

Extreme Precipitation Return Levels for Multiple Durations on a Global Scale

Gaby J Gründemann^{1,1,1}, Enrico Zorzetto^{2,2,2}, Hylke E Beck^{3,3,3}, Marc Schleiss^{4,4,4}, Nick Van de Giesen^{1,1,1}, Marco Marani^{5,5,5}, and Ruud J. van der Ent^{1,1,1}

¹Delft University of Technology

²Duke University

³Princeton University

⁴TU-Delft

⁵University of Padua and Duke University

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Abstract

Quantifying the magnitude and frequency of extreme precipitation events is key in translating climate observations to planning and engineering design. Past efforts have mostly focused on the estimation of daily extremes using gauge observations. Recent development of high-resolution global precipitation products, now allow estimation of global extremes. This research aims to quantitatively characterize the spatiotemporal behavior of precipitation extremes, by calculating extreme precipitation return levels for multiple durations on the global domain using the Multi-Source Weighted-Ensemble Precipitation (MSWEP) dataset. Both classical and novel extreme value distributions are used to provide an insight into the spatial patterns of precipitation extremes. Our results show that the traditional Generalized Extreme Value (GEV) distribution and Peak-Over-Threshold (POT) methods, which only use the largest events to estimate precipitation extremes, are not spatially coherent. The recently developed Metastatistical Extreme Value (MEV) distribution, that includes all precipitation events, leads to smoother spatial patterns of local extremes. While the GEV and POT methods predict a consistent shift from heavy to thin tails with increasing duration, the heaviness of the tail obtained with MEV was relatively unaffected by the precipitation duration. The generated extreme precipitation return levels and corresponding parameters are provided as the Global Precipitation EXtremes (GPEx) dataset. These data can be useful for studying the underlying physical processes causing the spatiotemporal variations of the heaviness of extreme precipitation distributions.

Highlights

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- Global precipitation return levels for 3-hour to 10-day durations are analysed
- Three different extreme value distributions are used to estimate the extremes
- The MEV distribution shows the most coherent spatiotemporal patterns
- Two distributions show a global shift from heavy to thin tails for longer durations

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Gaby J. Gründemann^{a,b}, Enrio Zorzetto^c, Hylke E. Beck^d, Marc Schleiss^e,
Nick van de Giesen^a, Marco Marani^f, Ruud J. van der Ent^{a,g,h}

^a*Department of Water Management, Faculty of Civil Engineering and Geosciences, Delft University of Technology, Delft, Netherlands*

^b*Centre for Hydrology, University of Saskatchewan, Canmore, Alberta, Canada*

^c*Program in Atmospheric and Oceanic Sciences, Princeton University, Princeton, New Jersey, USA*

^d*Department of Civil and Environmental Engineering, Princeton University, Princeton, New Jersey, USA*

^e*Department of Geoscience and Remote Sensing, Faculty of Civil Engineering and Geosciences, Delft University of Technology, Delft, Netherlands*

^f*Dipartimento di Ingegneria Civile, Edile ed Ambientale, Università degli Studi di Padova, Padova, Italy*

^g*Department of Physical Geography, Faculty of Geosciences, Utrecht University, Utrecht, Netherlands*

^h*Water Research Centre, School of Civil and Environmental Engineering, University of New South Wales, Sydney, Australia*

Abstract

Quantifying the magnitude and frequency of extreme precipitation events is key in translating climate observations to planning and engineering design. Past efforts have mostly focused on the estimation of daily extremes using gauge observations. Recent development of high-resolution global precipitation products, now allow estimation of global extremes. This research aims to quantitatively characterize the spatiotemporal behavior of precipitation extremes, by calculating extreme precipitation return levels for multiple durations on the global domain using the Multi-Source Weighted-Ensemble

Email address: g.j.gruendemann@tudelft.nl (Gaby J. Gründemann)

Precipitation (MSWEP) dataset. Both classical and novel extreme value distributions are used to provide an insight into the spatial patterns of precipitation extremes. Our results show that the traditional Generalized Extreme Value (GEV) distribution and Peak-Over-Threshold (POT) methods, which only use the largest events to estimate precipitation extremes, are not spatially coherent. The recently developed Metastatistical Extreme Value (MEV) distribution, that includes all precipitation events, leads to smoother spatial patterns of local extremes. While the GEV and POT methods predict a consistent shift from heavy to thin tails with increasing duration, the heaviness of the tail obtained with MEV was relatively unaffected by the precipitation duration. The generated extreme precipitation return levels and corresponding parameters are provided as the Global Precipitation Extremes (GPEX) dataset. These data can be useful for studying the underlying physical processes causing the spatiotemporal variations of the heaviness of extreme precipitation distributions.

Keywords: Precipitation extremes, MSWEP, Metastatistical extreme value distribution, Generalized extreme value distribution, Peaks-over-threshold, Global domain

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1. Introduction

Extreme precipitation events are a major contributor to natural disasters (CRED, 2019). Accurate estimates of the severity of intense precipitation events are needed for an enhanced disaster risk understanding, such as that

5 of floods and landslides. The urgency of this is indicated as the first priority
 6 of the Sendai Framework for Disaster Risk Reduction (UNSIDR, 2015). The
 7 accurate quantification of extremes is also necessary for infrastructure plan-
 8 ning and design. Some countries already provide spatiotemporal estimates
 9 of extreme precipitation based on extreme value distributions (EVDs), for
 10 example, for Australia (Ball et al., 2019), the Netherlands (Beersma et al.,
 11 2018), and the US (e.g., Perica et al., 2015, 2018). However, many countries
 12 and regions do not have sufficient local data available (Gründemann et al.,
 13 2018; Kidd et al., 2017; van de Giesen et al., 2014), such that spatially-
 14 distributed extreme precipitation estimates are not possible.

15 Several previous studies have developed global-scale datasets of extreme
 16 precipitation. Courty et al. (2019) calculated intensity-duration-frequency
 17 curves at the global domain and their scaling with different event durations
 18 using reanalysis data and the Generalized Extreme Value (GEV) distribu-
 19 tion with fixed tail behavior. Dunn et al. (2020) produced the HadEX3
 20 dataset, which contains 29 generic precipitation and temperature indices,
 21 although these indices are not based on EVDs. Furthermore, this dataset
 22 has a coarse 1.25° latitudinal \times 1.875° longitudinal resolution, with data-
 23 gaps due to insufficient available gauge data. Other global studies mostly
 24 focused on examining which type of distribution is most suitable to capture
 25 the tail behavior of extreme precipitation (Cavanaugh and Gershunov, 2015;
 26 Cavanaugh et al., 2015; Papalexiou et al., 2013). In addition, the spatial
 27 patterns of the parameter that controls the tail decay have been studied for
 28 the GEV distribution (Papalexiou and Koutsoyiannis, 2013; Ragulina and
 29 Reitan, 2017), and the Generalized Pareto (GP) distribution (Serinaldi and

30 Kilsby, 2014). However, several issues remain to be addressed in order to
 31 obtain global-domain extreme precipitation return levels: 1) the choice of
 32 the dataset with associated uncertainties, 2) the focus on daily durations,
 33 3) the choice of the time blocks over which block-maxima are determined,
 34 and 4) the exploration of possible alternatives to the classical EVDs and the
 35 associated uncertainty, especially with respect to the tail behavior.

36 1. Several (quasi-)global gridded precipitation datasets have been devel-
 37 oped in recent years, each with strengths, weaknesses, and uncertain-
 38 ties. See Sun et al. (2018), Beck et al. (2019a) and Rajulapati et al.
 39 (2020) for recent overviews of available datasets and their associated
 40 uncertainties. Most of these datasets are based on gauge, reanaly-
 41 sis, or satellite sensor data. Notable examples of gauge-based datasets
 42 include GPCC-FDR (Becker et al., 2013; Schneider et al., 2011) and
 43 REGEN (Contractor et al., 2020). However, gauges are extremely un-
 44 evenly distributed across the globe (Kidd et al., 2017; Schneider et al.,
 45 2014), and the number of active gauges has been declining in recent
 46 decades (Mishra and Coulibaly, 2009). Satellite-based products such
 47 as CMORPH (Joyce et al., 2004), GSMaP (Ushio et al., 2009), IMERG
 48 (Huffman et al., 2015), and PERSIANN (Hong et al., 2004) have a rel-
 49 atively high spatio-temporal resolution. However, they do not cover
 50 regions outside of 60°N/S , and are only available from 2000 onwards,
 51 which significantly hinders their use for extreme value analyses. Pre-
 52 cipitation products with a true global coverage and long records are
 53 reanalyses, such as ERA-5 (Hersbach et al., 2020), JRA-55 (Kobayashi
 54 et al., 2015), and MERRA-2 (Gelaro et al., 2017). However, reanaly-

sis products tend to exhibit strong systematic biases in the magnitude and frequency of precipitation (Decker et al., 2012; Liu et al., 2018; Ménégos et al., 2013).

2. Global-scale analyses of precipitation extremes are generally based on daily precipitation records (Cavanaugh et al., 2015; Koutsoyiannis, 2004a,b; Papalexiou and Koutsoyiannis, 2013; Papalexiou et al., 2013; Ragulina and Reitan, 2017; Serinaldi and Kilsby, 2014). In practice, however, multiple durations are needed for the design of infrastructure (e.g., Nissen and Ulbrich, 2017) or urban drainage networks (e.g., Mailhot and Duchesne, 2009). It is known that precipitation extremes of different durations scale differently with temperature (Wasko et al., 2015), but little is known about the variation of EVD properties (tail behavior) for different temporal resolutions. Studies that did derive extreme precipitation statistics for durations ranging from minutes to a few days have mostly focused on small regions (McGraw et al., 2019; Nissen and Ulbrich, 2017; Overeem et al., 2008).

3. Studies estimating return levels of extreme precipitation by using annual maxima typically use calendar years to delineate the annual periods from which maxima values are extracted (e.g., De Paola et al., 2018; Marani and Zanetti, 2015; Papalexiou and Koutsoyiannis, 2013; Ragulina and Reitan, 2017; Villarini et al., 2011). When the variable of interest is river discharge instead of precipitation, however, hydrological years are typically used instead of calendar years to determine the annual maxima (Ward et al., 2016). For discharge values this is important, since peak discharge and flooding could occur during 31

December to 1 January transition and one event would be included in two calendar years. Although not often considered, this could also happen for precipitation. The annual maxima method could pick multiple values from a single rainy season that may, for example, be highly influenced by the El Niño-/Southern Oscillation, which is known to impact precipitation extremes (Allan and Soden, 2008; Rasmusson and Arkin, 1993).

4. The Generalized Extreme Value (GEV) distribution, the most widely used EVD, is typically fitted through one of two approaches: a) using annual maximum precipitation series and maximum likelihood (Coles, 2001) or L-moment (Hosking, 1990) estimation approaches, or b) using a Peak-Over-Threshold (POT) method to fit a Generalized Pareto Distribution to excesses above the threshold and a Poisson process to the sequence of threshold exceedances (Coles, 2001). In contrast to GEV and POT, the recently developed Metastatistical Extreme Value (MEV) distribution is fitted using all events with recorded precipitation instead of only the most severe. The inclusion of more events reduces the uncertainty due to sampling effects, which is important when dealing with short time series (Hu et al., 2020; Marani and Ignaccolo, 2015; Marra et al., 2018, 2019a; Miniussi and Marani, 2020; Zorzetto et al., 2016; Zorzetto and Marani, 2019). This is particularly advantageous when analyzing short remote sensing precipitation products, as the commonly applied GEV requires many years of data to accurately estimate the tail of the distribution (Papalexiou and Koutsoyiannis, 2013). Additionally, GEV parameter estimation depends heavily on a

few large values, which makes it very sensitive to the possible presence of outliers, a relatively common occurrence in remote sensing estimates of precipitation amounts (Zorzetto and Marani, 2020). The GEV tail behavior is mostly controlled by its shape parameter, which is very sensitive to sampling effects and the choice of the method used for estimation. To overcome these problems, some studies have suggested to use one universal value of the shape parameter that is applicable to the whole world Koutsoyiannis (2004a,b), or a shape parameter value within a narrow range between exponential and heavy-tail behavior (Papalexiou and Koutsoyiannis, 2013), or one shape parameter per region, that is similar within climate types and elevation ranges (Ragulina and Reitan, 2017). The estimation of the shape parameter is particularly difficult with short data series, though crucial for the accurate estimation of extremes.

In this study we contribute to overcome these issues by 1) using a dataset that merges all three main sources of precipitation data, 2) estimating extremes for several event durations, 3) using hydrological years in our analyses, and 4) comparing results from three different extreme value methods (GEV, POT and MEV). Specifically, we are interested in quantitatively characterizing the behavior of extreme precipitation and the spatiotemporal variation of extreme value distributional tails at the global domain.

126 2. Material and Methods

127 2.1. Data

128 The global precipitation product used in this study is the Multi-Source
129 Weighted-Ensemble Precipitation (MSWEP-V2.2) dataset. MSWEP is par-
130 ticularly suited for our purpose due to its global coverage, long temporal
131 span, high spatial and temporal resolution. We used data from 1 January
132 1979 to 31 October 2017 at a 0.1° latitude \times 0.1° longitude resolution at
133 3-hourly time steps. We selected all land-cells between 90°N and 58°S for
134 our analysis. MSWEP precipitation estimates are derived by merging five
135 different satellite- and reanalysis-based global precipitation datasets. The
136 dataset is one of the few precipitation products with daily (as opposed to
137 monthly) gauge corrections, applied using a scheme that accounts for gauge
138 reporting times (Beck et al., 2019b). MSWEP has shown robust performance
139 compared to other widely used precipitation datasets (e.g., Alijanian et al.,
140 2017; Bai and Liu, 2018; Beck et al., 2017, 2019a; Casson et al., 2018; Hu
141 et al., 2020; Sahlu et al., 2017; Satgé et al., 2019; Zhang et al., 2019), thus
142 underlying its potential for improving the characterization of extreme pre-
143 cipitation worldwide. We refer to Beck et al. (2019b) for a comprehensive
144 description of the dataset.

145 2.1.1. Quality Control

146 The integration of erroneous gauge observations into MSWEP-V2.2 can
147 occasionally result in implausible precipitation values. Therefore, we imple-
148 mented a three-step quality control procedure of the 3-hourly data prior to
149 the analysis. We first discarded negative values, which are physically impos-

150 sible. The second step was to discard outliers, which we defined as values
 151 deviating from the mean by more than 30 standard deviations. We also dis-
 152 carded data surrounding the outliers for the same time step using a 11×11
 153 grid-cell window, as erroneous gauge observations may have influenced sur-
 154 rounding cells in the production of the MSWEP dataset. The third step was
 155 to remove years with > 30 discarded days or < 5 ‘wet’ 3-hourly periods,
 156 identified using a threshold of $0.2 \text{ mm } 3\text{h}^{-1}$ following Wasko et al. (2015).
 157 Finally, we only included in the analysis data from grid cells with at least
 158 30 years of data remaining, as a minimum record length of 30 years is cus-
 159 tomary and recommended to obtain reliable results (Arguez and Vose, 2011;
 160 Kendon et al., 2018; Westra et al., 2013).

161 *2.1.2. Durations and Identification of Independent Events*

162 The durations we selected for our analysis are 3, 6, 12 and 24 hours, and
 163 2, 3, 5 and 10 days. In order to create statistically-independent precipitation
 164 events for multiple durations, we first separated 3-hourly events following the
 165 declustering method to limit the autocorrelation of the samples described in
 166 Marra et al. (2018, their Section 3.1). For longer durations, independent
 167 events are the maximum intensities within each independent event and non-
 168 overlapping period using moving windows (Marra et al., 2020).

169 *2.1.3. Hydrological Year*

170 A common challenge in global-scale assessments is the delineation of the
 171 hydrological year, given the regional variability in the climatological precip-
 172 itation seasonality. We therefore developed an uniform way to define the
 173 hydrological year. To avoid splitting one rainy season over two different

174 years, we computed the median of the monthly precipitation for each grid-
 175 cell, and defined the start of the hydrological year to be the first day of the
 176 driest month. Supplementary Material Figure S1a shows the starting month
 177 of the hydrological year as determined by this method. These data are also
 178 available in the GPEX dataset (Gründemann et al., 2021). As MSWEP-V2.2
 179 spans the interval from 1 January 1979 to 31 October 2017, we discarded the
 180 data prior to the start of the first hydrological year, thus keeping 38 complete
 181 years. Only where the hydrological year starts in December there are just 37
 182 complete years, which occurs in 5.8 % of the grid cells.

183 We also investigated whether there is a significant difference between the
 184 use of calendar and hydrological years for the estimated daily extremes for
 185 GEV and MEV. The POT method is based on the values over a high thresh-
 186 old, irrespective of when they occurred. Therefore, there is by definition no
 187 difference in calculating the extremes using hydrological or calendar years
 188 for the POT method. To determine the difference for GEV and MEV, we
 189 first calculated the daily return levels for normal calendar years, using the
 190 MSWEP data from 1979 to 2016. Second, we calculated the return levels for
 191 the same distributions and the same years, by removing the months before
 192 the start of the hydrological year from the year 1979 and adding them to the
 193 year 2016. We did this in order to use the exact same data, so the differences
 194 in the return level estimates are solely due to a different starting month.

195 *2.2. Extreme Value Distributions*

196 Three extreme value distributions were fitted to the MSWEP data to
 197 calculate extreme precipitation return levels: the GEV, POT, and MEV
 198 distributions. Annual (hydrological year) maxima were used to estimate the

three parameters of the GEV using the L-moments approach, because of its robust performance for small samples (Hosking, 1990). The GEV cumulative distribution function (CDF) is given by:

$$G(z) = \begin{cases} \exp \left\{ - \left[1 + \xi \left(\frac{z-\mu}{\sigma} \right) \right]^{-\frac{1}{\xi}} \right\}, \xi \neq 0 \\ \exp \left\{ -\exp \left[- \left(\frac{z-\mu}{\sigma} \right) \right] \right\}, \xi = 0 \end{cases} \quad (1)$$

with location parameter $\mu \in (-\infty, \infty)$, scale parameter $\sigma > 0$, and shape parameter $\xi \in (-\infty, \infty)$. The annual extremes estimated by GEV are translated into those of the parent distribution, following Koutsoyiannis (2004a, equation 3).

As a second EV model we use a Peaks Over Threshold approach, describing precipitation accumulations exceeding a high threshold using a GP distribution, while modelling the frequency of threshold exceedances using a Poisson point process (Coles, 2001; Davison and Smith, 1990). This framework also yields GEV as the resulting extreme value distribution, which is then used to determine the quantile corresponding to a given return period. The GP CDF is given by:

$$H(y) = \begin{cases} 1 - \left(1 + \frac{\xi y}{\beta} \right)^{-\frac{1}{\xi}}, \xi \neq 0 \\ 1 - \exp \left(-\frac{y}{\beta} \right), \xi = 0 \end{cases} \quad (2)$$

where $y > 0$ are precipitation excesses over the threshold, with $\beta > 0$ and $\xi \in (-\infty, \infty)$ the GP scale and shape parameters respectively. A relevant aspect in applying the POT model is a suitable choice of the threshold used to define precipitation exceedances. Our global-scale application requires studying the distribution of precipitation extremes across markedly different climatic regions, thus excluding the adoption of a constant threshold value.

219 We studied the effect of the threshold choice using multiple threshold selec-
 220 tion methods on a global sample of grid cells (see Supplementary Material
 221 Section 2 and and Figure S3). Our results showed that this choice had a lim-
 222 ited effect on the estimated return levels (Figure S3a). We chose to perform
 223 our global analysis by selecting for each cell a threshold value such that it
 224 is exceeded on average 3 times each hydrological year. As a consequence of
 225 this choice, the sample size available for fitting the GP distribution remains
 226 constant across different precipitation durations. The method used to fit
 227 the GP distributions is the Probability Weighted Moments (PWM; e.g., see
 228 Hosking and Wallis, 1987).

229 The third model applied here is the MEV distribution (Hosseini et al.,
 230 2020; Hu et al., 2020; Marani and Ignaccolo, 2015; Miniussi et al., 2020a,b;
 231 Zorzetto et al., 2016). In the MEV framework, all “ordinary” precipitation
 232 events, i.e. all events above a small threshold, are used to infer this EV
 233 distribution. The threshold we applied is $0.2 \text{ mm } 3\text{h}^{-1}$, coinciding with
 234 the earlier defined ‘wet event’. Weibull parameters were estimated for each
 235 hydrological year separately, based on all wet events using the PWM method
 236 (Greenwood et al., 1979) as done in Zorzetto et al. (2016). The MEV-Weibull
 237 CDF is given by:

$$\zeta_m(x) = \frac{1}{M} \sum_{j=1}^M \left\{ 1 - \exp \left[- \left(\frac{x}{C_j} \right)^{w_j} \right] \right\}^{n_j} \quad (3)$$

238 where j is the hydrological year ($j = 1, 2, \dots, M$), $C_j > 0$ is the Weibull scale
 239 parameter, $w_j > 0$ is the Weibull shape parameter, and n_j is the number of
 240 wet events observed in hydrological year j (Marani and Ignaccolo, 2015).

241 It should be noted that the methods we applied in this study do not

242 use any parameter bounds. Although Papalexiou and Koutsoyiannis (2013)
 243 argued that the GEV shape parameter of daily precipitation lies between ex-
 244 ponential and heavy-tail behavior, this is not used as additional information
 245 to constrain our fits. Doing so would cause artificial breaks in the obtained
 246 spatiotemporal patterns, the analysis of which is the main objective of this
 247 study. Moreover, the scientific debate on bounds is not settled, especially
 248 for durations longer than a day, and different bounds are used in different
 249 studies (e.g., Blanchet et al., 2016; Yilmaz et al., 2017). In order to avoid
 250 underestimation of extremes in practical settings, our dataset (Gründemann
 251 et al., 2021) also includes the Gumbel estimates which may be used as a
 252 lower bound (see Supplementary Material Section 6).

253 *2.2.1. Observed Return Period*

254 The MSWEP dataset analyzed here has 38 complete years of data. There-
 255 fore, the empirical return period associated with the maximum value on
 256 record computed according to the Weibull empirical frequency estimate is
 257 $T_{\text{observed}} = 39$ years. However, only 91 % of all cells had 38 complete years
 258 of data, so the maximum observed return period is sometimes lower: for 7 %
 259 of the cells only 37 complete years were available, and for 2 % of the cells
 260 36 years or less were available. However, for simplicity we still refer to the
 261 corresponding maximum return level as T39 in the results.

262 *2.2.2. Tail Behavior*

263 Both the GEV and MEV distributions are flexible and can describe dif-
 264 ferent tail behaviors. They are, therefore, appropriate models to study the
 265 characteristics of local precipitation extremes. The tail behavior of the two

266 distributions differs, as illustrated in Figure S4 for different combinations of
 267 scale and shape parameters. The shape parameter ξ of the GEV distribution,
 268 obtained either through the annual maxima or POT approach, encodes the
 269 nature of the tail of the distribution. Based on the value of ξ , the GEV can
 270 take one of three forms: a positive GEV shape parameter ($\xi > 0$, “*Fréchet*”)
 271 corresponds to a power-law tail, i.e., to a slowly-decaying probability of large
 272 events. This heavy-tail behavior contrasts with the case of an exponential
 273 tail ($\xi = 0$, “*Gumbel*”), and with the case of a distribution with an upper end
 274 point, which corresponds to negative values of the shape parameter ($\xi < 0$,
 275 “*inverse Weibull*”).

276 The MEV distribution assumes that precipitation events are Weibull-
 277 distributed. The tail decay of this distribution is controlled by its shape
 278 parameter: for $w < 1$ its tail behavior is “sub-exponential”, i.e., heavier than
 279 that of an exponential (recovered for $w = 1$), albeit with a characteristic
 280 scale (Laherrere and Sornette, 1998; Wilson and Toumi, 2005). For $w > 1$
 281 the Weibull tail is super-exponential, with a fast decaying tail, while still
 282 retaining an infinite upper end point. Hence, the shape parameter of the
 283 Weibull distribution encodes the propensity of a site to be subjected to large
 284 extreme events (Wilson and Toumi, 2005; Zorzetto et al., 2016). However,
 285 the tail decay of the MEV distribution is not only dependent on that of
 286 ordinary values (through w) but is also affected by the yearly number of
 287 events (Marra et al., 2018) and by the inter-annual variations of C_j , w_j and
 288 n_j .

289 In an effort to compare the heaviness between the distributions, we have
 290 come up with a measure of heaviness that is based on the return levels them-

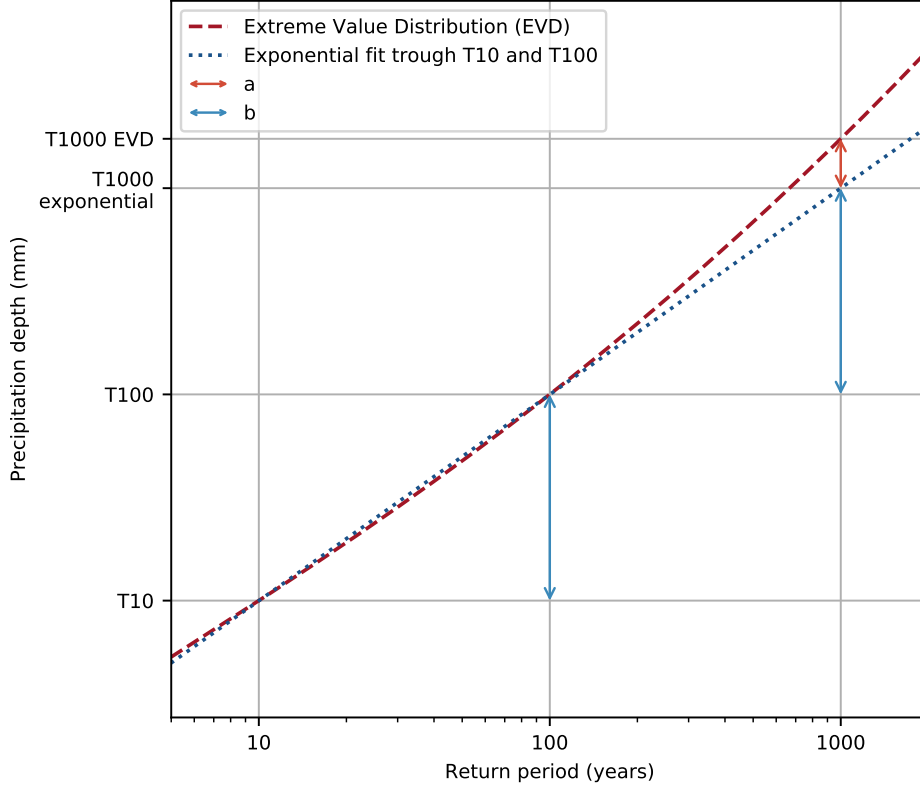


Figure 1: Illustration of our method to measure the tail heaviness for any distribution based on return levels only.

291 selves (Figure 1). The difference between the 1000-year return level and the
 292 10-year return level can be described as follows:

$$T1000 = T10 + b + b + a \quad (4)$$

293 Where b is the difference between the 100-year and 10-year return level,
 294 i.e.: $b = T100 - T10$, and a is the additional increase caused by the heaviness
 295 of the tail (Figure 1). A positive a is indicative of heavy tails and a negative

296 a of thin tails. For pure exponential tails it holds that $a = 0$. The value for
 297 a is highly dependent on the local precipitation systems, so we defined the
 298 heaviness amplification factor $h_{T10-T100-T1000}$ to be a normalization of a :

$$h_{T10-T100-T1000} = \frac{a}{b} = \frac{T1000 - 2 \times T100 + T10}{T100 - T10} \quad (5)$$

299 In words, the meaning of $h_{T10-T100-T1000}$ is the fractional additional in-
 300 crease between T1000 and T100 that is more than the increase that could be
 301 expected from a pure exponentially tailed distribution. A distribution has a
 302 heavy tail for $h > 0$ and a thin tail for $h < 0$. Here, we chose a range for the
 303 heaviness metric over large return periods from 10 to 1000 years, since the
 304 1000-year return levels are known to be influenced by the distribution choice
 305 (e.g., Rajulapati et al., 2020) and that is precisely what we wanted to com-
 306 pare. Yet, it should be noted that this metric may easily be adjusted to other
 307 return periods and other factors between the return periods. For GEV and
 308 POT the heaviness metric is independent of the return period range as long
 309 as the return periods are a factor 10 apart. Although for MEV this heaviness
 310 metric is only valid for the return period range over which it is computed,
 311 using other ranges (T2-T20-T200 and T5-T50-T500) did not yield significant
 312 differences (Figure S6).

313 **3. Results and Discussion**

314 *3.1. Hydrological Year*

315 Figure 2 shows the frequency distribution of 1000-year return levels esti-
 316 mated using calendar and hydrological years for GEV and MEV. The spatial
 317 distribution of the T1000 differences is presented in Supplementary Material

318 Figure S1b for GEV and Figure S1c for MEV. We found that in the case
319 of GEV quantiles, the fraction of sites characterized by differences within
320 $\pm 0.5\%$ is larger than that observed for MEV. When the hydrological year
321 starts in the winter months, the hydrological year is only shifted by a few
322 months. In such instances, the annual maxima mostly stay the same between
323 the calendar and hydrological years, though the included events could differ.
324 For GEV this means that for many cells there is almost no difference in the
325 T1000 estimates, whereas for MEV the difference is small.

326 On the other hand, when the offset with a calendar year is approximately
327 6 months, around June, there are many different events included in the hy-
328 drological years compared to the calendar years. This results in different
329 annual maxima and large differences in the estimated extremes for GEV
330 and MEV. The differences are most pronounced in the Southern hemisphere
331 and in locations where the hydrological year starts around June, e.g., in
332 the Mediterranean region, in the Middle-East, in Southern Africa, in Brazil,
333 around Indonesia, and in the western US (Figure S1a). For MEV the overall
334 sensitivity in T1000 estimates remains lower than that of GEV. In particular,
335 the distribution of differences in Figure 2 exhibits thicker tails for GEV (e.g.,
336 as measured by the wider 5th to 95th percentile interval). This suggests that
337 regional sensitivity to the definition of block maxima can be quite significant
338 for the GEV approach.

339 Figure S2 in the supplementary material presents the frequency distribu-
340 tions of all analyzed return levels. The lower return levels are less impacted
341 by the start of the hydrological year than the higher ones.

3.2. Extreme Precipitation Estimates

Figure 3 shows the 100-year precipitation return levels for a 24-hour duration. Extreme value estimates for other durations and return periods are featured in the Global Precipitation EXtremes (GPEX) dataset (Gründemann et al., 2021). The spatial patterns of the extremes estimated by GEV and MEV are similar to Zorretto and Marani (2020, their Figure 9), while the spatial pattern of the underlying GEV parameters are consistent with Courty et al. (2019, their Figure 1). The global spatial pattern of return levels for the three EV methods is similar, although large regional differences can be observed. The GEV and POT results are similar in magnitude and show similar differences when compared to MEV. The estimated precipitation extremes are generally lower for both GEV and POT compared to MEV quantiles. MEV estimates exhibit smooth spatial patterns, whereas the spatial patterns using GEV and POT are more irregular, consistent with the results of Zorretto and Marani (2020) for the conterminous US. The reduced spatial coherence in patterns of extremes for GEV and POT is particularly evident in the Great Plains of North America, and in Northern Russia, Southeast Asia, and Central Africa. Other extreme value approaches and distributions may also yield more coherent spatial patterns of precipitation extremes (e.g., Rajulapati et al., 2020), but comparison of all possible extreme value approaches was not the scope of this study. Furthermore, our analysis (Figure 3) reveals the presence of a large number of circular areas with heavier extremes, corresponding to the location of gauges used for correcting precipitation estimates in the MSWEP algorithm (Beck et al., 2019b). The effect of these local corrections is much larger for traditional EV models (POT and

367 GEV), while MEV appears less sensitive to these local corrections.

368 In order to study the ability of the three distributions to capture the
369 spatial coherence of precipitation extremes, we selected several case study
370 areas. They collectively cover a wide range of climates and domain sizes, the
371 locations of which can be found in Figure 3a. Within a single case study
372 area, we expect the precipitation estimates to be statistically homogeneous
373 because of their precipitation generating mechanisms (Cavanaugh and Ger-
374 shunov, 2015; Cavanaugh et al., 2015) or elevation (Ragulina and Reitan,
375 2017). Figure 4a shows the coefficient of variation (CV) of T100 extreme
376 precipitation estimates for these case studies. The CV is the ratio of the
377 standard deviation to the mean and is used to compare the relative variation
378 between the study areas. The higher the CV, the higher the relative spread
379 of the precipitation estimates within a spatial domain. This figure shows
380 quite similar behavior for GEV and POT, though POT has a slightly lower
381 spread. The CV for MEV is lower, which points to more spatially coherent
382 T100 precipitation estimates based on single point time series (with 38 years
383 of training data).

384 To further investigate the global differences in magnitude between the
385 three methods, we examine the extremes for each distribution using a spa-
386 tially weighted mean over the global land surface. This is displayed for
387 multiple return periods and durations as depth-duration-frequency curves
388 (Figure 5). We first compare the maximum precipitation observed in the
389 dataset to the precipitation predicted from each distribution. As there are
390 38 complete years of MSWEP data, the maximum empirically observed re-
391 turn level is 39 years (T39 observed, the black dotted line in Figure 5). While

392 locally the empirical T39 estimate could be very different from the true re-
 393 turn level, we expect the global average of this value to be representative
 394 of the true T39. For GEV and POT, we expected the estimated T39 to be
 395 close to the observed value since only the largest values are used to fit these
 396 distributions. For MEV, we did not necessarily expect a good agreement for
 397 T39, but its performance should be better for return levels greater than the
 398 length of the observation time series (Marra et al., 2018, 2019b; Schellander
 399 et al., 2019; Zorzetto et al., 2016). The results in Figure 5 show that for
 400 the short duration events, the observed T39 is close to the T39 for all three
 401 distributions. For increasing durations, the deviation between empirically
 402 observed and EV modeled T39 quantiles increases, particularly for MEV.
 403 This could be because a smaller number of events per year is used for the
 404 fit of MEV-Weibull, whereas the number of events used for the fit of GEV
 405 and POT remains constant for all durations. Both GEV and POT show an
 406 underestimation and MEV an overestimation. This figure also shows again
 407 that the differences between GEV and POT are small. The global average
 408 estimated extremes for GEV and POT are notably lower than for MEV, as
 409 was already visible from Figure 3. This difference is more pronounced for
 410 larger return periods and longer durations.

411 One reason the quantiles estimated using MEV are higher than using
 412 GEV and POT is related to the increase in estimation uncertainty of Weibull
 413 parameters when the number of events per hydrological year is low. This is
 414 especially relevant in arid regions and for long durations. For instance, for 5
 415 and 10-day durations the average annual number of events is 36 and 21 events
 416 respectively. It is therefore possible that this leads to an overestimation by

MEV. To overcome this, windows of two or more years could result in a better parameter estimation (Miniussi and Marani, 2020). A second factor which may be relevant for MEV quantile estimates is the use of a fixed threshold for defining a precipitation event.

3.3. Tail Behavior

To better understand the differences between extremes estimated using the three extreme value methods, we analyze their tail behavior using the heaviness amplification factor $h_{T10-T100-T1000}$ (Eq. 5). Figure 6 presents $h_{T10-T100-T1000}$ for a 24-hour duration worldwide for each of the three distributions. We refer to Figures S7-S13 in Section 4 in the supplementary material for maps of $h_{T10-T100-T1000}$ for the other durations. Both GEV (Figure 6a) and POT (Figure 6b) exhibit a large spatial variability in addition to a low spatial coherence. This makes it difficult to discern clear spatial patterns with the exception of a few notable regions. For instance, in the Amazon, $h_{T10-T100-T1000}$ is mostly negative, suggesting a tail with an upper limit, while in Eastern and Southern Australia $h_{T10-T100-T1000}$ it is strongly positive, denoting strong heavy tail behavior. This map roughly corresponds to the spatial patterns of the daily GEV shape parameter shown by Papalexiou and Koutsoyiannis (2013, their Figure 13) and Ragulina and Reitan (2017, their Figure 4). We also find that for the GEV and POT methods, grid cells associated with heavy tails can be adjacent to cells with thin tails. Furthermore, GEV and POT do not always show the same type of tail, heavy or thin, in the same grid cells. In 72 % of the cases the sign of the underlying shape parameter agrees, while in 28 % of the cases the signs are different for daily precipitation. This highlights the large uncertainty as-

sociated with estimating reliable tail parameters from short time series and the sensitivity of the GEV and POT methods to sampling effects.

The heaviness of the MEV distribution (Figure 6c) shows a more coherent spatial pattern. At virtually all grid cells the heaviness amplification factor $h_{T10-T100-T1000}$ (Eq. 5) indicates heavy tail behavior and there is a high consistency within geographic regions and for all durations (Figures S7-S13). Based on previous studies (Cavanaugh et al., 2015; Papalexiou and Koutsoyiannis, 2013; Papalexiou et al., 2013; Ragulina and Reitan, 2017), this predominantly heavy-tail behavior of daily precipitation was expected and is well captured by MEV. There are also topographical patterns visible in the heaviness amplification factor (Figure 6c), though they are not as clearly distinguishable as for the shape parameter itself (Figure S5). The heaviness tends to be higher in arid areas, and lower in mountainous areas. Examples of arid areas with high heaviness include the Sahara, the Namib and Kalahari in Africa, the Gobi, Thar and Taklamakan in Asia, the Atacama Desert in South America, large areas of Southwestern Australia, and the Arabian desert and other areas in the Middle East. This same pattern is to a lesser extent also visible for the heaviness of GEV (Figure 6a) and POT (Figure 6b).

At high elevations a small $h_{T10-T100-T1000}$ is usually found for MEV (Figure 6c). Examples include the Rocky Mountains and the Sierra Madres in North America, the northern Andes and large areas of the Brazilian Highlands in South America, the Ethiopian Highlands, the Scandinavian Mountains, and the Tibetan Plateau. These spatial patterns are in contrast with what Papalexiou et al. (2018, their Figure 6) found for hourly Weibull tails in

467 the USA, where the heaviest tails are in the mountainous areas, and the thin
468 tails are in the south-east. However, our results correspond well to Ragulina
469 and Reitan (2017, their Figure 4), who showed that heaviness decreases with
470 elevation.

471 A comparison of the heaviness for different distributions and durations
472 is presented as a boxplot in Figure 7. For spatial maps of the heaviness
473 for the different durations we refer to Figures S7-S13. For GEV and POT,
474 predominantly heavy tails are observed for short durations and thinner tails
475 for long durations. Furthermore, GEV and POT both show a decreasing
476 variability in the heaviness for longer durations, indicated by both shorter
477 whiskers and boxes. The decrease of the heaviness of the tails for increasing
478 durations is in line with the findings of Cavanaugh and Gershunov (2015),
479 who found that longer duration extremes exhibit thinner tails. For GEV
480 and POT the longer durations largely indicate tails with a finite upper end
481 point. This occurs for instance in half of the cases for a duration of 10
482 days for GEV, and more than half for POT. One implication of this finding
483 is that, when computing return levels for a single location (see Figures S3
484 and S12), it is possible for the very large return periods that the shorter
485 duration quantiles are more intense than the longer duration quantiles. This
486 is physically impossible (see Figure S14a,b,f and g), and we should thus be
487 extremely careful when interpreting such results.

488 MEV, on the other hand, shows different heaviness patterns than GEV
489 and POT (Figure 7 and Figures S7-S13). MEV shows almost entirely a
490 heavy-tail behavior, which remains consistent across the range of durations
491 examined. Furthermore, also the variability for MEV is constant across du-

492 rations, though with a slight increase for longer durations. The MEV distri-
 493 bution thus produces a spatially and temporally coherent heavy tail behavior
 494 based on a 38 years calibration sample and a single grid-cell analysis. This
 495 is a promising result, as MEV, in contrast to the traditional methods ana-
 496 lyzed, provides essential information on the spatial coherence of precipitation
 497 extremes without any prior hypothesis on its spatial structure, for example
 498 through a spatial clustering scheme (Demirdjian et al., 2018). In fact, the
 499 spatial structure of the tail heaviness obtained through the MEV analysis
 500 could be used as a measure of statistical homogeneity for regionalization
 501 studies.

502 In this work, we studied the global distribution of rainfall extremes based
 503 on stationary statistical models. It has long been recognized that climatic
 504 change as well as the inherent variability of the climate system modulate the
 505 frequency and intensity of heavy rainfall, prompting the adoption of non-
 506 stationary models in water resources management (Milly et al., 2008; Yilmaz
 507 and Perera, 2014; Gu et al., 2017). However, in addition to the uncertainty
 508 originating from inference on finite length measurements, adopting a non-
 509 stationary description of extreme rainfall leads to more complex statistical
 510 models, leading in turn to a potentially increased uncertainty of predicted
 511 return levels (Lins and Cohn, 2011; Montanari and Koutsoyiannis, 2014;
 512 Serinaldi and Kilsby, 2015; Milly et al., 2015). This issue can be mitigated
 513 by linking shifts in the rainfall process to mechanistic physical processes. For
 514 instance, this can be achieved by linking rainfall statistics to climate features
 515 that are more easily predicted by global circulation models (e.g., Grimm and
 516 Tedeschi, 2009; Whan et al., 2020; Zorzetto and Li, 2021; Fowler et al., 2021).

517 Elucidating these links at the global scale and for the broad range of rainfall
518 durations explored here remains a daunting task which we plan to explore in
519 future work.

520 4. Conclusions

521 The aim of this research was to quantitatively characterize the spatiotem-
522 poral variation of global precipitation extremes and their associated extreme
523 value distribution tails. We have fitted three different extreme value methods
524 (GEV, POT, and MEV) to a global precipitation dataset, MSWEP V2.2, to
525 estimate extreme precipitation return levels for several durations. In order
526 to compare the tails of the three distributions, we introduced a novel heav-
527 iness amplification factor $h_{T10-T100-T1000}$ (Eq. 5). Instead of using calendar
528 years to delineate between different years, we used hydrological years, the
529 start of which we defined as the driest month. To our knowledge, this is a
530 novel approach for analyzing precipitation extremes on the global domain.
531 We demonstrated that there is a substantial difference in the extremes de-
532 pending on the definition of yearly blocks used in the extreme value analysis
533 (Figure 2). These differences were most notable in the Southern hemisphere,
534 and in locations where the driest month occurs around June (Figure S1). Al-
535 though there is no systematic bias, we still recommend to apply the extreme
536 value analyses for estimating extreme precipitation based on hydrological
537 years in future studies. Our analysis indicates that this can be particularly
538 relevant in the Southern hemisphere and in regions characterized by marked
539 seasonal cycles.

540 It is well known that the traditional GEV and POT methods require very

541 long data series for accurate estimation of the tail behavior, and our study
 542 confirms that there is a low spatial coherence for the tail properties of both
 543 distributions (Figure 6a and b) using just 38 years of training data. The tail
 544 properties of the MEV distribution are spatially more coherent (Figure 6c)
 545 and hence the estimated return levels are more spatially coherent as well
 546 (Figure 3c). This spatially coherent behavior, consistent with previous results
 547 obtained over the conterminous US (Zorzetto and Marani, 2020), shows that
 548 the MEV distribution is able to capture spatially consistent tail behavior
 549 from short time series and by a single grid-cell analysis, without any prior
 550 information on the spatial precipitation structures. The analysis of the MEV
 551 tail behavior reveals distinct spatial patterns, as the heaviness appears to be
 552 controlled by climate zones and orography. Heavier tails are observed in arid
 553 areas, and thinner tails in mountainous regions. More in-depth analyses are
 554 necessary to draw definite conclusions on what exactly controls the heaviness
 555 of extreme value distribution tails. The performance of MEV is promising
 556 for regions without long local precipitation records. Furthermore, our study
 557 shows that the tail behavior captured by MEV is coherent and heavy both
 558 spatially and temporally (Figures 6, 7 and S7-S13). For GEV and POT, on
 559 the other hand, the tail behavior decreases with increased event duration,
 560 resulting in a thin tail with a finite endpoint for about half of the cells for a
 561 duration of 10 days.

562 We also conclude that both GEV and POT generally underestimate the
 563 observed extremes, whereas MEV overestimates them (Figure 5). This occurs
 564 particularly for long-duration extremes and large return periods. We do
 565 consider it likely, however, that the results could be improved, for instance

566 by changing the event threshold or by fitting the Weibull distribution over
567 two or more years for dry areas (Miniussi and Marani, 2020), so as to reduce
568 inter-annual variability of the parameters due to samples of limited length.
569 Our results suggest that this issue is particularly relevant at the longest
570 durations examined. For GEV and POT the results could also be improved
571 by adopting spatial models (Davison et al., 2012; Huser and Wadsworth,
572 2020).

573 The data generated for this study are openly available as the GPEX
574 dataset (Gründemann et al., 2021). These data include extreme precipitation
575 return levels and extreme value distribution parameters for durations between
576 3 hours and 10 days at a global gridded 0.1° resolution. They could be used
577 by engineers as a reference of precipitation extremes for data-scarce regions
578 in particular. For scientific purposes, all underlying parameters are also
579 available and can be used to answer several outstanding questions, such as:
580 what are the controls on the tail behavior of extremes, and what is driving
581 the different changes in tail heaviness with duration for GEV, POT, and
582 MEV?.

583 **Declaration of Interest**

584 The authors declare that they have no known competing financial inter-
585 ests or personal relationships that could have appeared to influence the work
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599 The GPEX dataset is available at the 4TU repository Gründemann et al.
600 (2021). The data included are the extremes estimated using the different
601 distributions, the observed extremes, and the parameters to estimate the ex-
602 tremes. These data are available for all durations included in this study. The
603 resolution of the dataset is 0.1° , the resolution of the MSWEP-V2.2 dataset.
604 For more information we refer to the Dataset Usage Notes in Section 5 of the
605 supplementary material.

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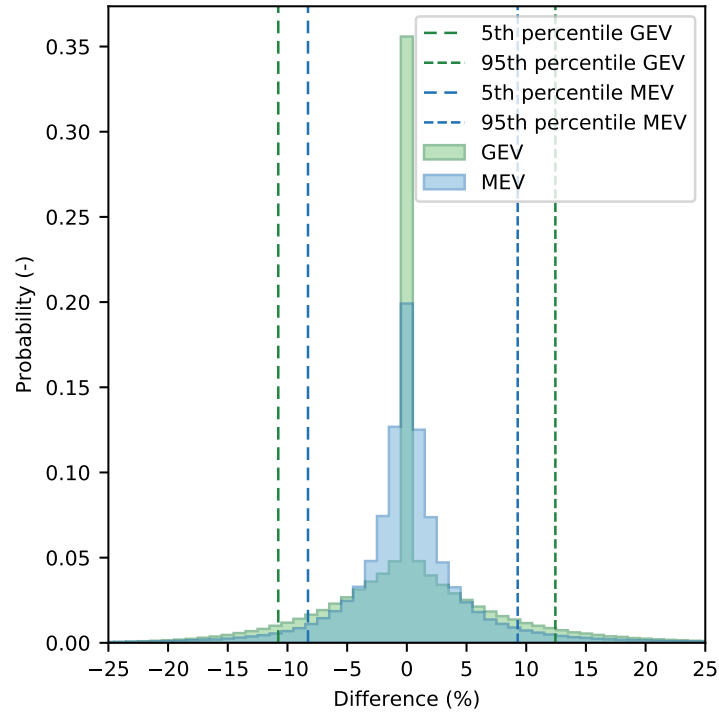


Figure 2: Weighted histogram showing the percentage difference in the values of T1000 quantiles calculated using calendar years and hydrological years. Included in the figure are all cells where the start of the hydrological year is different than the calendar year (i.e., the hydrological year does not start in January, see Supplementary Material Figure 1a). A negative difference indicates that the T1000 estimate is larger using hydrological years, whereas a positive difference indicates that the T1000 estimate is larger using calendar years.

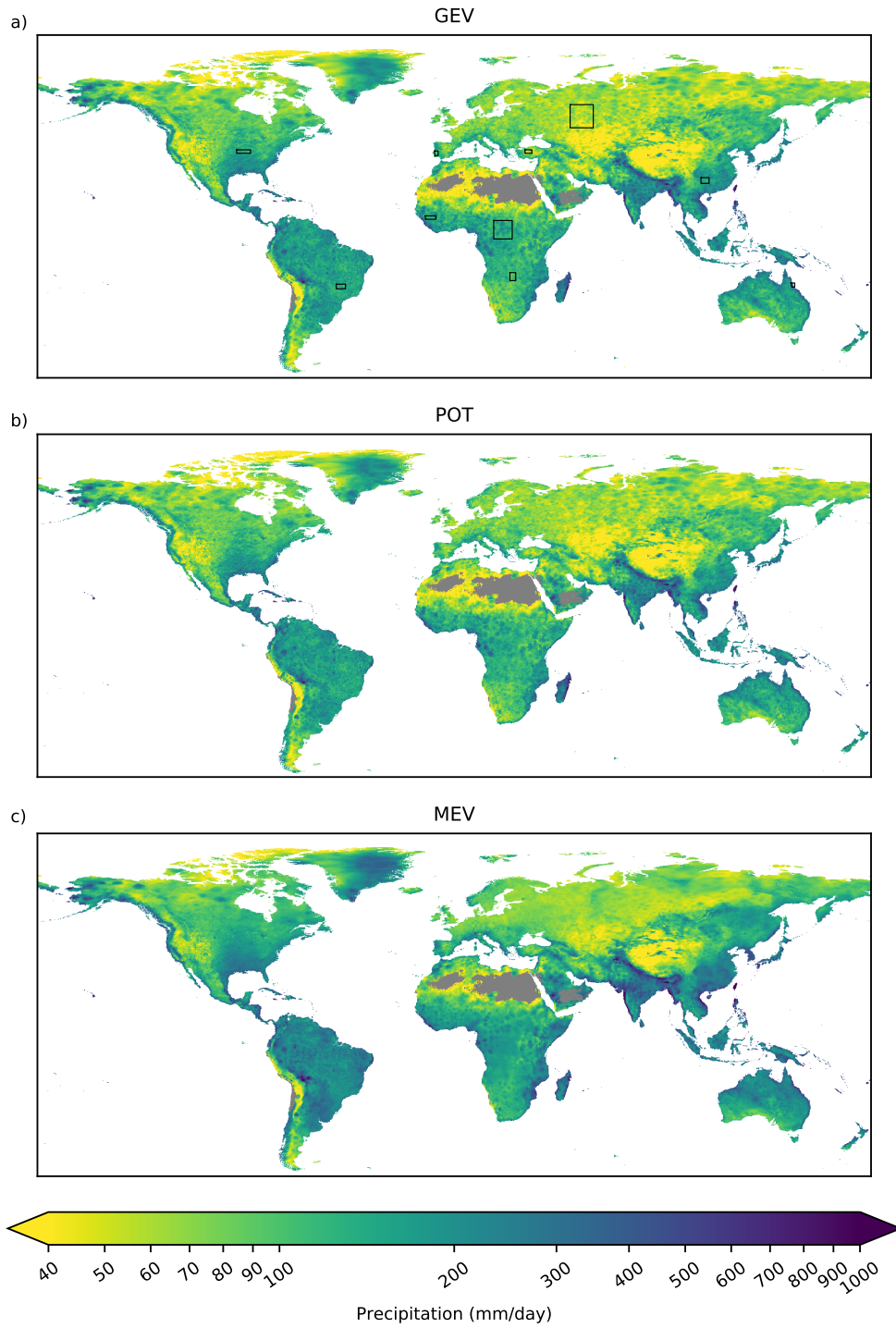


Figure 3: Precipitation return levels with a duration of 24-hours for a 100-year return period for different extreme value distributions: (a) the Generalized Extreme Value (GEV) distribution, (b) the Peak Over Threshold (POT) method, and (c) the Metastatistical Extreme Value (MEV) distribution. The black rectangles in panel a are the case studies corresponding to the areas in Figure 4.

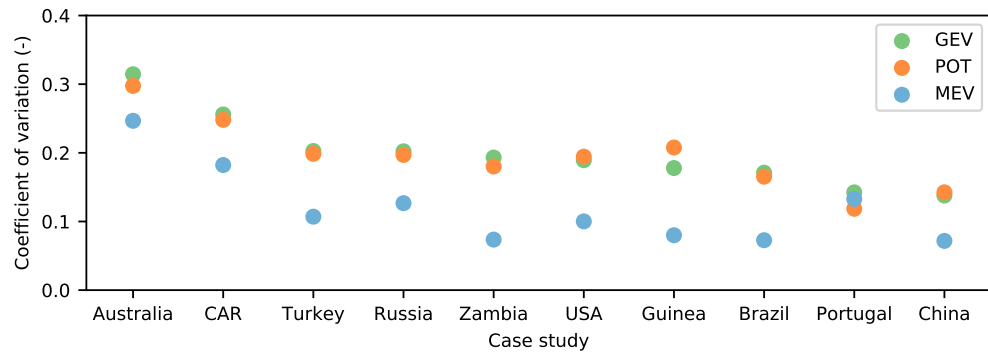


Figure 4: Coefficient of variation for the difference in estimated T100 quantiles for the three extreme value methods for 24-hour precipitation at selected case study areas. The coefficient of variation is the standard deviation of the precipitation divided by the mean precipitation. The locations of the case study areas are displayed in Fig 3a.

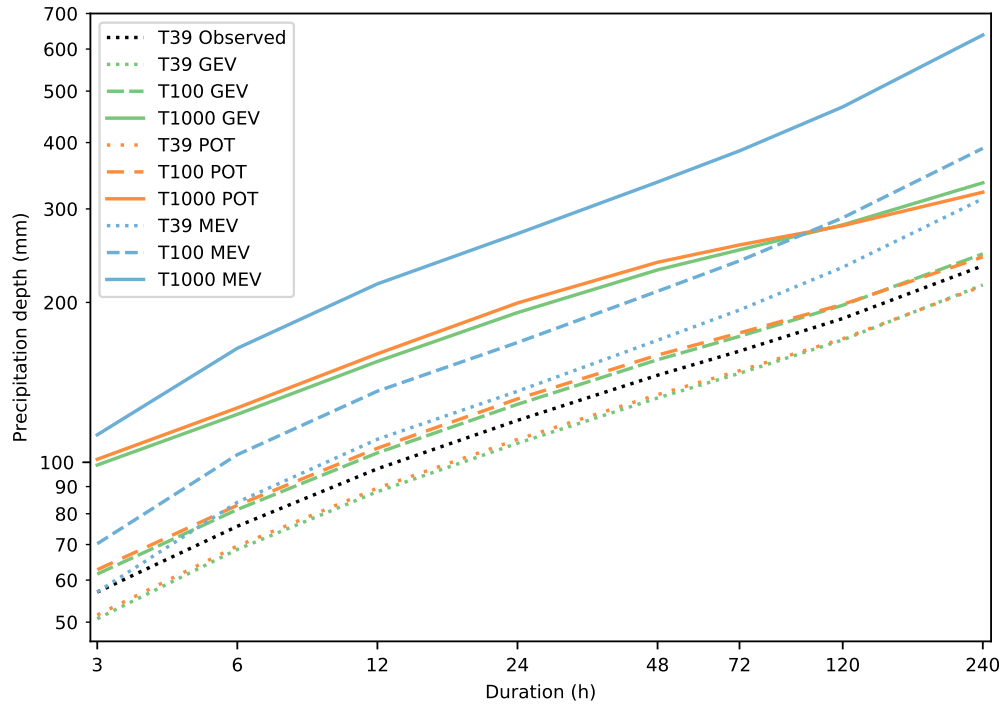


Figure 5: Area-weighted average depth-duration-frequency curves for the global land surface. T39 Observed is the mean spatially weighted maximum precipitation observed in the MSWEP-V2.2 dataset.

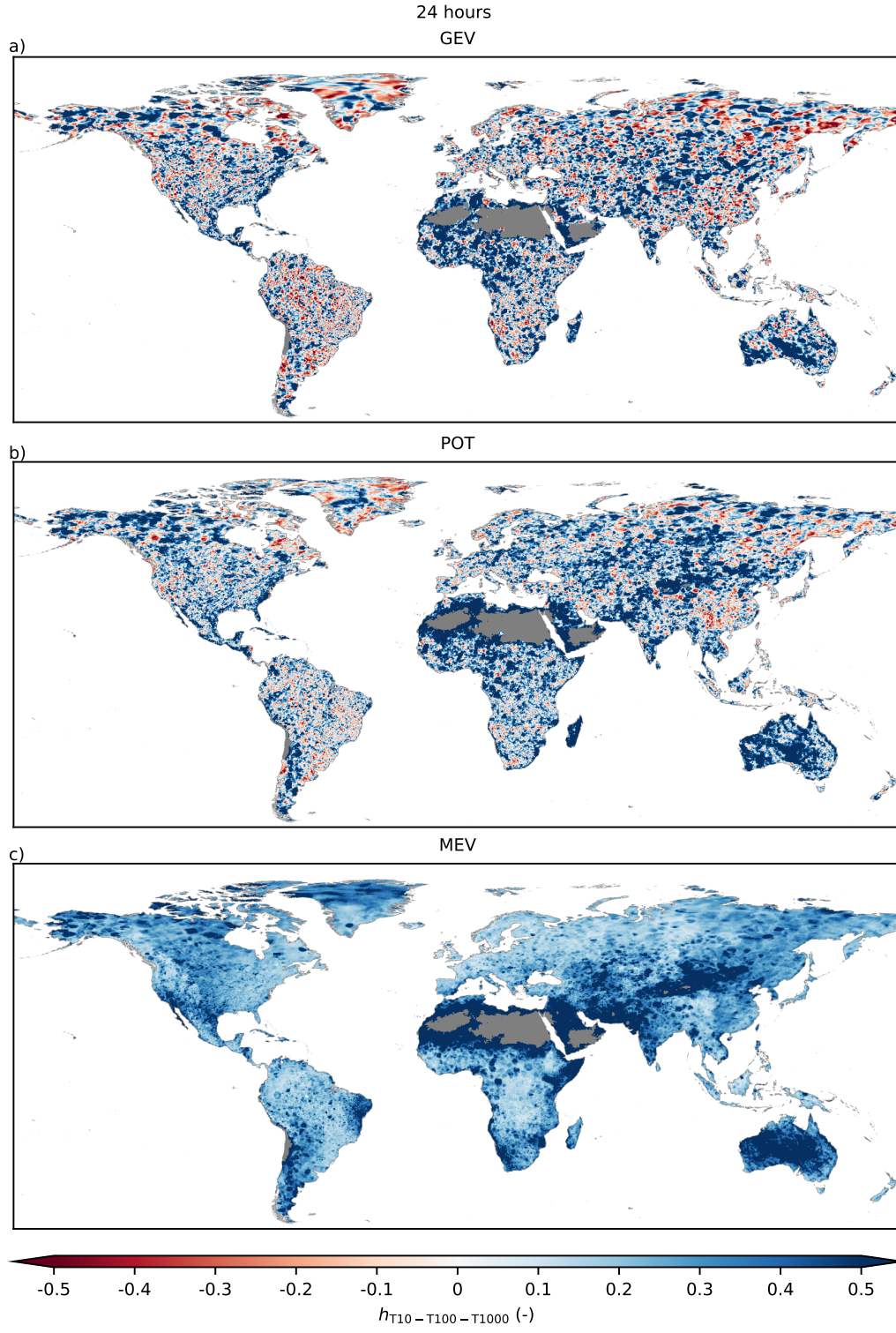


Figure 6: The heaviness amplification factor $h_{T_{10}-T_{100}-T_{1000}}$ (Eq. 5) for daily precipitation calculated for different extreme value methods: (a) GEV, (b) POT, (c) MEV. Red indicates a thin tail, white an exponential tail, and blue a heavy tail. See section 2.2.2 for more information on the heaviness metric, and Figures S7-S13 for maps of $h_{T_{10}-T_{100}-T_{1000}}$ for the other durations.

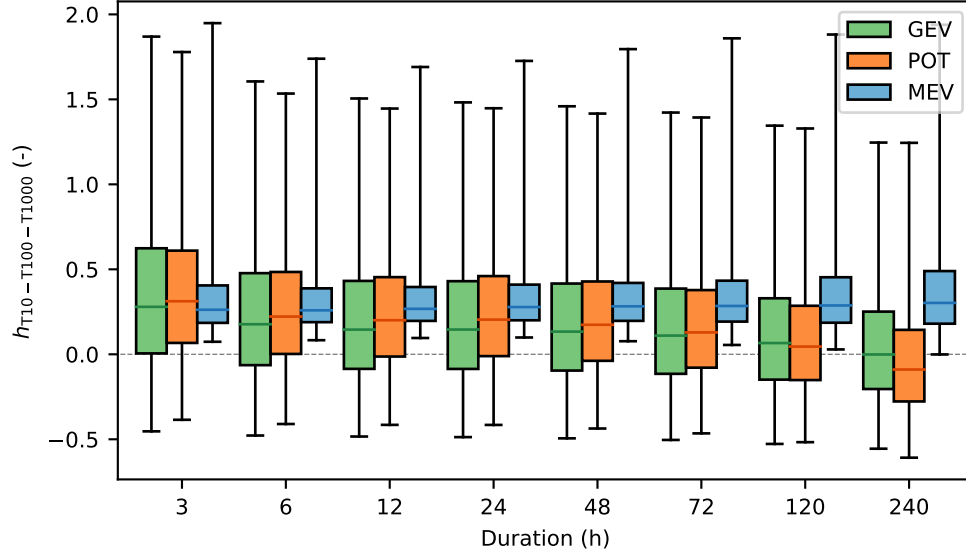


Figure 7: Boxplots showing the distribution of the heaviness amplification factor $h_{T10-T100-T1000} (-)$ for different durations and extreme value methods: (a) GEV and POT, and (b) MEV. The whiskers denote the 1st and 99th percentiles. The top and bottom of the boxes represent the 75th and 25th percentiles, respectively. The dashed gray horizontal lines indicate exponential tails. See section 2.2.2 for more information on the heaviness metric.

Supplementary Material for “Extreme Precipitation Return Levels for Multiple Durations on a Global Scale”

August 25, 2021

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1 Calendar and Hydrological Year

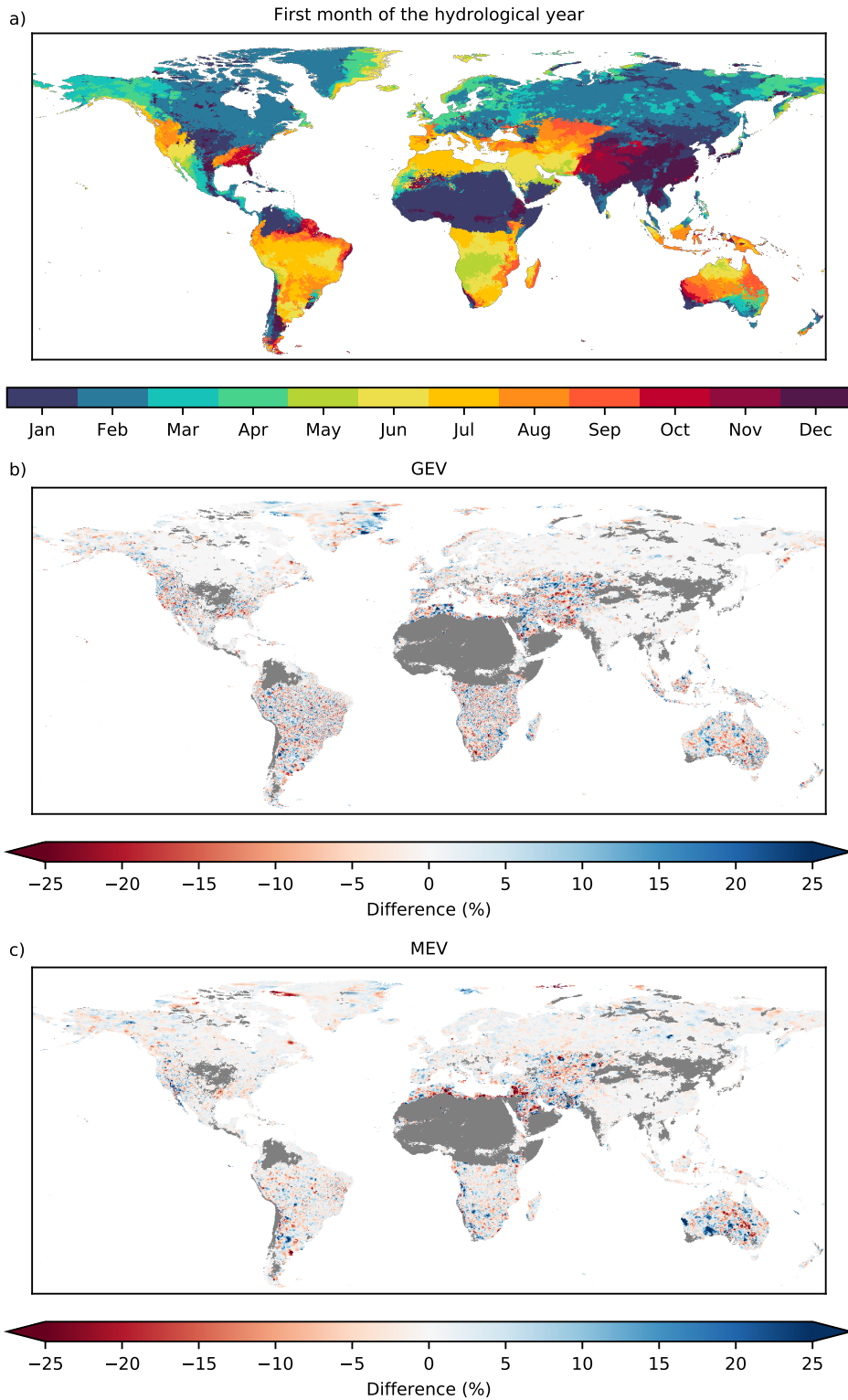


Figure S1: (a) The month indicating the start of the hydrological year. (b) The percentage difference of the daily 1000-year return levels of calendar and hydrological years for GEV. (c) The percentage difference of the daily 1000-year return levels of calendar and hydrological years for MEV. A negative difference indicates that the T1000 estimate is larger using hydrological years, whereas a positive difference indicates that the T1000 estimate is larger using calendar years.

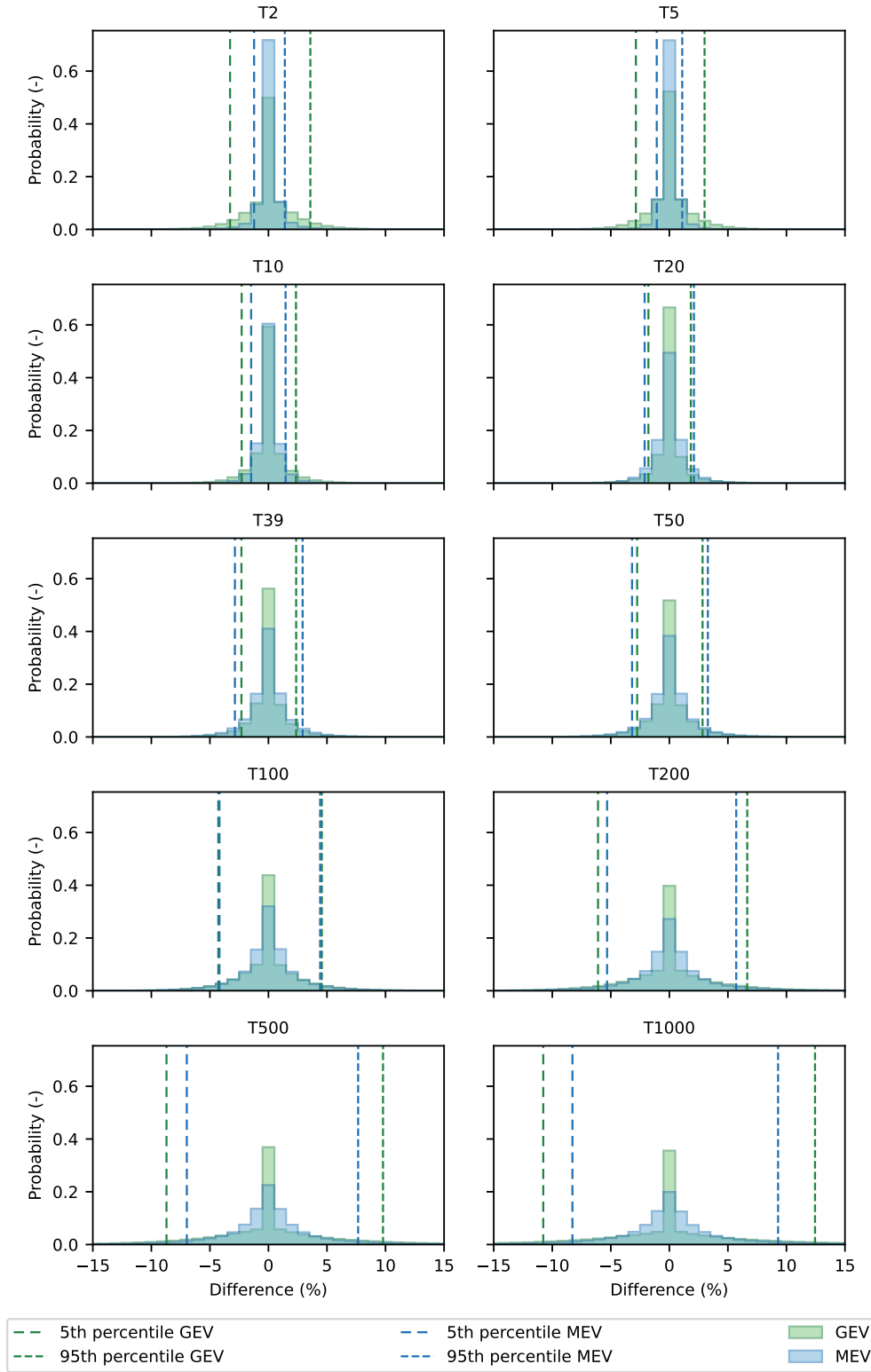


Figure S2: Weighted histograms showing the percentage difference in the values of several return periods (T, see titles) calculated using calendar years and hydrological years. Included in the figure are all cells where the start of the hydrological year is different from the calendar year (i.e., the hydrological year does not start in January, see Figure S1a). A negative difference indicates that the return period estimate is larger using hydrological years, whereas a positive difference indicates that the return period estimate is larger using calendar years.

2 Threshold Analysis for POT

There are many different ways to select an appropriate threshold for the Peak Over Threshold (POT) analysis, such as a value over a specific threshold, a percentage, or an average number of events per year. We refer to Caeiro and Gomes (2016); Langousis, Mamalakis, Puliga, and Deidda (2016) for recent overviews of such threshold selection. Which method is the most effective remains to be an unanswered question. For our global-scale application, we have analyzed to take on average 1 to 5 events per year. This fixed number of events, as opposed to the events over a predefined threshold, ensures that the number of events per year remains constant for all durations, and allows for analyses on the global domain.

The 100-year return levels for all durations and thresholds are shown in Figure S3a. The boxplot shows minimal difference in the T_{100} estimates for the different thresholds for the short durations, and a minor difference for the longer durations, namely that the estimated T_{100} are slightly larger for inclusion of less events per year. Overall the figure shows only slight differences, so more information has to be considered for accurate threshold selection.

A comparison of the three parameters for all durations and thresholds is shown in Figure S3b-d. The location and scale parameters in S3b and c show very similar results for the different thresholds. The shape parameter in S3d, however, has more variation for the different POT-thresholds. For all durations, the variability of the shape parameter decreases for the inclusion of more events per year, so more events to fit the distribution to. For durations between 3 and 48 hours, the shape slightly increases for the inclusion of more events per year. In other words, the shape is lower for on average 1 event per year, and higher for on average 5 events per year for durations between 3 and 48 hours. For 72 hours, the median shape parameter of the different thresholds is constant, though the variability decreases for the inclusion of more events. For the 5 and 10-day durations (120 and 240 hours, respectively), the shape parameter slightly decreases for the inclusion of more (non-extreme) events. The general pattern of the GP shape parameter is similar to that of the GEV, as they both show a decreasing shape parameter for increasing durations.

Previous studies that have looked into the GP shape parameter on the global domain focused on daily durations. Papalexiou and Koutsoyiannis (2013) estimated the mode of the shape for daily precipitation as 0.134, but displaying a large spread. Serinaldi and Kilsby (2014) estimated the shape for four different seasons based on 1898 stations with more than 100 years of data. They found the shape depends on the season and varies between 0.061 and 0.097, also displaying a large spread. Furthermore, they found that if you lower the threshold and include more non-extreme events, the shape parameter is higher. Our results are similar to that of Serinaldi and Kilsby (2014), as the median for daily precipitation for the different thresholds varies between 0.711 and 0.918, and we also found that the inclusion of more non-extreme events leads to a higher shape parameter for daily precipitation.

In short, the above analysis shows that the threshold selection is of minor influence on the estimation of the 100-year return level. Of the three underlying parameters, the threshold selection has minimal influence on the location and scale parameter, though a much greater influence on the shape parameter. The variability of the shape parameter is the highest for the threshold of on average 1 event per year, and decreases when more non-extreme events are used to fit the distribution to. Furthermore, the shape parameter increases for the inclusion of more events for the shorter durations (3-hours to 2-days), remains constant for a 3-day duration, and decreases for the 5 and 10-day durations. Based on these analyses, we chose to show the threshold of on average 3 events per year in the main manuscript, as that threshold has the right balance of a low variability and not as much of an underestimation of the 100-year return levels for the longest durations.

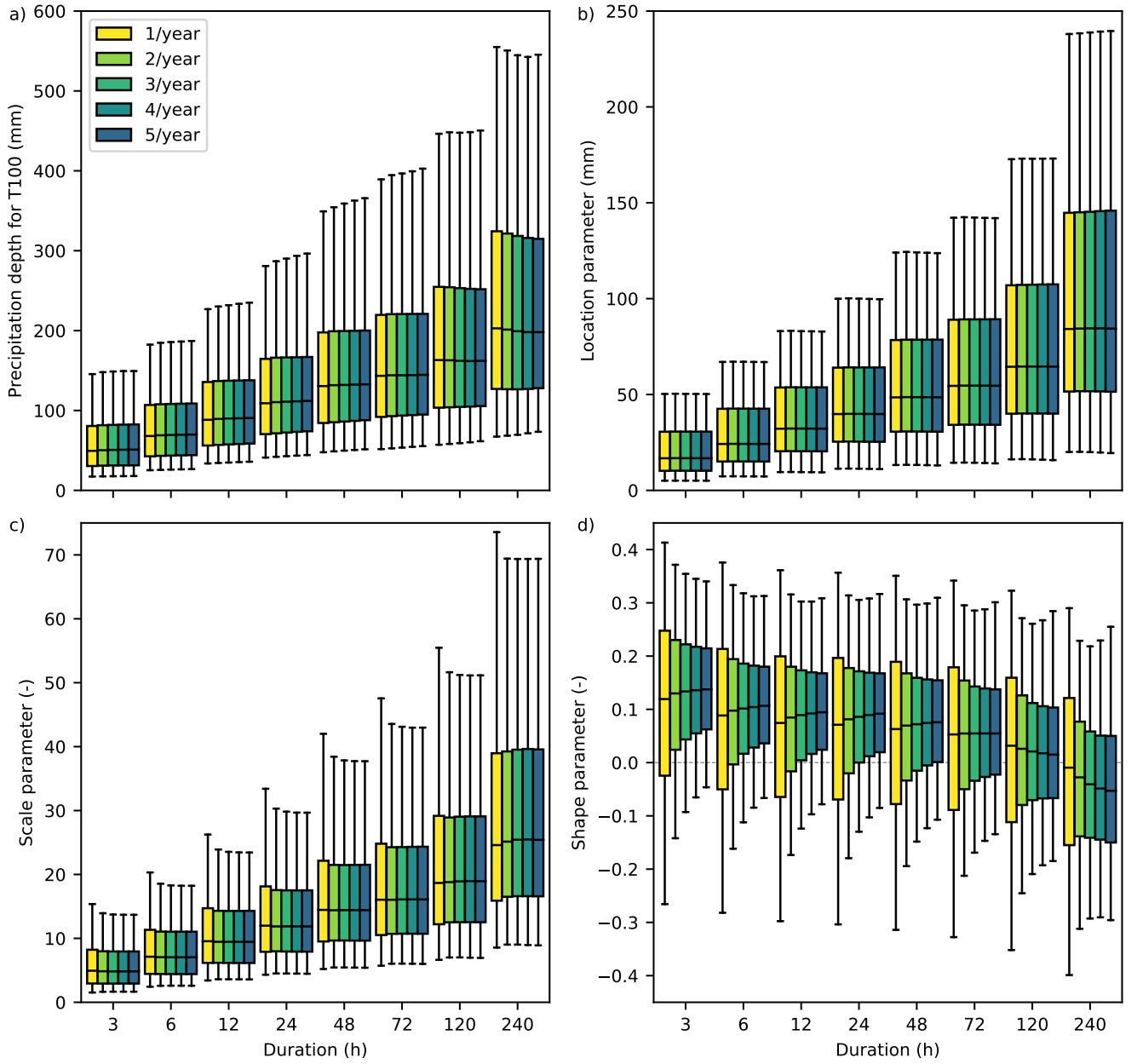


Figure S3: Weighted boxplot showing the distribution of (a) the 100-year precipitation return levels, and the (b) location, (c) scale, and (d) shape parameters for several durations and thresholds for the Peak Over Threshold method. Thresholds range from on average one to five events per year. The top and bottom of the boxes represent the 75th and 25th percentiles, respectively. The whiskers denote the 1st and 99th percentiles. The dashed gray horizontal line in (d) indicates exponential tails. A shape parameter smaller than zero indicates a thin tail with an upper limit, a shape parameter larger than zero indicates a heavy power-law tail.

3 Shape Parameter

The GEV and MEV distributions are both flexible and able to describe different tail behaviors. The tail behavior the two distributions varies, see Section 2.2.2 in the main manuscript for an overview and Figure S4 for different combinations of scale and shape parameters.

Maps of the shape parameter for the three distributions for daily duration are shown in Figure S5. The spatial patterns of the shape parameters are largely similar to those of the heaviness $h_{T10-T100-T1000}$, described in Section 3.3 and Figures 5 and 6 of the main manuscript. There are, however, some notable differences. For MEV, the mountainous areas are more defined, indicated by higher shape values (less heavy tail behavior). Looking at the shape parameter, we find, however, that the relationship of the Weibull shape parameter with elevation is more complicated. Lower shapes (heavier tails) are generally observed on the leeward side of large mountain ranges, and higher shapes (though still indicative of heavy-tail behavior) on the windward side that is dominated by orographically enhanced frontal precipitation. This is for example visible in the Rocky Mountains, Indonesia, and Norway, and corresponds to findings of Cavanaugh and Gershunov (2015, their Figure 5), who showed that exponential tails are observed in regions where extreme precipitation is predominantly generated by one type of system.

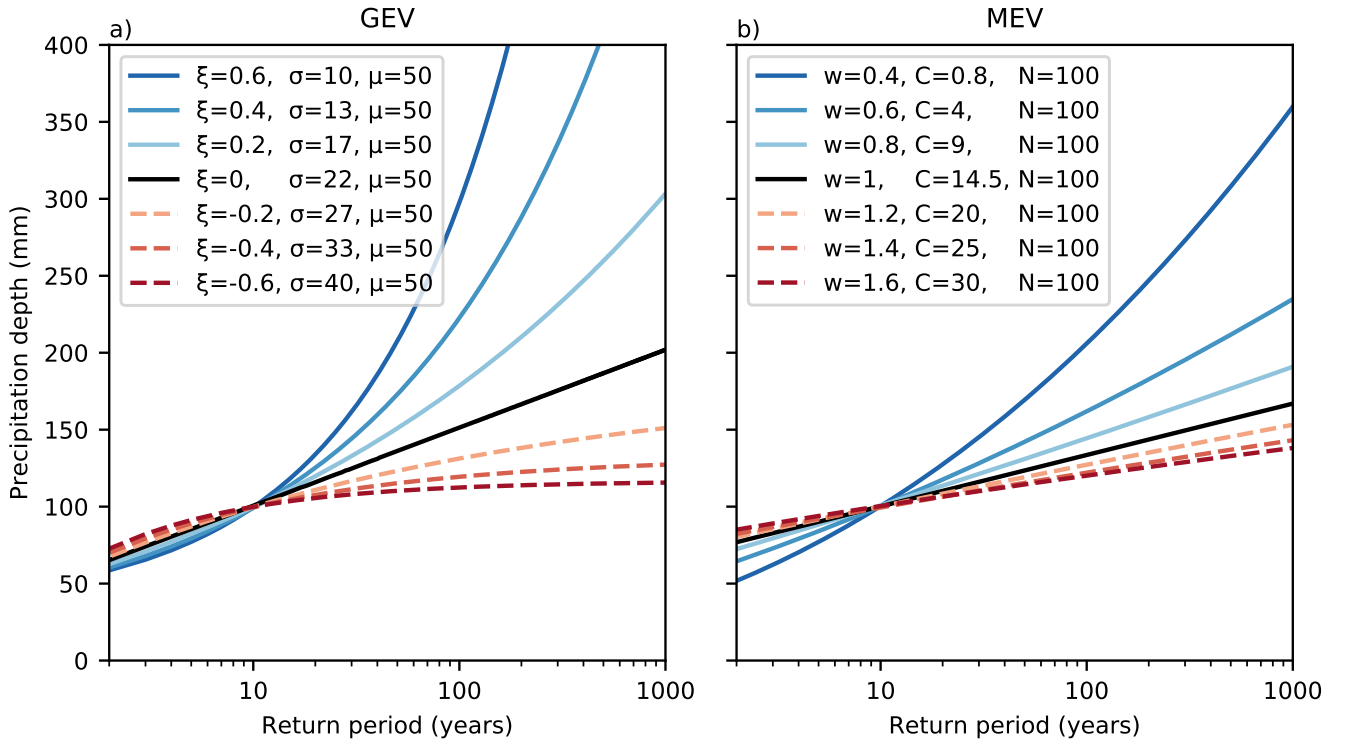


Figure S4: (a) Behavior of the shape (ξ) and scale (σ) parameters of the GEV distribution, with a constant location parameter (μ), and (b) behavior of the shape (w) and scale (C) parameters of the MEV-Weibull distribution, with a constant number of events (N). The results for MEV have been obtained with constant w , C and N parameters for each year. The values of the shape and scale parameter pairs have been chosen such that they all have a precipitation depth of approximately 100 mm for a 10-year return period.

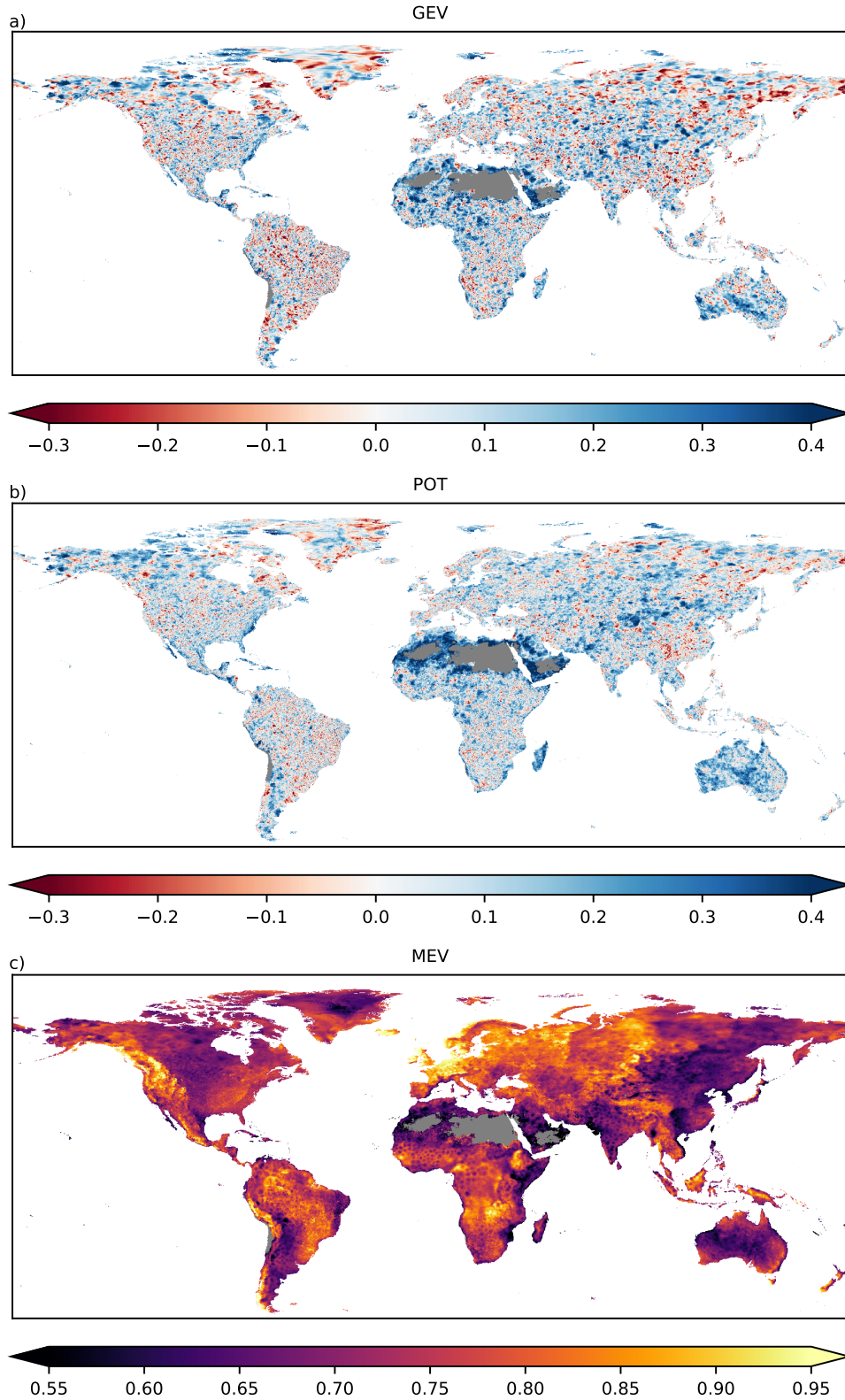


Figure S5: The shape parameter (—) for daily precipitation calculated for different extreme value methods: (a) GEV, equation 1 — ξ_{GEV} , (b) POT, equation 2 — ξ_{GP} and (c) MEV, equation 3 — w . For MEV, the mean shape parameter of all yearly Weibull distributions is displayed. The colorbar min and max are based on the 1st and 99th percentile. For GEV and POT, red indicates a thin shape with an upper limit, white an exponential shape, and blue a heavy power-law shape. For MEV all median shapes are indicative of sub-exponential heavy shapes, though darker colors are heavier than lighter colors.

4 Tail Behavior MEV for Different Return Level Combinations

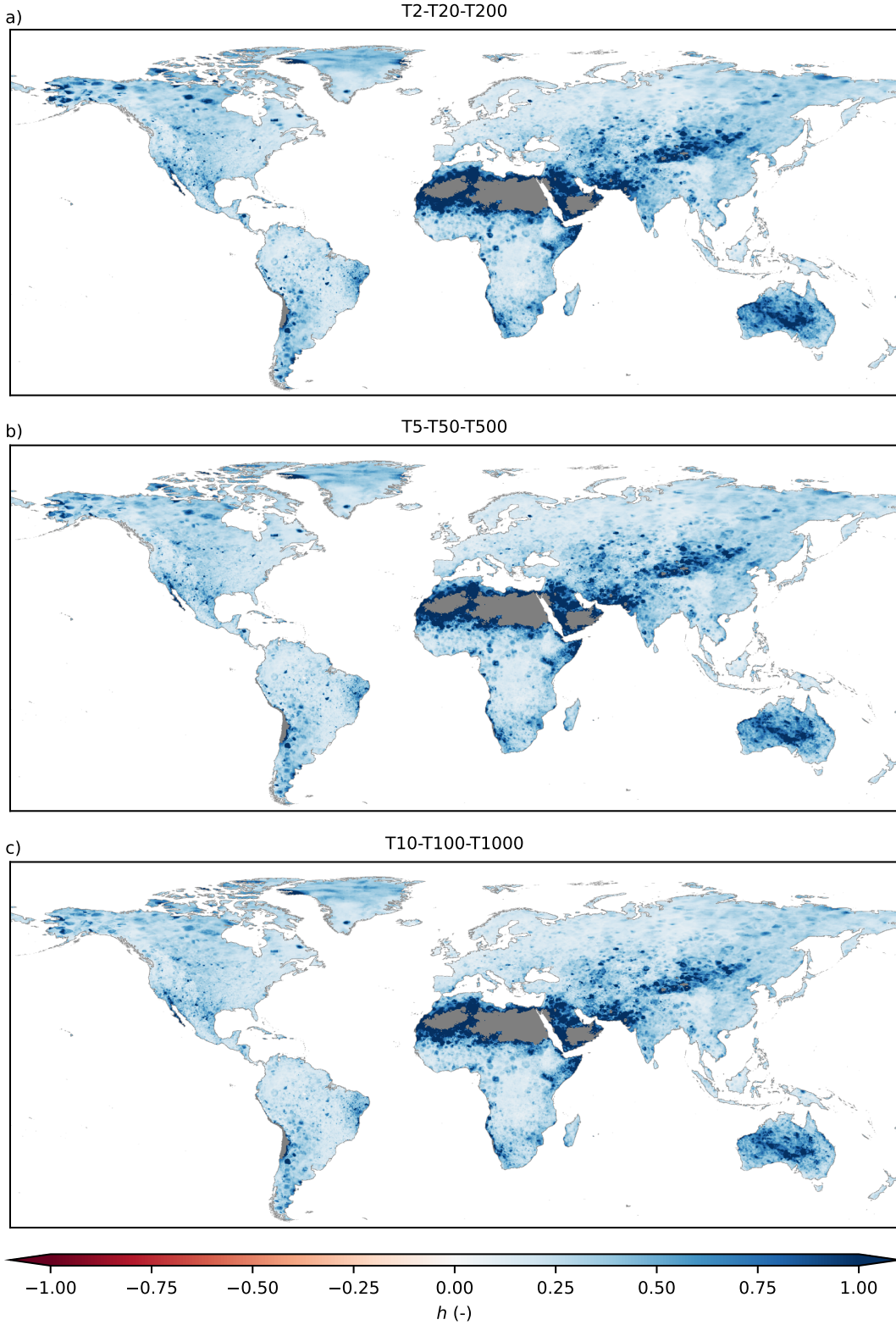


Figure S6: The heaviness amplification factor h (Eq. 5) for daily precipitation calculated for different combinations of return levels for the MEV distribution: (a) $h_{T2-T20-T200}$, (b) $h_{T5-T50-T500}$, (c) $h_{T10-T100-T1000}$. Red indicates a thin tail, white an exponential tail, and blue a heavy tail. See section 2.2.2 for more information on the heaviness metric.

5 Tail Behavior for Multiple Durations

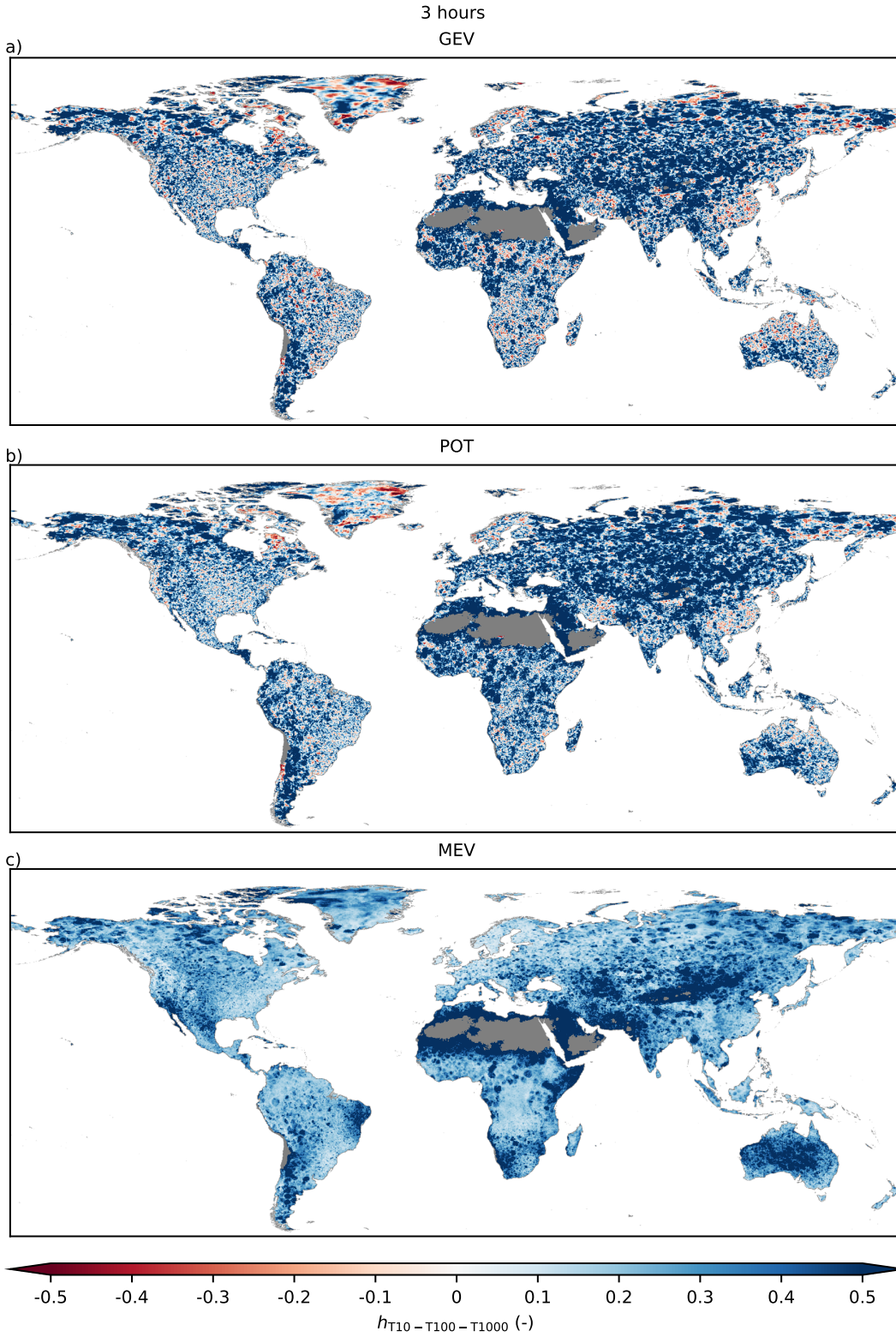


Figure S7: The heaviness amplification factor $h_{T_{10}-T_{100}-T_{1000}}$ (Eq. 5) for 3-hourly precipitation calculated for different extreme value methods: (a) GEV, (b) POT, (c) MEV. Red indicates a thin tail, white an exponential tail, and blue a heavy tail. See section 2.2.2 for more information on the heaviness metric.

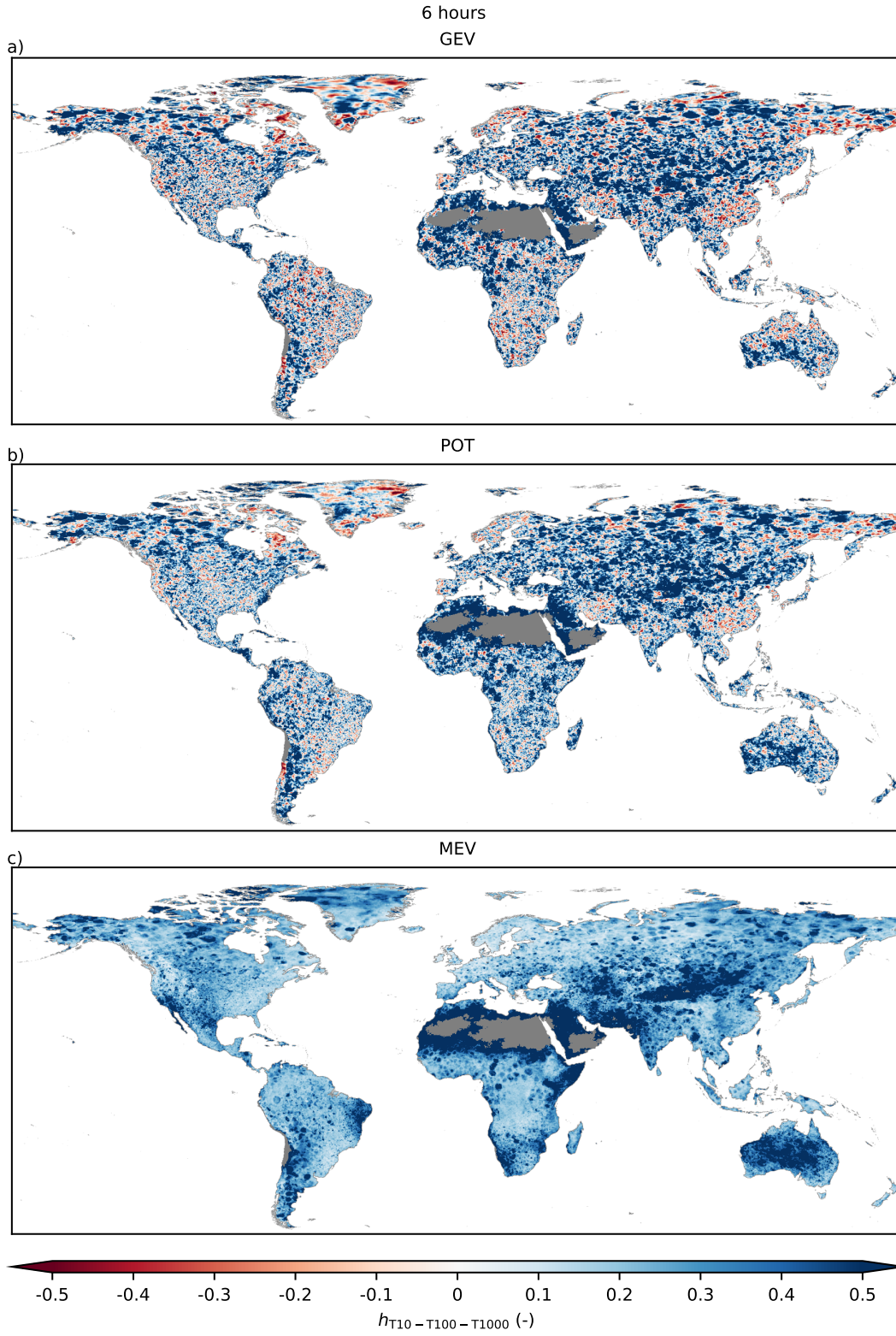


Figure S8: The heaviness amplification factor $h_{T10-T100-T1000}$ (Eq. 5) for 6-hourly precipitation calculated for different extreme value methods: (a) GEV, (b) POT, (c) MEV. Red indicates a thin tail, white an exponential tail, and blue a heavy tail. See section 2.2.2 for more information on the heaviness metric.

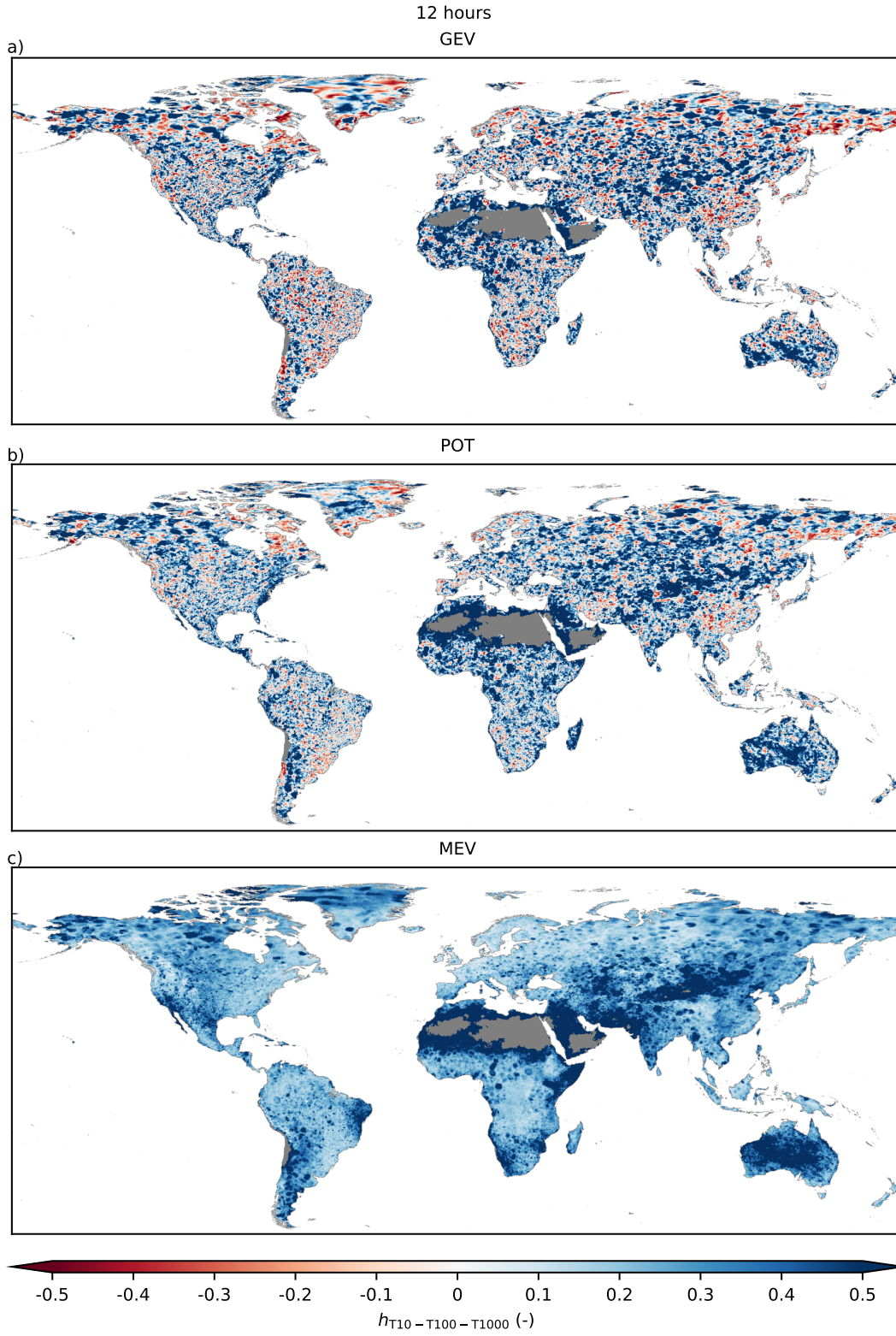


Figure S9: The heaviness amplification factor $h_{T10-T100-T1000}$ (Eq. 5) for 12-hourly precipitation calculated for different extreme value methods: (a) GEV, (b) POT, (c) MEV. Red indicates a thin tail, white an exponential tail, and blue a heavy tail. See section 2.2.2 for more information on the heaviness metric.

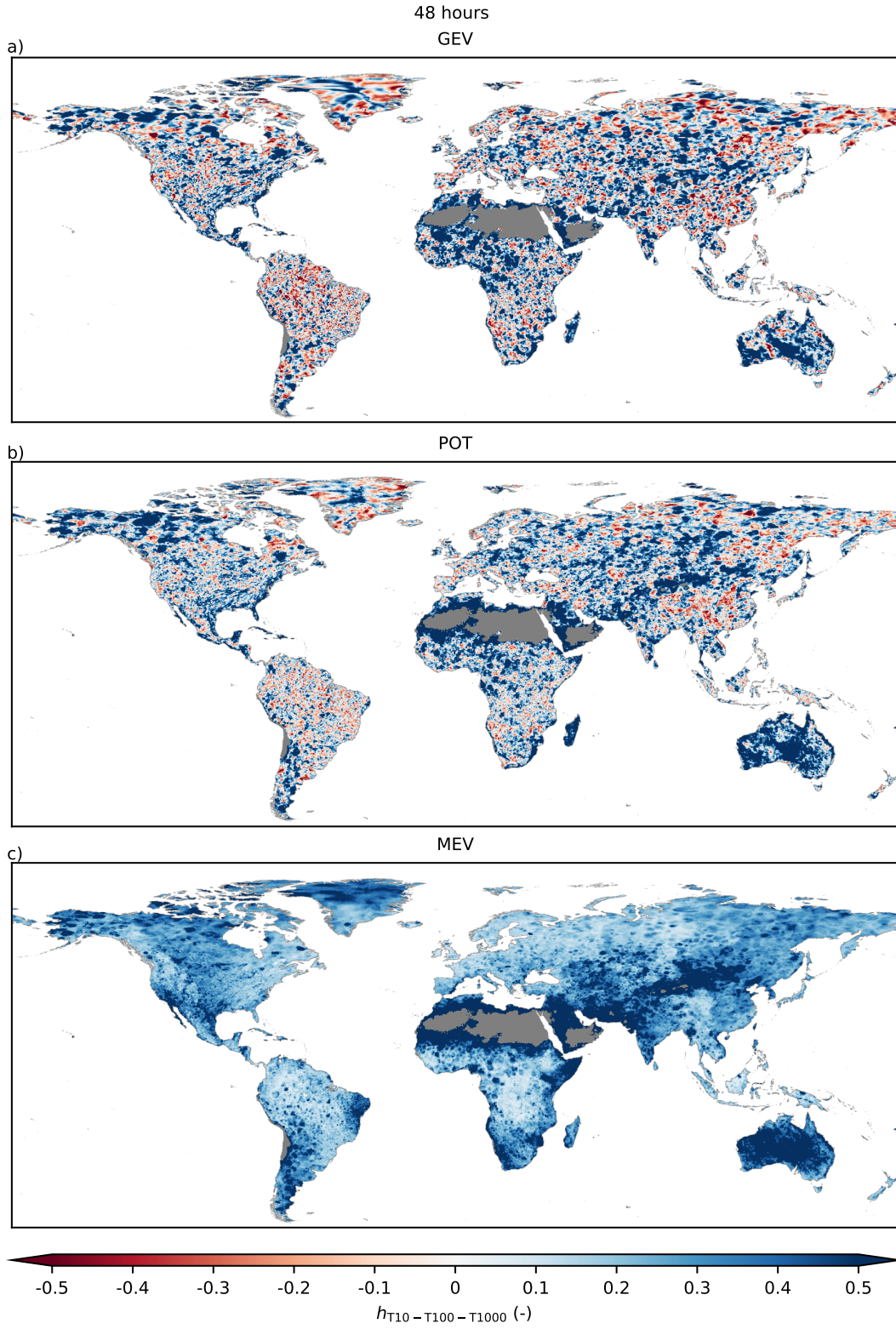


Figure S10: The heaviness amplification factor $h_{T10-T100-T1000}$ (Eq. 5) for 2-day precipitation calculated for different extreme value methods: (a) GEV, (b) POT, (c) MEV. Red indicates a thin tail, white an exponential tail, and blue a heavy tail. See section 2.2.2 for more information on the heaviness metric.

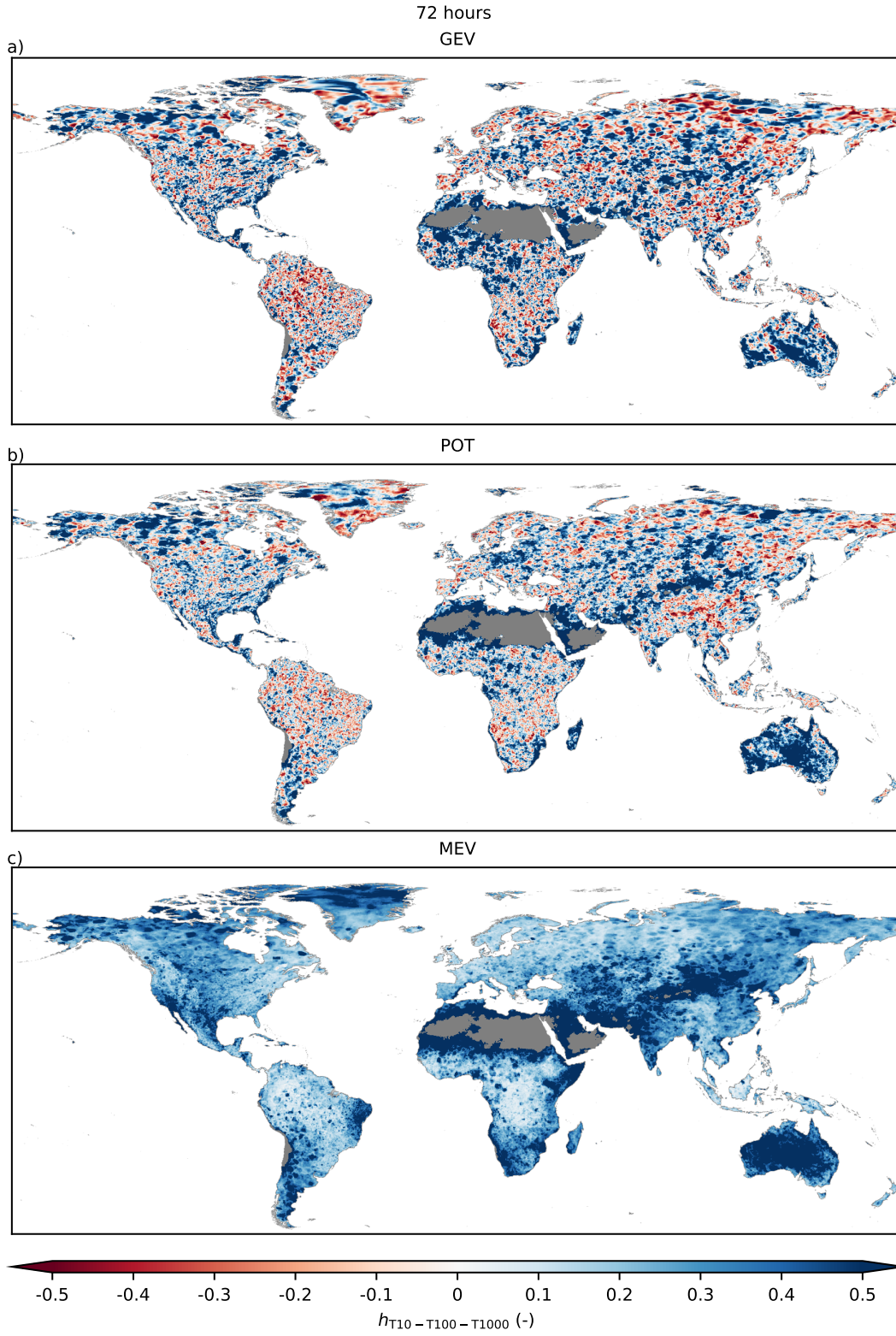


Figure S11: The heaviness amplification factor $h_{T10-T100-T1000}$ (Eq. 5) for 3-day precipitation calculated for different extreme value methods: (a) GEV, (b) POT, (c) MEV. Red indicates a thin tail, white an exponential tail, and blue a heavy tail. See section 2.2.2 for more information on the heaviness metric.

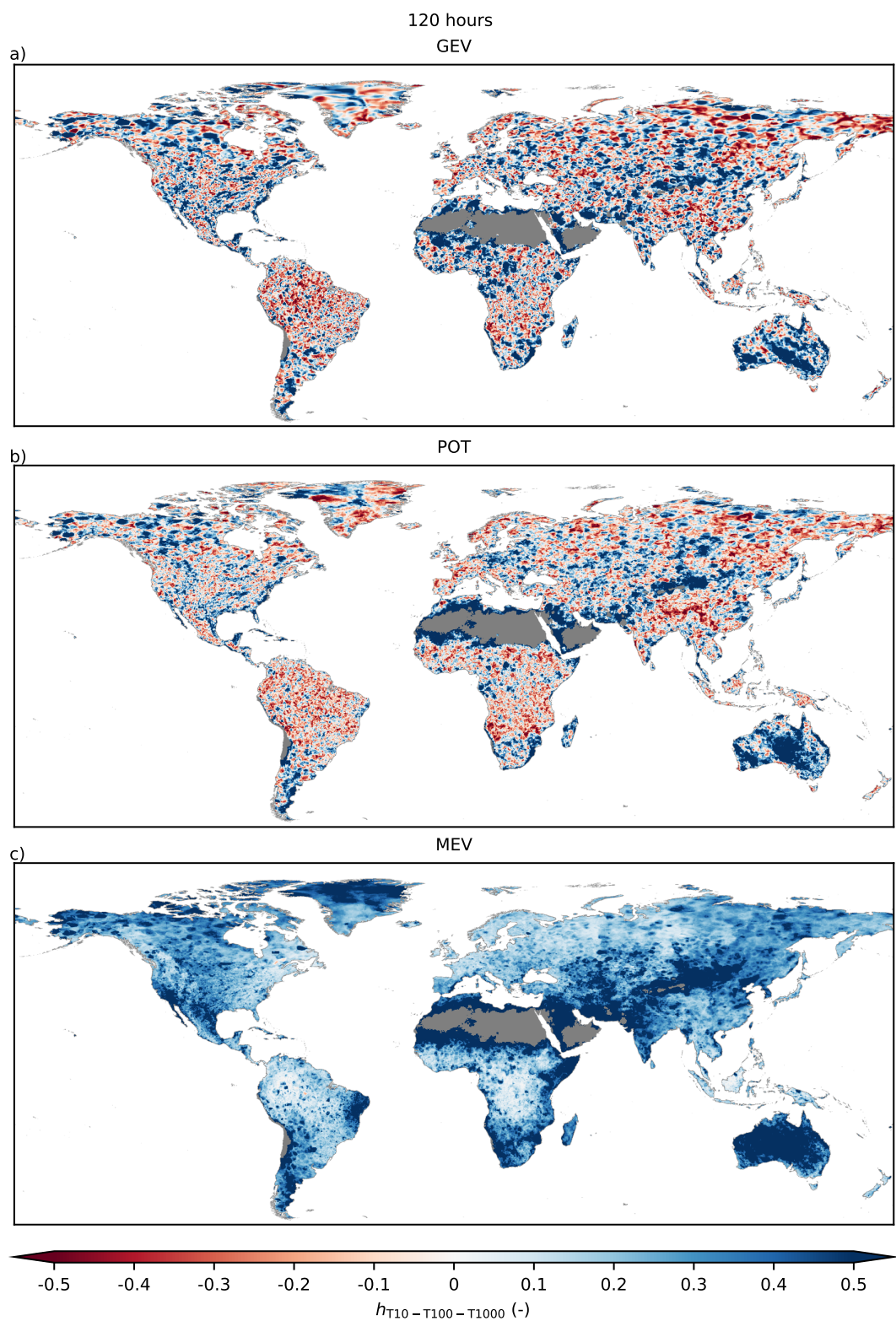


Figure S12: The heaviness amplification factor $h_{T_{10}-T_{100}-T_{1000}}$ (Eq. 5) for 5-day precipitation calculated for different extreme value methods: (a) GEV, (b) POT, (c) MEV. Red indicates a thin tail, white an exponential tail, and blue a heavy tail. See section 2.2.2 for more information on the heaviness metric.

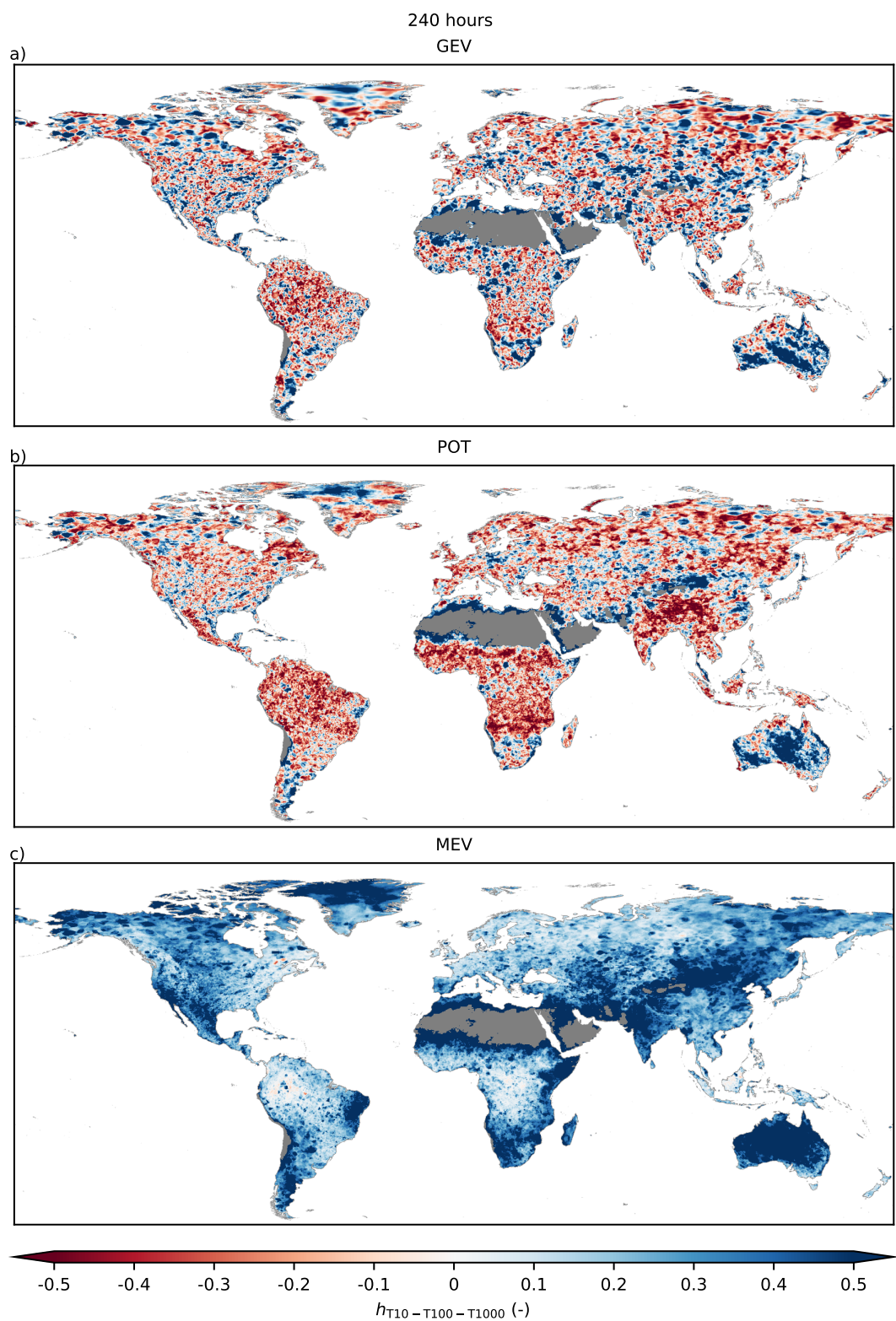


Figure S13: The heaviness amplification factor $h_{T10-T100-T1000}$ (Eq. 5) for 10-day precipitation calculated for different extreme value methods: (a) GEV, (b) POT, (c) MEV. Red indicates a thin tail, white an exponential tail, and blue a heavy tail. See section 2.2.2 for more information on the heaviness metric.

6 Dataset Usage Notes

The GPEX dataset created in this study is available at the 4TU repository (Gründemann et al., 2021). It provides openly accessible and readily available hydrologically relevant return levels of extreme precipitation estimates worldwide. It contains the precipitation estimates of the three extreme values distributions, the observed estimates, the parameters of the three distributions, as well as a few other variables used in this study (Table S1). Furthermore, the extreme precipitation estimates and parameters of the Gumbel distribution are included in the dataset. In this section we provide some possible uses of the dataset, and instructions and disclaimers for proper use, both for large or regional-scale usage as well as for a single cell or point-scale.

Table S1: Variables included in the GPEX dataset.

Variable	Description
GEV estimate	Extreme precipitation return levels estimated using GEV (mm)
POT estimate	Extreme precipitation return levels estimated using POT (mm)
MEV estimate	Extreme precipitation return levels estimated using MEV (mm)
Gumbel estimate	Extreme precipitation return levels estimated using Gumbel (mm)
Observed estimate	Observed extreme precipitation return levels (mm)
GEV location parameter	Location parameter of the GEV distribution
GEV scale parameter	Scale parameter of the GEV distribution
GEV shape parameter	Shape parameter of the GEV distribution
GEV heaviness	Heaviness amplification factor ($h_{T10-T100-T1000}$) of the GEV distribution
POT location parameter	Location parameter for a GEV distribution estimated by fitting the GP distribution
POT scale parameter	Scale parameter for a GEV distribution estimated by fitting the GP distribution
POT shape parameter	Shape parameter for a GEV distribution estimated by fitting the GP distribution
POT heaviness	Heaviness amplification factor ($h_{T10-T100-T1000}$) of the POT distribution
MEV number of events	n parameter of the MEV distribution, number of events per hydrological year
MEV scale parameter	Scale parameter of the MEV distribution
MEV shape parameter	Shape parameter of the MEV distribution
MEV mean number of events	Mean of the n parameter of the MEV distribution, mean number of events per hydrological year
MEV mean scale parameter	Mean of the scale parameter of the MEV distribution
MEV mean shape parameter	Mean of the shape parameter of the MEV distribution
MEV heaviness	Heaviness amplification factor ($h_{T10-T100-T1000}$) of the MEV distribution
Gumbel scale parameter	Scale parameter of the Gumbel distribution
Gumbel shape parameter	Shape parameter of the Gumbel distribution
Annual maxima	Annual maximum precipitation for each hydrological year (mm)
Start hydrological year	Number indicating the month in which the hydrological year starts
Running parameter	Running parameter (hours) of the declustering method by Marra et al. (2018)
Land mask	Mask used for this study to indicate land cells and ocean cells

6.1 Large-Scale Applications

The GPEX dataset contains global scale extreme precipitation estimates and its parameters at a spatial resolution of 0.1° , covering 3-hourly to 10-day durations. The dataset contains information about precipitation extremes for the entire Earth's land surface except Antarctica. The estimates of three distributions described in the main manuscript, as well as the return levels as observed, and estimated using the Gumbel distribution are included in the dataset. Among the three distributions, the traditional GEV and POT provide comparable large-scale average extremes, although differences can be substantial at smaller scales. When using the dataset at regional scales, we advise taking the average of the precipitation estimates, as neighboring cells could differ. Note that since only 38 years of data were available, the fitted model parameters and associated return values are subject to considerable uncertainty. Furthermore, we acknowledge that the use of just one dataset does not represent the true uncertainty in the generation of the dataset created. We do not think this affects our results for observed global spatial patterns significantly, but in a practical setting we recommend verifying the estimates with local observations if available, and to reproduce the precipitation return level estimates with a full uncertainty range estimation.

The novel MEV distribution provides more spatially coherent patterns of the extremes. Its mean shape parameter for daily events captures the (heavy-)tail behavior, and follows orographic patterns. The extremes estimated by MEV are higher than those estimated by GEV and POT. However, for large return periods and long durations, MEV can overestimate the

extremes, due to the small number of events available for the fitting. We, therefore, recommend analyzing the extremes of all distributions in this dataset to obtain an indication of the uncertainty.

6.2 Small-Scale Applications

The dataset is also suitable for small-scale applications either in comparative studies or for direct use in data sparse regions, but one should be aware of the different statistical characteristics of point-scale and grid-scale. Due to averaging effects in gridded datasets, precipitation extremes of point-scale observations are higher (Cavanaugh & Gershunov, 2015; De Michele, Kottegoda, & Rosso, 2001; Ensor & Robeson, 2008; Hu et al., 2020; Sivapalan & Blöschl, 1998; Zorzetto & Marani, 2019). Illustrative examples of two locations, Vienna and San Francisco, are included in Figure S14. Analysis of the return level plots shows the estimates of the three distributions discussed in the main manuscript, as well as the estimates using the Gumbel distribution compared to the observed ones. We converted the annual maximum precipitation to 'observed' return levels (Figure S14e+j). It should be kept in mind though that these 'observed' return levels are also different from the 'true' return levels. For (sub-)daily durations and low return periods, there is generally a good agreement between the observed return levels and the estimates of the three EVDs. For longer durations and return periods, however, the estimated extremes deviate from the observed extremes. This is seen in San Francisco (Figure S14f-j) where MEV overestimates and GEV, POT and Gumbel underestimate the extremes.

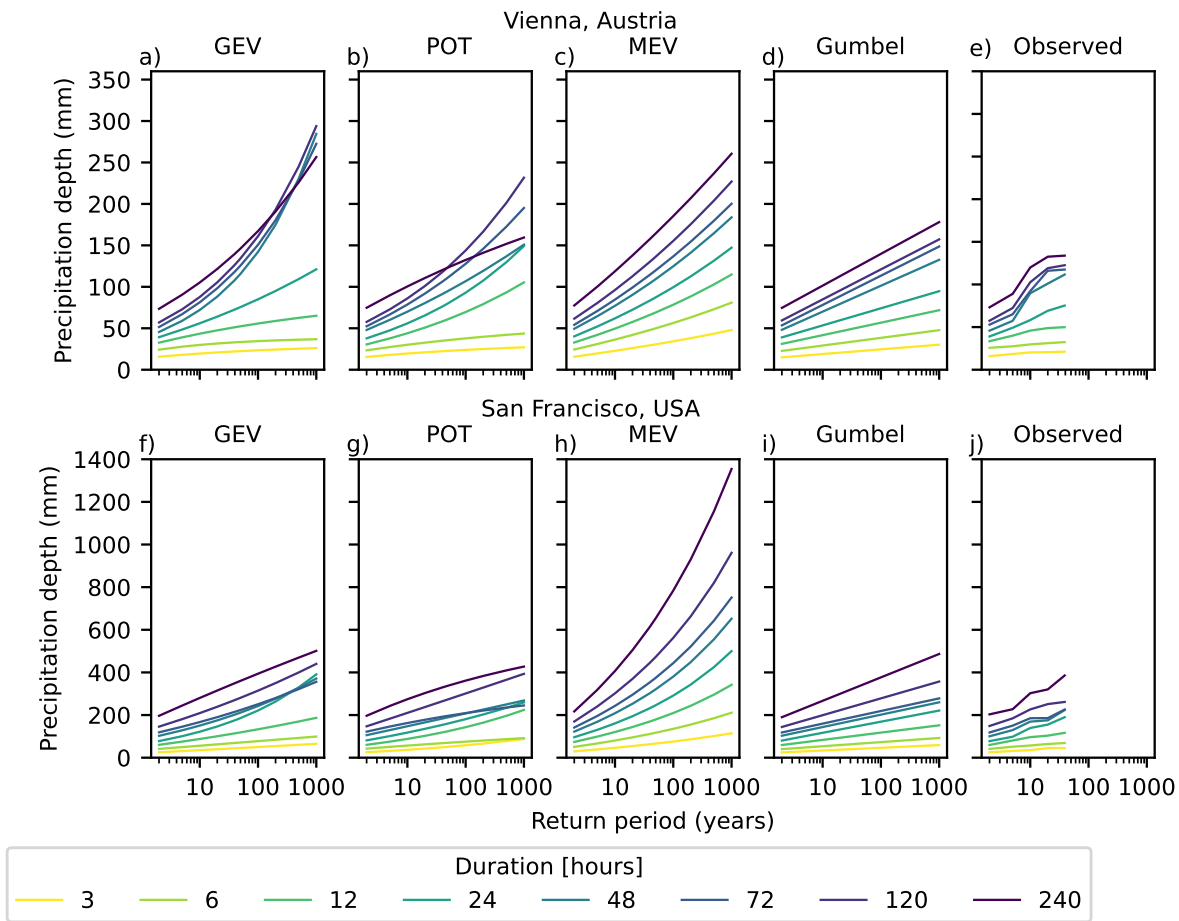


Figure S14: Return level plots for specific locations and different distributions. a–e) Vienna, Austria (48.234°N, 16.415°E) and f–j) San Francisco, California, USA (37.784°N, 122.400°W). Observed (e & j) are the annual maxima converted to return periods.

Furthermore, increasing event durations result in lower shape parameters (less heavy tails), which was seen for all three distributions discussed in the main manuscript. An implication of this is that for long-durations the shape parameter indicates a finite endpoint (GEV and POT), or a very thin tail (MEV), while heavier tails are generally observed for short-durations. When estimating very large return periods (e.g., T_{500}), it is therefore possible for shorter duration estimates to be more intense

than the corresponding quantiles computed for longer durations, which is physically impossible (see also Figure S14a,b,f and g). Additionally, Papalexiou and Koutsoyiannis (2013) argued based on long time series that the true population GEV shape parameter of daily precipitation lies between exponential and heavy-tail behavior, and thin-tails lead to an underestimation of the extremes. For applications where underestimation is undesirable, we have included the extreme precipitation estimates using the Gumbel distribution as well. The Gumbel distribution is an exponential distribution, and equal to the GEV distribution where the shape parameter equals zero. Therefore, if the shape parameter of GEV or POT is negative, the Gumbel estimates could be used instead in order to avoid underestimation.

To get a better understanding of the range and uncertainty of a single cell location, we recommend to look at return level plots of the four distributions at the cell of interest in combination with its neighboring cells. This is particularly important for GEV and POT, due to the absence of coherent spatial patterns and the erratic manifestation of the tail behaviors. Previous results (Zorzetto, Botter, & Marani, 2016) show that the benefits of MEV over GEV are greater for large return periods relative to the sample size available for the fit. Hence, for the estimation of large quantiles, MEV may be presumed to be more accurate. Depending on the practical application one could then choose to use the most extreme value, use the MEV value, use the Gumbel value in case of a negative shape for GEV or POT, or use a spatial average of the GEV and POT estimates.

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