Regional Flood Frequency Analysis Using Physics-based Hydrologic Modeling

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Abstract

We investigate the validity of implicit assumptions in regional flood frequency analysis (RFFA) using Monte Carlo-style simulations of three distributed hydrological models forced with rainfall events generated using stochastic storm transposition. We test the long-standing assumption that for a set of sites within a region, *physical homogeneity* — defined in terms of the variability of meteorological inputs, the physics of runoff generation, and runoff routing — implies *statistical homogeneity* of peak flows defined in terms of the existence of a common underlying statistical distribution with parameters that can be inferred using information from neighboring sites. Our modeling results do not support this assumption, with potentially important implications for RFFA methodologies and for the very definitions of homogeneity. We show that statistically homogeneous rainfall does not translate into predictable peak flow distribution parameters across drainage scales. Specifically, we show that changes in the coefficient of variation and skewness of peak flows cannot be inferred from upstream area alone, making popular regionalization techniques such as the index-flood method and quantile regression inadequate approximations for flood frequency estimation. Our findings are consistent across the three hydrological model formulations, lending confidence that our conclusions are not an artifact of epistemological model decisions. Finally, we argue that our methodology can serve as a framework to test new proposed empirical RFFA methods, and that it opens the door to a unified physics-informed framework for prediction of flood frequencies in ungauged basins embedded in gauged regions.

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2	Using Physics-based Hydrologic Modeling					
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13	Key Points					
14	1. We test the validity of implicit assumptions in empirical regional flood frequency analysis					
15	using three hydrological models and stochastic storm transposition					
16	2. We show that a link between physical homogeneity in hydrological processes and					
17	homogeneity of statistical distributions in peak flows cannot be assumed					
18	3. Our work provides a framework to establish a physics-based regional flood frequency					
19	analysis					
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42 Keywords: Regional flood frequency analysis (RFFA), peak flow scaling, stochastic storm
43 transposition, physical homogeneity, statistical homogeneity, hydrologic modeling, Bulletin 14D.

44 Plain Language Summary

45 So-called "regionalization" techniques are useful for estimating design peak river flows such as46 the "100-year flood" at locations that lack flow measurements. Although these strategies are

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47 widely used in flood protection, several associated assumptions are unproven because of 48 limitations in data availability in terms of spatial sparsity and record lengths. This study tests some 49 of the long-standing assumptions in regional flood frequency analysis (RFFA) using hydrologic 50 simulations that include not only a large number of observed rainfall scenarios, but also by the 51 support of three different hydrologic model formulations. Our findings indicate that several of the 52 assumptions used in RFFA are incorrect, highlighting the need to consider new strategies that 53 benefit from physics-informed hydrologic process modeling.

54 1. Introduction

55 The systematic power-law scaling of riverine flood peak flows (Q) with respect to drainage area (A), first identified by Fuller, (1914), has since become a well-established feature in hydrological 56 57 research and practice. This scaling describes the regional change of characteristic flows (e.g. mean annual floods, bank-full flows, 100-year floods) across spatial scales as $Q = \alpha A^{\theta}$, with α and θ 58 59 known as the scaling intercept and scaling exponents, respectively. This type of scaling arises in 60 scale-invariant systems with self-similarity properties (V. K. Gupta et al., 2007) and seems to 61 reflect a fundamental symmetry of nature (Schroeder, 2012). This result gave rise to empirical 62 frameworks for Regional Flood Frequency Analysis (RFFA) to predict peak flow distributions at 63 ungagged locations. RFFA, in its many forms, is the most widely accepted technique for estimation 64 of annual flood quantiles (or also known as peak flow quantiles) in the United States and 65 elsewhere, with results being used for infrastructure design, land use planning, and insurance.



Figure 1. a) Location of long-term streamflow gauges in the state of Iowa, USA, color coded according to three homogeneous regions defined by Eash, (2001), and b) estimated 100-year flood peak for each gauge location. Both panels also show the locations of two gauges in the Turkey river watershed (dark outlined circles).

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67 RFFA techniques generally employ both sparse data sets and unproven assumptions that seem to 68 hold in idealized individual rainfall-runoff events (Furey et al., 2016; Perez et al., 2019b) and for 69 highly idealized river network structures (e.g. Gupta et al., 1996; Menabde & Sivapalan, 2001). In 70 Iowa, for example, 307,714 km of rivers are gauged at only 140 unevenly distributed locations 71 (Figure 1a), which tend to coincide with the most densely populated municipalities. This leaves 72 the large majority of the state's waterways—and riverine communities—ungauged and dependent 73 on empirical RFFA techniques to determine flood protection strategies. The power law scaling 74 pattern in Figure 1b between the estimated 100-year flood using Bulletin 14B methodologies is 75 evident, but it's also remarkable that for a typical ungauged watershed the smallest estimate for 76 the 100-year flood is 6 times smaller than the largest estimate. Using a regression line through the cloud of points to estimate the 100-year flood for a generic 100 km² basin would certainly 77 78 underestimate flood levels half of the time and overestimate the other half.

While the steps and statistical inferential tools within particular empirical RFFA methodologies
vary from country to country (England et al., 2019; Kjeldsen, 2011; Meigh et al., 1997), they all

81 rely on a central assumption: the peak flow distribution at an ungauged site can be inferred from 82 observations at nearby locations that are hydrologically similar. The statistical inference process 83 also relies on an underlying set of assumptions that remain largely untested including 84 independence of peak flow data, stationarity, statistical homogeneity, scaling of distribution 85 parameters, and the appropriateness of statistical estimators. In Canada, the United Kingdom, and 86 the European Union, the preferred RFFA methodology is the index flood method, or simple-87 scaling as defined by Gupta et al., (1994; see Section 2.1), and the preferred underlying distribution 88 of peak flows is the generalized extreme value (GEV) distribution (Reed & Robson, 2008; Zhang 89 & Stadnyk, 2020). These methodologies are detailed in national reports such as the Review of 90 Applied European Flood Frequency Analysis Methods (Kjeldsen, 2011). In the United States, the 91 US Geological Survey (USGS) recommends the use of the Log-Pearson Type III (LP3) 92 distribution, and the more flexible regional quantile regression method (IACWD, 1982), which 93 has been described as exhibiting multi-scaling (Gupta et al., 1994; see Section 2.1). More 94 sophisticated geostatistical methods have been proposed but are note widely used (e.g. Ouarda, 95 (2013); Salinas et al., (2013)).

96 We argue in this paper that three modern developments in theoretical and computational hydrology 97 have opened the door to investigating the validity of the empirical formulas employed in RFFA in 98 a systematic fashion. First is the identification of power laws that connect the peak flows from 99 extreme rainfall-runoff events in nested catchments as a function of drainage area (Ayalew et al., 100 2014; Gupta, 2004; Gupta et al., 2010; Ogden & Dawdy, 2003; Perez, et al 2019a; Perez, et al 101 2019b). Second is the development of simplified but realistic distributed hydrological models, 102 along with efficient algorithms to integrate mass balance and water transport differential equations, 103 that can predict, simultaneously, peak flows resulting from storm events at multiple locations. The third is the deployment of weather radars that observe precipitation fields at high spatial and
temporal resolution, which in turn can inform stochastic methods—such as the Stochastic Storm
Transposition (SST) framework used in this study—to generate many realizations of extreme
storms.

108 Based on these three developments, recent work by (Perez, et al 2019b) suggested that even in a 109 medium-sized catchment (~4,000 km²) there is a complex spatial pattern in the third moment of 110 the peak flow distribution (skewness) that cannot be explained by upstream area alone. This 111 finding implies that most of the common RFFA models (e.g. the so-called Index Flood Method 112 and regional quantile regressions) contain significant epistemic errors, thus casting doubt on 113 several RFFA assumptions. Although these results highlighted the need to consider new strategies 114 to better capture peak flow distributions across scales, Perez, et al (2019b) was limited to a single 115 hydrologic model formulation and thus contained inherent epistemic errors in the representation 116 of hydrologic processes; it has been argued that this can give rise to misleading conclusions 117 (Beven, 2018).

118 Our aim in this paper is not only to generalize the findings of Perez et al., (2019b) by limiting these 119 epistemic errors and to correct several more minor methodological shortcomings, but also to 120 further investigate the implicit and explicit assumptions of existing empirical RFFA methods. In 121 particular, we seek to test the long-standing assumption that for a set of sites within a region, 122 physical homogeneity — defined in terms of the spatial variability of meteorological inputs, and the physics of runoff generation and runoff routing — implies statistical homogeneity of peak 123 124 flows— defined in terms of the existence of a common underlying statistical distribution, and the 125 ability to infer local distribution parameters (e.g. distribution moments). To this end we 126 implemented three distributed hydrological models in a medium size basin (~4,000 km²). The three

127 models are all able to simulate streamflow dynamics at multiple internal sub-catchments 128 simultaneously but vary in the level of complexity of their physical processes, spatial resolution, 129 and spatial variability of model parameters. The three models are forced with a very large number 130 of extreme storms created using SST to provide an approximation of the population of realistic 131 peak flows, which we analyze statistically (We henceforth refer to this very large sample as a 132 population). The use of SST helps us address issues related to sample size, while the use of the 133 three hydrological models helps reduce biases caused by epistemic errors. Three overarching 134 questions drive this analysis: (i) What is the link between physical homogeneity and statistical 135 homogeneity among locations within a watershed or region?, (ii) what are appropriate 136 simplifications for the scaling of the statistical moments of peak flow distributions across scales?, 137 (*iii*) can the peak flow distribution parameters at an ungauged site be inferred from measurements 138 at neighboring sites?

139 In Section 2 (*Background*), we briefly review the most relevant research that led to the assumptions 140 that we are testing. In Section 3 (*Study Area*), we describe the study watershed and the parameters 141 used by each model. In Section 4 (*Methodology*) we describe our experimental setup including the 142 three hydrological models, the SST methodology, and the methods that we use to analyze model 143 outputs. In Section 5 (Results), we analyze the rainfall, runoff production, and peak flow 144 distributions generated by our numerical experiment. In Section 6 (Discussion), we detail how our 145 results inform the assumptions we are testing and the potential implications. Finally, in Section 7 146 (*Conclusions*) we summarize our major findings and suggest a road map to replace the current 147 empirical RFFA framework by one informed by decades of accumulated hydrological knowledge 148 of the physical characteristics and processes involved in the rainfall-runoff and routing 149 transformations responsible for riverine flooding.

150 **2. Background**

151 Dawdy et al., (2012) details the origins of RFFA equations in the United States, including the 152 adoption of both power laws that relate flood quantiles to explanatory variables and the LP3 153 distribution for estimating flood quantiles at gauged sites. Their review reveals that the steps 154 involved in RFFA are largely ad hoc but justified from the available empirical observations of 155 available flood data. They also explain RFFA's reliance on the existence of statistical homogeneity 156 (defined in Sec. 2.1) of flood data. They emphasized the need of a more physically-based 157 framework to connect event scale observations of peak flow scaling to the scaling of annual flood 158 quantiles. Their rationale is that unlike at the event scale—where the connection between the 159 physical phenomena that give rise to peak flows along a river network can be revealed—flood 160 quantiles and the physical phenomena that control them have proven more elusive due to issues of 161 statistical estimators, sample sizes, and other procedural issues. From our point of view, a well-162 developed physically-based framework of RFFA should provide estimates of uncertainty 163 introduced by epistemic errors in the description of physical processes, sample size, 164 nonstationarity, and techniques for including all available hydrological data in a region.

165 **2.1. Statistical Homogeneity**

At the core of RFFA lies the definition of statistical homogeneity of peak flow data and under what circumstances it exists. The simplest example of statistical homogeneity can be given in the context of the index flood method (IFM). In it, the underlying assumption about the distribution of peak flows is that the ratio of a peak flow quantile to an index flood (e.g. mean or median annual flood, bank full flood) is constant over a homogeneous geographic region (Kinnison & Colby, 1945). This feature is known as simple scaling of statistical distributions. In this context, a set of catchments is said to be homogeneous if the coefficients of variation $CV = \sigma/\mu$ of their peak flow

173 distributions are the same and if higher order moments such as skewness are constant spatially. 174 Gupta et al. (1994) developed a more general framework for the analysis of peak flow data, called 175 multi-scaling, that did not require a constant coefficient of variation. Instead, they defined 176 statistical homogeneity more generally: "A homogeneous region has, for all gaging locations 177 within the region, flood peaks that follow probability distributions which are rescaled versions of 178 one another, based only on drainage area and nothing else." Gupta et al. (1994) specifically showed 179 that under the assumption that the skewness is constant in a region, regressions of flood quantiles 180 with different exceedance probabilities p would scale with respect to drainage area with exponents that depend on the quantile value, that is $Q_p = \alpha_p A^{\theta_p}$, and that CV would systematically increase 181 182 for small catchments up to an area threshold and then decrease beyond it, consistent with USGS 183 flood data. The results by Gupta et al. (1994) provided a theoretical footing for the methodologies 184 in Bulletin 17 (WRC, 1977) and Bulletin 17B (IACWD, 1982) of using similarity in regional 185 skewness to define homogeneous regions that continue to be used today in the most recent release 186 of the Bulletin 14C (England et al. 2019).

187 While the assumptions of simple or multi-scaling have been adopted to describe peak flow 188 variability over a region of interest, the reality is that these assumptions cannot be rejected by 189 observations, mainly because of large sampling errors—that is, severely limited sample sizes— 190 that obscure the true variability of the peak flow distributions over a region. (Perez et al., 2019b) 191 showed this via minimization of sampling errors through the generation of a large number of 192 realistic rainfall-runoff events. Their work demonstrated that skewness exhibits a complex spatial 193 structure that cannot be explained solely by drainage area, in contrast with the definition of simple 194 scaling and multi-scaling. Our study builds on that approach.

195 **2.2. Physical Homogeneity**

196 A parallel definition-physical homogeneity-is also prevalent in RFFA literature. It refers to 197 similarity in the geophysical characteristics of catchments and the meteorological drivers such as 198 rainfall and evapotranspiration. For, example, the extent of Region 1 in Iowa (Fig. 1) coincides 199 with the Des Moines Lobe, a landform resulting from the last glacial episode. Similar examples 200 are easy to find in other states' RFFA studies. A connection between physical homogeneity and 201 statistical homogeneity is widely assumed and has been spelled out in the empirical region of 202 influence (ROI) RFFA framework introduced by Burn (1990) and tested in Zrinji & Burn (1994) 203 and Burn (1997). Burn (1997) used a dissimilarity measure introduced by (Webster & Burrough, 204 (1972) and tested if the peak flow distributions of the catchments could be considered statistically 205 homogeneous using the H metric introduced by Hosking & Wallis, (1993). The methodology 206 requires "the choice of a threshold value that functions as a cut-off point for a dissimilarity 207 measure. All sites that have a dissimilarity measure with the catchment of interest that is greater 208 than the threshold value are excluded from the region of influence for that particular catchment" 209 (Burn, 1997). Catchments that pass the test are expected to be statistically homogeneous and are 210 considered appropriate for pooling data.

ROI methods have been refined to better characterize catchment dissimilarity and the definition of statistical homogeneity using L-moments. Hall & Minns, (1999) used non-linear clustering methods to define homogeneous regions using a variety of geomorphological descriptors. Viglione et al., (2007) investigated statistical homogeneity measures and showed that L-moments based tests are more powerful when the samples are slightly skewed, while rank tests have better performance in cases of high skewness. They also show that many of the indices used to determine statistical homogeneity lack the power to discriminate between homogeneous and heterogeneous

218 regions. Recently, Ilorme & Griffis, (2013) introduced more sophisticated measures for physical 219 homogeneity and statistical similarity H, while Basu & Srinivas, (2014) used non-linear kernels to 220 divide the attribute space into groups that are not necessarily continuous in space. The attributes 221 include physiographic, land use/land cover, drainage, climate and flood seasonality descriptors of 222 the watersheds contributing to flow at the sites, and geographical location attributes. Underlying 223 all of this work is the assumption that physical homogeneity implies statistical homogeneity. Even 224 more recently (Zhang et al., 2020) proposed "a novel region revision procedure to complement the 225 well-known region of influence and L-Moments techniques that automates the identification of 226 homogeneous regions across continental domains" that could be used in the Canadian statistical 227 flood estimation guideline under the FloodNet project (www.nsercfloodnet.ca).

228 **3.** Study Area and Data Sources

We select the Turkey River watershed (4,385 km²), located in northeast Iowa in the midwestern 229 230 United States, as a case study to address our objectives. Turkey River is located in Iowa's Region 231 2 according to the USGS RFFA report by Eash (2001) (Figure 2). This watershed is the same one 232 used by Perez et al. (2019b) and it also has been widely used in hydrologic modeling frameworks 233 to understand the control of different hydrologic process in the peak flow response. Examples 234 include the evaluation of the importance of the level of storm spatial and temporal detail (Zhu et 235 al., 2018) and climate-driven shifts in the seasonality of snowmelt and soil moisture (Yu et al., 236 2019) in peak flow distributions.



Figure 2. The Turkey River basin and the corresponding variability of soil types and landcover in the region.

238 4. Methodology

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239 Our experimental design seeks to generalize the findings of Perez et al. (2019b) by addressing four

- 240 RFFA-relevant questions about the distributions of peak flows along a river network:
- 1. Can the moments of the peak flow distribution for an arbitrary location in the river network be
- 242 inferred using data from other sites within the network?
- 243 2. Are power-law equations sufficient to describe the scaling of peak flow distribution moments
- 244 with respect to drainage area?

3. Is the coefficient of variation ($CV = \sigma/\mu$) or the skewness (γ) of peak flows constant across drainage scales and throughout a region under conditions of physical homogeneity in meteorology, runoff generation, and runoff transport mechanisms?

4. Are the IFM (simple scaling) or the quantile regression (multi-scaling) frameworks appropriate
for describing flood data in a purportedly physically homogenous region?

250 To this end we used three different hydrological models to simulate peak flows across the river 251 network associated to an equal number of storms derived using the SST framework described 252 below. Our experiment is designed to ensure that rainfall distributions are homogeneous across the 253 river network. To accomplish this, our experiment creates an unprecedented data set of 50,000 254 spatiotemporally realistic rainstorms, comprising 10,000 synthetic years that assume that 5 storms 255 occur every year, which are used to simulate annual peak flows at many locations in the river 256 network using each hydrological model. This is a five-fold increase over the number of storms and 257 flood events considered in Perez et al. (2019b), and it was selected to ensure homogeneity in 258 rainfall inputs and to minimize sampling uncertainty in order to approximate the population of the 259 peak flows.

260 4.1. Stochastic Storm Transposition (SST) Framework

To address our research questions, we created 50,000 rainfall scenarios comprising 10,000 synthetic years. The needs of our study impose two basic requirements: *i*) individual rainstorms must have realistic spatiotemporal structure within rainstorms to ensure realistic multiscale depiction of flood events across the river network, and *ii*) each portion of the watershed must "see" approximately identical rainfall distributions so that sampling error associated with rainfall inputs is minimal. Creating such an event set is nontrivial—existing high-resolution precipitation datasets generally span only several decades and thus suffer from high sampling error (Wright, 2018), while 268 most stochastic rainfall generation techniques either lack spatiotemporal detail or struggle to
269 reproduce observed extreme rainfall rates (Furrer & Katz, 2008; Wright et al., 2020).

270 To address these issues, we couple the open-source SST software RainyDay (Wright et al., 2017) 271 with 17 years (2002-2018) of Stage IV gage-corrected radar-based precipitation (Du, 2011; Lin, 272 n.d.) to generate 50,000 storm scenarios for Turkey River. SST generates these scenarios via 273 temporal resampling and spatial transposition of observed precipitation events from the 274 surrounding transposition domain. Wright 2020 reviewed the development and applications of 275 SST since its inception. SST, as implemented in RainyDay, ensures that the aforementioned first 276 requirement (e.g. realistic intrastorm spatiotemporal structure), to the extent that the Stage IV 277 observations reflect the true precipitation structure (resolution-related effects remain, e.g. Zhu et 278 al., (2018). The second requirement is satisfied through the following steps:

The transposition domain (Figure 2a) refers to a region over which observed extreme
 precipitation should be statistically similar. While we do not formally verify this regional
 rainfall homogeneity here, it's existence or lack thereof will have no influence on the
 results of this study due to other aspects of the SST methodology. The transposition domain
 used here spanned 94.25-89.25°W and 40.5-45.5°N.

2. RainyDay used April-November Stage IV precipitation to create a "storm catalog"
285 consisting of the 320 most intense precipitation events within the domain. "Most intense"
286 is defined with respect to 72-h duration rainfall accumulations over areas the size and shape
287 of the Turkey River watershed. This same duration, as well as similar a time period and
288 storm catalog size, has been used in previous SST-based FFA studies in Turkey River
289 (Perez et al., 2019b; Yu et al., 2019, 2020).

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290 3. RainyDay randomly selected *k* storms from the catalog to represent a synthetic year. The 291 number of storm arrivals *k* was calculated by RainyDay using the Poisson distribution with 292 rate parameter λ , defined as the ratio of total number of events in the storm catalog to the 293 number of years in the record (i.e., $\lambda = 320/17 = 18.82$ storms yr⁻¹).

4. RainyDay transposed the selected *k* storms randomly (with uniform probability) within the
transposition domain and computed the precipitation field over the watershed, thus creating
a synthetic year of precipitation scenarios for the study watershed. These RainyDay-based
precipitation events have the identical spatial and temporal resolution and structure as the
input Stage IV data (hourly, ~5 km).

299 RainyDay repeated steps 3-4 to create 10,000 synthetic years, each comprised of k precipitation 300 events in Turkey River. Within each synthetic year, the largest five of the k 72-h rainfall events 301 with respect to the size and shape of the watershed were retained for hydrologic simulations. 302 Testing showed that five events per synthetic year (i.e. 50,000 storms in total) was sufficient to 303 generate the simulated annual flood peaks across the entire range of subbasin scales represented 304 (i.e. up to 4,385 km²). This constitutes a methodological improvement over Perez et al (2019b), 305 which retained only the largest single event per year, for a total of 10,000 events. Follow-up testing 306 revealed that using only one event per year could lead to spatial biases in the resulting flood 307 simulations, since the storms that produce annual flood peaks on small headwaters tributaries tend 308 to have different characteristics from those that produce annual peaks along the river's mainstem.

4.2. Description of Three Hydrological Models

310 **4.2.1. HLM**

311 The Hillslope-Link Model (HLM) is a distributed hillslope-scale rainfall-runoff model that 312 partitions the landscape into individual control volumes following the landscape decomposition 313 into hillslope areas outlined in Mantilla & Gupta, (2005). The model is parsimonious, using 314 ordinary differential equations to describe transport between adjacent control volumes (Krajewski 315 et al., 2017; Mantilla, 2007; Mantilla et al., 2006). All flows in HLM are downstream, all the 316 closure relations are unique storage-flow relations, and there are no feedback loops of surface and 317 subsurface flows. In addition, the parameters describing vertical and horizontal flows in hillslopes 318 are assumed to be the same throughout the watershed. HLM is the simplest and more spatially 319 homogenous of the three models used.

320 4.2.2. WRF-Hydro

321 The physics-based distributed Weather Research and Forecasting Hydrological modeling system 322 (Gochis et al., 2020) calculates energy and moisture fluxes over square grid cells using the Noah-323 MP land surface model (Niu et al., 2011). Noah-MP is coupled to modules for surface and 324 subsurface routing over the terrain and through river channels. While WRF-Hydro is highly 325 flexible, we use the same configuration as in National Water Model (NOA, 2016) which is used 326 by National Oceanic and Atmospheric Administration (NOAA) for operational flood forecasting. 327 One exception is that snow processes are not used in this study to render WRF-Hydro results more 328 comparable to the other two models, which lack snow process representations. The model was 329 calibrated manually with respect to long-term water balances, infiltration processes, and flood 330 hydrographs. See (Yu et al., 2020) for more information on this WRF-Hydro configuration and 331 calibration procedures. From here on, we will use the abbreviation "WRF" to refer to WRF-Hydro 332 in figures and tables.

333 4.2.3. Watershed Modeling Framework (WMF)

334 The Watershed Modeling Framework (WMF) is a Python-Fortran programable hydrological 335 modeling and analysis tool (Velásquez et al., 2020; Velez et al., 2018). The hydrological model of 336 WMF is a modification of the hydrological TETIS model (Francés et al., 2007; Velez, 2001). WMF 337 represents the hydrological processes using grid elements. Each element has five storages 338 representing the runoff, canopy, subsurface, aquifer, and channel processes. The water moves 339 vertically across storages and horizontally from an upstream element to a downstream element. 340 The vertical movement is mainly governed by the evapotranspiration rate, the infiltration rate, and 341 the percolation rate. The downstream movement uses the rill approximation to represent the runoff 342 (Foster et al., 1984) and a non-linear equation to represent the subsurface (Kubota & Sivapalan., 343 1995). The channel routing follows the GKW (Geomorphological Kinematic Wave) approach 344 developed by Velez (2001). We use the information of POLARIS (Chaney et al., 2016) to 345 distribute the parameters of the soil at a resampled resolution of 150m.

Table 1. Comparison of main attributes for the three hydrological models used in our study.

	Equations	Spatial resolution	Rainfall-	Subsurface	River Network	
	Class		Runoff		Transport	
HLM	ODE	Hillslopes	Top-layer	Linear Reservoir	Non-linear	
		~ 0.018 km ²			reservoirs	
WRF-	ODE-PDE	Square Cells	Noah-MP	Vertical Richard's	Muskingum-Cunge	
Hydro		1 km ² (land column); ~		Equation (upper); 2D-		
-		0.063 km ² (overland flow)		Darcy's Flow (lower)		
WMF	DE	Square Cells	Phi-index +	Kubota & Sivpalan	Kinematic Wave	
		$\sim 0.022 \text{ km}^2$	exfiltration	(1995)		

347 **4.3.** Process-based framework for flood frequency analysis

Continuous simulations were conducted with all three models for the period 2002-2018 to (*i*) determine their ability to simulate streamflow variability at five gauged locations in Turkey River, and (*ii*) to create a daily-scale "database" of initial conditions to be used to initialize the models for each of the 50,000 flood events. These continuous simulations, as well as the SST framework described in Section 4.1, were forced using Stage IV precipitation and other meteorological
forcings (which varied by model) from NLDAS-2 (Mitchell et al., 2004).

The process-based flood frequency simulation framework is the same as described in Yu et al. (2019) and Yu et al. (2020); it is only briefly described here. The initial condition for each event was randomly selected amongst daily model snapshots. The selection of the date corresponding to the initial condition was preserved for the three models (i.e. all three models start each event with similar initial conditions). The goal of this process was to have all the models on a similar footing at the beginning of each rainfall-runoff event to minimize the effects of differences in initial conditions in our analysis.

361 **4.3 Evaluation of physical and statistical homogeneity**

362 **4.3.1.** Meteorological homogeneity and realism

363 *Meteorological homogeneity*, or more specifically, homogeneity in the probability distributions of 364 extreme rainfall across a range of durations, is a primary focus of regional rainfall frequency 365 analysis (Svensson & Jones, 2010), a branch of research and practice that shares a common lineage 366 and methods with RFFA. Meteorological homogeneity is also a prerequisite for physical 367 homogeneity in runoff production. The SST methodology (Sec. 4.1) was deliberately conducted 368 to ensure that the populations of rainfall events that are experienced by each computational unit 369 (e.g. grid cell or hillslope) in the hydrologic models are effectively identical. Furthermore, by 370 transposing observed storms, SST preserves the spatial and temporal structure of observed rainfall, 371 meaning that homogeneity will be consistent across scales. It should be noted that the relevance of 372 existing metrics and tests for evaluating rainfall homogeneity (e.g. Hosking & Wallis, 1993) is

questionable, since such tests were developed under the presumption of limited sample sizes, asopposed to the very large samples used in this study.

In addition to homogeneity, the population of rainfall events used in our hydrological simulations must be relatively realistic in terms of frequency and magnitude. Otherwise, conclusions arising from the simulations would have limited relevance to real-world conditions. This realism is evaluated by creating rainfall intensity-duration-frequency (IDF) curves at the pixel scale, following Wright et al. (2017). These SST-based IDF curves are then compared against published rain gage-based IDF estimates from NOAA's Atlas 14 project (Bonnin et al., 2006).

381 **4.3.2.** Physical homogeneity of runoff generation

382 Physical homogeneity of runoff generation is typically approached from the point of view of the 383 spatial variability of physical properties associated to runoff production (e.g. hydraulic 384 conductivities, terrain roughness, land use, land cover, and soil types) and their representation 385 through relevant model parameters. This can be misleading, however, because it is difficult to 386 determine what is an acceptable range of variability in parameter values that can influence the 387 behavior of nonlinear equations that can lead to large differences in runoff production in the 388 models. Instead of examining homogeneity in the context of model parameters, therefore, we will 389 examin it in the context of the event-based runoff coefficients RC, defined by the ratio between total event rainfall and total surface runoff during the event $RC = \sum q_s / \sum P$. The runoff coefficient 390 391 will vary from event to event and amongst model control volumes (e.g. hillslope to hillslope or 392 grid cell to grid cell). While it will depend primarily on the total amount of rainfall and on 393 antecedent soil moisture conditions, other factors play a role including the rainfall variability 394 during the event, differences in slope, infiltration capacities, soil depth, and porosity.

Recently Jadidoleslam et al. (2019) introduced the deficit-based rainfall metric P_{θ} = 395 $\sum P/(D_s(\phi - \theta_o))$ to categorize runoff coefficients observed in catchments in Iowa, where $\sum P$ is 396 397 the total event precipitation, ϕ is the porosity, and θ_o is the antecedent soil moisture, and D_s is the 398 integration depth of soil moisture values. They found that the index collapsed the variability of 399 runoff coefficients significantly, providing predictability for the runoff coefficient with spearman 400 correlation coefficients of 0.6. This metric is ideal to contrast runoff production from event to 401 event and from one control volume to another, but it also gives us an objective comparison between 402 our three different models. Note that in the HLM, parameters related to the rainfall-runoff 403 transformation are constant in space, though a dependency on hillslope length introduces spatial 404 variability. In WRF-Hydro, the control volumes are square grid cells, but there is significant 405 variability in parameter values specified from land use and soil maps. In WMF, the two sources of 406 variability are combined, because the control volumes are irregular in size and have variable 407 parameters

408 **4.3.3.** Statistical Homogeneity of Peak Flow Distributions

409 After establishing the range of homogeneity (or lack thereof) in runoff processes, we can establish 410 the level of statistical homogeneity amongst peak flow distributions for locations along the river 411 network. Peak flows are a scale-dependent random variable, which makes establishing 412 homogeneity more complicated than rainfall and runoff. Here, we use the spatial variability of the 413 coefficient of variation and skewness to determine if the distributions can be rescaled using this 414 information alone. As in the case of rainfall, the 50,000 events are treated as 10,000 synthetic years 415 and all peak flow statistics are calculated using the most extreme 10,000 events for each channel 416 link in the river network.

417 **5. Results**

418 **5.1. Model Performance Evaluation**

419 The three models were set up for Turkey River. Parameters in WRF-Hydro and WMF were 420 determined using data for the period 2002 to 2018 via calibration, while HLM was parameterized 421 using values from the operational IFC model for the Iowa domain (Quintero et al., 2016). We 422 determine the ability of each of the models to capture peak flows for a total of 782 rainfall runoff 423 events. Simulated peak flows were compared against observations, as were observed and simulated 424 hydrographs at five gauged locations within the watershed (Figure 3). While bias and random error 425 are evident in all three models, no single model is clearly superior and all three models exhibit 426 reasonable performance at capturing both peak flows and other streamflow dynamics.

427 It is hard to determine if any model is superior. To evaluate the models, we compute the peak flows correlation coefficient (R^2), and the streamflow KGE (Kling Gupta Efficiency) index (H. V. Gupta 428 et al., 2009) and Pbias (volume bias). R^2 represents the simulated peak flow accuracy. The KGE 429 430 summarizes the correlation, the mean bias, and the deviation bias to determine the performance, and the Pbias represents the total percentage at the simulated volume. WMF shows the higher peak 431 flow correlation coefficient with an R^2 equals to 0.87 (Figure 3c). Followed by HLM with an R^2 432 equal to 0.85 and WRF with an R^2 of 0.72. According to Figure 3 d to h, there is no model with a 433 434 dominant KGE or Pbias. Regardless of the observed performances, we accept the three models for 435 the purpose of the current work.



Figure 3. A comparison of model performance to simulate peak flows and hydrographs. Panels a,b, and c show the simulated vs the observed peakflows for the HLM, WRF, and WMF respectively. Panels d to h show the observed streamflow (black) and the simulated streamflow for the HLM (green), WRF (orange), and WMF (blue) models at five USGS gauges.

436 **5.1 Realism of Rainfall Intensity Duration Frequency (IDF) Curves**

437 Rainfall intensity-duration-frequency (IDF) curves at the pixel scale, following Wright et al.

- 438 (2017) are calculated for the Turkey river domain. In the Figure 4, we compare the SST-based IDF
- 439 curves against published rain gage-based IDF estimates from NOAA's Atlas 14 project (Bonnin
- 440 et al., 2006).



Figure 4. SST Intensity Duration Frequency (IDF) curves for accumulations ranging between 1h and 72h. And, Atlas-14 IDF curves for accumulations of 24h.

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442 **5.2 Physical Homogeneity of Runoff Generation**

Violin plots show the variability in simulated flood event runoff coefficients of individual control 443 444 volumes (i.e. hillslopes or grid cells) as a function of P_{θ} (Figure 5). The variability of RC vs. P_{θ} is analogous to the concept of curve number, as it represents the expected runoff coefficient for a 445 446 given precipitation total and the antecedent soil moisture, and it represents the prediction of each 447 one of the models for the statistics of runoff regeneration in the region. In the case of the HLM 448 model (green violins) we observe the largest values of runoff coefficient as the model structure 449 assumes that the majority of quick runoff entering the channels is surface runoff. The extent of the variability of RCs for different values of P_{θ} is small, as expected, because all the hillslopes in the 450 451 domain share the same values. In the case of RCs predicted by WRF-Hydro we see the smallest 452 values of runoff coefficient, which indicates that a significant fraction of the quick runoff into the 453 channels is provided by subsurface flows. Finally, the WMF model predicts an intermedia 454 prediction for surface runoff, but it exabits the largest variability or RCs. These results reveal that even in landscapes that appear to have very similar land cover, land use, and soil types, such as
those found in the Midwest, significant spatial and temporal heterogeneity can be introduced by
the physical processes that control runoff production.



Figure 5. Runoff ratios calculated using three hydrological models as a function of deficitbased precipitation for 50,000 simulated rainfall runoff events.

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459 **5.2. Statistical Homogeneity of Peak Flow Distributions**

Scaling plots show the mean μ , standard deviation σ , coefficient of variation $CV = \sigma/\mu$, skewness 460 γ , and kurtosis κ with respect to drainage area for the three models' simulated peak flow 461 462 distributions using the 10,000 peak flow samples at each location (Figure 6). Recall from Sec. 2.1 463 that the homogeneity requirement for the index flood method (i.e. simple scaling) is for CV to be 464 constant across scales and for every location in the region. For the three models this assumption is 465 clearly violated for the population of peak flows (Figure 6g-i). It could be argued that the variations 466 are within acceptable margins of error for any of the given empirical RFFA techniques; however, 467 as is shown by Perez et al (2019b), an incorrect representation of the skewness in the estimation 468 of the index flood method will add systematic bias to resulting peak flow estimates that changes 469 across scales. In addition, the index flood method requires that the index flood be predicted using 470 a power law regression equation with respect to drainage area. Residuals of power law fits 471 (supplemental Fig. S1) show the existence of one or more scale breaks between 100 and 1000 km²,



472 depending on the model.

Figure 6. Scaling plots for the mean μ , standard deviation σ , coefficient of variation $CV = \sigma/\mu$, skewness γ , and kurtosis κ of peak flow distributions with respect to upstream drainage area A, from the three hydrological models considered in this study.

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The multi-scaling framework (Sec. 2.1) relaxes the requirement that CV is constant across scales, but it requires that all distribution moments scale as a power-law of drainage area. While this appears reasonable for the mean and standard deviation, all three models show significant variability in skewness that is not explained by area alone (Figure 6j-1).

The spatial distribution of skewness in the drainage network offers an additional view into the limits of the concept of a regional skew value and its existence (Figure 7). The maps of skewness for the three models reveal systematic variations of skew across the drainage network, which are undoubtedly related at least in part to the interaction of storm patterns with network geometry, with parameter variability related to land use and soil properties likely playing roles in WRF-Hydro and WMF results. The map generated by WRF-Hydro exhibits the most unstructured pattern of skewness in space and the widest range of variation of its value.

It is important to note that the spatial variability of skewness shown in Figure 7 is substantial, on the order of magnitude of the skewness values across Iowa in the most recent USGS statewide flood frequency study (Eash et al., 2013), and as large as those found for the entire Midwest region of the United States in Bulletin 17B (IACWD 1982) (see supplemental Fig. S2).

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Figure 7. Spatial distribution of skewness γ of the peak flow distribution with respect to upstream drainage area *A* from the three hydrological models considered in this study.

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491 **5.3 Physics-based Predictions in the Context of Observed Flood Data**

Each of the three models provides a deterministic representation of hydrologic response to a given rainfall input, and therefore a unique prediction of the peak flow distribution for every computational node in the river network. These distributions and particular quantiles from them can be compared against observations. In Figure 8, we focus on simulated 100-year peak flows and compare them against USGS estimated 100-year events (Eash et al. 2001) at gauged locations in Iowa's Region 2 (see Figure 1). In addition to the population estimate using 10,000 annual flood
peaks, Figure 8 also shows the values that would be estimated if we used smaller samples of 100,
and 1,000 annual peaks. Variability in these estimates provides a depiction of plausible deviations
from true population values due to sample size.



Figure 8. Predictions of the 100-year peak low in Turkey River. Colored dots represent simulated peak flows quantiles, with different colors highlighting how uncertainty varies with sample size. Black dots show the 100-year peak flow estimates (Eash et al. 2001) using observations in Region 2 in Iowa (see Fig.1 for map of RFFA regions). Blue vertical lines denote the 16 km² spatial resolution of Stage IV precipitation used in this study. Predictions below that scale are questionable, and rainfall resolution effects may explain the apparent bias relative to observations at small drainage scales. The effect of resolution scale is less relevant for larger scales (e.g. Cunha et al., 2012; Mandapaka, P. V. et al., 2009).

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The simulated results in Figure 8 show predictions that emerge from the assumptions that are built into model equations and parameterizations as well as the rainfall fields generated through SST. This result is in high contrast to the purely empirical RFFA approach of fitting power law regressions and using these regressions to estimate values at ungauged sites. The physically based estimated frequencies are calculated for each channel in the river network, and they include the available information about the hydrology of the region that is encoded by each one of the hydrologic models. The magnitude of variability of the 100-year flood estimate as a function of sample size is a consequence of the rate of convergence of the moments of the peak flow distribution. In Figure 9 we illustrate this issue by showing the results of a bootstrap analysis of estimation for the mean, standard deviation, skewness and kurtosis, as a function of sample size at three different locations in the basin that drain different watershed areas.



Figure 9. Estimated peak flow distribution moments (mean μ , standard deviation, σ , skewness, γ , and kurtosis, κ) at three different catchment scales for varying sample sizes.

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515 Figure 9 shows that peak flow distribution moments can be accurately estimated for sample sizes 516 larger than 1,000. In the case of small sample sizes, the estimation error can be larger than 50% 517 for all the moments. In the case of the mean and standard deviation, the small sample estimates 518 seem to be unbiased, however, for the skewness and kurtosis, it can be seen that the estimates using 519 small sample sizes are biased toward low values. The rates of convergence appear to be scale 520 independent. The need for a sample size of 1,000 peak flows to estimate population moments 521 offers a glaring contrast with the length of datasets that are typically used in RFFA that typically 522 range between 30 to 100 years.

523 6 Discussion

524 The three hydrological models represent different approximations of the true hydrological cycle 525 in the Turkey River study watershed. There is no doubt that the actual flow paths are more complex 526 than those represented by any of the models. The models appear to capture key aspects of hillslope 527 and channel processes in the basin (Figure 3), suggesting that the simulated variability of peak 528 flows depict a plausible representation of reality. Furthermore, the three models provide estimates 529 for the 100-year flood for a catchment in Region 2 of Iowa that are relatively consistent (i.e. within 530 an order of magnitude) of estimates made using flood peak observations (Figure 8). We 531 acknowledge that this simulated variability has its limits, however. While all three models offer a 532 variety of optional process configurations and virtually infinite parameter spaces, only a single 533 configuration and parameter set were used for each model in this study. Furthermore, the SST 534 methodology's reliance on transposing observed storms limits the variability that can arise from 535 precipitation inputs (Wright et al., 2020). Finally, we totally neglect snowmelt and rain-on-snow 536 flooding, which can be relevant for the study region. Thus, it is reasonable to expect that real peak 537 flows would exhibit more, not less, variability in space and time than the variability modeled here. 538 While it is thus difficult to make definitive statements based on our results, but they serve as a 539 useful starting point to discuss the magnitude of peak flow variability in space and across scales. 540 In addition, some of the limitations of our study (e.g. parameter equifinality, snowmelt processes) 541 are well within the capabilities of many contemporary distributed modeling frameworks.

542 Our findings are consistent with and help to generalize the results of Perez et al., (2019b). Namely, 543 simulated peak flow quantiles exhibit a level of variability that cannot be simplified into a simple-544 or multi-scaling framework, indicating that the underlying assumptions in widely-used RFFA 545 methods such as IFM and quantile regressions will introduce estimation biases. Simple scaling is

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546 rejected by the spatial variability of the coefficient of variation, while multi-scaling is rejected by 547 the variability of skewness in space and across scales (Figure 7Figure 8, and S2). Furthermore, 548 scaling of the mean annual flow and standard deviation with respect to drainage area does not 549 collapse to a simple power law across scales (Figure 9). In fact, we observe a scaling break in all 550 three models ranging from 100-1,000 km², possibly tied to a scale at which river routing and 551 network topology effects start exerting control over peak flow variability. Further assessment of 552 the errors introduced by assuming simple- or multi-scaling of peak flow data is outside the scope 553 of our work and is instead left as an important future research task.

In general, the simulated peak flows reject the premise that data from neighboring sites provide information about the specific attributes of the peak flow distribution at an ungauged site. We found strong patterns of variability in skewness (Figure 7Figure 8) that would introduce biases when interpolating or extrapolating from one site to another. Furthermore, the degree of variability in skewness within the 4,385 km² study watershed is comparable to the variability of published estimates of flood peak skewness for Iowa (Eash et al. 2013) and for the entire midwestern US (IACWD, 1982) shown in supplemental Fig. S2.

561 The spatial structure of skewness predicted by the three models is remarkable in several ways. 562 First, the differences in values estimated by the three models suggests that skewness is very 563 sensitive to the underlying local physical processes controlling the magnitude of flood peaks. We 564 attempted several analyses to identify the variables involved in explaining the variability of 565 skewness, but no simple predictive equation has emerged (results not shown). It would thus appear 566 that every aspect of the physics of peak flow generation matters when it comes to determining 567 skewness. Second, the value of local skewness differ from the average value of skewness in the 568 catchment across all scales, indicating that the distribution of peak flows is sensitive to both small-

569 scale (e.g. infiltration dynamics; Figure 7) and large-scale (e.g. river network topology; see 570 Mantilla (2010) for a theoretical context on the role of river networks) sources of variability in 571 runoff production and transport. Third, variability in skewness implies that the notion of a single 572 regional skew value, which is central to RFFA practice in the United States and also implicit in 573 other empirical RFFA methodologies, is an incorrect assumption. It is difficult to pinpoint what 574 aspects of each model lead to a particular value of skewness; variability arises not only from 575 parametric differences but also dynamics that include non-linearities and threshold values and the 576 spatial and temporal variability of the rainfall fields.

577 Another important insight from our simulation exercise is just how slowly the peak flow sample 578 moments converge to the "population" moments (Figure 9). This implies that sample size 579 dominates the uncertainty in estimates of flood quantiles, even using regionalization techniques, 580 confirming Perez et al. (2019b). Our estimates of the 100-year flood using 100 synthetic years in 581 Fig. 10 can easily double or halve the "true" population value, and it takes roughly 1,000 synthetic 582 years of simulated data to reliably obtain estimates that are within 20% of the population values. 583 The consistency in this finding across the three models suggests that this feature is a property of 584 the peak flow random variable, which is greatly influenced by both initial states (e.g. soil moisture, 585 storage in channels) and the spatiotemporal structure of rainfall (Zhu et al., 2018). These results 586 should raise concerns about both data-driven studies that try to estimate nonstationary trends in 587 peak flows over short timespans and studies that aim to project changes in flood regimes under 588 different climate change scenarios.

A frequent criticism of the usage of hydrologic models, including physics-based ones, in flood frequency studies is the perception that model uncertainties are prohibitively high. As argued in Wright et al., (2020), such criticisms generally presume the primacy of data while glossing over 592 the numerous and poorly-understood uncertainties in statistically-based RFFA. We agree 593 wholeheartedly with the critical role of data, but argue that statistically-based RFFA makes poor 594 use of the wealth of relevant datasets currently available—including distributed high-resolution 595 rainfall, soils, land cover, and hydrography, not to mention the vast majority of streamflow 596 observations that do not rise to the level of annual maxima. Physically-based approaches such as 597 the one used here not only leverages these rich resources, but also our understanding (which is 598 admittedly imperfect, and further simplified within models) of the movement of water through the 599 landscape. This makes the estimates made through our process more robust and less susceptible to 600 outliers in the data.

601 **7 Conclusions**

602 Three hydrological models and stochastic storm transposition (SST) were used to investigate the 603 validity of implicit assumptions in the empirical methodology of regional flood frequency analysis 604 (RFFA) for prediction of flood quantiles at ungauged locations. This combined physics-based 605 framework allows us to estimate peak flow frequencies throughout the river network by calculating 606 thousands of realistic dynamic flooding scenarios. These scenarios give us an unprecedent look 607 into the nature of the "annual peak flow random variable," which is at the heart of the empirical 608 RFFA. The framework allowed us to investigate the four questions, reprinted from Section 2.3: (i) 609 what is the link between physical homogeneity and statistical homogeneity among locations within 610 a watershed or region?, (ii) what are appropriate simplifications for the scaling of the statistical 611 moments of peak flow distribution across scales?, (iii) can the peak flow distribution parameters 612 at an ungauged site be inferred from measurements at neighboring sites?, and (iv) what are the 613 necessary sample sizes to characterize peak flow distributions.

614 Our results do not support the link between *physical homogeneity*, defined as the spatial variability 615 in the physics of runoff generation and runoff routing, and *statistical homogeneity*, defined as the 616 existence of a common underlying statistical distribution of peak flows and the possibility of 617 inferring the distribution parameters using regional observations of peak flows. However, our 618 results cannot ultimately reject the link, instead they show that more general and restrictive 619 definitions of homogeneity are needed. Specifically, our simulated peak flow results show that the 620 local value of the third moment in the population distribution (skewness) cannot be inferred from 621 regional values, and instead depends on very specific local conditions including but not limited to 622 land use and soil distribution and their effect on runoff production, river network topology and 623 geometry, and local hydraulic geometry. While not explicitly evaluated here, basin shape and orientation relative to prevailing storm motions and space-time structures likely to exert important 624 625 controls on skewness; this was partially addressed by Perez et al. (2019b), but it remains an open 626 problem.

627 The spatial variability of skewness predicted by the three hydrologic models indicates that simple-628 scaling (i.e. the index flood method, generally used in Europe and Canada) and multi-scaling (e.g. 629 quantile regressions used in the United States) frameworks provide at best first-order 630 approximations of peak flow variability at ungauged locations. Those scaling frameworks require 631 that the moments of a peak flow distribution be inferred from basin area alone through power laws 632 or other deterministic functions. Instead, our results suggest that well-validated distributed 633 hydrological models may be needed to estimate local moments for peak flow distributions at 634 ungauged locations, due to the complexity of rainfall-runoff transformations and river routing. In 635 addition, the methodology presented here can be used to quantify the magnitude of sampling error 636 involved in the inference of particular peak flow quantiles such as the 100-year flood.

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A careful inspection of differences between the three models allow us to identify key features in the rainfall-runoff and runoff-transport processes that play important roles in determining the degree of spatial heterogeneity of peak flows distributions. In particular, in terms of sources of spatial variability of skewness, we can list four factors:

River Network Topology: the connectivity of side tributaries creates consistent spatial
patterns in the variability of skewness along the river network. Although the three models
used here employ somewhat different depictions of the river network (e.g. different
drainage densities) in their formulations, they all show consistent patterns of variation
along the major rivers in the watershed, although different amongst models, reflecting the
complex aggregation and attenuation of flows that determine peak flow quantiles.

Routing: The routing schemes in our paper range from kinematic based ODE formulations
to one dimensional PDEs, some of which include channel and terrain slope information.
These different formulations give rise to different patterns of variability in skewness along
a particular channel.

Infiltration Processes at the characteristic scale of the model: the three models use different
discretization scales (e.g. grid structures) to describe physical processes. This translates
into differences in sub-catchment scale responses and suggests that spatial variability in
infiltration processes in the real world will leave a fingerprint on peak flow distributions.

4. Routing of Runoff into River Network Channels: akin to infiltration processes, the routing
schemes for surface and subsurface runoff toward river channels differs among the three
models. Those differences are reflected in the skewness maps, indicating that the routing
of water over and through real-world hillslopes will have a signature in peak flow
distributions across spatial scales within a catchment.

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660 Which physical processes control the ultimate distribution of peak flows, and how they control 661 them, remains to be revealed and quantified. The models used in this paper are gross 662 simplifications of real-world physical processes, which are no doubt far more complicated. In 663 addition, other flood generation mechanisms such as snowmelt would add a new dimension to the 664 problem. There is little reason to believe, however, that more complex and realistic model physics 665 would lead to smoother spatial fields of peak flow distribution moments. In addition, our storm 666 resampling procedure in SST does not fully reflect the extent of the real rainfall spatial and 667 temporal variability of a true population of storms. For example, a recent study shows that the 668 orientation of the catchment in relation to the directionality of storms is a relevant control on flood 669 quantiles (Perez et al., 2020). We have made an effort to provide all the data and models needed 670 to replicate the results of this study in public repositories, however we are well aware that 671 replicating a scientific study of this magnitude can be daunting (Hutton et al., 2016). However, we 672 believe that our results are generic enough that they can be generated with any other high-673 resolution distributed hydrological model. The refinement and generalization of our ideas would 674 require the engagement of other research groups interested in addressing these questions in other 675 regions in the world.

Although SST is in its early stages and the physical models employed in this work have limited levels of sophistication, our results show the urgent need to transition from purely statisticallybased empirical RFFA to approaches that are informed by the best available local characterization of the physical processes that determine peak flows.

The results presented in this paper lead us to argue that the tools and theoretical underpinnings to establish a physics-based RFFA framework are within reach. We invite the hydrological community and the relevant agencies to test our conclusions using their own hydrological models. 683 If our findings hold under further scrutiny, we urge reconsideration of the current path of empirical 684 RFFA, which consists of periodic updating using recent observations, along with incremental 685 refinement of statistical methods. Our results show that the variability of peak flows cannot be 686 captured with the relatively short records in the data sets available today (generally less than 100 687 years). Instead, it appears that about 1,000 independent peak flows at multiple sites would be 688 needed to approximate the true quantile values within acceptable margins of error (see Fig. 10). It 689 seems difficult to believe that a new statistical technique, based on the same limited data, could 690 overcome such a hurdle. On the other hand, the path anticipated by (Eagleson, 1972; V. K. Gupta 691 et al., 2007; Klemeš, 1987) for a dynamical framework for prediction of flood quantiles based on 692 our understanding of the physical processes that determine peak flow magnitude is within reach 693 and it can provide a robust estimation framework for the spatial and temporal variability of peak 694 flows frequencies. It is our hope that in the future, "Bulletin 17D" will mark a reinvention of flood 695 frequency estimation that truly reflects the monumental advances in distributed 696 hydrometeorological observations and modeling of the last several decades and that is supported 697 by universally agreed scientific theory.

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Figure 1.



Figure 2.

(a) Study Area

SST Domain Watershed

(b) gSSURGO Soil Class





Figure 3.



Figure 4.



Figure 5.



Figure 6.



Upstream area [km²]

Figure 7.





Figure 8.



Figure 9.



















