). All MMLEA members are from the CMIP5 era and use historical forcing. Data 122 from each member is cut off at the year 2020 to match the observed period. We also use each 123 member's corresponding pre-industrial control run (PI), which are separate from the MMLEA. 124 One advantage of large ensembles is their capability to extract the forced signal from each 125 member by subtracting the ensemble mean (Kay et al. 2015). More details on the MMLEA are 126 provided in Table S1 (adapted from Deser et al. 2020).

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129 Mode Definitions

Our mode definitions for each basin loosely follow the conventions for the AMV, PDO, 130 and VM (Enfield et al. 2001, Mantua et al. 1997, and Bond et al. 2003 respectively). Instead of 131 applying unique methods (EOF analysis for the NP, regional area mean for the NA) to each 132 basin, we apply EOF analysis to each basin in an attempt to better capture individual modes of 133 variability. We compute the first three EOFs over the NP region of 20°N-65°N, 120°E-100°W 134 and the first EOF over the NA region of 0°N-65°N, 120°W-0°. We also compute further EOFs 135 for the NA, however these do not have as clear of a physical interpretation and do not affect our 136 results in any meaningful way. They are not included except in Supplemental Figure 1, which 137 shows a matrix of relationships between the first three EOFs of each basin, with no significant 138 relationships involving any NA EOFs except the first. Figure 1 shows the spatial patterns of 139 these EOF modes in the first row, while the second row shows the corresponding time series. 140 NA-EOF1 represents the AMV, characterized by a tripole spatial pattern and a predominantly 141 multi-decadal time series. NP-EOF1 represents the PDO, characterized by a dipole spatial pattern 142 with warming (cooling) in the central NP and cooling (warming) along the eastern boundary, 143 along with a mostly decadal time series. NP-EOF3 represents the VM, also characterized by a 144 dipole pattern offset to the west from the PDO's, as well as a mostly decadal time series. 145

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A major challenge in understanding climate mode drivers is separating internal variability 147 from externally forced variability. This is typically achieved by removing the estimated forced 148 trend. MMLEA data is detrended by removing the ensemble mean, following Deser et al. (2020). 149 150 We define full variability (hereafter full) as the unmodified output of each member of the MMLEA, and we define internal variability as the full minus the ensemble mean. Observations 151 are detrended linearly at each grid point, following convention. Figure 2 shows the effects of 152 linear detrending on the EOF analysis in the NP. When no detrending occurs, the first EOF 153 captures the externally forced global warming signal based on its uniform warming pattern, 154 while the second EOF is clearly the PDO. Linear detrending causes the PDO to become the first 155 EOF, while the second EOF still resembles the externally forced signal. Comparing the time 156 series of these two apparent global warming signals reveals that they are remarkably similar. The 157 second EOF of linearly detrended NP SSTs resembles the non-linear features of the global 158 159 warming signal. Similar results exist for the NA (not shown). Other more complex detrending methods exist and offer different interpretations about what is forced versus what is internal, 160 especially in the Atlantic (e.g. Frankignoul et al. 2017, Qin et al 2020). However, these 161 differences in forced signals are not substantial enough to affect our results qualitatively. 162 163

Note that the composite mean and the ensemble mean are distinct and computed differently, namely in the order of operations because EOF analysis is a non-linear computation. Here, the composite mean is defined as EOF analysis being done first, followed by averaging of the EOFs. In contrast, the ensemble mean is where averaging is done first, and then EOF analysis is performed on the averaged SSTs. Physically, the ensemble mean represents the externally forced signal in each model, while the composite mean represents the average internal modes across all members.

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All time series are normalized and smoothed with a 10-year low-pass Lanczos filter to focus on decadal to multi-decadal variability, following classical methods.

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176 177 Figure 1. EOF modes and normality analysis of observed HadISST linearly-detrended North Atlantic and North Pacific SSTs. 178 From left to right: NA-EOF1, NP-EOF2, and NP-EOF3. The percentage at the top of each column represents the 179 variance explained by each EOF. NA-EOF1 is analogous to the AMV, NP-EOF1 to the PDO, and NP-EOF3 to the VM. NP-180 EOF2 is analogous to the non-linear global warming signal (more detail in text). Rows from top to bottom: 1) Spatial patterns of 181 each EOF mode. Red (blue) corresponds to warming (cooling) when time series is positive (negative). 2) Filtered time series of 182 each EOF mode. 3) Probability distribution of each time series are shown with shaded green curves; corresponding standard 183 normal distributions are shown with black curves. 4) Quantile-quantile plots with data shown in blue points. A standard normal 184 distribution is shown in red for comparison. Three quantitative assessments of normality are also applied: 1) The Shapiro-Wilk 185 Test (SW). The corresponding p-value is listed, with p-values greater than 0.05 (95% confidence) implying a normal distribution. 186 2) D'Agostino's K-Squared Test (K^2). The corresponding p-value is listed, with p-values greater than 0.05 (95% confidence) 187 implying a normal distribution. 3) The Anderson-Darling Test (AD). The first value is the resulting measure of normality, and the 188 second value is the 95% critical value for the AD test. If the first value is smaller than the critical value, the distribution can be 189 assumed to be normal. All tests should be considered equally in determining whether or not to assume a normal distribution.

190 Normality of Modes

191 To determine whether standard parametric statistics can be used or not, the normality of

our time series must be assessed first. While non-Gaussian parametric statistics exist, Gaussian
 assumptions are common with geophysical time series analysis and are used by the previous

studies on the NA-NP inter-basin relationships (d'Orgeville and Peltier, 2007, Zhang and

Delworth 2007, Wu et al. 2011, Marini and Frankignoul 2014, and Nigam et al. 2020).

We test the four modes' filtered time series with five normality tests as recommended by 197 Yap and Sim (2011) and Ghasemi and Zahediasl (2012). Two are qualitative assessments: a 198 histogram with a standard normal curve fitted to the data (Figures 1I-L), and a quantile-quantile 199 plot (Figures 1M-P). Three are quantitative: the Shapiro-Wilk test (SW), the D'Agostino 200 skewness test (K²), and the Anderson-Darling test (AD) (shown as text in Figures 1M-P). Only 201 NP-EOF3 passes more than a single quantitative test. The combination of all tests generally 202 suggests that only NP-EOF3 of the filtered time series can be described as normal, so Gaussian 203 assumptions cannot be made for analysis of mode relationships. Therefore, non-parametric 204

statistics are required for significance testing.

206

207 Significance Testing

We use a non-parametric bootstrapping method for significance testing. Our primary 208 statistical tool is cross-correlation, so we build this method to evaluate the significance of a given 209 "real" cross-correlation. We create sets of random white-noise time series by shuffling each 210 observed unfiltered time series 1000 times. We also used sets of AR1 red noise (Katz 1982) and 211 quantile-mapped sets (Maraun 2013) and found no qualitative differences in our results. Each 212 random time series is filtered, and each possible pair of modes between each basin is cross-213 correlated for the entire random set. These 1000 cross-correlations are then used to compute 95% 214 significance thresholds for the corresponding observed cross-correlation. This method is similar 215 to the bootstrap used by Wu et al. (2011), although our method differs slightly. Wu et al. (2011) 216 calculate the 95th percentile at each specific lag in their cross correlation (hereafter the "point 217 test"). 218

218 219

Their interest, however, is not on a particular lag, but instead of the peaks of the crosscorrelation that are above the significance threshold. The specific lag at which these peaks occurred was unimportant - whether it occurred with 0 year lag or 30 years lag, their conclusions would remain the same. By definition, their statistical test requires that a particular lag be of interest, meaning that the lag at which the peaks occur *is* important, contrary to their conclusions. This can be viewed as an *a priori* test with an *a posteriori* conclusion, which suggests their significance thresholds may not be appropriate.

Alternatively, a "peak test" can be used. Instead of calculating the significance thresholds at each lag, we choose the maximum value of each random cross-correlation to compute the thresholds from. The result is that at 95% confidence, 5% of random cross-correlations have any points that are significant when using the peak test, while ~50% have significant points when using the point test. This shows that using an improper significance test can result in many spurious significant points on any given cross-correlation. All significance thresholds shown here are computed using the peak test.

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227



1900 1920 1940 1960 1980 2000 2020 *Figure 2.* The effect linear detrending has on observed EOFs in the NP. For observations where no detrending occurs, A) and C)
show the first EOF spatial pattern and time series respectively. B) and D) show the same but for the second EOF. For linearly
detrended observations, E) and G) show the first EOF spatial pattern and time series respectively, while F) and H) show the same
for the second EOF. I) shows the time series from C) (linearly detrended) and H) together to show their similarity (93%
correlation, significant at the 99% threshold). The red curves are the same time series, while the blue curve in I) is equivalent to
the time series in C) but with the linear slope removed.



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Figure 3. EOF variance explained of MMLEA ensemble members. Top row form left to right: NA-EOF1 and NP-EOF1. Bottom row form left to right: NP-EOF2 and NP-EOF3. Black lines represent observed values. Colors correspond to models as follows:
dark red (CANESM2), orange (CESM), gold (CSIRO MK36), green (GFDL CM3), blue (GFDL ESM2M), and purple (MPI).
For each model, two sub-columns are shown: the full, or total, variability on the left and the internal variability on the right. Each circle/star represents a single member. Xs represent pre-industrial control values.

252 Results: Model Assessment

Here, we want to assess how realistic the simulated modes are relative to the observed, 253 including how much of the observed variability can be explained by forcing and internal 254 variability. One way we can analyze this is to compare the fraction of explained variance for 255 each EOF to the corresponding value for observations, as shown in Figure . EOF 1 in each basin 256 shows some inter-model spread, but the observed value falls within the internal range for each 257 model. For both basins, EOF 1 generally explains between 30-40% variance, while EOF 2 258 explains 15-20%, and EOF 3 around 10-15%. In some cases for each member, the full variability 259 explains more variance than its corresponding internal variability. This can be attributed to the 260 external forcing that is present in the full but not the internal. Generally, the internal variability 261 should agree with the observations better than the full variability, although both appear relatively 262 similar to the observed values. The MMLEA looks qualitatively similar enough to observations 263 for this metric to proceed. 264

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266 We also assess the realism of the MMLE subjectively by looking at the EOF spatial patterns. Supplemental Figure 2 shows the observed EOFs and the composite mean EOFs for 267 each model. The MMLEA NA EOFs all roughly share a similar tripole spatial pattern to the 268 observed first NA-EOF. The magnitude of the gradients across all ensembles differs from the 269 observations, however. Extending from the Gulf Stream region, the MMLEA models generally 270 show an opposite trend to the rest of the basin (cooling in the Gulf Stream region when the rest 271 272 of the basin is warming), whereas the observations show uniform warming or cooling. In the NP, NP-EOF1 shows good agreement, all showing the classic PDO pattern. The second observed 273

EOF differs significantly from those in the MMLEA. The observed pattern shows uniform

warming or cooling across the entire basin, while the third observed EOF shows a typical VM

pattern. The MMLEA appears to skip over the uniform warming/cooling spatial pattern, showing
 only the PDO as EOF1 and the VM as EOF2 in most members. One possible explanation for this

only the PDO as EOF1 and the VM as EOF2 in most members. One possible explanation for thi is that the observed EOF2 represents non-linear features of external forcing, which would

successfully be removed in an MMLEA model by subtracting the ensemble mean, but not in

observations through the linear detrending method. It is interesting that this is different in the

Atlantic and Pacific, suggesting that method for removal of the forced signal is basin dependent.

Generally, the MMLEA members do show internal modes similar to the observed modes (when

comparing observed NP-EOF3 to MMLEA NP-EOF2), suggesting that the models simulate

variability realistically enough to analyze potential mode interactions.

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Figure 4. Cross-correlations of NA-NP mode relationships. From left to right: AMV (NA-EOF1) vs. PDO (NP-EOF1)
relationship, AMV vs. NP-EOF2 relationship, and AMV vs. VM (NP-EOF3) relationship. From top to bottom: Observations
with no detrending, CESM-LE (from MMLEA) full variability, CESM-LE ensemble mean, Observations with linear detrending,
CESM-LE internal variability, and CESM pre-industrial control variability. Black horizontal lines are 95% statistical significance
thresholds as calculated by the "peak" test. Vertical ticks show where cross-correlations are significant. For positive (negative)
lags, the NA (NP) mode leads. All othe MMLEA large ensembles show qualitatively similar results.



294
 295 Figure 5. Cross-correlations of MMLEA internal only NA-NP mode relationships. From left to right: AMV (NA-EOF1) vs. PDO
 296 (NP-EOF1) relationship, AMV vs. NP-EOF 2, and AMV vs. VM. Black horizontal lines are 95% statistical significance
 297 thresholds as calculated by the "peak" test. Vertical ticks show where cross-correlations are significant. For positive (negative)
 298 lags, the NA (NP) mode leads. 269 MMLEA members are utilized for a total of 807 relationships. Only 11 (<2%) have
 299 statistically significant points.

300 Results: Relationship Analysis

Figure 4 shows a series of NA-NP relationships from observations and the CESM Large Ensemble (CESM-LE), one of the MMLEA ensembles. The top row shows the non-detrended (or full) observed EOF cross-correlations. Without detrending, the first EOF of each basin captures the global warming or externally forced signal. The signal in each basin is clearly connected, with significant points along a peak at 0 lag. This can be interpreted as global warming affecting each basin in a very similar way. The other relationships, which capture the internal modes in the NP, show no statistically significant connection.

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309 The next row shows the relationships for full members of the CESM-LE. Again, most of 310 these relationships show a peak near 0 lag for all three relationships. This suggests that external forcing affects all of the first three NP EOFs, such that a nearly significant peak appears near 0 311 lag for most members. When this global warming signal is shown by itself as the ensemble mean 312 in the third row, the same connection as in the observed forced signals appears, with the forced 313 signal in each basin being significantly correlated. Note that the ensemble mean EOFs 2 and 3 do 314 not have a physical meaning and can be neglected. Also, averaging across all ensemble members 315 may mutually cancel out polarized clusters of members, so caution must be used when 316 computing ensemble means (Bellucci et al. 2017). 317

So far, we have shown that the combined internal and forced relationships are statistically 319 significant, particularly only the forced relationships. Rows four through six analyze only the 320 internal relationships to determine whether connections between the internal modes exist or not. 321 Row four shows the linearly detrended observed relationships. The AMV-PDO relationship 322 (hereafter all relationships referred to as AMV-PDO, AMV-VM, etc.) shows similar results as 323 Wu et al. (2011), Marini and Frankignoul (2014), and Nigam et al. (2020), and both the AMV-324 PDO and AMV-NP-EOF2 show similar results as d'Orgeville and Peltier (2007) and Zhang and 325 Delworth (2007). However, we show different significance thresholds as per the peak test. The 326 sign of the correlation and the precise lag at which the maximum correlation occurs may slightly 327 vary from study to study due to differing methods, particularly the sign of the EOF output and 328 329 filter used. Linear detrending is also not ideal for removing the observed forced signal (Frankignoul et al. 2017), which may allow the forced signal to remain and cause a spurious 330 higher correlation as seen in AMV-NP-EOF2. 331

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To account for the inadequacy of linear detrending, the MMLEA internal relationships and the 333 CESM PI are shown in rows five and six respectively. Only one of these 120 members show a 334 significant relationship, suggesting that any internal connection between the NA and NP basins is 335 indistinguishable from random noise. All other MMLEA ensembles show qualitatively similar 336 results. Figure 5 shows the internal relationships for all six MMLEA ensembles analyzed, with 337 only 11 out of 807 (<2%) relationships having any statistically significant points. The same 338 concept of external forcing driving significant relationships appears in all of the ensembles, 339 despite their various differences. 340

341 Discussion 342

Our results suggest that, for the NA and NP basins, an internal connection between the 343 two does not need consideration as a potential source of variability. These findings may also 344 have broader implications regarding the roles of external forcing and internal variability as 345 346 drivers of climate modes. Present theories on climate mode drivers focus on varying roles for the ocean, atmosphere, internal variability, and external forcing (e.g. Clement et al. 2015, Newman 347 et al. 2016, Wills et al. 2018, Zhang et al. 2018, O'Reilly et al. 2019, Zhang et al. 2019, etc.). 348 External forcing is particularly challenging, due to the direct/linear and indirect/non-linear 349 effects on climate variability (Frankignoul et al. 2017, Li et al. 2020). Additionally, there is 350 debate about the role and importance of these internal modes in an increasingly forced global 351 352 climate (Ting et al. 2009, DelSole et al. 2011, Haustein et al. 2019, Mann et al. 2020, etc.). 353

354 Additionally, our results suggest that traditional observed AMV and PDO definitions contain at least some external forcing. Care must be taken when using these mode definitions to 355 properly remove the forced signal so as to isolate the internal variability. However, this is a 356 challenging task outside of the realm of large ensembles. Tools such as the MMLEA will be vital 357 358 in making progress toward isolating the forced response in observations, as they can possibly average out various model differences and provide a closer analog to the observed forced signal. 359 These modeling tools are especially useful due to the relatively short observed period, which 360 may not be sufficient to adequately observe variability on multi-decadal timescales, such as those 361 studied here. 362

363

Finally, our findings do not rule out other regions or modes driving variability in the NA 364 and NP. Other mode relationships have been shown to exist, such as how ENSO helps drive the 365

- 366 PDO (Newman et al. 2003). Further work can include a complex matrix of potential relationships
- between global modes as in Shin et al. (2010) or using global EOFs such as Nigam et al. (2020).
- 368 Methods presented here can assist in a thorough decomposition of sources of variability in a
- 369 particular region. This can lead to better understanding of variability drivers, which can
- 370 ultimately result in improved climate models and more accurate climate forecasts.
- 371

372 Acknowledgments, Samples, and Data

- The authors gratefully acknowledge funding from the NSF Climate and Large Scale Dynamics
- program and NOAA Climate Program Office that supported this work. We also acknowledge all
- modelling groups and the US CLIVAR Working Group on Large Ensembles for making their
- data available in the Multi-Model Large Ensemble data repository. We thank Brian Mapes and
- 377 his Applied Data Analysis class for inspiration and discussions on proper statistical analysis.
- 378
- 379 Data Access: <u>https://www.cesm.ucar.edu/projects/community-projects/MMLEA/</u> (MMLEA);
- 380 <u>https://www.metoffice.gov.uk/hadobs/hadisst/</u> (HADISST).
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