# Towards narrowing uncertainty in future projections of local extreme precipitation

Francesco Marra<sup>1</sup>, Moshe Armon<sup>2</sup>, Ori Adam<sup>3</sup>, Davide Zoccatelli<sup>4</sup>, Osama Moh'd Najeeb Gazal<sup>5</sup>, Chaim I Garfinkel<sup>2</sup>, Dorita Rostkier-Edelstein<sup>6</sup>, Uri Dayan<sup>7</sup>, Yehouda Enzel<sup>2</sup>, and Efrat Morin<sup>8</sup>

<sup>1</sup>Institute of Atmospheric Sciences and Climate, National Research Council
<sup>2</sup>Hebrew University of Jerusalem
<sup>3</sup>Hebrew University
<sup>4</sup>University of Padova
<sup>5</sup>Faculty of Agricultural and Environmental Sciences, Szent Istvan University
<sup>6</sup>Department of Environmental Physics, Environmental Sciences Division, IIBR
<sup>7</sup>Department of Geography, Hebrew University of Jerusalem
<sup>8</sup>Institute of Earth Sciences, The Hebrew University of Jerusalem

November 28, 2022

#### Abstract

Projections of extreme precipitation based on modern climate models suffer from large uncertainties. Specifically, unresolved physics and natural variability limit the ability of climate models to provide actionable information on impacts and risks at the regional, watershed and city scales relevant for practical applications. Here we show that the interaction of precipitating systems with local features can constrain the statistical description of extreme precipitation. These observational constraints can be used to project local extremes of low yearly exceedance probability (e.g., 100-year events) using synoptic-scale information from climate models, which is generally represented more accurately than the local-scales, and without requiring climate models to explicitly resolve extremes. The novel approach offers a path for improving the predictability of local statistics of extremes in a changing climate, independent of pending improvements in climate models at regional and local scales.

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6 7	<sup>1</sup> The Fredy and Nadine Herrmann Institute of Earth Sciences, The Hebrew University of Jerusalem, Israel		
8 9	<sup>2</sup> Institute of Atmospheric Sciences and Climate, National Research Council of Italy, CNR-ISAC, Bologna, Italy		
10	<sup>3</sup> Faculty of Agricultural and Environmental Sciences, Szent Istvan University, Hungary		
11 12	<sup>4</sup> Department of Environmental Physics, Environmental Sciences Division, IIBR, Ness-Ziona, Israel		
13	<sup>5</sup> Department of Geography, The Hebrew University of Jerusalem, Israel		
14			
15	Corresponding author: Francesco Marra ( <u>f.marra@isac.cnr.it</u> )		
16	*Via Gobetti, 101, 40129, Bologna, Italy		
17	Key Points:		
18 19	• Local observations can be used to constrain the intensity distribution of precipitation events associated to different synoptic systems		
20 21	• The constraints allow projecting extreme return levels at scales relevant for impact studies from synoptic information from climate models		
22 23 24	• The approach improves the predictability of local extremes, independent of improvements in climate models at regional and local scales		

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- 26 Projections of extreme precipitation based on modern climate models suffer from large
- 27 uncertainties. Specifically, unresolved physics and natural variability limit the ability of climate
- models to provide actionable information on impacts and risks at the regional, watershed and city
- 29 scales relevant for practical applications. Here we show that the interaction of precipitating
- 30 systems with local features can constrain the statistical description of extreme precipitation.
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- 35 predictability of local statistics of extremes in a changing climate, independent of pending
- 36 improvements in climate models at regional and local scales.

### 37 Plain Language Summary

Climate change impact studies are currently restrained by the limited accuracy of climate models 38 in resolving precipitation extremes and by the uncertainties characterizing their analysis. We use 39 here a novel approach which permits to project extreme precipitation for future climatic 40 scenarios based on the combination of coarse-scale information from climate models with local 41 42 observations. Focusing on the south-eastern Mediterranean, we provide projections of precipitation extremes which could not yet be derived using traditional methods, such as the 43 events occurring on average once in 100 years. The combined effect of changes in intensity and 44 average yearly number of two dominant synoptic systems is projected to increase the intensity of 45

- the 100-year events in the coast and in the desert areas of the region, and to decrease it
- 47 elsewhere. The novel approach offers a path for improving the predictability of extremes in a
- 48 changing climate, independent of pending improvements in climate models.

## 49 **1 Introduction**

50 In recent decades, natural hazards associated with extreme precipitation, such as floods

- and landslides, claimed thousands of lives and billions of US\$ in damages every year (NOAA,
- 52 2020; Paprotny et al., 2018). These numbers are expected to grow in response to an expansion of
- 53 population and wealth towards hazard-prone areas and to modifications in the hydrological cycle
- induced by climate change (Ceola et al., 2014; Fischer and Knutti, 2016; Winsemius et al.,
- 55 2016). Quantifying climate change impact on extremes is thus a major challenge for the research
- community (Blöschl et al, 2019). Hydrological design and risk management, particularly relevant
- 57 for adaptation efforts, require information on low *yearly* exceedance probabilities (Chow et al.,
- 1988), such as the events exceeded on average once in 100 years (hereon 100-year *return levels*,
- 59 with 1% yearly exceedance probability). To directly quantify return levels, long data series are
- 60 required, several times longer than the exceedance probability timescale. Since observational
- 61 records rarely exceed 50-100 years, some form of statistical extrapolation is generally required
- 62 (Coles, 2001).

Earth system models (ESMs) are commonly used to guide impact studies. However, 63 current ESMs are not able to explicitly resolve convective and microphysical processes critical 64 for precipitation extremes, and rely instead on parameterizations (Wilcox and Donner, 2007). 65 Additionally, their output is most relevant at scales which are too coarse for many practical 66 applications (Fischer et al., 2013; Hausfather et al., 2019). Dynamical downscaling methods can 67 provide projections for a region of interest, but are sensitive to the boundary conditions provided 68 by global models (Shepherd, 2014; Keller et al., 2018). Furthermore, their application is limited 69 by computational requirements so that, currently, only few regions are covered with 10-20-year 70 simulations (Kendon et al., 2014; Fosser et al., 2020), which are too short to reliably estimate 10-71 years return levels, let alone 100-year events. Alternatively, statistical models are combined with 72 variables that are strongly related to extreme precipitation but more reliably reproduced in ESMs, 73 such as temperature (Snippel et al. 2015; Pfahl et al., 2017). 74

The methods currently adopted to quantify return levels, however, heavily rely on 75 extremes, such as the maxima values in each year or the values exceeding high thresholds 76 (Coles, 2001). As these are rare and subject to large uncertainties, the applicability of these 77 methods in a changing climate is limited (Serinaldi and Kilsby, 2015). In fact, stochastic climate 78 variability sets a lower bound on the uncertainty in observed and modelled extremes (Fatichi et 79 al., 2016). Reliable projections of extreme return levels for future climate scenarios thus 80 necessarily entail either intensive dynamical downscaling of ESMs with convection-permitting 81 models, or novel statistical approaches able to better exploit the available information. 82

It is shown here that the interaction of precipitating systems with local features, such as coastlines or orography, can constrain the statistical description of precipitation intensity. These constraints, derived from in-situ observations, permit predicting future extreme return levels at the local-scales based on coarse-resolution global climate model projections, and without requiring models to explicitly resolve the extremes.

88 **2 Study area and data** 

89 The south-eastern Mediterranean is regarded as a climate change hotspot, highly vulnerable to water scarcity and precipitation-induced hazards (Alpert et al., 2002; Giorgi, 2006). 90 Strong spatial gradients in precipitation climatology (Fig. S1 and S2 in the Supporting 91 Information) emerge from the interactions of two main types of precipitating systems with 92 coastline and orography (Diskin, 1970): (i) low-pressure systems moving inland along westerly 93 94 tracks (Mediterranean cyclones, hereon *Type-1*), and (ii) low-pressure systems mainly extending from the south (active Red Sea troughs, *Type-2*). These are characterized by distinct spatial 95 patterns and both yield extreme precipitation amounts (Armon et al., 2018; Marra et al., 2019a). 96

97 ESMs predict substantial changes in the intensity and occurrence frequency of both systems

(Hochman et al., 2018a; Hochman et al., 2018b; Zappa et al., 2015), implying non-linear changes
in the compound extremes, which can be further complicated by local effects.

100 2.1 Precipitation data

Daily precipitation data, summed up to 6:00UTC, were provided by the Ministry of 101 Water and Irrigation of Jordan (97 stations between 1980-1981 and 2017-2018) and the Israel 102 Meteorological Service (>1300 stations between 1948-1949 and 2017-2018). Data from Israeli 103 stations flagged as missing, inaccurate, interpolated or obtained from multi-day accumulations 104 were excluded from the analysis. Jordanian data were supplied with no quality indicators; we 105 therefore rely on quality controls by the data provider. Separate records measured in proximity of 106 up to 1 km distance and 50 m elevation were merged. Records were organized by hydrologic 107 years (September 1 to August 31). For each station, years with more than 14 unavailable days 108 and records with less than 30 hydrological years were discarded. The final dataset consists of 459 109 stations (404 from Israel, 55 from Jordan, average spatial density of ~1/75 km<sup>-2</sup>) with 30-70 110 complete years of record (50.1±13.3 years). Stationarity of the annual maxima at each station is 111 ensured using the Phillips and Perron (1988) test (5% significance level), indicating that the data 112 adequately represent extremes under present conditions. 113

### 114 2.2 Local groups of stations

Groups of stations in which distinct local features dominate the interaction with the 115 116 precipitating systems are identified using a *kmeans* clustering algorithm based on geographical (latitude, longitude, elevation) and precipitation (average wet-day amount, and standard 117 deviation of the wet-day amounts) properties, without any direct use of extreme precipitation 118 properties or classification of the precipitating systems. The variables are normalized to zero-119 mean equi-dispersed distributions; the algorithm is iterated 99 times to ensure stable results. 120 Following the Calinski and Harabasz (1974) criterion, six groups are obtained, roughly 121 identifiable as: mountains, northern coast, lowlands, coast, deserts west of the Dead Sea rift, and 122 deserts east of the rift. The last two groups are characterized by similar climatic conditions and 123 are likely separated primarily due to the geographical distance, although differences in other 124 125 aspects may exist, such as elevation and distance from the sea. These two groups, which are sparsely populated (only 21 stations in one group), were merged. The classification used in the 126 analysis consists of five groups: mountains, northern coast, lowlands, coast, and deserts (Fig. 1a). 127

### 128 **3 Methods**

129 Extreme precipitation events were shown to emerge from underlying distributions of

ordinary events (Marani and Ignaccolo, 2015; Zorzetto et al., 2016), whose tails are generally

described by two parameters (e.g., stretched-exponential or power-type) (Cavanaugh et al., 2015;

132 Papalexiou et al., 2018; Marra et al., 2020b). By relying on ordinary events, for which more data

is available, this approach decreases the stochastic uncertainties inherent in the realization of

extremes (Zorzetto et al., 2016; Marra et al., 2018). Events generated by different types of

- 135 processes and thus described by distinct distributions such as mid-latitude vs. tropical cyclones
- 136 (or, in our case, *Type-1* vs. *Type-2*) can be combined to derive a compound distribution for
- extreme return levels (Marra et al., 2019a; Miniussi et al., 2020). This distribution quantifies the
- 138 yearly exceedance probability  $\zeta$  associated with the precipitation amount x as a function of the
- intensity distributions of the ordinary events ( $F_{i=1,...,S}$ , where *i* represents the type of process) and
- 140 the expected value of their yearly number of occurrences  $(n_i)$  such that:  $\zeta(x) \simeq F_1^{n_1} \cdot F_2^{n_2} \cdot ... \cdot$
- 141  $F_s^{n_s}$  (Marra et al., 2019a). In this framework, changes in extreme return levels can be expressed
- 142 as functions of the projected changes in the intensity distributions of the ordinary events and in
- the expected value of their yearly occurrences. While the occurrence frequency of synoptic
- events in the region can be resolved by ESMs (Hochman et al., 2018a; Cavicchia et al., 2020),
- 145 precipitation intensity requires information on *two* degrees of freedom (i.e., the two parameters
- 146 describing the distribution).
- 147 3.1 Ordinary events distributions and return levels
- Ordinary events are defined as non-zero (i.e.,  $\ge 0.1$  mm) daily precipitation amounts (Zorzetto et al., 2016) associated with a precipitation type based on a semi-automatic, dailybased, synoptic classification (Alpert et al., 2004). Wet days corresponding to systems that are expected to be dry may have been wrongly classified; for example, synoptic conditions in the aftermath of Mediterranean cyclones are easily misinterpreted by the semi-automatic method. These were individually examined and labelled as *Type-1* if occurring up to 2 days after a *Type-1* day, and as *Type-2* in the remaining cases (Table S1).
- Previous studies show that a Weibull distribution (stretched-exponential) in the form 155  $F(x; \lambda, \kappa) = 1 - e^{-\left(\frac{x}{\lambda}\right)^{\kappa}}$ , where  $\lambda$  is the scale and  $\kappa$  the shape parameter, well describes the tail of 156 the two types of ordinary events in the region (Marra et al., 2019a). The shape parameter 157 determines the tail heaviness, with heavier tails for smaller shapes and vice versa. These 158 parameters are estimated left-censoring the lowest 75% of the observations while keeping their 159 weight in probability, and using a least-square linear regression in Weibull-transformed 160 coordinates (Marra et al., 2019a). The left-censoring prevents contaminations from the lower tail 161 of the distribution, which may require more general formulations (Papalexiou et al., 2018; 162 Cavanaugh et al., 2015) and is sensitive to the accuracy of the measurement device (Marra et al., 163 2019a). After left-censoring, the number of data points used for the parameter estimation in each 164 of the stations is  $426 \pm 175$  for Type-1 (minimum 66), and  $153 \pm 68$  (minimum 30) for Type-2. 165 166 The expected number of yearly ordinary events is computed, for each type, as the mean of the yearly number of wet days. Extreme return levels are computed numerically by inverting the 167 formulation  $\zeta(x) \simeq F_1^{n_1} \cdot F_2^{n_2}$ . Sample uncertainty in parameters and return levels is quantified 168 via bootstrap with replacement ( $10^3$  repetitions) among the years in the record (Overeem et al., 169 2008). The resulting return levels (Fig. 1; Fig. S2) are consistent with traditional methods based 170 on the observed annual maxima (Fig. S3), but have significantly smaller uncertainty (22%, as 171 opposed to 39%, median uncertainty on 100-year return levels). 172





Figure 1. Extremes emerge from the interaction of precipitation systems with local features. (a) Map of 174 175 the study region showing the local terrain elevation, the main precipitating systems tracks, and the 176 location of the daily precipitation stations used in the study, coloured according to groups in which different local features dominate the interaction with the precipitating systems. (b) Distribution of 2-year 177 and 100-year return levels (50% and 1% yearly exceedance probability, respectively) of the five groups 178 shown in panel a (with matching colours) displayed as transposed cumulative distributions. The 179 respective uncertainty (shading) is calculated as the 90% confidence interval from  $10^3$  bootstrap samples 180 with replacement among the years in the record. The median uncertainty across all groups is 14% (22%) 181 for 2-year (100-year) return levels. (c, d) Map of the 2-year (c) and 100-year (d) return levels (colours 182 183 indicate daily precipitation intensity).

### 184 3.2 Local constraints of the intensity distributions

A robust relationship between the scale  $\lambda$  and shape  $\kappa$  parameter of the ordinary events 185 distributions would reduce the representation of precipitation intensity to *one* degree of freedom, 186 enabling us to provide projections of extremes based only on changes in the mean intensity of 187 ordinary events. The significance of the relationship between the parameters describing the two 188 types of ordinary events at each of the five groups of stations is tested using the rank correlation 189 (10<sup>4</sup> Monte Carlo reshuffling realizations). The coefficient  $\alpha$  of the relations in the form  $\kappa = \alpha$ . 190  $\log \lambda + C$  is derived for each of the five groups and the two event types using a linear regression 191 model based on a  $\chi^2$  minimization and considering parameter estimation errors in a Monte Carlo 192 framework ( $10^3$  realizations). 193

194 The coefficients  $\alpha$ , calculated for each group, represent the local constraints on the 195 intensity distribution. Higher  $\alpha$  implies a stronger decrease in tail heaviness in response to an 196 increase in the median intensity, and vice versa. Under these constraints the distribution has one

- 197 degree of freedom, meaning that any quantity not orthogonal to the constraint (e.g., mean, 198 median, standard deviation, etc.) is sufficient to describe the distribution. Here, we use the 199 median intensity, hereon denoted *I*, as it is less sensitive than the mean to the stochastic 200 uncertainty in the realization of extremes:  $F_i(x; \lambda_i, \kappa_i) = F(x; I_i)$ . The return level *x* associated 201 with the yearly exceedance probability *p* can be written as a function of median intensity and 202 expected number of yearly occurrences of the two types of ordinary events by inverting the 203 extreme value distribution  $\zeta(x)$ :  $x(p) = \zeta^{(-1)}(p; I_1, n_1; I_2, n_2)$ .
- We assume that temporal changes in the distribution of ordinary events will preserve 204 these local observational constraints. This resembles the assumptions behind regionalization 205 approaches in which spatial information is traded for record length (Buishand, 1991), but extends 206 its meaning in that (i) temporal changes are allowed, and (ii) the information on the interaction 207 between precipitating systems and local features provided by each individual station is fully 208 exploited (e.g. Marra et al., 2020a). To support our assumption, we test the significance of the 209 constraints in historical observations in a Monte Carlo framework by examining groups of non-210 consecutive years with consistently different median intensity (see Fig. S4) along the following 211 steps: (1) at each station and precipitation type, years are ranked according to the median 212 ordinary events intensity; (2) six 5-year subsets of non-consecutive years are created by selecting 213 three groups (15 years) from the largest intensity years and three from the smallest intensity; (3) 214 Weibull parameters are estimated at each station for the 5-year subsets; (4)  $10^3$  *m*-elements 215 216 synthetic samples, where *m* is the number of wet-days in the observed 5-year subsets, are generated according to the obtained distributions and the parameters describing the samples are 217 estimated to quantify the impact of parameter estimation uncertainty; (5) logarithmic relations 218 between the parameter pairs are derived for each subset; (6) the  $\alpha$  coefficient representing the 219 local constraint is compared to the distribution of coefficients of the logarithmic relations at (5). 220
- 221 3.3 Climate projections

Projected changes in median intensity and expected number of yearly occurrences of the 222 two precipitation types are obtained by examining the difference between the ends of the 21st 223 century (~2080-2100) and the 20<sup>th</sup> century (~1980-2005) under the RCP8.5 emission scenario 224 225 (Riahi et al., 2011). We estimated these differences using the data presented in Hochman et al. (2018a) and Zappa et al. (2015), calculated for 8 and 17 CMIP5 models, respectively. We choose 226 the changes in occurrence and median intensities from these two studies, as they are produced for 227 the desired time period and emission scenario, and because these parameters are considered more 228 robust than the changes in extremes that can be derived from the CMIP5 models themselves 229 (Fatichi et al., 2016). In particular, the changes in synoptic circulation over the study region 230 derived from CMIP5 ensembles were shown to be robust (Hochman et al., 2017; Hochman et al., 231 2018a; Zappa et al., 2015). 232

The acquired changes we used are: *Type-1*: expected number of yearly occurrence is projected to decrease by 15-35% ( $-25 \pm 10$  %); median intensity is projected to decrease by 20-25% ( $+22.5 \pm 5$  %); *Type-2*: expected number of yearly occurrence is projected to increase by 13% (+13  $\pm$  5 %); annual *Type-2* precipitation amounts are projected to remain unchanged,

which leads to a 12% decrease in the median intensity  $(-12 \pm 5 \%)$ . These numbers result in a

238 20-30% decrease in mean annual precipitation, which is consistent with the AR5 IPCC report

239 (IPCC, 2014). As based on relative differences between historic and future simulations, we

expect these projections to be less sensitive to systematic biases in the quantification of wet days

from CMIP5 models (e.g., too many drizzle days).

Changes in extreme return levels are computed in a Monte Carlo framework considering 242 uncertainties in the projections and in the local constraints (i.e., the  $\alpha$  coefficients), as follows. 243 At each station,  $10^3$  projections are created by (1) sampling the projected change in number and 244 median intensity of the two ordinary events types from normal distributions, and (2) sampling 245 the  $\alpha$  coefficient of the local constraint relations from the Monte Carlo realizations. Note that, 246 since the ratio between median and mean of Weibull distributions smoothly depends on the 247 shape parameter  $\kappa$  and is independent from the scale  $\lambda$ , one can safely assume a one-to-one 248 correspondence between projected changes in the mean and in the median (e.g., a 5% change in 249

the mean corresponds to  $\sim$ 5% change in the median). This is useful since the median is a better

descriptor for observed data whereas the mean is commonly provided by ESMs output.

### 252 **4** Application to the south-eastern Mediterranean

253

### 4.1 Local constraints on the distribution of ordinary events

While relations between scale  $\lambda$  and shape  $\kappa$  parameters of the ordinary events 254 distributions are not expected *a priori*, statistically significant relations (>3 $\sigma$  significance level) 255 are found for the given data when focusing on local groups of stations in which distinct local 256 features dominate the interactions with precipitation systems (Fig. 2a-c; Fig. S4). Dependence of 257 the form  $\kappa = \alpha \cdot \log \lambda + C$ , where  $\alpha$  and C are empirically-determined, was found to 258 approximate these relations in each group, generally explaining most of the observed variance 259 (Fig. S4). The hypothesis of a local constraint  $\alpha$  being significantly different from the 260 coefficients obtained from temporally splitting the records is rejected in all the cases (5% 261 significance level). Thus, the local values of  $\alpha$  indeed reflect historical changes in the median 262 intensity of ordinary events at each station, supporting the validity of the approach under 263 changing conditions (Fig. 2c; Fig. S4). It is worth noting that these relations are based on 264 historical observations and thus comprise observed changes in both dynamics and 265 thermodynamics. 266

The observed constraints imply that changes in the median intensity are linked to contrasting changes in extremes, i.e., decreasing median intensity decreases the precipitation amount yielded by typical ordinary events (Fig. 2d, e) but increases the probability associated with the largest events, and *vice versa* (Fig. 2f). This counter-intuitive behaviour is consistent with previous theory and observations of extreme precipitation, and supports the local constraints approach as a framework for quantifying changes in extremes (O'Gormann and Schneider, 2009; Pendergrass, 2018; Pendergrass and Knutti, 2018; Myhre et al., 2019; Wasko et al., 2018).



Figure 2. Local constraints on the intensity distribution of ordinary precipitating events. (a, b) Scatter 275 plots of the shape ( $\kappa$ ) and scale ( $\lambda$ ) parameters of the observed distributions for the two types of ordinary 276 events; colours refer to the five groups of stations as in Fig. 1a. (c) Example of the local constraint (Type-277 278 *I*, northern coast); triangles represent the median (among stations) parameters obtained in the split-sample 279 test using, for each station, groups of five non-consecutive years with increasing median intensity of the ordinary events; triangles thus represent historical variations of intensity. Local constraints for all cases 280 are shown in Fig. S4. (d) Schematic of the projection of changes in the intensity distribution of the 281 ordinary events along the constraints ( $\alpha = 0.3$ ,  $\lambda = 11.0$ , 9.0, 7.0 mm day<sup>-1</sup> and  $\kappa = 0.8$ , 0.74, 0.66; 282 black, blue and cyan, respectively); (e) event exceedance probability distributions associated with the 283 three pairs of scale and shape parameters shown in (d); (f) the largest 1% of the events in these 284 distributions. 285

286 4.2 Projections of future extremes

274

The sensitivity of extreme return levels to changes in the ordinary events (Fig. S5; Fig. S6) highlights that different return levels can have different responses, and that the local sensitivities associated with each event type can differ significantly. For example, in most of the region return levels are tied to changes in intensity and number of *Type-1* events, while changes in *Type-2* are crucial drivers for extreme return levels in the desert areas (Fig S5). Local changes in extreme return levels are thus related to mean (or median) changes in precipitation in a complex manner.

The projected changes in occurrence frequency of the two types (25% decrease and 13% 294 increase, respectively) and intensity (20-25% and 12% decrease, respectively), yield the changes 295 in the extreme return levels shown in Fig. 3 (see Fig. S7 for more details). An overall 5-20% 296 decrease of the 2-year return levels is seen, driven by the decrease in the occurrence frequency of 297 298 Mediterranean cyclones and in the median intensity of both types of systems. Since in the climatological setting of the region 2-year return levels roughly correspond to 99<sup>th</sup> wet-day 299 percentiles, this is consistent with previous results based on downscaling methods (Hochman et 300 al, 2018b). The picture is drastically different for the 100-year return levels which could not be 301 assessed in previous studies. Along the coast and in the southern desert, the negative sensitivity 302 to changes in the median intensity (Fig. S5; Fig. S6) dominates, and the rarest extremes are 303 projected to increase, consistently with Fig 2f. These results imply two adverse effects: (i) 304 amplified water scarcity and reduced flood and landslide risks in most of the region (Alpert et al., 305 2002; Samuels et al., 2009; Peleg et al., 2015); and (ii) increased intensity of the most severe 306 events along the coast and southern deserts, associated with augmented risk of extreme pluvial 307 308 flooding in coastal cities, and of flash floods, debris flows and geomorphic responses in the southern deserts (Shmilovitz et al., 2020; Rinat et al., 2020). 309

310





**Figure 3**. Projected changes in extreme precipitation return levels. Projected changes in 2-year and 100-



314 (difference between ~2080-2100 and ~1980-2005) under the RCP8.5 emission scenario. (a) Distribution

of the projected change and relative uncertainty (90% confidence interval considering uncertainties both

in climate projections and local constraints) shown as transposed cumulative distributions; colours refer to

the five groups of stations as in Fig. 1a. (b-c) Map of the projected changes for the 2-year (b) and 100-

318 year (c) return levels.

#### 319 **5 Discussion and conclusions**

320 In the south-eastern Mediterranean, the dominance of two precipitating systems and the availability of high-density local data makes it possible to simplify the statistical description of 321 ordinary precipitation events, and therefore of extreme events that emerge as the tails of their 322 distributions. Previous studies on the water resources of the region projected a "less rainfall, 323 more extremes" situation, with increased extremes insufficient to impact water resources in 324 generally drving conditions. However, these previous studies could not quantify changes in 325 extreme return levels and therefore risk (Alpert et al., 2002; Peleg et al., 2015). Combining 326 information on the occurrence frequency and intensity of the two dominant precipitation types 327 from ESM projections and observational constraints from rain stations, we show that the changes 328 in extreme return levels strictly depend on the sought probability. A tendency towards a general 329 decrease in the intensity of the 2-year events is found, together with an increase of the most 330 severe (100-year) events along the coast and in the desert areas. 331

The robustness of the synoptic variations in the RCP8.5 scenario in the region (Hochman 332 333 et al., 2018a; Zappa et al., 2015) and of the local constraints (Fig. S6), demonstrate the reliability of the proposed approach and the local projected response. Nevertheless, our predictions may be 334 refined by analysing additional scenarios and local data. It is plausible that similar improvements 335 in the projection of extremes can be made in other regions, even though projected changes in the 336 337 synoptic circulation systems might me less robust (Shepherd, 2014), calling for specific efforts to narrow this source of uncertainty. Additionally, future climate might reach some tipping point 338 after which the observational local constraints may no longer hold, a possibility that could be 339 tested using long simulations from convection-permitting models. For example, new synoptic 340 systems could be introduced in the region (such as tropical-like cyclones), or the track of existing 341 systems could change to such a degree that the interactions with local features might change 342 substantially, thus deviating from the observed constraints (e.g. northward shift of Mediterranean 343 cyclones track). Our results, which pertain to daily precipitation, assume no change in the spatial 344 345 structure of precipitation events at scales smaller than the resolutions of the used climate models. Improvements in the statistical description of the precipitating systems at multiple temporal and 346 spatial scales derived from observations and/or convection permitting models could fill this gap 347 by quantifying their structural response to external forcing (Cannon and Innocenti, 2019; Wasko 348 et al., 2016; Peleg et al., 2018; Marra et al., 2020b). 349

In contrast to traditional methods, the local constraints approach does not require long records; rather, it only requires local observations of ordinary events to constrain the intensity distributions. To this end, remotely sensed precipitation datasets represent a promising source of information for ungauged areas (Marra et al., 2019b). While uncertainty in ESMs remains a

- significant challenge to the community (Palmer and Stevens, 2019), our results point to increased
   investment in local measurements as an actionable and promising path to reduced uncertainty in
   the projection of extremes, independent of climate modelling efforts.
- The framework can be extended to other processes whose extremes emerge from
- 358 underlying distributions of ordinary events, such as extremes emerging from the combination of
- different physical phenomena, e.g. winds and storm surges from different types of cyclones
- 360 (Miniussi et al., 2020; Cavicchia et al., 2020)0. Similarly, it can be applied to phenomena whose
- intensity and occurrence may change independently, e.g. occurrence and maximum lifetime
   intensity of tropical cyclones (Knutson et al., 2010). In regions where local constraints can be
- intensity of tropical cyclones (Knutson et al., 2010). In regions where local constraints can be
   obtained, the approach proposed here can improve the predictability of climate change impact on
- extremes at scales relevant for impact studies, whose uncertainty was previously considered
- 365 irreducible due to modelling uncertainty and natural variability.

#### 366 Acknowledgements

- 367 The authors declare no conflict of interests. Precipitation data were provided by Israel
- 368 Meteorological Service (<u>https://ims.gov.il/en</u>, September 2018; data is freely available in Hebrew
- only) and Ministry of Water and Irrigation of Jordan, Technical Affairs, Studies Directorate,
- 370 Hydrological and Meteorological Information Systems (rainfall stations archives files, October
- 2018; available upon request to the data providers). The authors thank Prof. Pinhas Alpert for the
- 372 synoptic classification. Data and codes supporting the results are available at:
- 373 <u>https://doi.org/10.5281/zenodo.4286160</u> and <u>https://doi.org/10.5281/zenodo.3971558</u>. The linear
- 374 fit with uncertainty in x and y was performed based on the function by J. Browaeys, MATLAB
- 375 Central File Exchange (<u>https://www.mathworks.com/matlabcentral/fileexchange/45711-linear-</u>
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- the Israel Ministry of Science and Technology [grant no. 61792], Israel Science Foundation
- 378 [grant no. 1069/18], NSF-BSF [grant no. BSF 2016953], JNF [grant no. 90-01-550-18] and
- 379 Google [gift grant]. It is a contribution to the HyMeX program. Authors contribution:
- 380 Conceptualization: FM, MA, OA, DZ, EM; Formal analyses: FM; Data curation: FM, OG;
- Funding acquisition: EM, OA, CG, UD, DRE, YE; Paper writing: FM; Paper review and editing:
- all authors.

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#### Geophysical Research Letters

Supporting Information for

#### Towards narrowing uncertainty in future projections of local extreme precipitation

Francesco Marra<sup>1,2,\*</sup>, Moshe Armon<sup>1</sup>, Ori Adam<sup>1</sup>, Davide Zoccatelli<sup>1</sup>, Osama Gazal<sup>3</sup>, Chaim I. Garfinkel<sup>1</sup>, Dorita Rostkier-Edelstein<sup>1,4</sup>, Uri Dayan<sup>5</sup>, Yehouda Enzel<sup>1</sup>, and Efrat Morin<sup>1</sup>

<sup>1</sup>The Fredy and Nadine Herrmann Institute of Earth Sciences, The Hebrew University of Jerusalem, Israel <sup>2</sup>Institute of Atmospheric Sciences and Climate, National Research Council of Italy, CNR-ISAC, Bologna, Italy <sup>3</sup>Faculty of Agricultural and Environmental Sciences, Szent Istvan University, Hungary <sup>4</sup>Department of Environmental Physics, Environmental Sciences Division, IIBR, Ness-Ziona, Israel <sup>5</sup>Department of Geography, The Hebrew University of Jerusalem, Israel

#### Contents of this file

Tables S1 Figures S1 to S7 References

Code	Alpert et al. (2004)	Precipitation type in this study
1	Red Sea Trough with the Eastern axis	2
2	Red Sea Trough with the Western axis	2
3	Red Sea Trough with the Central axis	2
4	Persian Trough (Weak)	3
5	Persian Trough (Medium)	3
6	Persian Trough (Deep)	3
7	High to the East	3
8	High to the West	3
9	High to the North	3
10	High over Israel (Central)	3
11	Low to the East (Deep)	1
12	Cyprus Low to the South (Deep)	1
13	Cyprus Low to the South (Shallow)	1
14	Cyprus Low to the North (Deep)	1
15	Cyprus Low to the North (Shallow)	1
16	cold Low to the West	1
17	Low to the East (Shallow)	1
18	Sharav Low to the West	2
19	Sharav Low over Israel (Central)	2

**Table S1.** Synoptic systems in the semi-automatic classification by Alpert at al. (2004) and corresponding precipitation types used in this study, as follows: (1) *Type-1* (Mediterranean cyclones), (2) *Type-2* (other type of systems), (3) individually examined and labelled as *Type-1* if occurring up to 2 days after a *Type-1* wet day, and as *Type-2* in the remaining cases



**Figure S1.** Mean annual precipitation. (a) Mean annual precipitation in the south-eastern Mediterranean. Precipitation is mainly contributed by Mediterranean cyclones (*Type-1*) as shown in the other panels. (b) Relative contribution to the mean annual precipitation from *Type-2* events. Average yearly number of *Type-1* (c) and *Type-2* (d) wet days.



**Figure S2**. (a) Distribution of return levels and relative uncertainty (90% confidence interval from 103 bootstrap samples with replacement among the years in the record) shown as transposed cumulative distributions; colors refer to for the five groups of stations in Fig. 1a. (b, c) Map of the 10-year (c) and 50-year (d) return levels (10% and 2% yearly exceedance probability).



**Figure S3**. Extreme return levels computed using traditional approaches. Two-year and 100-year return levels (50% and 1% yearly exceedance probability) for daily precipitation amounts computed using traditional methods based on extreme value theory: a Generalized Extreme Value distribution is fitted to the annual maxima series using the method of the L-moments (Hosking, 1990). (a) Distribution of return levels and relative uncertainty (90% confidence interval from 103 bootstrap samples with replacement among the years in the record) shown as transposed cumulative distributions; median uncertainty of 17% (39%) for 2-year (100-year) return levels. The colors of the five groups of stations are as in Fig. 1a. (b, c) Map of the 2-year (c) and 100-year (d) return levels (50% and 1% yearly exceedance probability). Note the larger spatial variability (noise) of the return levels relative to Fig. 1, reflecting the largely increased uncertainty characterizing traditional methods based on the analysis of observed extremes (here, the annual maxima).



**Figure S4**. Local constraints of the parameters of the intensity distribution of the ordinary events. Scatter plot of the scale ( $\lambda$ ) and shape ( $\kappa$ ) parameters of the observed Weibull distributions for the two types of ordinary events and for the five groups of stations; colors as in Fig. 1a. The regressions used to define the local constraints ( $\alpha$  coefficients) are shown as dashed lines. Triangles represent the median (among stations) parameters obtained using, for each station, groups of five non-consecutive years with increasing median intensity of the ordinary events.



**Figure S5**. Sensitivity of extreme return levels to changes in the average characteristics of ordinary events. Sensitivity of 2-year (a-d) and 100-year (e-h) return levels (50% and 1% yearly exceedance probability) to changes in median intensity (*I*) and expected number of yearly occurrence (*n*) of the two types of ordinary events. The sensitivity or extreme return levels x(p) to changes in the median intensity and in the expected number of ordinary events is computed numerically using the partial derivatives of  $x(p) = \zeta^{(-1)}(p; I_1, n_1; I_2, n_2)$  with respect to the four variables, i.e.,  $\frac{\partial x(p)}{\partial I_1}, \frac{\partial x(p)}{\partial I_2}, \frac{\partial x(p)}{\partial n_1}$ , and  $\frac{\partial x(p)}{\partial n_2}$ . Units of change in x per 1% change in the predictor are quantified. The figure highlights two important points. First, different return levels can have different responses. Since practical applications rely on different return levels, this implies that the potential impact of climate change strongly depends on the application of interest. For instance, sewer systems are generally designed to be overtopped not more than once in 2-5 years. In contrast, dams, bridges and river dikes are designed for much lower probabilities of failure (e.g., 100- or even 1000-year). Second, the local sensitivities associated with each type of event can differ significantly. Therefore, local changes in extremes are related to mean (or median) changes in precipitation in a complex manner.



**Figure S6**. Sensitivity of extreme return levels to changes in the ordinary events. Distribution of the sensitivity of 2-year and 100-year return levels to changes in the average characteristics of the two types ordinary events shown as transposed cumulative distributions; uncertainties in the sensitivity to the median intensities are computed as 90% confidence interval of the local constrains (90% confidence interval of the  $\alpha$  coefficients). (a) Sensitivity to the *Type-1* intensity. (b) Sensitivity to the *Type-1* yearly number of events. (c) Sensitivity to the *Type-2* intensity. (d) Sensitivity to the *Type-2* yearly number of events.



**Figure S7**. Projected changes in extreme precipitation return levels (RCP8.5 emission scenario; 10-year and 50-year return levels). Projected changes in 10-year and 50-year return levels (10% and 2% yearly exceedance probability) for the end of the century (difference between ~2080-2100 and ~1980-2005) under the RCP8.5 emission scenario. (a) Distribution of the projected change and relative uncertainty (90% confidence interval considering uncertainties in climate projections and local constraints) shown as transposed cumulative distributions; colors as in Fig. 1a. (b-c) Map of the projected change for the 10-year (b) and 50-year (c) return levels.

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