Exploring the environmental drivers of global terrestrial CO2 fluxes inferred from OCO-2 and a geostatistical inverse model

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Abstract

The carbon cycle displays strong sensitivity to short term variations in environmental conditions, and it is key to understand how these variations are linked with variations in CO2 fluxes. Previously, atmospheric observations of CO2 have been sparse in many regions of the globe, making it challenging to evaluate these relationships. However, the OCO-2 satellite, launched in July 2014, provides new insight into global CO2 fluxes, particularly in regions that were previously difficult to monitor. In this study, we combine OCO-2 observations with a geostatistical inverse model to explore data-driven relationships between inferred CO2 flux patterns and environmental drivers. We further use year 2016 as an initial case study to explore the applicability of the geostatistical approach to large satellite-based inverse problems. We estimate daily, global CO2 fluxes at the model grid scale and find that a combination of air temperature, daily precipitation, and photosynthetically active radiation (PAR) best describe patterns in CO2 fluxes in most biomes across the globe. PAR is an adept predictor of fluxes across mid-to-high latitudes, whereas a combined set of daily air temperature and precipitation shows strong explanatory power across tropical biomes. However, we are unable to quantify a larger number of relationships between environmental drivers and CO2 fluxes using OCO-2 due to the limited sensitivity of total column satellite observations to detailed surface processes. Overall, we estimate a global net biospheric flux of -1.73 ± 0.53 GtC in year 2016, in close agreement with recent inverse modeling studies using OCO-2 retrievals as observational constraints.

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31 Key points:

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 34 (2) A combination of photosynthetically active radiation, air temperature and precipitation 35 best describe variations in CO₂ fluxes in most biomes across the globe. 36 	on
 37 (3) The geostatistical approach yields flux totals that are consistent with an OCO-2 inverse 38 modeling inter-comparison. 	rse
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61 Abstract

The carbon cycle displays strong sensitivity to short term variations in environmental conditions, 62 and it is key to understand how these variations are linked with variations in CO₂ fluxes. 63 Previously, atmospheric observations of CO_2 have been sparse in many regions of the globe, 64 making it challenging to evaluate these relationships. However, the OCO-2 satellite, launched in 65 July 2014, provides new insight into global CO_2 fluxes, particularly in regions that were 66 previously difficult to monitor. In this study, we combine OCO-2 observations with a 67 geostatistical inverse model to explore data-driven relationships between inferred CO₂ flux 68 69 patterns and environmental drivers. We further use year 2016 as an initial case study to explore the applicability of the geostatistical approach to large satellite-based inverse problems. We 70 estimate daily, global CO₂ fluxes at the model grid scale and find that a combination of air 71 temperature, daily precipitation, and photosynthetically active radiation (PAR) best describe 72 patterns in CO₂ fluxes in most biomes across the globe. PAR is an adept predictor of fluxes 73 across mid-to-high latitudes, whereas a combined set of daily air temperature and precipitation 74 75 shows strong explanatory power across tropical biomes. However, we are unable to quantify a larger number of relationships between environmental drivers and CO₂ fluxes using OCO-2 due 76 to the limited sensitivity of total column satellite observations to detailed surface processes. 77 Overall, we estimate a global net biospheric flux of -1.73 ± 0.53 GtC in year 2016, in close 78 79 agreement with recent inverse modeling studies using OCO-2 retrievals as observational constraints. 80

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83 1. Introduction

The carbon cycle is closely linked with short term environmental variations, and it is critical to 84 explore the connections between these variations and variability in CO₂ fluxes. However, surface 85 86 atmospheric CO₂ observations in many regions outside of North America and Europe are sparse, making it difficult to constrain CO₂ fluxes or to investigate the sensitivity of the CO₂ fluxes to 87 underlying environmental drivers across many broad regions of the globe (e.g., *Peylin et al.*, 88 2013; Crowell et al., 2019). We define the term "environmental drivers" as any meteorological 89 variables or characteristics of the physical environment that can be modeled or measured and 90 91 may correlate with net ecosystem exchange (NEE). Across North America, dense, continuous atmospheric CO₂ observations from ground-based towers and aircraft make it possible to 92 extensively study the relationships between CO₂ fluxes and these environmental drivers at 93 regional and continental levels (e.g., Peters et al., 2007; Gourdji et al., 2012; Fang and 94 Michalak, 2015; Shiga et al., 2018); by contrast, a paucity of in situ observations for many 95 regions, including the tropics and the Southern Hemisphere, makes it difficult to conduct 96 comparable studies in these regions. 97

- 98 CO₂-observing satellites, including GOSAT (*Kuze et al.*, 2009), TanSat (*Yang et al.*, 2018) and
- 99 OCO-2/OCO-3 (*Crisp*, 2015; *Eldering et al.*, 2019) observe CO₂ broadly across the globe and
- 100 can provide new insight into CO_2 fluxes, particularly in regions that were previously difficult to
- 101 monitor (e.g., the tropics and Southern Hemisphere). To date, observations from the OCO-2
- satellite, launched in July 2014, have been used to constrain CO₂ flux variability at point source
- 103 (e.g., Nassar et al., 2017), regional (e.g., Liu et al., 2017; Chatterjee et al., 2017; Palmer et al.,
- 104 2019) and global scales (e.g., *Crowell et al.*, 2019; *Miller et al.*, 2018; *Wang et al.*, 2019).
- 105 The next challenge is to connect OCO-2 observations and estimated CO₂ fluxes with
- 106 environmental drivers (e.g., *Liu et al.*, 2017; *Chevallier et al.*, 2018). Understanding this
- 107 connection between fluxes and environmental drivers is critical for improving bottom-up or
- 108 process-based flux models (e.g., *Huntzinger et al.*, 2017).
- 109 A geostatistical inverse model (GIM) provides a unique lens to explore these connections.
- 110 Specifically, a GIM does not prescribe or rely on a traditional prior flux model. The choice of
- prior fluxes in a classical inverse model is often subjective, and this choice can impact the
- posterior flux estimate (e.g., *Peylin et al.*, 2013; *Houweling et al.*, 2015; *Phillip et al.*, 2019). For
- example, existing state-of-art terrestrial biosphere models (TBMs) provide divergent flux
- estimates at regional to global scales and display very different sensitivities to environmental
- drivers (e.g., *Huntzinger et al.*, 2017). By contrast, a GIM can assimilate a wide range of
- 116 environmental drivers, making it possible to evaluate data-driven relationship between these
- drivers and the CO₂ fluxes inferred from atmospheric observations (see Sect. 2). Existing GIM
- studies have investigated connections of CO₂ fluxes and environmental drivers for North
- 119 America (Gourdji et al., 2010, 2012; Commane et al., 2017; Shiga et al. 2018) and the globe
- 120 (*Gourdji et al.*, 2008) using a variety of *in situ* CO₂ observations.
- 121 New satellite observations of CO_2 provide a novel opportunity to expand this analysis across the 122 globe. However, the sheer number of observations from satellites like OCO-2 also present novel
- 123 computational and statistical challenges for GIMs that were originally designed for far smaller *in*
- situ CO₂ datasets (e.g., *Miller et al.*, 2019). To overcome this challenge, we combine the GIM
- 125 with the adjoint of a global chemical transport model. Using this framework, we not only
- estimate daily global CO₂ fluxes at the model grid scale (4° latitude $\times 5^{\circ}$ longitude) but also
- 127 quantify posterior uncertainties in the estimated fluxes. The primary purpose of this study is to

128 couple a GIM to a global adjoint model and use this framework to evaluate the relationships

between the environmental drivers and the CO₂ fluxes inferred from OCO-2. We focus on a

single year (i.e., 2016) as an initial case study -- to explore the applicability of the geostatistical

approach to large satellite-based inverse problems. We first describe the implementation of the

132 GIM for OCO-2 observations; we then evaluate and discuss the results of this approach (e.g.,

inferred flux estimates and relationships with environmental drivers) using the 2016 case study.

134 **2. Data and Methods**

135 **2.1 Approach overview**

136 We design a framework that couples the GIM to a global adjoint model (version v35n of the GEOS-Chem adjoint, Henze et al., 2007) and explore the applicability of the geostatistical 137 approach to inverse problems with a large number of flux grid boxes (i.e., $\sim 1.2 \times 10^6$) and a large 138 number of OCO-2 satellite observations (i.e., $\sim 9 \times 10^4$). We use year 2016 as an initial case 139 140 study, as there is better temporal coverage of good-quality data from OCO-2 throughout the entire year relative to years 2015 and 2017. For example, there are 7 week-long gaps in the 141 OCO-2 data in year 2015 and a 1.5-month gap in the OCO-2 data in year 2017, whereas there are 142 no such gaps in year 2016. This time period also overlaps with an OCO-2 inverse modeling 143 inter-comparison (MIP) study, enabling direct comparison with those results (Crowell et al., 144 2019). We specifically estimate CO₂ fluxes for September 1, 2015 to December 31, 2016 but 145 discard the first four months as a spin-up time period. We also offer up a wide range of 146 environmental drivers and allow the GIM to select a subset that best predicts spatiotemporal 147 patterns in CO₂ fluxes at the model grid scale, described in detail below (Sects. 2.2-2.4). 148

149 **2.2. OCO-2** satellite observations

We utilize 10-s average X_{CO2} generated from version 9 of the satellite observations for the period from September 1, 2015 through the end of year 2016 (e.g., *Chevallier et al.*, 2019). We use both land nadir- and land glint-mode retrievals in the inverse model. Recent retrieval updates have eliminated biases that previously existed between land nadir and land glint observations (*O'Dell et al.*, 2018). Moreover, *Miller and Michalak* (2020) evaluated the impact of these updated OCO-

155 2 retrievals on the terrestrial CO_2 flux constraint in different regions of the globe; the authors

found that the inclusion of both land nadir and land glint retrievals yielded a stronger constrainton CO₂ fluxes relative to using only a single observation type.

158 **2.3 Geostatistical inverse model**

159 A GIM does not require an emission inventory or a bottom-up model as an initial guess of

- 160 fluxes; instead, a GIM can leverage a wide range of environmental driver datasets to help predict
- spatial and temporal patterns in the CO₂ fluxes (e.g., *Gourdji et al.*, 2008, 2012; *Shiga et al.*,
- 162 2018). We further pair the GIM with a statistical approach known as model selection to
- 163 objectively determine which set of drivers can best reproduce CO_2 observations from OCO-2.
- 164 This setup makes it feasible to both estimate CO_2 fluxes and to explicitly quantify the
- relationships between the fluxes and the underlying environmental drivers.

The fluxes, as estimated by the GIM, consist of two components. First, the GIM will scale the environmental drivers to match patterns in the atmospheric observations, and this component of the flux estimate is referred to as the 'deterministic model'. Second, the GIM will model spacetime patterns in the CO_2 fluxes that are implied by the atmospheric observations but not explained by any environmental drivers, and this component of the fluxes is referred to as the 'stochastic component'. The best flux estimate is a sum of the deterministic model and the stochastic component:

$$173 \qquad \boldsymbol{s} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\zeta} \tag{1}$$

174 where s are $m \times l$ unknown fluxes, **X** is a $m \times p$ matrix of environmental drivers (see Sect. 2.4), 175 β are $p \times 1$ unknown scaling factors or drift coefficients. These coefficients quantify the relationships between each of the *p* environmental drivers (i.e., each column of **X**) and the 176 estimated CO₂ fluxes. The product of **X** and $\boldsymbol{\beta}$ is the deterministic model (**X** $\boldsymbol{\beta}$). The stochastic 177 component ($\boldsymbol{\zeta}$) is zero-mean with a pre-specified spatial and/or temporal correlation structure; it 178 describes spatial and temporal patterns in the fluxes that are not captured by the deterministic 179 model. For the setup here, the drift coefficient (β) associated with each environmental driver is 180 constant in space and time, while the stochastic component (ζ) varies at the model grid scale. 181 We estimate both the fluxes (s) and the drift coefficients (β) by minimizing the GIM cost 182 function (e.g., Kitanidis and Vomvoris, 1983; Kitanidis, 1995; Michalak et al., 2004): 183 $L_{s,\beta} = \frac{1}{2} (\mathbf{z} - h(\mathbf{s}))^T \mathbf{R}^{-1} (\mathbf{z} - h(\mathbf{s})) + \frac{1}{2} (\mathbf{s} - \mathbf{X}\boldsymbol{\beta})^T \mathbf{Q}^{-1} (\mathbf{s} - \mathbf{X}\boldsymbol{\beta}) \quad (2)$ 184

The cost function includes two components: the first component indicates that the fluxes (s), 185 when run through an atmospheric model, h(s), should match the observations (z) to within a 186 specific error tolerance $(\mathbf{z} - h(\mathbf{s}))$ that is prescribed by the covariance matrix **R** $(n \times n)$. **R** 187 describes model-data mismatch errors, including errors from the atmospheric transport model 188 and the OCO-2 retrievals, among other errors. The second component of Eq. 2 stipulates that the 189 structure of the stochastic component $(\mathbf{s} - \mathbf{X}\boldsymbol{\beta})$ is described by the covariance matrix \mathbf{Q} $(m \times m)$. 190 **O**, like **R**, must be defined by the modeler before estimating the fluxes; it represents the 191 192 variances and spatiotemporal covariances of the stochastic component. We estimate Q using a statistical approach known as Restricted Maximum Likelihood (RML; e.g., Kitanidis, 1997; 193 194 Gourdji et al., 2012; Miller et al., 2016). Q includes both diagonal and off-diagonal elements; the latter decay with the separation time and distance between two model grid boxes. We 195 196 construct **R** as a diagonal matrix with constant elements on the diagonal. Supporting Information Text S1 provides a detailed explanation of the approach used here to estimate the covariance 197 matrix parameters. 198

199 After estimating the covariance matrix parameters, we then estimate the CO_2 fluxes by iteratively

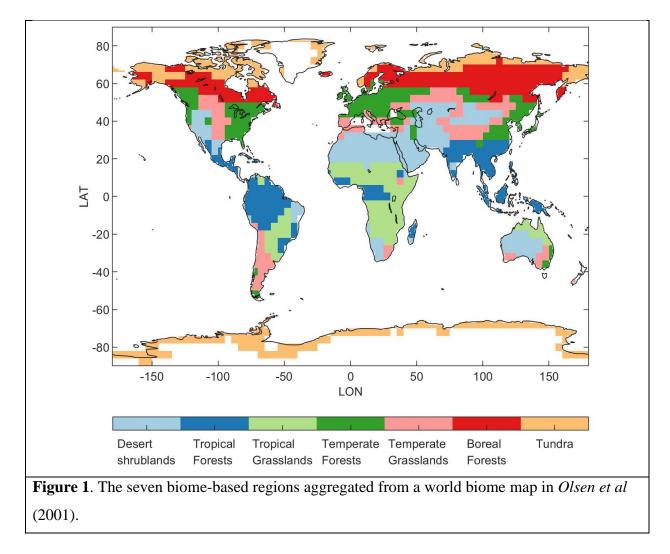
200 minimizing Eq. 2 using the Limited-memory Broyden-Fletcher-Goldfarb-Shanno algorithm (L-

BFGS, *Liu and Nocedal*, 1989). We use this approach to simultaneously estimate both s and β .

202 *Miller et al* (2019) describe this iterative approach to minimize Eq. 2 in detail.

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204 **2.4 Auxiliary environmental drivers**



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206 We consider a wide range of environmental drivers (X). These are meteorological variables

207 primarily related to heat, water, and radiation, available from NASA's Modern-Era Retrospective

Analysis for Research and Applications, Version 2 (MERRA-2, *Rienecker et al.*, 2011).

209 Specifically, we consider daily 2-m air temperature, daily precipitation, 30-day average

210 precipitation, photosynthetically active radiation (PAR), surface downwelling shortwave

radiation, soil temperature at 10-cm depth, soil moisture at 10-cm depth, specific humidity, and

- relative humidity. We also include a non-linear function of 2-m air temperature as an
- environmental driver (refer to hereafter as scaled temperature). This function is from the
- 214 Vegetation Photosynthesis and Respiration Model (VPRM, Mahadevan et al., 2008) and

describes the non-linear relationship between temperature and photosynthesis (e.g., *Raich et al.*

216 1991, see Supporting Information Text S2).

217 Note that we do not include any remote sensing indices (e.g., solar-induced chlorophyll

218 fluorescence (SIF) or leaf area index (LAI)) in the present study. Rather, the focus of this study

is to explore environmental drivers of CO_2 fluxes, not remote sensing proxies for CO_2 fluxes.

220 We group the globe into seven biome-based regions and allow the GIM to use different

environmental drivers in different biomes. The seven-biome map (Fig. 1) is derived from the

biomes in *Olson et al* (2001), aggregated to form larger regions. As a result of this setup, each

column of **X** includes a single environmental driver for a single biome. Therefore, each

environmental driver is represented by a total of seven columns in **X**. Within each column, all

elements are zeros except for elements that correspond to a single biome.

226 We also include several constant columns of ones in **X**. These columns are analogous to the

227 intercept in a linear regression. Existing GIM studies always include one or more constant

columns within X (e.g., Gourdji et al. 2008; Gourdji et al., 2012; Miller et al., 2016). In this

study, we specifically use a total of seven constant columns, one for each biome. We also include

a constant column for the ocean.

231 We further consider non-biospheric fluxes in the **X** matrix, including fossil fuel emissions from

the Open-source Data Inventory for Anthropogenic CO₂ monthly fossil fuel emissions

233 (ODIAC2016, Oda et al., 2018), climatological ocean fluxes from Takahashi et al. (2016), and

biomass burning fluxes from the Global Fire Emissions Database (GFED) version 4.1

(*Randerson et al.*, 2018). We only allocate a single column of **X** for fossil fuel, biomass burning,

and ocean fluxes, respectively, because these fluxes are not the focus of this study.

In total, we consider a total of 81 columns for the X matrix: 8 constant columns of ones, 70

columns associated with environmental drivers, and 3 columns associated with anthropogenic,

239 ocean, and biomass burning fluxes.

240

241 **2.5 Model selection**

242 We utilize a model selection framework to evaluate which subset of the environmental drivers

- 243 (i.e., columns of **X**) best describe variations in CO₂ fluxes as inferred from the OCO-2
- observations. The inclusion of additional environmental drivers or columns in **X** will always
- improve the model-data fit, but the inclusion of too many variables in X can yield an overfit of
- the OCO-2 observations or can yield unrealistic drift coefficients (β) (e.g., Zucchini, 2000).

Instead of including all environmental drivers in X, we use model selection to decide which set 247 of environmental drivers to include in **X**. In this study, we implement a type of model selection 248 249 known as the Bayesian Information Criterion (BIC; Schwarz, 1978), which has been extensively used in recent GIM studies (e.g., Gourdji et al., 2012; Miller et al. 2013; Fang and Michalak, 250 2015). Using the BIC, we score different combinations of environmental drivers that could be 251 252 included in **X** based on how well each combination reproduces the OCO-2 observations. We calculate these scores using the following equation for the implementation here (Miller et al. 253 2018; Miller and Michalak, 2020): 254 $BIC = L + p \ln(n^*)$ (3)255 where L is log likelihood of a particular combination of environmental drivers (i.e., columns of 256 **X**), p is the number of environmental drivers in this particular combination, and n^* is the 257 258 effective number of independent observations. The first component (L) rewards combinations that are a better fit to the observations, whereas the second component in Eq. 3 $(p\ln(n^*))$ 259

260 penalizes models with a greater number of columns to prevent overfitting. The best combination

of environmental drivers for \mathbf{X} is the combination that receives the lowest score (Supporting

Information Text S3 and Table S2). We implement the BIC using a heuristic branch and bound

algorithm (*Yadav et al.*, 2013) to reduce computing time. *Miller et al* (2018) describes this model

selection procedure in greater detail, including the specific setup and equations for the BIC.

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266 **2.6 Posterior uncertainties**

In a GIM, the direct solution to calculate the posterior covariance matrix V_s (dimensions $m \ge m$) can be computed as (e.g., *Saibaba and Kitanidis*, 2014; *Miller et al.*, 2019):

$$269 V_s = V_1 + V_2 V_3 V_2^T (4)$$

270
$$\mathbf{V}_1 = (\mathbf{Q}^{-1} + \mathbf{H}^{\mathrm{T}}(\mathbf{R}^{-1}\mathbf{H}))^{-1}$$
 (5)

271
$$V_2 = V_1 Q^{-1} X$$
 (6)

272
$$\mathbf{V}_3 = (\mathbf{X}^T \mathbf{Q}^{-1} \mathbf{X} - (\mathbf{Q}^{-1} \mathbf{X})^T \mathbf{V}_1 \mathbf{Q}^{-1} \mathbf{X})^{-1}$$
 (7)

where the j	posterior error o	covariance matrix	V _s is the sum of	V ₁ and `	$V_2V_3V_2^T$, and H is a r	ı ×m
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- 275 matrix describing the footprint sensitivity of the observations (z) to the fluxes (s). Note that V_1 is
- the posterior error covariance matrix in a classic Bayesian inverse model (e.g., *Rodgers*, 2000;
- 277 Brasseur and Jacob, 2017). $V_2V_3V_2^T$ accounts for the additional uncertainty in the fluxes due to
- 278 the unknown drift coefficients (β).
- 279
- 280 The calculations in Eq. 5 are not computationally feasible for most inverse problems with very
- 281 large datasets; the matrix sum in V_1 is often too large to invert, and we do not explicitly construct
- an **H** matrix or its transpose \mathbf{H}^{T} . Instead, we employ a low-rank approximation of \mathbf{V}_{1} that
- 283 circumvents these problems. Specifically, we approximate the matrices in V_1 as a low rank
- update to **Q** using a limited number of eigenpairs (i.e., eigenvectors and eigenvalues). *Miller et*
- *al* (2019) and the Supporting Information Text S4 describe the uncertainty quantification in
- 286 greater detail.

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288 **3. Results and Discussion**

3.1 Connections between CO₂ fluxes and environmental drivers

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299	Table 1 . Estimated drift coefficients (β) and associated uncertainties in β for environmental
300	drivers selected using the BIC

Biomes	Selected environmental drivers	Drift coefficients (β)	Uncertainties in β , with 95% confidence interval [*]
Boreal forests	PAR	-1.59	0.16
Temperate grasslands	Daily precipitation	-0.15	0.05
	PAR	-0.29	0.04
Temperate forests	Daily precipitation	-0.36	0.03
	PAR	-0.81	0.03
Tropical grasslands	Daily precipitation	-0.55	0.06
	Scaled temperature	-0.35	0.04
Tropical forests	Daily precipitation	-0.23	0.05
	PAR	0.27	0.05
	Scaled temperature	-0.04	0.02
Desert and	Daily precipitation	-0.27	0.03
shrublands	Scaled Temperature	-0.07	0.01

301 *Supporting Information Text S5 provides detail on the calculations of uncertainties in β .

A small number of environmental drivers can describe most spatiotemporal variability in CO₂

fluxes as estimated in the GIM. In this study, we define spatiotemporal variability as any spatial

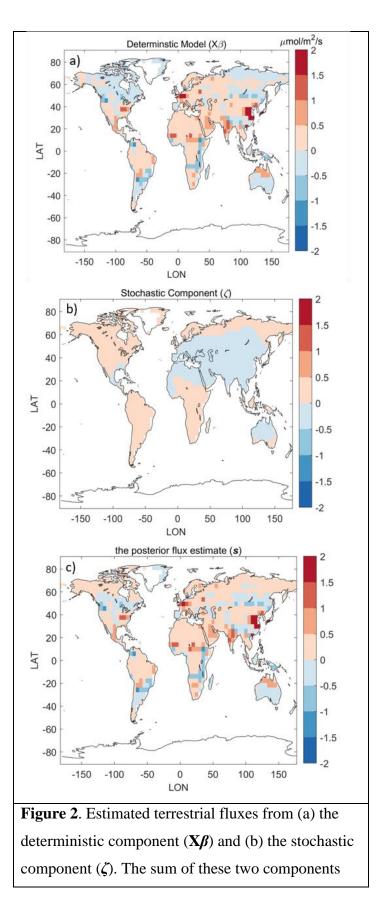
or temporal patterns in CO₂ fluxes that manifest at the daily, 4° (latitude) $\times 5^{\circ}$ (longitude)

resolutions of the GEOS-Chem model during the one-year study period (year 2016). The

deterministic model accounts for ~89.6% of the variance in the estimated fluxes (Fig. 2a), and

the stochastic component conversely accounts for only 10.4% of the flux variance (Fig. 2b).

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equals (c) the posterior flux estimates (*s*). Here the posterior flux estimates include contributions from all source types, including flux patterns that map onto fossil fuels from ODIAC2016.

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A combination of PAR, daily temperature, and daily precipitation best describe patterns in CO₂ fluxes in most biomes across the globe (Table 1). PAR is an adept predictor of fluxes across midto-high latitudes, whereas a combined set of daily air temperature and daily precipitation are better predictors across tropical biomes.

313

The deterministic model also includes fossil fuel emissions from ODIAC2016 but not biomass 314 315 burning fluxes from GFED or ocean fluxes from Takahashi et al., (2016). Fossil fuel emissions from ODIAC2016, when passed through the GEOS-Chem model, help describe enough 316 variability in the OCO-2 observations to be selected using the BIC. By contrast, neither biomass 317 318 burning fluxes from GFED nor ocean fluxes from Takahashi et al. (2016) help reproduce the OCO-2 observations more than the penalty term in the BIC, and these fluxes are therefore not 319 selected using the BIC. Specifically, biomass burning and ocean fluxes may not have been 320 selected for several reasons: either those fluxes are small relative to fossil fuel emissions and 321 322 NEE, the land OCO-2 observations from 2016 are not sensitive to biomass burning and ocean fluxes, and/or the flux patterns in GFED and Takahashi et al., (2016) are not consistent with the 323 324 OCO-2 observations. Instead, biomass burning and ocean fluxes are included within the stochastic component of the flux estimate. 325

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327 Overall, we only select a limited number of environmental drivers (12 out of 70, ~18%) using 328 model selection. Specifically, we never select more than 3 environmental drivers in any individual biome (Table 1). This result indicates two likely conclusions. First, a few simple 329 330 linear relationships may adeptly describe flux variability at the scale and resolution of a global 331 gridded atmospheric model, although the underlying leaf- and organism-level processes are admittedly more complex. Indeed, previous top-down studies (e.g., Gourdji et al., 2008, 2012; 332 Fang and Michalak, 2015; Shiga et al., 2018) also found that simple linear relationships can 333 effectively describe broad spatial and temporal patterns in CO₂ flux variability across North 334

America and across the globe. Such simple linear relationships allow for a straightforward

assessment of the explanatory power of environmental drivers, and make it possible to compare

these relationships inferred from atmospheric observations against the relationships used in

TBMs (e.g., *Huntzinger et al.*, 2013; *Fang and Michalak*, 2015).

339 Second, additional environmental drivers, when run through an atmospheric transport model and

interpolated to the times and locations of OCO-2 observations, are not sufficiently unique to

parse out their differing relationships with CO_2 fluxes. Model selection ensures that we only

include environmental drivers that contribute unique information to the flux estimate and do not

343 overfit the OCO-2 observations. If multiple environmental drivers are highly correlated or

colinear, then the inclusion of more than one of these drivers will not contribute unique

information. As a result, we are unable to quantify a larger number of environmental driver

relationships using OCO-2. Fig. 3 illustrates an example of air temperature and PAR. In most of

347 the biomes, there is a weak correlation (R < 0.4; left column) between 2-m air temperature and

PAR; however, the correlation is much stronger (R > 0.8; right column) when these

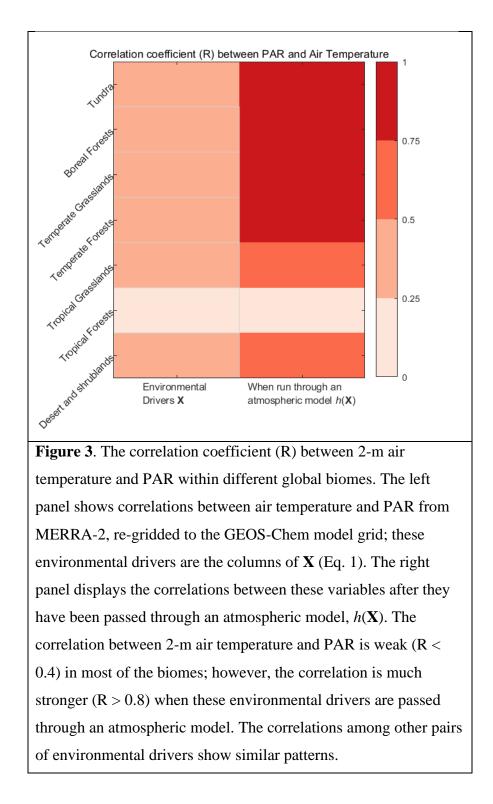
environmental drivers are passed through an atmospheric model $(h(\mathbf{X}))$. A larger number of

environmental drivers is not selected due to this high level of correlation or collinearity among

the columns in $h(\mathbf{X})$. This collinearity, not errors in the OCO-2 retrievals or atmospheric model,

appears to be a limiting factor in the model selection results.

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357 3.1.1 PAR shows stronger explanatory power than temperature or precipitation in mid-to 358 high latitudes

PAR is selected for four biomes: temperature forests, temperate grasslands, boreal forests and 359 tropical forests (Table 1). In the middle and high latitudes, PAR, rather than temperature or 360 361 precipitation, appears to better reproduce seasonal patterns in CO₂ fluxes. This result reflects the 362 fact that light availability is likely an important factor that drives CO₂ flux variability in mid-tohigh latitudes (e.g., Fang and Michalak, 2015; Baldocchi et al., 2017). The *β* values for PAR 363 indicate a strong to moderate negative correlation with estimated CO₂ fluxes, suggesting that an 364 increase (or decrease) in PAR is associated with a decrease (or increase) in NEE, and an increase 365 366 (or decrease) in carbon uptake; this β value is larger in boreal and temperate forests relative to grasslands, indicating a stronger relationship between PAR and net biosphere CO₂ fluxes in those 367 368 biomes (Table 1; Fig. 4a). 369 Indeed, previous studies also indicate that PAR and similar environmental drivers (e.g., 370 shortwave radiation) are closely associated with CO_2 fluxes. For example, a top-down study of North America (Fang and Michalak, 2015) found that shortwave radiation is more adept than 371 372 other environmental variables in reproducing spatiotemporal variability of NEE, particularly

across the growing season. Moreover, several site-level studies have reached parallel conclusions

374 (e.g., *Mueller et al.*, 2010; *Yadav et al.*, 2010); these studies indicated that PAR is strongly

375 correlated with photosynthesis, consistent with current mechanistic understandings of the light

limitation on photosynthesis (e.g., *Gough et al.*, 2007).

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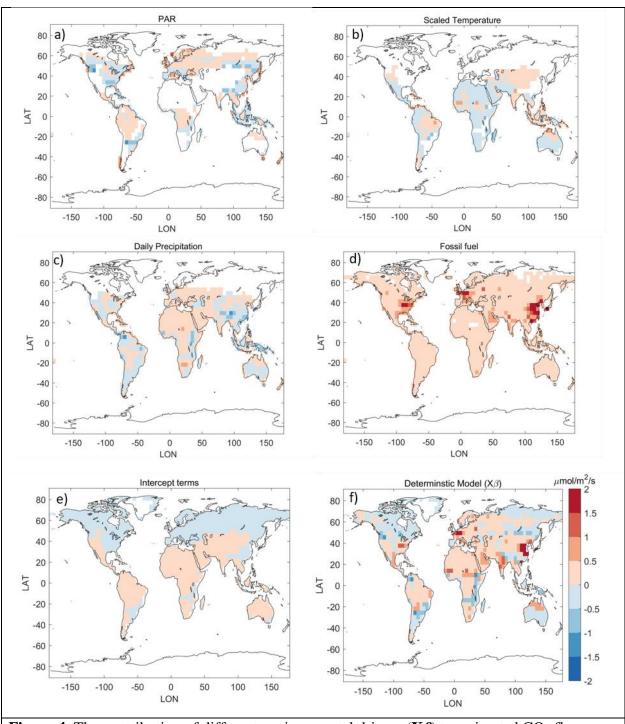


Figure 4. The contribution of different environmental drivers $(\mathbf{X}\boldsymbol{\beta})$ to estimated CO₂ fluxes from the GIM. The individual panels display the contribution of a) PAR, b) scaled temperature, c) daily precipitation, d) fossil fuel, e) the intercept terms, and f) the full deterministic model $(\mathbf{X}\boldsymbol{\beta})$. White colors in panels (a-c) reflect the fact that not all environmental drivers are selected in all biomes.

379 **3.1.2 Drought is likely associated with flux variations across tropical forests**

A composite of PAR, scaled temperature, and daily precipitation adeptly describe variability in

381 CO₂ fluxes across tropical forests, as seen through the OCO-2 observations. PAR in tropical

forests is usually a function of the presence or absence of clouds (e.g., *Baldocchi et al.*, 2017;

383 *Zeri et al.*, 2014); cloudiness is also associated with rainfall. Therefore, low PAR over tropical

forests is likely an indicator of cloud presence and rainfall. A positive β estimated for PAR

suggests that a decrease in PAR, indicative of enhanced precipitation, is associated with

increased carbon uptake. Furthermore, the negative β value assigned to scaled temperature

387 (Supporting Information Text S2) implies that an increase in air temperature, which often

exceeds optimal temperature over tropical forests, is associated with reduced carbon uptake.

389 Recent studies (e.g., Jiménez-Muñoz et al., 2016; Liu et al., 2017; Palmer et al., 2019) indicated

that tropical droughts associated with the 2015-2016 El Niño events likely resulted in above

average carbon release. Indeed, the combination of high values of PAR, high air temperature,

and low precipitation may be a manifestation of these drought impacts.

Indeed, multiple lines of evidence indicate that drought is associated with diminished carbon
uptake in tropical forests (e.g., *Philips et al.*, 2009; *Brienen et al.*, 2015; *Baccini et al.*, 2017).

For example, *Gatti et al* (2014) suggested that a suppression of photosynthesis during tropical

drought may cause a reduction in carbon uptake. *Brienen et al* (2015) added that tropical drought

is often associated with higher-than-normal temperature, which may further contribute to

reducing gross primary production (GPP) and carbon uptake. Overall, this GIM study supports

the conclusion that environmental conditions indicative of drought are associated with net carbonemissions from tropical forests.

3.1.3 CO₂ fluxes, as inferred from OCO-2, are closely correlated with temperature and precipitation in tropical grasslands

403

Temperature and precipitation closely correlate with variability in CO₂ fluxes across tropical
grasslands. This result suggests that heat and water availability are likely associated with carbon
fluxes across this biome.

A negative β value for precipitation indicates that an increase in precipitation is associated with 407 an increase in carbon uptake, which is in line with current knowledge that water availability 408 409 facilitates photosynthesis, especially in arid or semi-arid regions. In addition, a negative β value for scaled temperature (Supporting Information Text S2) indicates that an increase in air 410 temperature is associated with a reduction in carbon uptake. Specifically, high temperatures in 411 412 the tropics often exceed the optimal temperature for photosynthesis (e.g., *Baldocchi et al.*, 2017), which can suppress GPP (e.g., Doughty and Golden, 2008). Overall, a combined set of air 413 temperature and precipitation adeptly describes CO_2 flux variability in tropical grasslands, 414 rendering it a net source in year 2016. 415

416

417 3.2 Estimated biospheric flux totals for different global regions

We estimate a global terrestrial biospheric CO₂ budget of -1.73 ± 0.53 GtC (Uncertainties listed 418 are the 95% confidence interval. Supporting Information Text S5 provides detail on the posterior 419 uncertainty estimate for biospheric fluxes.). Among the seven biomes, middle to high latitudes 420 (primarily temperate, boreal and tundra biomes) act as a significant carbon sink; tropical biomes 421 are a net source; desert and shrubland regions play a small, neutral role (Table 2). Note that we 422 subtract flux patterns that map onto fossil fuels ($X\beta$, Fig. 4d) from the posterior flux estimate (s, 423 Fig. 2c) to obtain an estimate for biospheric fluxes (including terrestrial NEE and biomass 424 425 burning fluxes). We estimate a β value of 1.09 ± 0.05 (95% confidence interval) for the fossil fuel emissions from ODIAC2016, indicating that the overall global magnitude of ODIAC2016 is 426 consistent with OCO-2 observations. We therefore assume that ODIAC2016 is a reasonable 427 428 global estimate for fossil fuel emissions.

Table 2. *Biospheric CO*₂ *flux totals estimated for different global biomes*

biomes	Tundra	Boreal	Temperate	Temperate	Tropical	Tropical	Deserts/shrublands
		forests	grasslands	forests	grasslands	forests	
Flux budget (Gt	-0.01 ±	-0.62	-1.71 ±	-1.78 ±	1.21 ±	1.16 ±	0.02 ± 0.30
C yr ⁻¹ , with	0.31	± 0.25	0.43	0.27	0.44	0.76	
95% confidence							
interval)							

431 These flux totals are broadly consistent with a recent MIP of different inverse models that

432 assimilate OCO-2 observations (*Crowell et al.*, 2019). The inverse modeling teams that

433 participated in the MIP employed different transport models, inverse modeling approaches, and

434 prior flux assumptions. The total global terrestrial biospheric flux, averaged across all models,

435 was -1.4 ± 0.7 GtC for the year of 2016. The MIP fluxes assimilate v7 of land nadir-mode X_{CO2}

436 retrievals; unlike this study in which we use v9 of land nadir- and glint-mode retrievals. In spite

437 of this difference, the averaged global flux from the MIP study and the estimate reported here are

438 very similar.

439 In order to provide an additional comparison with the MIP results, we group the estimated fluxes

440 into TRANSCOM land regions (*Gurney et al.*, 2002). We split the classic TRANSCOM regions

441 at the Equator to avoid regions that encompass parts of both the northern and southern

hemisphere, as in *Crowell et al* (2019). In most of the regions, the fluxes estimated using the

GIM are very similar to those reported in the MIP (Fig. 5); however, the fluxes estimated here

are significantly different in a limited number of regions (e.g., tropical Australia and northern

tropical Africa), a possible reflection of differences between the v9 and v7 OCO-2 retrievals

446 (*O'Dell et al.*, 2018; *Miller et al.*, 2019). For example, we estimate a smaller CO₂ source for

447 northern tropical Africa relative to the MIP study. However, previous studies (e.g., *Wang et al.*,

448 2019) indicated that existing satellite-based estimates of CO₂ fluxes for this region may be too

high. OCO-2 collects far more observations across northern Africa during the dry season than the

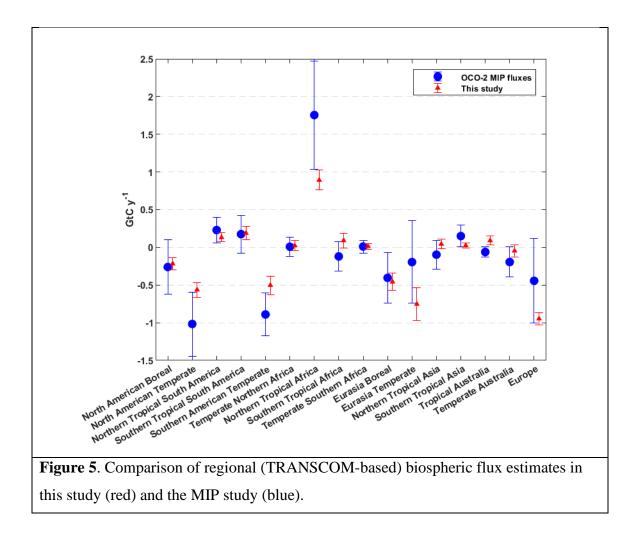
450 wet season due to persistent cloudiness in the wet season, and existing studies have postulated

451 that this difference in data availability may be to blame for a high bias in CO₂ fluxes estimated

452 from OCO-2 (*Crowell et al.* 2019; *Wang et al.* 2019).

The fluxes estimated here are also broadly consistent with aircraft-based CO₂ observations, a
topic discussed in the Supporting Information Text S6.

In prep. for JGR Atmos.



456

457 **3.3 Estimated posterior uncertainties**

The posterior uncertainties for individual biomes range from 0.25 to 0.76 GtC yr⁻¹. Estimated fluxes for tropical forests have higher uncertainties than any other biome (0.76 GtC yr⁻¹), likely a consequence of poor observational coverage due to persistent cloudiness. By contrast, a large number of good-quality OCO-2 retrievals provides robust constraints over temperate forests, yielding a small posterior uncertainty (0.27 GtC yr⁻¹) in the estimated flux.

It is important to note that the posterior uncertainties calculated in most classical Bayesian or

- 464 geostatistical inverse models account for many but not all possible sources of uncertainty. For
- 465 example, the posterior uncertainties presented here account for the sparsity of the OCO-2
- 466 observations, random observational or atmospheric transport errors, and uncertainties due to
- 467 uncertain drift coefficients (β) (e.g., *Kitanidis and Vomvoris*, 1983; *Michalak et al.*, 2004).

However, these calculations do not fully account for bias-type errors: regional- or continental-468 scale biases in the OCO-2 observations, biases in modeled atmospheric convection (e.g., Basu et 469 470 al., 2018; Schuh et al., 2019), or biases in modeled interhemispheric transport, among other possible biases. Most classical Bayesian and geostatistical inverse models assume that the 471 observational or model errors are Gaussian with a mean of zero (e.g., Kitanidis and Vomvoris. 472 1983; Michalak et al., 2004; Tarantola, 2005), making it challenging to account for the types of 473 biases listed above. As a result, the posterior uncertainties estimated in this study are typically 474 smaller than the range of flux estimates produced from the recent MIP study (Fig. 5; Crowell et 475 al., 2019). 476

477

478 4. Conclusions

In this study, we couple a GIM to a global adjoint model and evaluate the data-driven

relationships between environmental drivers and CO₂ fluxes inferred from OCO-2. Using year

481 2016 as an initial case study, we explore the applicability of the geostatistical approach to large

482 satellite-based inverse problems. We find that

(1) A combination of air temperature, daily precipitation, and PAR best describe patterns in
CO₂ fluxes in most biomes across the globe;

(2) PAR is an adept predictor of fluxes across mid-to-high latitudes, whereas a combination
of daily air temperature and daily precipitation shows strong explanatory power across
tropical biomes;

488 (3) A larger number of environmental driver datasets is not selected because they are not

489 sufficiently unique to parse out their differing relationships with CO_2 fluxes using OCO-2.

This high collinearity, not errors in the OCO-2 retrievals or atmospheric model, appears to bea limiting factor;

(4) We estimate a global terrestrial biospheric budget of -1.73 ± 0.53 GtC in year 2016, in

493 close agreement with recent inverse modeling studies that use OCO-2 retrievals as

494 observational constraints.

495

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- *Data availability*
- 503 The version 9 of 10-s average OCO-2 retrievals are available at
- 504 ftp://ftp.cira.colostate.edu/ftp/BAKER/; data information of the OCO-2 MIP is available at
- 505 https://www.esrl.noaa.gov/gmd/ccgg/OCO2/; data information of the ObsPack data product is
- 506 available at *http://www.esrl.noaa.gov/gmd/ccgg/obspack/*.

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